Discrete Probability for Discrete Quantum Computing

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1 Classical Probability Spaces

We review the conventional presentation of probability spaces and then discuss several variations that avoid using the real interval $[0, 1]$.

1.1 Real-Valued Probability Spaces

A *probability space* [1, 2, 3] specifies the necessary conditions for reasoning coherently about collections of uncertain events. It consists of a *sample space* Ω, a space of *events* $\mathcal{E}$, and a *probability measure* μ. In this paper, we will only consider *finite* sets of events: we therefore define a sample space Ω as an arbitrary non-empty finite set and the space of events $\mathcal{E}$ as 2^Ω, the powerset of Ω. Given the set of events $\mathcal{E}$, a *probability measure* is a function $\mu : \mathcal{E} \to [0, 1]$ such that:

- $\mu(\Omega) = 1$, and
- for a collection E_i of pairwise disjoint events, $\mu(\bigcup_i E_i) = \sum_i \mu(E_i)$.

Example 1 (Two-coin probability space)**.** Consider an experiment that tosses two coins. We have four possible outcomes that constitute the sample space $\Omega = \{HH, HT, TH, TT\}$. There are 16 total events including for example the event $\{HH, HT\}$ that the first coin lands heads up, the event $\{HT, TH\}$ that the two coins land on opposite sides, and the event $\{HT, TH, TT\}$ that at least one coin lands tails up. Here is a possible probability measure for these events:

$$\begin{aligned}
\mu(\emptyset) &= 0 & \mu(\{HT, TH\}) &= 2/3 \\
\mu(\{HH\}) &= 1/3 & \mu(\{HT, TT\}) &= 0 \\
\mu(\{HT\}) &= 0 & \mu(\{TH, TT\}) &= 2/3 \\
\mu(\{TH\}) &= 2/3 & \mu(\{HH, HT, TH\}) &= 1 \\
\mu(\{TT\}) &= 0 & \mu(\{HH, HT, TT\}) &= 1/3 \\
\mu(\{HH, HT\}) &= 1/3 & \mu(\{HH, TH, TT\}) &= 1 \\
\mu(\{HH, TH\}) &= 1 & \mu(\{HT, TH, TT\}) &= 2/3 \\
\mu(\{HH, TT\}) &= 1/3 & \mu(\{HH, HT, TH, TT\}) &= 1
\end{aligned}$$

The assignment satisfies the two constraints for probability measures: the probability of the entire sample space is 1, and the probability of every collection of disjoint events (e.g., $\{HT\} \cup \{TH\} = \{HT, TH\}$) is the sum of the individual probabilities. The probability of collections of non-disjoint events (e.g., $\{HT, TH\} \cup \{TH, TT\} = \{HT, TH, TT\}$) may add to something different than the probabilities of the individual events. It is useful to think that this probability measure is completely induced by the two coins in question and their characteristics in the sense that each pair of coins induces a measure, and each measure must correspond to some pair of coins. The measure above is induced by two coins such that the first coin is twice as likely to land tails up than heads up and the second coin is double-headed. □

In a strict computational or experimental setting, one may question the reliance of the definition of probability space on the uncountable and uncomputable real interval $[0,1]$. This interval includes numbers like $0.h_1h_2h_3\ldots$ where h_i is 1 or 0 depending on whether Turing machine M_i halts or not. Such numbers cannot be computed. This interval also includes numbers like $\frac{\pi}{4}$ which can only be computed with increasingly large resources as the precision increases. Therefore, in a resource-aware setting, it is more appropriate to consider probability measures that map events to a finite set of elements computable with a fixed set of resources. We will consider two approaches: set-valued probability measures [4, 5] and interval-valued probability measures [6, 7, 8, 9].

