

Stats205_HW3

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Problem 1

```
options(digits=5)
Prior.1 <- epi.betabuster(mode = 0.75, conf = 0.05, greaterthan = F, x = 0.6)
Prior.2 <- epi.betabuster(mode = 0.01, conf = 0.99, greaterthan = F, x = 0.02)
Prior.3 <- epi.betabuster(mode = 1, conf = 0.01, greaterthan = F, x = 0.8)

capture.output(as.data.frame(Prior.1, row.names = 'Prior One'))

## [1] "          shape1 shape2 mode    mean  median   lower  upper  variance"
## [2] "Prior One 23.567 8.5223 0.75 0.73442 0.73934 0.57161 0.8696 0.0058946"

capture.output(as.data.frame(Prior.2, row.names = 'Prior Two'))

## [1] "          shape1 shape2 mode    mean  median   lower  upper  variance"
## [2] "Prior Two 11.035 994.47 0.01 0.010975 0.010652 0.0055035 0.018276 1.0784e-05"

capture.output(as.data.frame(Prior.3, row.names = 'Prior Three'))

## [1] "          shape1 shape2 mode    mean  median   lower  upper  variance"
## [2] "Prior Three 20.638      1    1 0.95379 0.96697 0.83632 0.99877 0.0019471"
# knitr::kable(rbind(Prior.1, Prior.2, Prior.3) , digits = rep.int(4, 8))
df <- rbind(Prior.1, Prior.2, Prior.3)
pander(df)
```

Table 1: Table continues below

	shape1	shape2	mode	mean	median	lower
Prior.1	23.567	8.5223	0.75	0.73442	0.73934	0.57161
Prior.2	11.035	994.47	0.01	0.010975	0.010652	0.0055035
Prior.3	20.638	1	1	0.95379	0.96697	0.83632

	upper	variance
Prior.1	0.8696	0.0058946
Prior.2	0.018276	1.0784e-05
Prior.3	0.99877	0.0019471

```
panderOptions('digits', 4)
```

Piror One is

$Beta(23.567, 8.5223)$

Piror two is

$Beta(11.035, 994.47)$

Piror three is

$Beta(20.638, 1)$

Problem 2

##(1)

$$y_i \sim \text{Bin}(n_i, \theta_i) \quad \text{ind}$$
$$\theta_i \sim \text{Beta}(\alpha, \beta) \quad \text{i.i.d}$$
$$\alpha, \beta \sim \mu, \eta$$
$$\mu, \eta \sim \text{LN}(m, \frac{1}{c}), \text{Beta}(a, b)$$

##(2)

```
ESP.data=read.csv("./GanzStudiesUsed-56.csv", header=T)
head(ESP.data)
```

```
##   i..n hits
## 1   32   14
## 2    7    6
## 3   30   13
## 4   30    7
## 5   20    2
## 6   10    9
```

```
jags_model ="model{
for( i in 1 : N ) {
Y[i] ~ dbin(theta[i], n[i])
theta[i] ~ dbeta(alpha, beta)
}
alpha = eta * mu
beta = eta * (1-mu)
eta ~ dlnorm(m, 1/C)
mu ~ dbeta(a, b)
}"
jags.data = list(Y = ESP.data$hits, n = ESP.data$i..n, N = dim(ESP.data)[1], a=20,b=40,m=0,C=3)
jags.param <- c("theta", "alpha", "beta","eta","mu")

jagsfit <- jags(data=jags.data, n.chains=5, inits=NULL,parameters.to.save =jags.param, n.iter=2000, n.burnin=1000)

## module glm loaded

## Compiling model graph
##   Resolving undeclared variables
##   Allocating nodes
## Graph information:
##   Observed stochastic nodes: 56
##   Unobserved stochastic nodes: 58
```

```

## Total graph size: 180
##
## Initializing model
##(3)
op.prior <- epi.betabuster(mode = 0.25, conf = 0.95, greaterthan = F, x = 0.3)
pb.prior <- epi.betabuster(mode = 0.33, conf = 0.95, greaterthan = F, x = 0.36)
ps.prior <- epi.betabuster(mode = 0.25, conf = 0.95, greaterthan = F, x = 0.255)

capture.output(as.data.frame(op.prior, row.names = 'prior Open minded', optional = TRUE))

## [1] " shape1 shape2 mode mean median lower upper variance"
## [2] "prior Open minded 58.825 174.48 0.25 0.25214 0.25143 0.19863 0.30968 0.00080481"
capture.output(as.data.frame(pb.prior, row.names = 'Prior Psi Beliver', optional = T))

## [1] " shape1 shape2 mode mean median lower upper variance"
## [2] "Prior Psi Beliver 100 202 0.33 0.33113 0.33075 0.27924 0.38513 0.00073096"
capture.output(as.data.frame(ps.prior, row.names = 'Prior Psi skeptics', optional = T))

## [1] " shape1 shape2 mode mean median lower upper variance"
## [2] "Prior Psi skeptics 100 298 0.25 0.25126 0.25084 0.20992 0.29496 0.0004715"
df <- rbind(op.prior, pb.prior, ps.prior)
panderOptions('digits', 4)
pander(df)

