

Confidence

two things affect our "confidence" in an estimate

- 1) how spread out the samples are
- 2) how many sample there are

Confidence Interval

Data $\{x_1, \dots, x_n\}$

Sample mean $\hat{\mu} = \bar{x} = \frac{1}{n} \sum x_i$

sample variance $\hat{\sigma}^2 = \frac{1}{n-1} \sum (x_i - \bar{x})^2$

95% confidence interval $[\hat{\mu} - 1.96 \cdot \frac{\hat{\sigma}}{\sqrt{n}}, \hat{\mu} + 1.96 \cdot \frac{\hat{\sigma}}{\sqrt{n}}]$

Handwritten notes in red:
- An arrow points from the text "95% confidence interval" to the term $\hat{\sigma}$ in the formula, with the note "分散の幅" (spread width).
- Another arrow points from the text "95% confidence interval" to the term \sqrt{n} in the formula, with the note "分散の大きさ" (size of spread).

Sample mean

$$\hat{\mu} = \bar{x} = \frac{1}{n} \sum_{i=1}^n x_i$$

functions of RVs are RVs

ex) x_1, x_2 : result of coin toss

$$Y = x_1 + x_2$$

$\{0, 1\} \in x_1, \{0, 1\} \in x_2$ etc

$\rightarrow Y$ is a random variable

Sums of Normal

$$X_1 \sim N(\mu_1, \sigma_1^2)$$

$$X_2 \sim N(\mu_2, \sigma_2^2)$$

$$Y = X_1 + X_2 \sim N(\mu_1 + \mu_2, \sigma_1^2 + \sigma_2^2)$$

↓

we can extend this to N IID

$$X \sim N(\mu, \sigma^2)$$

$$X_1 + X_2 + \dots + X_n \sim N(n\mu, n\sigma^2)$$

Mean of the Estimate

$$E(\hat{\mu}) = E\left(\frac{1}{n} \sum_{i=1}^n x_i\right)$$

$$= E\left(\frac{1}{n} (x_1 + x_2 + \dots + x_n)\right)$$

$$= \frac{1}{n} E(x_1) + \frac{1}{n} E(x_2) + \dots + \frac{1}{n} E(x_n)$$

$$= \frac{1}{n} \mu + \frac{1}{n} \mu + \dots + \frac{1}{n} \mu$$

$$= \mu$$

Variance of Estimate

$$\text{var}(\hat{\mu}) = \text{var}\left(\frac{1}{N} \sum_{i=1}^N x_i\right)$$

$$= E\left[\left(\frac{1}{N} \sum_{i=1}^N x_i - \mu\right)^2\right]$$

$$= \frac{1}{N^2} E\left[\left(\sum_{i=1}^N x_i - N\mu\right)^2\right]$$

$$= \frac{1}{N^2} \text{var}\left(\sum_{i=1}^N x_i\right)$$

$$= \frac{1}{N^2} N \sigma^2$$

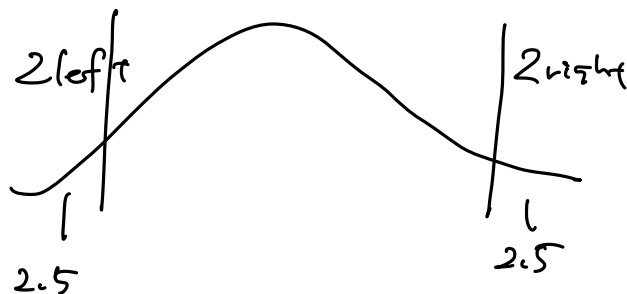
$$= \frac{\sigma^2}{N}$$

$$\sigma^2 = \text{var}[X] = E[(X - \mu)^2]$$

$$\hat{\mu} \sim N\left(\mu, \frac{\sigma^2}{N}\right)$$

when σ^2 is larger, variance of estimate is larger
 " N " " smaller

10-24 点関数



$$Z \sim N(0, 1)$$

$$0.025 = \int_{-\infty}^{z_{left}} \frac{1}{\sqrt{2\pi}} e^{-\frac{1}{2}z^2} dz = \Phi(z_{left})$$

cdf

$$z_{left} = \Phi^{-1}(0.025) = -1.96$$

Standardization

$$X \sim N(\mu, \sigma_0^2) \rightarrow Z = \frac{X - \mu}{\sigma_0} \sim N(0, 1)$$