1.2 Set-valued Probability Measures

Instead of using every point in the real interval $[0,1]$ we can partition the interval into disjoint sets and only consider probability measures up to set membership. The simplest such situation is to partition the interval $[0,1]$ into $\{0\}$ (which we will call *impossible*) and the half-open interval $(0,1]$ (which we will call *possible*). The addition that was used to aggregate probabilities is now abstracted to $\vee$ such that $x \vee y = \textit{impossible}$ if and only if $x = y = \textit{impossible}$. We will call the resulting set $\{\textit{impossible}, \textit{possible}\}$ together with the with associated operation $\vee$, the set $\mathscr{L}_2$. The definition of a probability measure in this case is modified to a function $\mu : \mathcal{E} \to \mathscr{L}_2$ such that:

- $\mu(\Omega) = \textit{possible}$, and
- for a collection E_i of pairwise disjoint events, $\mu(\bigcup_i E_i) = \bigvee_i \mu(E_i)$.

Example 2 (Two-coin probability space with finite set-valued probability measure)**.** Under the new set-valued requirement, the probability measure in the first example becomes:

$$\begin{aligned}
\mu(\emptyset) &= \textit{impossible} & \mu(\{HT,TH\}) &= \textit{possible}\\
\mu(\{HH\}) &= \textit{possible} & \mu(\{HT,TT\}) &= \textit{impossible}\\
\mu(\{HT\}) &= \textit{impossible} & \mu(\{TH,TT\}) &= \textit{possible}\\
\mu(\{TH\}) &= \textit{possible} & \mu(\{HH,HT,TH\}) &= \textit{possible}\\
\mu(\{TT\}) &= \textit{impossible} & \mu(\{HH,HT,TT\}) &= \textit{possible}\\
\mu(\{HH,HT\}) &= \textit{possible} & \mu(\{HH,TH,TT\}) &= \textit{possible}\\
\mu(\{HH,TH\}) &= \textit{possible} & \mu(\{HT,TH,TT\}) &= \textit{possible}\\
\mu(\{HH,TT\}) &= \textit{possible} & \mu(\{HH,HT,TH,TT\}) &= \textit{possible}
\end{aligned}$$

Despite the fact that we have lost all numeric information, the probability measure still reveals that the second coin is double-headed. We have however lost the information regarding the bias in the first coin. This information can be recovered with a more refined probability measure as we show next. □

1.3 Interval-valued probability measures

A natural generalization of the disjoint set-valued measure above is to allow the sets to overlap. In this case, we split the interval $[0,1]$ in a collection of *overlapping* closed sub-intervals. First we illustrate the main ideas using a simple example.

Example 3 (Two-coin probability space with four intervals)**.** We split the unit interval $[0,1]$ in the following four closed sub-intervals: $[0,0]$ which we call *impossible*, $[0,\frac{1}{2}]$ which we call *unlikely*, $[\frac{1}{2},1]$ which we call *likely*, and $[1,1]$ which we call *necessary*. Using these new values, we can modify the probability measure of Ex. 1 by mapping each numeric value to the smallest sub-interval containing

it to get the following:

$$\begin{array}{rcl@{\qquad}rcl}
\mu(\emptyset) & = & \textit{impossible} & \mu(\{HT,TH\}) & = & \textit{likely} \\
\mu(\{HH\}) & = & \textit{unlikely} & \mu(\{HT,TT\}) & = & \textit{impossible} \\
\mu(\{HT\}) & = & \textit{impossible} & \mu(\{TH,TT\}) & = & \textit{likely} \\
\mu(\{TH\}) & = & \textit{likely} & \mu(\{HH,HT,TH\}) & = & \textit{necessary} \\
\mu(\{TT\}) & = & \textit{impossible} & \mu(\{HH,HT,TT\}) & = & \textit{unlikely} \\
\mu(\{HH,HT\}) & = & \textit{unlikely} & \mu(\{HH,TH,TT\}) & = & \textit{necessary} \\
\mu(\{HH,TH\}) & = & \textit{necessary} & \mu(\{HT,TH,TT\}) & = & \textit{likely} \\
\mu(\{HH,TT\}) & = & \textit{unlikely} & \mu(\{HH,HT,TH,TT\}) & = & \textit{necessary}
\end{array}$$