```

Table 3: Table continues below

	shape1	shape2	mode	mean	median	lower	upper
op.prior	58.825	174.48	0.25	0.25214	0.25143	0.19863	0.30968
pb.prior	100	202	0.33	0.33113	0.33075	0.27924	0.38513
ps.prior	100	298	0.25	0.25126	0.25084	0.20992	0.29496

	variance
op.prior	0.00080481
pb.prior	0.00073096
ps.prior	0.0004715

Priors are the Beta distributions with above parameters.

```

##(4)
jags.param <- c("alpha", "beta", "mu")
jags.data = list(Y = ESP.data$hits, n = ESP.data$i..n, N = dim(ESP.data)[1], a = op.prior$shape1, b=op.prior$shape2)
jagsfit.op <- jags(data=jags.data, n.chains=5, inits=NULL, parameters.to.save =jags.param, n.iter=3000, n.burnin=1000)

## Compiling model graph
## Resolving undeclared variables
## Allocating nodes
## Graph information:

```

```
## Observed stochastic nodes: 56
## Unobserved stochastic nodes: 58
## Total graph size: 180
##
## Initializing model
```

```
jagsfit.op$BUGSoutput$summary
```

```
##           mean      sd      2.5%      25%      50%      75%      97.5%
## alpha      9.04655  3.644474  4.19639   6.54019   8.32165  10.61217  18.46498
## beta      19.33049  7.645174  9.07227  14.11267  17.78209  22.90983  39.03660
## deviance 273.85172 10.998389 253.83401 265.99247 273.39450 281.06173 296.57446
## mu         0.31851  0.014091  0.29061   0.30926   0.31844   0.32795   0.34575
##           Rhat n.eff
## alpha      1.0064   510
## beta      1.0064   510
## deviance 1.0032  1100
## mu         1.0012  4100
```

```
jags.param <- c("alpha", "beta", "mu")
```

```
jags.data = list(Y = ESP.data$hits, n = ESP.data$i..n, N = dim(ESP.data)[1], a = pb.prior$shape1, b=pb.prior$rate1)
```

```
jagsfit.pb <- jags(data=jags.data, n.chains=5, inits=NULL, parameters.to.save =jags.param, n.iter=3000, n.burnin=500)
```

```
## Compiling model graph
## Resolving undeclared variables
## Allocating nodes
## Graph information:
## Observed stochastic nodes: 56
## Unobserved stochastic nodes: 58
## Total graph size: 180
##
## Initializing model
```

```
jagsfit.pb$BUGSoutput$summary
```

```
##           mean      sd      2.5%      25%      50%      75%      97.5%
## alpha      9.35126  3.564911  4.53192   6.93848   8.69858  10.95789  18.12435
## beta      18.54913  7.177079  8.89970  13.66466  17.16668  21.85097  35.67414
## deviance 272.70758 10.694880 253.57577 265.29017 272.26007 279.58315 295.34388
## mu         0.33579  0.014199  0.30869   0.32621   0.33535   0.34495   0.36467
##           Rhat n.eff
## alpha      1.0046   740
## beta      1.0047   720
## deviance 1.0021  1800
## mu         1.0022  1800
```

```
jags.param <- c("alpha", "beta", "mu")
```

```
jags.data = list(Y = ESP.data$hits, n = ESP.data$i..n, N = dim(ESP.data)[1], a = ps.prior$shape1, b=ps.prior$rate1)
```

```
jagsfit.ps <- jags(data=jags.data, n.chains=5, inits=NULL, parameters.to.save =jags.param, n.iter=3000, n.burnin=500)
```

```
## Compiling model graph
## Resolving undeclared variables
## Allocating nodes
## Graph information:
## Observed stochastic nodes: 56
```

```
## Unobserved stochastic nodes: 58
## Total graph size: 180
##
## Initializing model
```

```
jagsfit.ps$BUGSoutput$summary
```

```
##           mean          sd      2.5%      25%      50%      75%      97.5%
## alpha      8.73259    3.696050   4.04038   6.31175   8.01466   10.22420   17.92420
## beta      19.45671    8.044204   9.16774  14.19470  17.87163  22.69193  39.85249
## deviance 274.37528  11.119937 254.11952 266.69560 273.72615 281.41633 298.00059
## mu         0.30921    0.013566   0.28216   0.30004   0.30952   0.31858   0.33516
##           Rhat n.eff
## alpha      1.0018  3000
## beta       1.0017  2700
## deviance   1.0016  2900
## mu         1.0011  5000
```

Posterior of Open-minded: mean is 0.3189 0.2899 0.34821 0.33566 0.30754 0.36332 0.30943 0.28373 0.33596

