

Example

	CA	US	MX
Buy = True	20	50	10
Buy = False	300	500	200

$$P(CA) = 210 / (210 + 550 + 210) = 0.19$$

$$P(US) = 550 / (210 + 550 + 210) = 0.51$$

$$P(MX) = 210 / (210 + 550 + 210) = 0.30$$

$$P(\text{country}, \text{bool}) = P(\text{bool} | \text{country}) P(\text{country})$$

$$P(\text{Buy} = \text{True}, \text{Country} = CA) = 20 / \underline{1080} = 0.019$$

→ very small numbers → *sum of all values*

32-bit number means 32 binary digits

→ log probability

Independence

$$A \perp B \text{ iff } p(A, B) = p(A) p(B)$$

$$\begin{aligned} \text{Buy} \perp \text{Country} &\Leftrightarrow p(\text{Buy}, \text{Country}) \\ &= p(\text{Buy}) p(\text{Country}) \end{aligned}$$

$$p(\text{Buy} | \text{Country}) = \frac{p(\text{Buy}, \text{Country})}{p(\text{Country})}$$

$$= \frac{P(\text{Buy}) P(\text{Country})}{P(\text{Country})}$$

$$= P(\text{Buy})$$

IID (independent and identically distributed)

— the outcome is independent and identically distributed

The Gambler's Fallacy ($\frac{1}{2} \times \frac{1}{2} \times \frac{1}{2} \times \frac{1}{2}$)

— It is a false believe that things will "balance out" in the end

eg) gambler who has just lost a bunch of times before it is more likely to they will win next

Monty Hall

assume you choose # 1

C : which door the car is behind ($C = 1, 2, 3$)

H : which door Monty Hall opens (assume $H \neq 2$)

$$P(H=2 | C=1) = 0.5$$

$$P(H=2 | C=2) = 0$$

$$P(H=2 | C=3) = 1$$

want $P(C=3 | H=2)$ $P(C=1 | H=2)$

$$\begin{aligned} P(C=3 | H=2) &= \frac{P(H=2, C=3)}{P(H=2)} \\ &= \frac{P(H=2 | C=3) P(C=3)}{\sum P(H=2 | C=c) P(C=c)} \\ &= \frac{P(H=2 | C=3) P(C=3)}{P(H=2 | C=1) P(C=1) + P(H=2 | C=2) P(C=2) + P(H=2 | C=3) P(C=3)} \\ &= \frac{1 \cdot (1/3)}{(1/2)(1/3) + 0(1/3) + 1(1/3)} \\ &= 2/3 \end{aligned}$$

$$\begin{aligned} P(C=1 | H=2) &= \frac{(1/2)(1/3)}{(1/2)(1/3) + 0(1/3) + 1(1/3)} \\ &= 1/3 \end{aligned}$$

→ should always switch door

Bernoulli Distribution

$$p(x) = \theta^x (1-\theta)^{1-x}$$

$$\text{ex) } p(1) = \theta^1 (1-\theta)^{1-1} = \theta$$

$$p(0) = \theta^0 (1-\theta)^{1-0} = 1-\theta$$

$$p(x) = p(0) + p(1) = 1$$

Maximum Likelihood Estimation

$$\text{data} = \{x_1, x_2, \dots, x_n\}$$

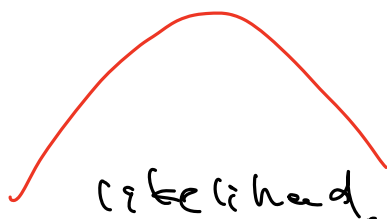
$$L(\theta) = p(\text{data} | \theta)$$

$$= \prod_{i=1}^n p(x_i | \theta)$$

$$= \prod_{i=1}^n \theta^{x_i} (1-\theta)^{1-x_i}$$

Why is it called maximum likelihood?

- What value of θ makes the data we collected most probable?
- What value of θ maximizes the likelihood?



Log - Likelihood

$\log(\cdot)$ is monotonically increasing

whatever θ maximizes L also maximizes $\log L$

$$\begin{aligned} \ell(\theta) &= \log L(\theta) = \log \prod_{i=1}^N \theta^{x_i} (1-\theta)^{1-x_i} \\ &= \sum_{i=1}^N \{ x_i \log \theta + (1-x_i) \log (1-\theta) \} \end{aligned}$$

$$\begin{aligned} \frac{d\ell}{d\theta} &= \sum_{i=1}^N \left(x_i \frac{1}{\theta} - (1-x_i) \frac{1}{(1-\theta)} \right) \\ &= \frac{1}{\theta} \sum_{i=1}^N x_i - \frac{1}{(1-\theta)} \sum_{i=1}^N (1-x_i) \end{aligned}$$

(数据个数) (数据个数)

C L T (central limit theorem)

- observation may due to sum of many random sources

ubiquitous (普遍存在)