This probability measure is more informative than the one in Ex. 2: not only does it reveal that the second coin is double-headed but it also reveals the bias in the first coin. □

The probability measure above appears quite intuitive but it is not really evident that it is well-defined. For example, how do we justify the following combination of assignments:

$$\mu(\{HH\}) = \textit{unlikely}, \quad \mu(\{TH\}) = \textit{likely}, \quad \mu(\{HH\} \cup \{TH\}) = \textit{necessary}$$

which assert that an *unlikely*-event whose probability is in the range $[0, \frac{1}{2}]$ and a *likely*-event whose probability is in the range $[\frac{1}{2}, 1]$ combine to a *necessary*-event whose probability is in the range $[1, 1]$. To understand this mystery, we look at another more general example and then give the formal mathematical definition of interval-valued probabilities.

Example 4 (Dempster-Shafer Theory of Evidence)**.** We have three employees whose precise ages A_1, A_2, and A_3 are not known. All is given is a range of ages for each employee: $A_1 \in \{23, 24\}$, $A_2 \in \{20, 21, 22\}$, and $A_3 \in \{21, 22\}$. The sample space in this case is $A_1 \times A_2 \times A_3$ which represents all the possible combinations of ages for the three employees:

$$\begin{array}{rcl}
\Omega & = & \{(23,20,21),(23,20,22),(23,21,21),(23,21,22),(23,22,21),(23,22,22), \\
& & (24,20,21),(24,20,22),(24,21,21),(24,21,22),(24,22,21),(24,22,22)\}
\end{array}$$

Subsets of Ω represent events as usual. Consider the event Ω that *some* employee's age is in the range $\{20, 21, 22\}$: since that event covers the entire sample space its probability must be 1. We can however produce a more informative answer by reasoning as follows: it is *impossible* for the first employee's age to be in the range $\{20, 21, 22\}$; it is *certain* that the second employee's age is in that range; and it is *possible* that the third's employee age is in that range. Aggregating the results, we see that it is *necessary* for one out of three employees to have their age in the required range, and it is *possible* for an additional employee to have their age in the required range. We summarize this information by reporting that the probability of this event is $[\frac{1}{3}, \frac{2}{3}]$ where the first number reports the *certainty* of the event and the second reports the *possibility* of the event. We could also have reasoned about the dual event that *no* employee's age is in the range $\{20, 21, 22\}$ to get the probability $[\frac{1}{3}, \frac{2}{3}]$

dually about the evidence *against* the event. The probability in this case

Now consider the event that some employee's age is in the range $\{23, 24\}$. Reasoning as above, the probability of this event is $[\frac{1}{3}, \frac{1}{3}]$ as only the first employee qualifies. Clearly if we were to ask the probability that some employee's age is in the range $\{20, 21, 22, 23, 24\}$ the result should be $[1, 1]$ as this range covers all the possibilities. The way to calculate this result from the two previous ones is as follows: the evidence for the combined event to be necessary is at least as strong as the evidence that each disjoint event that contributes to it is necessary. So the lower bound probability for the combined event is at least $\frac{1}{3}$. Similarly the upper bound probability is at least 1. But now we can reason using the complements of the events about the necessity and possibility of refuting each event. For the event that some employee's age is *not* in the range $\{20, 21, 22\}$ we have a probability $[\frac{1}{3}, \frac{1}{3}]$ because it is necessary that A_1 is not in that range. Similary, the event that some employee's age is *not* in the range $\{23, 24\}$ is $[\frac{2}{3}, \frac{2}{3}]$. Thus although the positive evidence for necessity