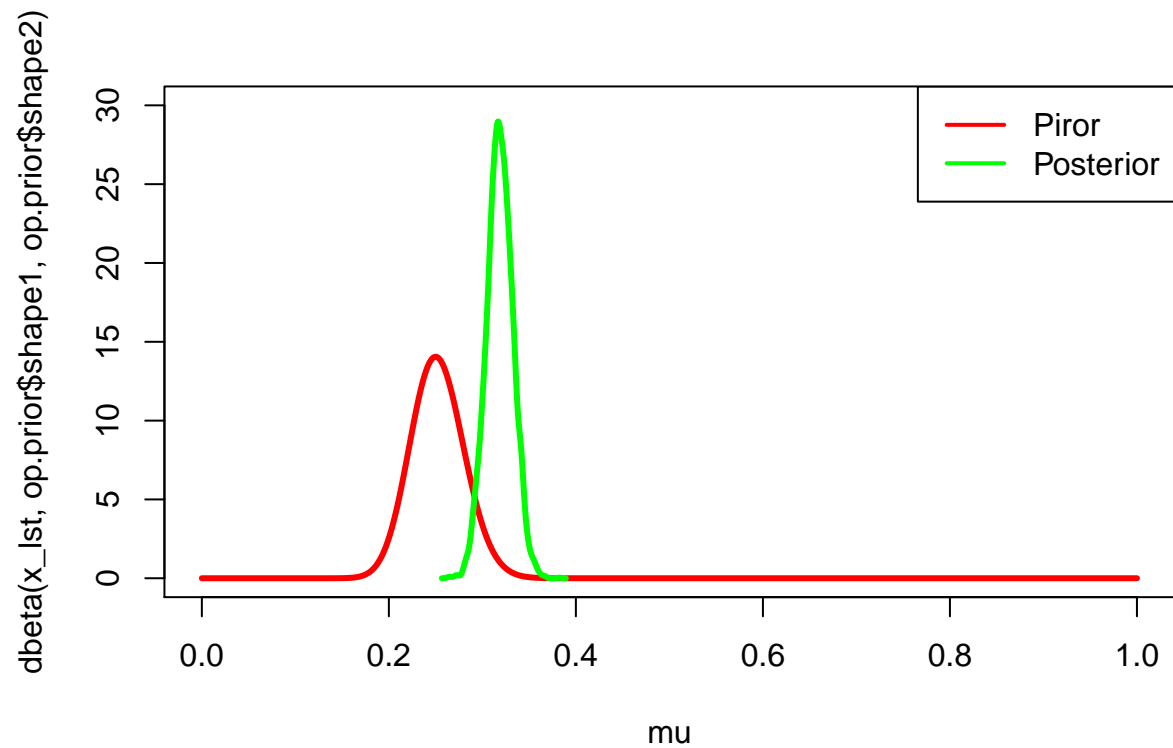
```
qtls = matrix(NA,3,3)
qtls[1,] = c(0.3189,0.2899,0.34821)
qtls[2,] = c(0.33566,0.30754,0.36332)
qtls[3,] = c(0.30943,0.28373,0.33596)
df = data.frame(qtls)
row.names(df) = c('Open-minded posterior','Psi believer posterior','Psi skeptics posterior')
colnames(df)= c('mean','2.5%','97.5%')
knitr::kable(df, format = "markdown")
```

	mean	2.5%	97.5%
Open-minded posterior	0.31890	0.28990	0.34821
Psi believer posterior	0.33566	0.30754	0.36332
Psi skeptics posterior	0.30943	0.28373	0.33596