Confidence Interval

$$-1.96 \leq Z \leq 1.96$$

$$-1.96 \leq \frac{\hat{\mu} - \mu}{\sigma_0} \leq 1.96$$

$$-1.96 \sigma_0 \leq \hat{\mu} - \mu \leq 1.96 \sigma_0$$

$$-1.96 \sigma_0 - \hat{\mu} \leq -\mu \leq 1.96 \sigma_0 - \hat{\mu}$$

$$\hat{\mu} - 1.96 \sigma_0 \leq \mu \leq \hat{\mu} + 1.96 \sigma_0$$

we can estimate

$$\sqrt{\frac{1}{n-1} \sum_{i=1}^n (x_i - \bar{x})^2}$$

$$95\% \text{ CI} = \left[\hat{\mu} - 1.96 \frac{\hat{\sigma}}{\sqrt{n}}, \hat{\mu} + 1.96 \frac{\hat{\sigma}}{\sqrt{n}} \right]$$

$$V \text{ CI} = \left[\hat{\mu} + \Phi^{-1}\left(\frac{1-\alpha}{2}\right) \frac{\hat{\sigma}}{\sqrt{n}}, \hat{\mu} + \Phi^{-1}\left(1 - \frac{1-\alpha}{2}\right) \frac{\hat{\sigma}}{\sqrt{n}} \right]$$

Ex. $V = 90\% \alpha = 0.1$

$$\hat{\mu} + \underbrace{\Phi^{-1}\left(\frac{1-0.1}{2}\right)}_{\Phi^{-1}\left(\frac{0.05}{2}\right)} \frac{\hat{\sigma}}{\sqrt{n}}, \hat{\mu} + \underbrace{\Phi^{-1}\left(1 - \frac{1-0.1}{2}\right)}_{\Phi^{-1}(1-0.05)} \frac{\hat{\sigma}}{\sqrt{n}}$$

CI : if I do some experiment many times, then 95% of those times the true μ

will be contained in the CI

CIT

coin flip : $X = \{0, 1\}$

$X \sim \text{Ber}(p)$ iid

$$\hat{p} = \bar{x}$$

$$E(\hat{p}) = E(X) = p$$

$$\text{var}(X) = p(1-p)$$

$$\hat{p} \rightarrow N(p, \frac{p(1-p)}{N})$$

$$t = \frac{\hat{\mu} - \mu}{\hat{\sigma} / \sqrt{N}} \times \frac{1/\sigma}{1/\sigma} = \left(\frac{\hat{\mu} - \mu}{\hat{\sigma} / \sqrt{N}} \right) / \left(\frac{\hat{\sigma}}{\sigma} \right)$$

$$\frac{1}{n-1} \sum_{i=1}^n \left(\frac{x_i - \bar{x}}{\hat{\sigma}} \right)^2$$

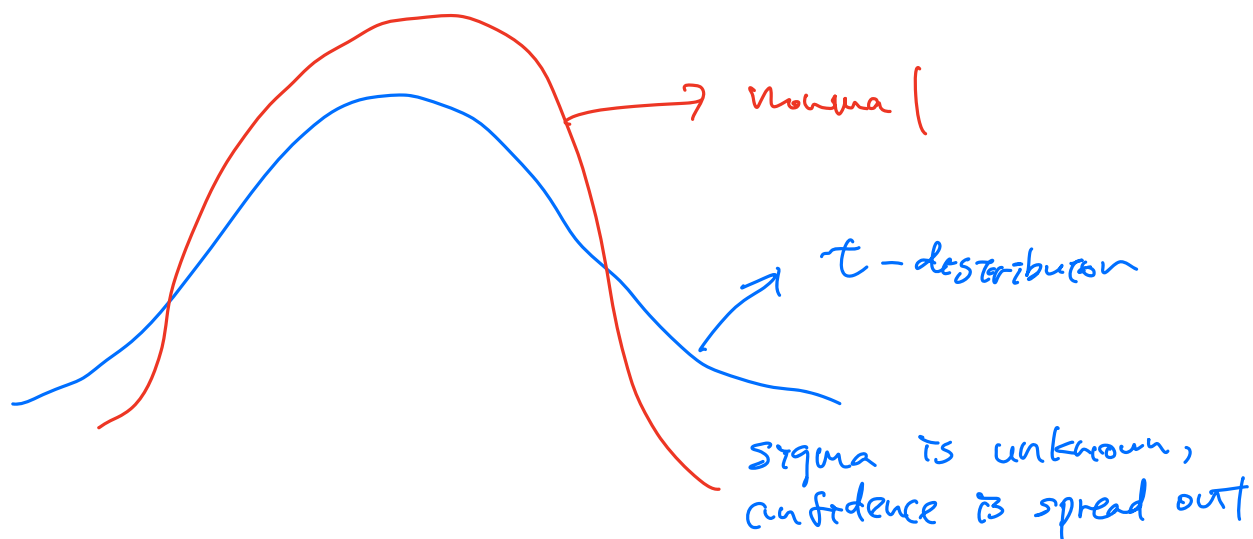
$$(n-1) \left(\frac{\hat{\sigma}}{\sigma} \right)^2 \sim \chi_{n-1}^2$$

$$Z \sim N(0, 1) \rightarrow \sum_{i=1}^n (Z_i - \bar{Z}) \sim \chi_{n-1}^2$$

T - Distribution

$$\frac{\text{standard normal}}{\sqrt{\text{chi-square / deg of freedom}}} \sim t \text{ deg of freedom}$$

when you standardize \bar{X} or $\hat{\mu}$ with the sample standard deviation, it's gonna be t-distribution