□

Example 5 (Dempster-Shafer Theory of Evidence)**.** We have five employees whose precise age is not known. All is given is a range for each employee:

$$\begin{array}{lcl} \text{Age}(M_1) & \in & \{23,24\} = D_1 \\ \text{Age}(M_2) & \in & \{20,21,22\} = D_2 \\ \text{Age}(M_3) & \in & \{20,21\} = D_3 \end{array}$$

What is the probability that an employee's age is in the range $\{22,23,24\}$? Looking at the data, it is *possible* for M_2's age to be within the range, it is *not possible* for M_3's age to be within the range, and it is *certain* or *necessary* that M_1's age is in the range. Clearly saying that an event is *necessary* is equivalent to saying that its complement is *not possible*. Aggregating the results for the three employees, we can calculate that is necessary that 1 employees have their ages within the range, and it is possible that 2 employees have their ages within the range. We express this formally as saying that the probability of the event is $\left[\frac{1}{3}, \frac{2}{3}\right]$. Now consider another event asking whether the ages are some other disjoint range $\{20,21\}$. Reasoning in a similar way we calculate that the probability for this event is $\left[\frac{1}{3}, \frac{2}{3}\right]$. Now let's take the two events together and ask about the possibility of the ages to be in the range $\{20,21,22,23,24\}$. Clearly that probability must be $[1,1]$ as every employee's age is in that range. This problem looks puzzle if we want to attack it directly, but will be more clear if we think in terms of conditional probability. We will compute the conditional probability for the usual real-valued probability first and for the interval-valued probability later.

- We randomly draw an employee among the three, so each employee has probability $\frac{1}{3}$ to be drawn, i.e.,

  $$\mu(\{i=1\}) = \mu(\{i=2\}) = \mu(\{i=3\}) = \frac{1}{3}$$

  In order to compute the usual real-valued probability, we need a probability distribution within the possible range D_1, D_2, and D_3. Let's assume they are equally probable, i.e.,

  $$\begin{array}{lcl} \mu(\{\text{Age}(M_1)=23\}) & = & \frac{1}{2} \\ \mu(\{\text{Age}(M_1)=24\}) & = & \frac{1}{2} \end{array} \qquad \begin{array}{lcl} \mu(\{\text{Age}(M_2)=20\}) & = & \frac{1}{3} \\ \mu(\{\text{Age}(M_2)=21\}) & = & \frac{1}{3} \\ \mu(\{\text{Age}(M_2)=22\}) & = & \frac{1}{3} \end{array} \qquad \begin{array}{lcl} \mu(\{\text{Age}(M_3)=20\}) & = & \frac{1}{2} \\ \mu(\{\text{Age}(M_3)=21\}) & = & \frac{1}{2} \end{array}$$

  Then, the probability that an employee's age is in the range $\{22,23,24\}$ can be computed as follow:

  $$\begin{array}{rl} & \mu(\{i|\text{Age}(M_i) \in \{22,23,24\}\}) \\ = & \mu(\{i=1\})\,\mu(\{\text{Age}(M_1) \in \{22,23,24\}\}) \\ & +\mu(\{i=2\})\,\mu(\{\text{Age}(M_2) \in \{22,23,24\}\}) \\ & +\mu(\{i=3\})\,\mu(\{\text{Age}(M_3) \in \{22,23,24\}\}) \\ = & \frac{1}{3}\cdot 1 + \frac{1}{3}\cdot\frac{1}{3} + \frac{1}{3}\cdot 0 = \frac{4}{9} \end{array}$$

- Assume we don't know the probability distributions within the possible range D_1, D_2, and D_3. All we know is whether they are *possible* or *necessary*. Then we replace the above computation from exact value to interval and consider

  $$\begin{array}{rl} & \mu(\{i|\text{Age}(M_i) \in \{22,23,24\}\}) \\ = & \mu(\{i=1\})\,\mu(\{\text{Age}(M_1) \in \{22,23,24\}\}) \\ & +\mu(\{i=2\})\,\mu(\{\text{Age}(M_2) \in \{22,23,24\}\}) \\ & +\mu(\{i=3\})\,\mu(\{\text{Age}(M_3) \in \{22,23,24\}\}) \\ = & \frac{1}{3}\cdot[1,1] + \frac{1}{3}\cdot[0,1] + \frac{1}{3}\cdot[0,0] = \left[\frac{1}{3}, \frac{2}{3}\right] \end{array}$$

  Yu-Tsung says: Check the conditional probability rule in the Dempster-Shafer Theory!