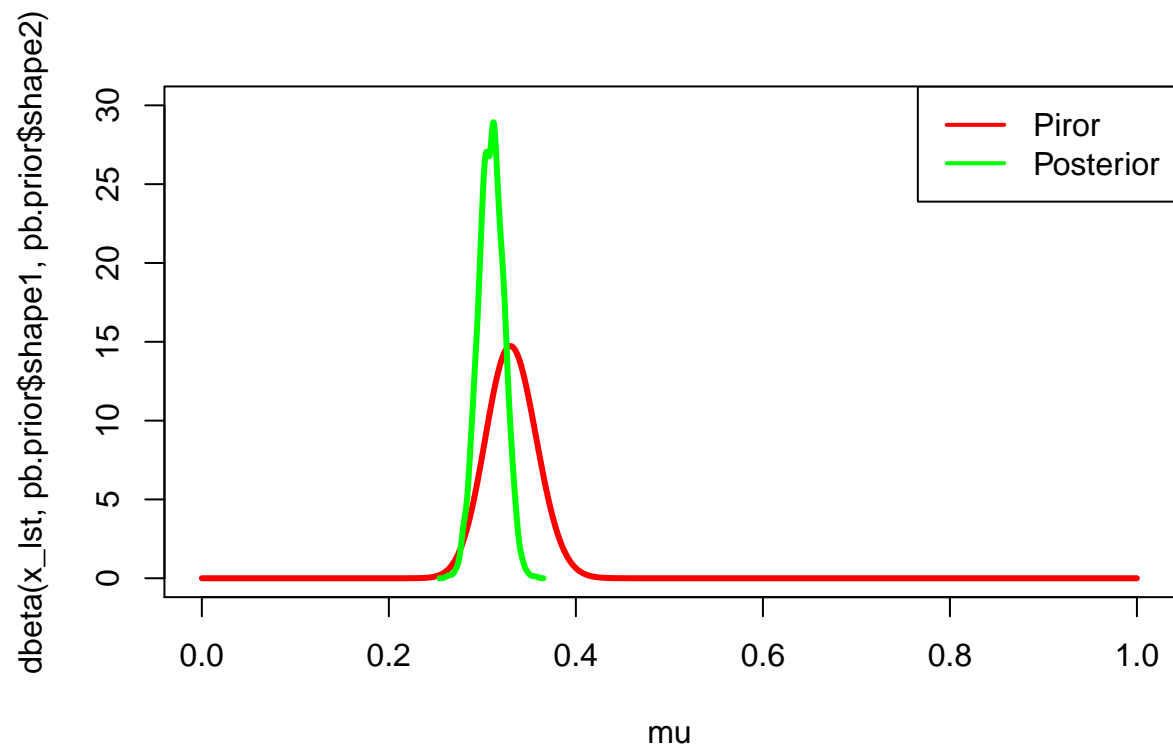
```
##(5)
```

Note that for the Posterior Distribution. The left and right curve should be converges to 0. Since we use MCMC, they don't show in those figures but does not affect the analysis much.

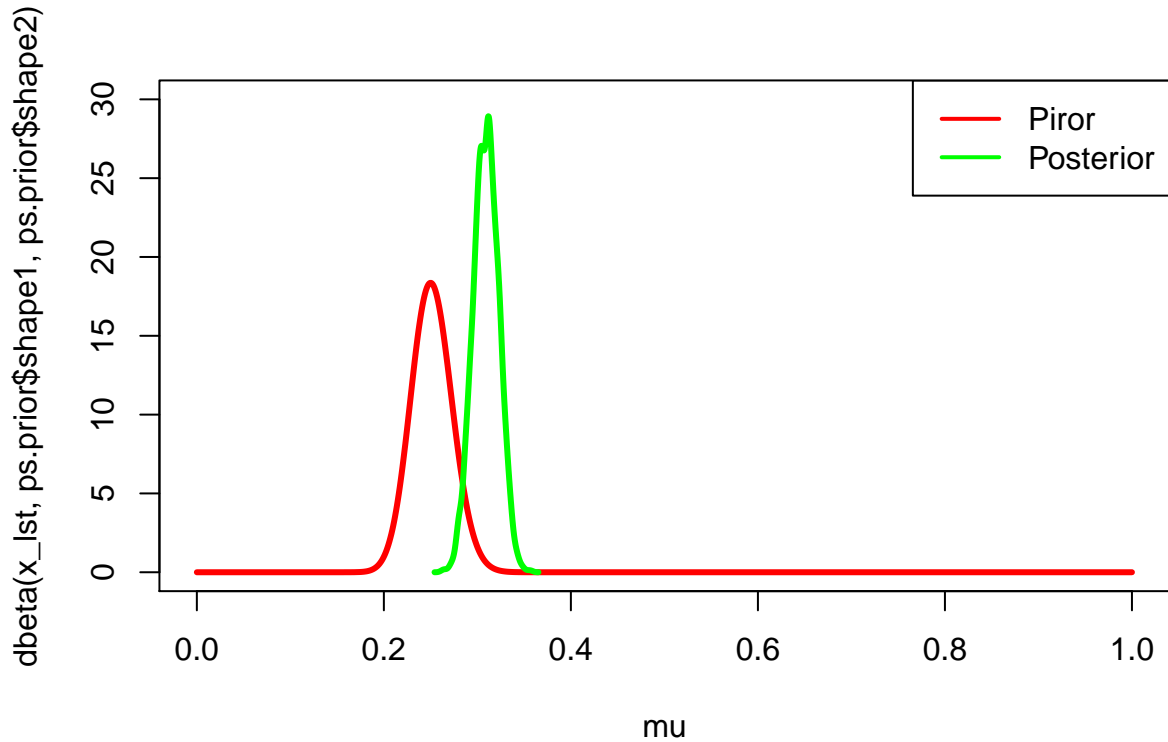
```
x_lst = seq(0,1,by=0.001)
plot(x_lst,dbeta(x_lst,op.prior$shape1,op.prior$shape2),col="red", type="l", xlab = 'mu',lwd=3,ylim = c(
lines(density(jagsfit.op$BUGSoutput$sims.matrix[,4]),col="green", type="l",ylab = 'pdf',lwd=3)
legend('topright',col=c('red','green'),legend=c('Piror','Posterior'),lwd=c(2,2))
```



```
x_lst = seq(0,1,by=0.001)
plot(x_lst,dbeta(x_lst,pb.prior$shape1,pb.prior$shape2),col="red", type="l", xlab = 'mu',lwd=3,ylim = c(0,30))
lines(density(jagsfit.ps$BUGSoutput$sims.matrix[,4]),col="green", type="l",ylab = 'pdf',lwd=3)
legend('topright',col=c('red','green'),legend=c('Piror','Posterior'),lwd=c(2,2))
```



```
x_lst = seq(0,1,by=0.001)
plot(x_lst,dbeta(x_lst,ps.prior$shape1,ps.prior$shape2),col="red", type="l", xlab = 'mu',lwd=3,ylim = c(0,30))
lines(density(jagsfit.ps$BUGSoutput$sims.matrix[,4]),col="green", type="l",ylab = 'pdf',lwd=3,ylim = c(0,30))
legend('topright',col=c('red','green'),legend=c('Piror','Posterior'),lwd=c(2,2))
```



For open-minded it's posterior mode and mean get larger with less std. THis infer that they under estimated the power of Psi.