Log-Likelihood of Normal distribution

$$l(\mu, \sigma^2) = \log L(\mu, \sigma^2) = \log \prod_{i=1}^N \frac{1}{\sqrt{2\pi\sigma^2}} e^{-\frac{1}{2}\left(\frac{x_i - \mu}{\sigma}\right)^2}$$

$$= \sum_{i=1}^N \log \frac{1}{\sqrt{2\pi\sigma^2}} e^{-\frac{1}{2}\left(\frac{x_i - \mu}{\sigma}\right)^2}$$

$$= \sum_{i=1}^N -\frac{1}{2} (\log 2\pi\sigma^2 - \frac{1}{\sigma^2} \left(\frac{x_i - \mu}{\sigma}\right)^2)$$

$$= -\frac{N}{2} \log 2\pi\sigma^2 - \frac{1}{2} \sum_{i=1}^N \left(\frac{x_i - \mu}{\sigma}\right)^2$$

$$\frac{\partial l}{\partial \mu} = \sum_{i=1}^N \left(\frac{x_i - \mu}{\sigma}\right) \frac{1}{\sigma}$$

$$\left(\left(\frac{x_i - \mu}{\sigma}\right)' = \left(\frac{x_i}{\sigma} - \frac{\mu}{\sigma}\right)' \right.$$

$$\left. = -\frac{1}{\sigma} \right.$$

$$= 0$$

$$\mu = \frac{1}{N} \sum_{i=1}^N x_i$$

$$l(\mu, \sigma^2) = -\frac{N}{2} \log(2\pi\sigma^2) - \frac{1}{2} \sum_{i=1}^N \left(\frac{x_i - \mu}{\sigma}\right)^2$$

$$v = \sigma^2 \text{ and } \sigma = \sqrt{v}$$

$$= -\frac{N}{2} \log(2\pi v) - \frac{1}{2} \frac{1}{v} \sum_{i=1}^N (x_i - \mu)^2$$

$$\frac{\partial l}{\partial v} = -\frac{N}{2} \cdot \frac{(2\pi v)'}{2\pi v} - \frac{1}{2} \cdot (-1) \frac{1}{v^2} \sum_{i=1}^N (x_i - \mu)^2$$

$$= -\frac{N}{2} \cdot \frac{1}{v} + \frac{1}{2} \frac{1}{v^2} \sum_{i=1}^N (x_i - \mu)^2$$

$$= 0$$

$$- \frac{N}{2} \cdot \frac{1}{\sigma} + \frac{1}{2} \frac{1}{\sigma^2} \sum_{i=1}^N (x_i - \mu) = 0$$

$$\sigma = \frac{1}{N} \sum_{i=1}^N (x_i - \mu)^2$$

Functions of random variables are random variable

$$Z = X_1 + X_2$$

$$X_1 \sim \text{Be}(p_1) \quad X_2 \sim \text{Be}(p_2)$$

$$Z = 0, 1, 2 \quad - \text{random}$$

MLE estimates are also random

$$\hat{\mu} = \frac{1}{N} \sum x_i = f(x_1, x_2, \dots, x_n)$$

$$\hat{\sigma}^2 = \frac{1}{N} \sum (x_i - \mu)^2 = g(x_1, x_2, \dots, x_n)$$



if $\hat{\mu} \in \mathcal{F}$ Random Variable, we can ask,

- what is their distribution
- what is their expectation

$$\hat{\mu} \sim \text{Normal} \left(\mu, \frac{\sigma^2}{N} \right)$$

$$E(\hat{\mu}) = \mu$$

$$E(\hat{\sigma}^2) \neq \sigma^2 \quad E(\hat{\sigma}^2) = \frac{n-1}{N} \sigma^2$$

expected value of the unbiased estimate is equal to the true variance

$$\text{unbiased estimate} : \hat{\sigma}^2 = \frac{1}{n-1} \sum_{i=1}^n (x_i - \bar{x})^2$$

(in ML, data is really large that \bar{x} and \bar{x}_{-1} doesn't make difference.

CDF (cumulative distribution function)

- CDF : $F(x) = P(X \leq x)$
- CDF calculated from PMF
 $P(X \leq x) = \sum_{k=-\infty}^x p(k)$, where $p(k) = \text{Prob}(X=k)$
- ex. $F(3) = p(1) + p(2) + p(3)$

CDF & PMF is both probabilities

continuous .

- CDF : $F(x) = P(X \leq x) = \int_{-\infty}^x f(u) du$

- pdf $f(x) = \frac{dF(x)}{dx}$

CDF is probability but pdf is not.

- cdf is always 0 at $-\infty$, and 1 at ∞

point estimate

- MLE - single value
- How confidence?

cf) two estimates (CTR A vs CTR B)

if we are not confident about our estimates, can we be confident that A is better than B?