□

In interval-valued probability measures, if the probability of an event E is $[a, b]$, we think of the left-endpoint a as representing the strength of the evidence that supports E, and the right-endpoint b as the strength of the evidence that contradicts E.

Thus if we have an event E with probability $[a, b]$ where $a = 0.1$ and $b = 0.7$, we have that:

- the strength of evidence supporting E is 0.1; since either E or its complement must happen, we conclude that there is 0.9 evidence supporting the complement of E;
- the strength of evidence contradicting E is 0.7; again since either E or its complement must happen, we conclude that there is 0.3 evidence contradicting the complement of E.

Yu-Tsung says: Do we use the law of excluded middle here? You remind me Agda :) Did Homotopy Type Theory people said anything about the probability?

Turning things around, the strength of evidence that contradicts E is evidence supporting the complement of E. The complement of E must therefore have probability $[1 - b, 1 - a]$ which we abbreviate $1 - [a, b]$:

$$\begin{aligned} \mu(\emptyset) &= \textit{impossible} \\ \mu(\Omega) &= \textit{necessary} \\ \mu(\Omega \backslash E) &= 1 - \mu(E) \end{aligned}$$

Next, if we define $\sum_i [a_i, b_i] = [\sum_i a_i, \sum_i b_i]$, then

- for a collection E_i of pairwise disjoint events, we have $\mu(\bigcup_i E_i) \subseteq \sum_i \mu(E_i)$.

Notice that the equality may not hold in general. This statement says that the evidences of $\bigcup_i E_i$ is at least as strong as putting all the evidences of E_i together, but some evidence may only be acquired for $\bigcup_i E_i$ as the whole. Therefore, $\mu(\bigcup_i E_i)$ is a subset of $\sum_i \mu(E_i)$, but may not equal. In our example,

$$\begin{aligned} & \mu(\{HH, TH\}) = \textit{necessary} = [1, 1] \\ \subseteq \quad & \left[\frac{1}{2}, \frac{3}{2}\right] = \left[0, \frac{1}{2}\right] + \left[\frac{1}{2}, 1\right] = \textit{unlikely} + \textit{likely} = \mu(\{HH\}) + \mu(\{TH\}) \ . \end{aligned}$$

However, $\mu(\{HH, TH\})$ is a proper subset of $\mu(\{HH\}) + \mu(\{TH\})$ because if we check the complement of $\{HH, TH\}$, we have

$$\mu(\{HH, TH\}) = 1 - \mu(\{HT, TT\}) = 1 - \textit{impossible} = \textit{necessary} \ ,$$

and $\mu(\{HT, TT\}) = \textit{impossible}$ cannot be used to reasoning the probability of $\{HH\}$ and $\{TH\}$ individually.

2 Quantum Probability Spaces

The mathematical framework above assumes that one has complete knowledge of the events and their relationships. However, in many practical situations, the structure of the event space is only partially known and the precise dependence of two events on each other cannot, a priori, be determined with certainty. In the quantum case, this partial knowledge is compounded by the fact that there exist non-commuting events which cannot happen simultaneously. To accommodate these more complex situations, we abandon the sample space Ω and reason directly about events. A quantum probability space therefore consists of just two components: a set of events $\mathcal{E}$ and a probability measure $\mu : \mathcal{E} \to [0, 1]$. We give an example before giving the formal definition.