For those Psi believer, it's posterior mode and mean get smaller with less std. THis infer that they over estimated the power of Psi.

For those Psi Skeptic, it's posterior mode and mean get larger with less std. THis infer that they under estimated the power of Psi.

Problem 4

##(1)

$$z_{s|c} \sim \text{Bin}(N_{s|c}, \theta_{s|c}) \quad \text{i.i.d}$$

$$\theta_{s|c} \sim \text{Beta}(\alpha_c, \beta_c) \quad \text{i.i.d}$$

$$\alpha_c, \beta_c \sim \mu, \eta$$

$$\mu \sim \text{Beta}(a, b)$$

$$\eta \sim \text{LN}(m, C)$$

##(2)

```
data=read.csv("./BattingAverage.csv", header=T)
head(data)
```



```
##      Player      PriPos Hits AtBats PlayerNumber PriPosNumber
## 1 Fernando Abad    Pitcher    1      7          1          1
## 2  Bobby Abreu Left Field   53    219          2          7
## 3   Tony Abreu   2nd Base   18     70          3          4
## 4 Dustin Ackley   2nd Base  137    607          4          4
## 5   Matt Adams   1st Base   21     86          5          3
## 6 Nathan Adcock   Pitcher    0      1          6          1
```

```
jags.data=list(Z = data$Hits, n = data$AtBats, Pos=data$PriPosNumber,N =nrow(data),m=0, C=3,a =23,b=77)
```

```
jags_model = "model{
  for (i in 1 : N)
  {
    Z[i] ~ dbin(theta[i], n[i])
    theta[i] ~ dbeta(alpha[Pos[i]], beta[Pos[i]])
  }
  for(j in 1:9)
  {
    alpha[j] =eta[j]*mu[j]
    beta[j] = eta[j]*(1-mu[j])
    eta[j] ~ dlnorm(m, 1/C)
    mu[j] ~ dbeta(a, b)
  }
  compare1 = theta[573]-theta[143]
  compare2 = theta[142]-theta[921]
  compare = mu[8]-mu[2]
}"
```

```
jags.param <- c("theta", "mu","compare1","compare2","compare")
```

```
jags_fit <- jags(data = jags.data, n.chains = 5, inits = NULL, parameters.to.save = jags.param,n.iter=30000)
```

```
## Compiling model graph
##   Resolving undeclared variables
##   Allocating nodes
## Graph information:
##   Observed stochastic nodes: 948
##   Unobserved stochastic nodes: 966
##   Total graph size: 3847
##
## Initializing model
```

```
# jags_fit$BUGSoutput$summary
jags.mcmc = as.mcmc(jags_fit)
# mcmc_trace(jags.mcmc, pars = c("mu[2]"))
```

```
##(3)
```

```
Number_lst = unique(data$PriPosNumber)
Name_lst = unique(data$PriPos)
capture.output(as.data.frame(Number_lst))
```

```
## [1] " Number_lst" "1          1" "2          7" "3          4" "4          3"
## [6] "5          5" "6          2" "7          6" "8          8" "9          9"
```