Hypothesis Testing

1-sample test : input data is 1-D array

2-sample test : " 2 1-D arrays

scipy, statsmodels or ~~python~~ ~~library~~

output : (test statistic, p-value)

the prob. of observing a result as extreme or more extreme than that was observed assuming the null hypothesis is true

ex) $H_0: \mu = 0$

Observed $\bar{x} = 100$, var = 10

→ unlikely, p value will be small

CI is looking inside the middle 95%
hypothesis test is looking outside at the 5%.

ex) significance threshold is 1%.

→ p-value is 0.1

$0.1 > 0.01$ (not statistically significant)

if the cost of a false alarm is high,
set a strict significance threshold

Terminology

if $p\text{-value} < \text{significance threshold}$
→ reject null hypothesis

EL, $p\text{-value} > \text{significance threshold}$
→ not binary (H0 or H1)
only reject or fail to reject

Examples

- 正片が写るのを判定
- 写らないかを?

How do we calculate the p-value?

$$P_{\text{right}} = 1 - \Phi(|z|)$$

$$P_{\text{left}} = \Phi(-|z|)$$

$$p = P_{\text{left}} + P_{\text{right}}$$

2 sample Z-test

convert it into a 1-sample test

Group 1 : $\{x_1, \dots, x_{N_1}\}$ with mean μ_1

Group 2 : $\{x_1, \dots, x_{N_2}\}$ with mean μ_2

$$H_0 : \mu_1 = \mu_2 \iff H_0 : \mu_1 - \mu_2 = 0$$

$$\text{Let } \gamma = \mu_1 - \mu_2$$

$$H_0 : \gamma = 0$$

$$H_1 : \gamma \neq 0$$

$$H_0 : \hat{\gamma} \sim N(0, \sigma_{\hat{\gamma}}^2) \quad \text{or} \quad H_0 : \hat{\gamma} \sim N(\mu_0, \sigma_{\hat{\gamma}}^2)$$

$$\hat{\gamma} = \frac{x_1 + x_2 + \dots + x_{N_1}}{N_1} - \frac{x'_1 + x'_2 + \dots + x'_{N_2}}{N_2}$$

$$\begin{aligned} \text{var}(\hat{\gamma}) &= \text{var}\left(\frac{x_1}{N_1}\right) + \dots + \text{var}\left(\frac{x_{N_1}}{N_1}\right) + \text{var}\left(\frac{x'_1}{N_2}\right) + \dots + \text{var}\left(\frac{x'_{N_2}}{N_2}\right) \\ &= \frac{\sigma_1^2}{N_1^2} + \dots + \frac{\sigma_1^2}{N_1^2} + \frac{\sigma_2^2}{N_2^2} + \dots + \frac{\sigma_2^2}{N_2^2} \\ &= \frac{\sigma_1^2}{N_1} + \frac{\sigma_2^2}{N_2} \end{aligned}$$

$$\sigma(\hat{\gamma}) = \sqrt{\frac{\sigma_1^2}{N_1} + \frac{\sigma_2^2}{N_2}}$$

Z - statistic

$$Z = \frac{\hat{\gamma} - \mu_0}{\sigma_{\hat{\gamma}}} = \frac{\hat{\gamma} - \mu_0}{\sqrt{\frac{\sigma_1^2}{N_1} + \frac{\sigma_2^2}{N_2}}}$$

検定統計量 : 帰無仮説が正しいと仮定したときに、
観測された事象がもたらす確率は
確率 α を計算する必要がある。

┌

有意水準

← 下 ; 上 →

「検定に成功した」と「検定に失敗した」という

検定に成功したとは、検定に失敗したとは、このように

└

p 値 : 帰無仮説に正しいと仮定した上で、
観測したような結果が極端なほど少ない確率

T-Test

- stop making assumption that we know std.

$$t = \frac{\hat{\mu} - \mu}{\sigma / \sqrt{n}}$$

Chi-square Test

- 0, 1 の値しかとる変数に適用して仮定する

→ この仮定を破る場合 chi-square test

paired sample

- instead of having two groups,
use only 1 group and test twice.

↔ independence sample Test

More than 2 groups.

- what if I test drug A, then B then drug C

$$\alpha \text{ Bonferroni} = \frac{\alpha}{\# \text{ tests}}$$

Frequentist Concerns.

- no peeking allowed (~~Let's peek~~)

Bayesian Approach

- we can answer, what is the probability that $\theta_A > \theta_B$?