Example 6 (One-qubit quantum probability space)**.** Consider a one-qubit Hilbert space with states $\alpha|0\rangle+\beta|1\rangle$ such that $|\alpha|^2+|\beta|^2=1$. The set of events associated with this Hilbert space consists of all projection operators. Each event is interpreted as a possible post-measurement state of a quantum system in current state $|\phi\rangle$. For example, the event $|0\rangle\langle 0|$ indicates that the post-measurement state will be $|0\rangle$; the event $|1\rangle\langle 1|$ indicates that the post-measurement state will be $|1\rangle$; the event $|+\rangle\langle +|$ where $|+\rangle = \frac{1}{\sqrt{2}}(|0\rangle + |1\rangle)$ indicates that the post-measurement state will be $|+\rangle$; the event $\mathbb{1} = |0\rangle\langle 0| + |1\rangle\langle 1|$ indicates that the post-measurement state will be a linear combination of $|0\rangle$ and $|1\rangle$; and the empty event $\mathbb{0}$ states that the post-measurement state will be the empty state. As in the classical case, a probability measure is a function that maps events to $[0,1]$: here is a partial specification of a possible probability measure:

$$\mu(\mathbb{0}) = 0, \quad \mu(\mathbb{1}) = 1, \quad \mu(|0\rangle\langle 0|) = 1, \quad \mu(|1\rangle\langle 1|) = 0, \quad \mu(|+\rangle\langle +|) = 1/2, \quad \ldots$$

Note that, similarly to the classical case, the probability of $\mathbb{1}$ is 1 and the probability of collections of orthogonal events (e.g., $|0\rangle\langle 0| + |1\rangle\langle 1|$) is the sum of the individual probabilities. In contrast, a collection of non-orthogonal events (e.g., $|0\rangle\langle 0|$ and $|+\rangle\langle +|$) is not itself an event. In the classical example, we argued that each probability measure is uniquely determined by two actual coins. A similar (but much more subtle) argument is valid also in the quantum case. By postulates of quantum mechanics and Gleason's theorem, it turns out that for large enough quantum systems, each probability measure is uniquely determined by an actual quantum state. □

To properly explain the previous example and generalize to arbitrary quantum systems, we formally discuss projection operators and then define a quantum probability space.

Definition 1 (Projection Operators; Orthogonality; Commutativity [10, 11, 12, 13])**.** Given a Hilbert space $\mathcal{H}$, a projection operator P is a linear transformation from $\mathcal{H}$ to itself such that $P^2 = P = P^\dagger$. Projection operators have the following properties:[1]

- Projection operators P_1 and P_2 are orthogonal if $P_1P_2 = P_2P_1 = \mathbb{0}$;
- Projection operators P_1 and P_2 commute if $P_1P_2 = P_2P_1$;
- If the projections P_1 and P_2 are orthogonal then $P_1 + P_2$ is also a projection;
- If the projections P_1 and P_2 commute then P_1P_2 is also a projection.

Amr says: Here it would be good to refer to the notion of "quantum test" and define events as sums of quantum tests. This will automatically include everything except the products of commutative projections which we will have to explain that they can be expressed as sums of orthogonal projections.

Definition 2 (Quantum Probability Space [15, 16, 11, 17, 14])**.** Given a Hilbert space $\mathcal{H}$, a *quantum probability space* consists of a set of events $\mathcal{E}$ and a probability measure $\mu : \mathcal{E} \to [0,1]$ such that:[2]

- The set of events consists of all projections. This set includes the empty projection, projection operators $|\psi\rangle\langle\psi|$ for each state $|\psi\rangle$, sums of *orthogonal* projections, and products of *commuting* projections;
- $\mu(\mathbb{1}) = 1$, and
- for mutually orthogonal projections E_i, we have $\mu\left(\sum_i E_i\right) = \sum_i \mu(E_i)$.

□

[1] "Projection" is sometimes called "orthogonal projection" or "self-adjoint projection" to emphasize $P^\dagger = P$ [14].