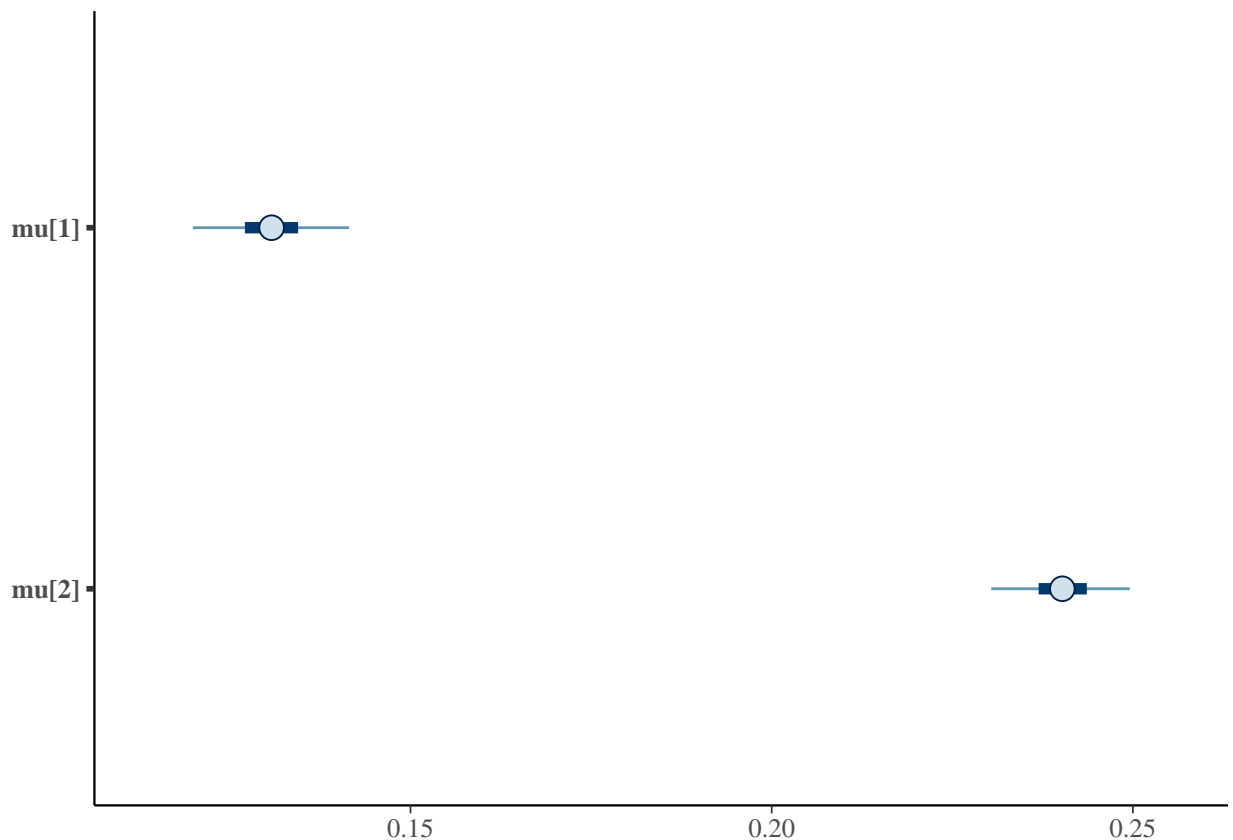
```
capture.output(as.data.frame(Name_lst))
```

```
## [1] " Name_lst" "1 Pitcher" "2 Left Field" "3 2nd Base"
```

```
## [5] "4      1st Base" "5      3rd Base" "6      Catcher" "7      Shortstop"
## [9] "8 Center Field" "9 Right Field"

# df <- rbind(Number_lst, Name_lst)
# pander(df)

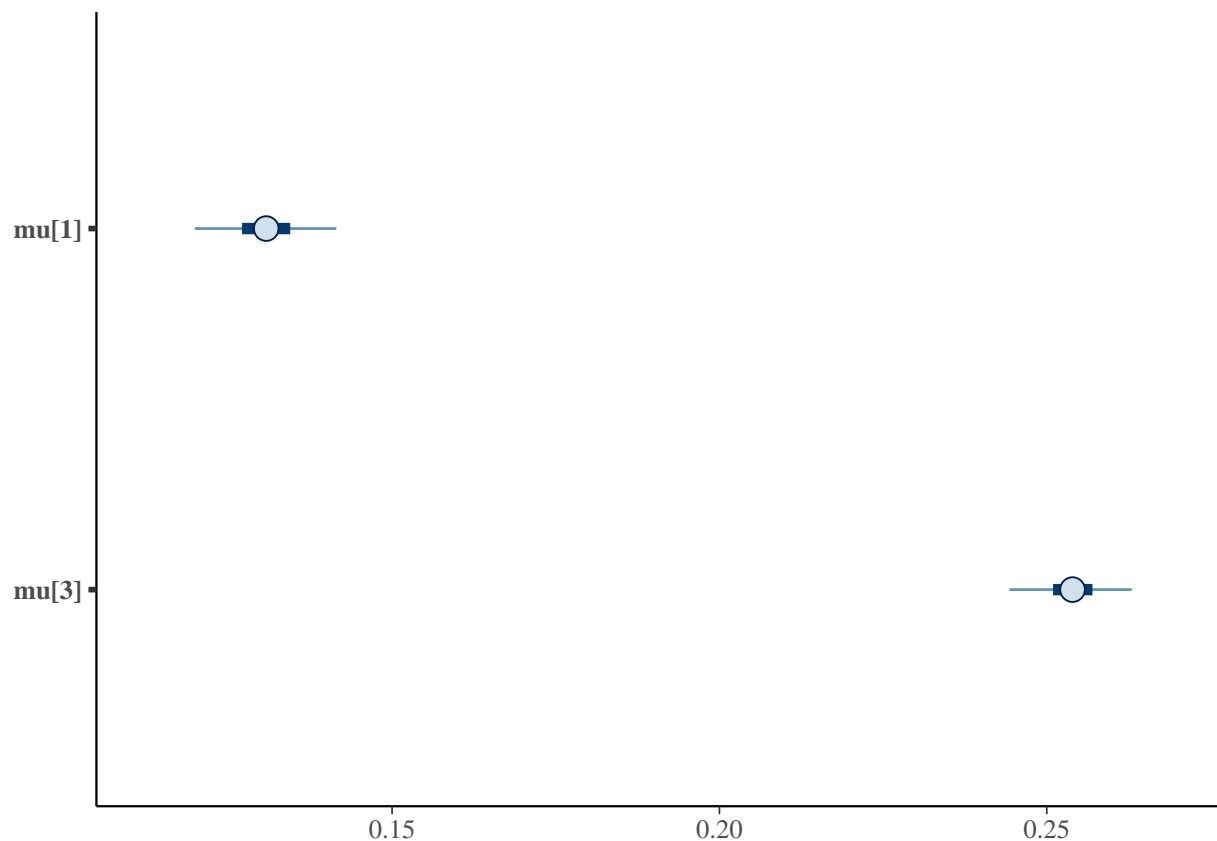
mcmc_intervals(jags.mcmc, pars=c("mu[1]", "mu[2]"), prob = 0.5, # 80% intervals - inner
prob_outer = 0.95, # 95% - outer
point_est = "mean")
```



By the table, $\mu[1]$ is pitcher and $\mu[2]$ is catcher. From above, we can infer that catcher (mean of at around 0.24) is better than pitcher (mean at bat around 0.13) at bat with credible interval 95%.

```
##(4)
```

```
mcmc_intervals(jags.mcmc, pars=c("mu[1]", "mu[3]"), prob = 0.5, # 80% intervals - inner
prob_outer = 0.95, # 95% - outer
point_est = "mean")
```



By the table, $\mu[1]$ is pitcher and $\mu[3]$ is First Base. From above, we can infer that First base (mean of at around 0.26) is better than pitcher (mean at bat around 0.13) at bat with credible interval 95%.

```
##(5)
```

```
which("Wellington Castillo" == data$Player)
```

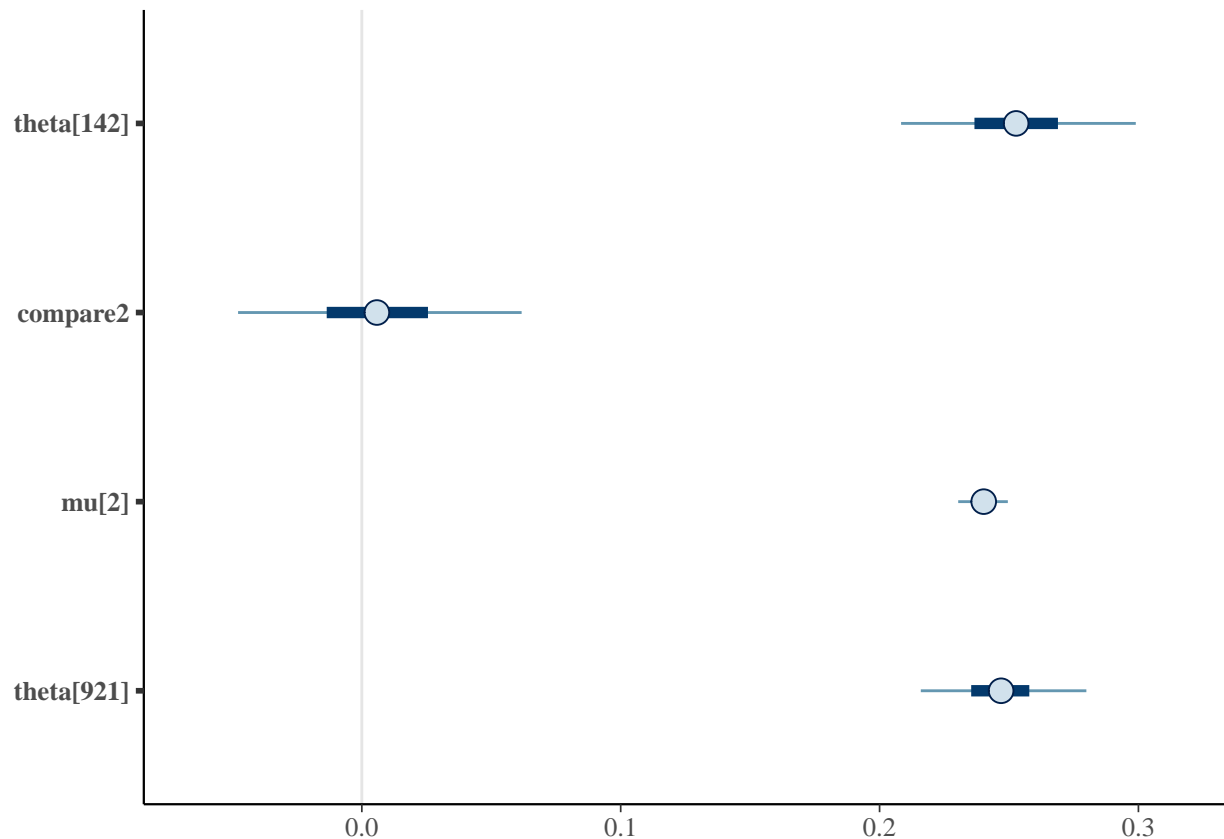
```
## [1] 142
```

```
which("Matt Wieters" == data$Player)
```

```
## [1] 921
```

```
142 catcher mu2 921 catcher mu2 573 central mu8 143 catcher mu2
```

```
mcmc_intervals(jags.mcmc, pars=c("theta[142]", "compare2", "mu[2]", "theta[921]"), prob = 0.5, # 80% intervals
prob_outer = 0.95, # 95% - outer
point_est = "mean")
```



```
p921 = 2/8
p142 = 45/170
print(p921)
```

```
## [1] 0.25
```

```
print(p142)
```

```
## [1] 0.26471
```