[2] It is possible to define a more general space of events consisting of all operators $\mathcal{A}$ on $\mathcal{H}$ and consider $\mu : \mathcal{A} \to \mathbb{C}$ [14, 13]. When an operator $A \in \mathcal{A}$ is Hermitian, $\mu(A)$ is the expectation value of A. We does not take this approach because we want to focus only on probability.

2.1 Quantum Probability Measures

For a given set of events $\mathcal{E}$, there are many possible probability measures $\mu : \mathcal{E} \to [0, 1]$. The Born rule, a postulate of quantum mechanics, states that each quantum state $|\phi\rangle$ induces a probability measure μ_ϕ as follows:

$$\mu_\phi(E) = \langle\phi|E\phi\rangle$$

Conversely, Gleason's theorem states that given a probability measure μ, there exist a quantum state $|\phi\rangle$ that induces such a measure using the Born rule. The theorem is only valid in Hilbert spaces with dimension $d \geq 3$. It is instructive to study counterexamples in $d = 2$, i.e., the case of a one-qubit system. Consider five states $|\psi_0\rangle$ to $|\psi_4\rangle$ that form five orthogonal bases $\{|\psi_0\rangle, |\psi_1\rangle\}$, $\{|\psi_1\rangle, |\psi_2\rangle\}$, $\{|\psi_2\rangle, |\psi_3\rangle\}$, $\{|\psi_3\rangle, |\psi_4\rangle\}$, and $\{|\psi_4\rangle, |\psi_0\rangle\}$ and consider the probability measure defined as follows. For all $i \in \{0, 1, 2, 3, 4\}$, we have $\mu_X(|\psi_i\rangle\langle\psi_i|) = 1/2$. For each orthogonal basis, the probability is 1 as desired and yet it is impossible to find a single quantum state that realizes such a probability measure (see `http://tph.tuwien.ac.at/~svozil/publ/2006-gleason.pdf`)

Amr says: the rest needs cleaning up and perhaps does not even belong in this section

Although it seems that we need an infinite long table to specify the quantum probability measure μ, our μ is actually given by a simple formula $\langle 0|E|0\rangle$. In general, Born discovered each quantum state $|\psi\rangle \in \mathcal{H}\backslash\{0\}$ induces a probability measure $\tilde{\mu}_\psi : \mathcal{E} \to [0,1]$ on the space of events defined for any event $E \in \mathcal{E}$ as follows [18, 19]:

$$\tilde{\mu}_\psi(E) = \frac{\langle\psi|E|\psi\rangle}{\langle\psi|\psi\rangle} \tag{1}$$

The Born rule satisfies the following properties:

- It can be extend to mixed states. Given a mixed state represented by a density matrix $\rho = \sum_{j=1}^N q_j \frac{|\psi_j\rangle\langle\psi_j|}{\langle\psi_j|\psi_j\rangle}$, where $\sum_{j=1}^N q_j = 1$, i.e., $\mathrm{Tr}(\rho) = 1$, then the Born rule can be extended to ρ by

$$\tilde{\mu}_\rho(E) \quad = \quad \mathrm{Tr}(\rho E) = \sum_{j=1}^N q_j \tilde{\mu}_{\Psi_j}(E) \ . \tag{2}$$

  Notice that $\left(\{1,\ldots,N\}, 2^{\{1,\ldots,N\}}, \mu(J) = \sum_{j\in J} q_j\right)$ is a classical probability space. Therefore, when we discretize the Hilbert space later, we may need to discretize this probability space as well.

- $\tilde{\mu}_\rho$ is a probability measure for all mixed state ρ.
- $\langle\psi|\phi\rangle = 0 \Leftrightarrow \tilde{\mu}_\psi(|\phi\rangle\langle\phi|) = 0$.
- $\tilde{\mu}_\psi(E) = \tilde{\mu}_{\mathbf{U}|\psi\rangle}\left(\mathbf{U}E\mathbf{U}^\dagger\right)$,where $\mathbf{U}$ is any unitary map, i.e., $\mathbf{U}^\dagger\mathbf{U} = \mathbb{1}$.

Naturally, we may ask: is every probability measure induced from a state by the Born rule? The answer is yes by Gleason's theorem when the dimension ≥ 3 [16, 12, 11]. Furthermore, a simple corollary of Gleason's theorem can show the Born rule is the unique function satisfying conditions 1. to 3.

Corollary 1. The Born rule is the unique function satisfying conditions 1. to 3.

Proof. Assume there is another function $\tilde{\mu}'$ such that $\tilde{\mu}'_\rho$ is a quantum probability measure for all mixed state ρ. We are going to prove $\tilde{\mu}' = \tilde{\mu}$.

Fix a pure normalized state ϕ, $\tilde{\mu}'_\phi$ is a quantum probability measure by condition 2. By Gleason's theorem, there is a mixed state ρ', such that $\tilde{\mu}'_\phi(E) = \mathrm{Tr}(\rho' E) = \sum_{j=1}^N q_j \tilde{\mu}_{\psi_j}(E)$ for all event E.

Consider the event $E' = \mathbb{1} - |\phi\rangle\langle\phi|$, we have

$$\begin{aligned} 0 \quad &\overset{\text{Condition 3}}{=} \quad \tilde{\mu}_\phi(E') \\ &= \quad \sum_{j=1}^N q_j \tilde{\mu}_{\psi_j}(E') \end{aligned}$$

Because $q_j > 0$, we have $\tilde{\mu}_{\psi_j}(E) = 0$, i.e., ψ_j is orthogonal to a co-dimension-1 subspace E'. However, the only subspace orthogonal to E' is span by $|\phi\rangle$. Hence, $\tilde{\mu}'_\phi = \tilde{\mu}_\phi$. □

2.2 Plan

In the remainder of the paper, we consider variations of quantum probability spaces motivated by computation of numerical quantities in a world with limited resources:

- Instead of the Hilbert space $\mathcal{H}$ (constructed over the uncountable and uncomputable complex numbers $\mathbb{C}$), we will consider variants constructed over finite fields [20, 21, 22].
- Instead of real-valued probability measures producing results in the uncountable and uncomputable interval $[0,1]$, we will consider finite set-valued probability measures [4, 5].

We will then ask if it is possible to construct variants of quantum probability spaces under these conditions. The main question is related to the definition of probability measures: is it possible to still define a probability measure as a function that depends on a single state? Specifically,

- given a state $|\psi\rangle$, is there a probability measure mapping events to probabilities that only depends on $|\psi\rangle$? In the conventional quantum probability space, the answer is yes by the Born rule [18, 19] and the map is given by: $E \mapsto \langle\psi|E\psi\rangle$.
- given a probability measure μ mapping each event E to a probability, is there a *unique* state ψ such that $\mu(E) = \langle\psi|E\psi\rangle$? In the conventional case, the answer is yes by Gleason's theorem [16, 12, 11].

3 All Continuous or All Discrete

Before we turn to the main part of the paper, we quickly dismiss the possibility of having one but not the other of the discrete variations. Specifically, it is impossible to maintain the Hilbert space and have a finite set-valued probability measure and it is also impossible to have a vector space constructed over a finite field with a real-valued probability measure.

3.1 Hilbert Space with Finite Set-Valued Probability Measure

However, there is a $\mathscr{L}_2$-valued probability measure

$$\hat{\mu}_1(E) = \begin{cases} \textit{impossible} & \text{, if } E = |+\rangle\langle+|; \\ \bar{\mu}(E) & \text{, otherwise.} \end{cases}$$

such that $\hat{\mu}_1 \neq \bar{\mu}_\psi$ for all mixed state $|\psi\rangle$.

3.2 Discrete Vector Space with Real-Valued Probability Measure

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