By the table, theta[142] is W.C. and M.W. From individuals above, we can infer that W.C. is slightly better than M.W. for at bat by comparing mean. however, M.W. performance is more stable than W.C. since M.W. has smaller credible interval than W.C. By individual shrinkage, we see that compare between two players has mean 5.4755e-03 with 95 credible interval (-5.0248e-02 6.1392e-02) which means that two players differs only 5 at bats for 100 games. Thus, we can conclude that two players has similar abilities at bat. there is an evidence of shrinkage in the estimates. Since both of the mean of the at bat for players comes close to mean from their all probability of at bat.

```
##(6)
```

```
which("Andrew McCutchen" == data$Player)
```

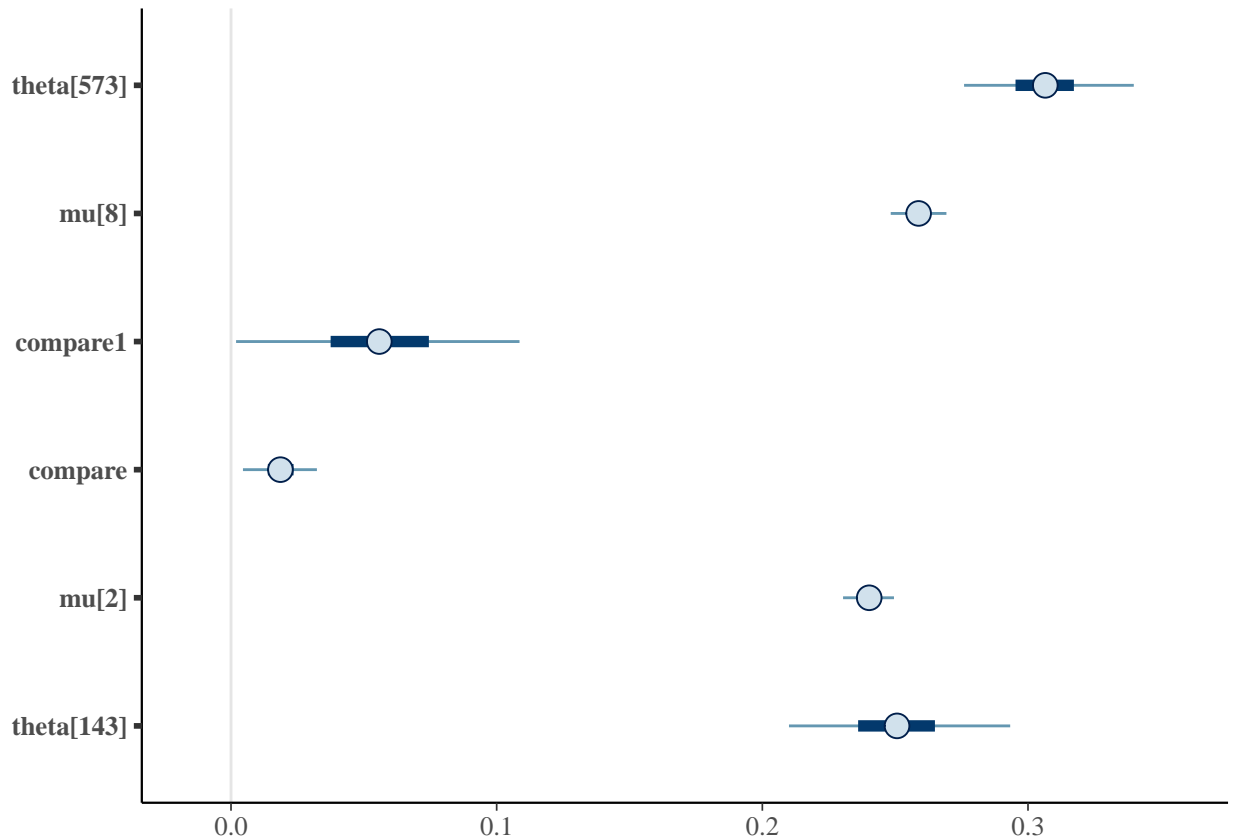
```
## [1] 573
```

```
which("Jason Castro" == data$Player)
```

```
## [1] 143
```

```
mcmc_intervals(jags.mcmc, pars=c("theta[573]", "mu[8]", "compare1", "compare", "mu[2]", "theta[143]"), prob
prob_outer = 0.95, # 95% - outer
```

```
point_est = "mean")
```



```
print(66/257)
```

```
## [1] 0.25681
```

```
print(194/593)
```

```
## [1] 0.32715
```

By the table, theta[573] is J.C. and A.M. From above, we can infer that J.C. is better than A.M. for at bat by comparing mean around 0.31 to 0.25. And J.C. performed more stable than A.M. since we can see a shrinkage from distribution of J.C. than A.M., which means J.C. can make 0 to 10 more balls average 5 balls than A.M. Thus, we conclude that J.C. is better. Thus, there is evidence of shrinkage in the estimates. Since both of the mean of the at bat for players comes close to mean from their all probability of at bat.