

The two opposing forces

#1 collect data (exploration)

#2 select choice with highest win rate (exploitation)

Algorithm

- Epsilon-greedy
- optimistic initial values
- UCB1
- Thompson sampling

CTR of 13%

2 million, 5% to 7%?

→ 100 million, 1000 million CTR or CTR?

100, 1000 billion (infinite number of experiments)

Epsilon-Greedy

greedy - method

picking the bandit with highest MLE win rate with no regard to confidence in prediction or amount of data.

ex) bandits with 90% and 80%

$$E(R) = (1-\epsilon) 0.9 + \epsilon \left( \frac{0.8 + 0.9}{2} \right)$$

Sample mean

χ<sup>2</sup> test for independence,  $\bar{x}_n = \frac{1}{n} \sum_{i=1}^n x_i$  by empirical mean

→  $O(n)$  is  $O(1)$  is  $O(n)$

$$\bar{X}_N = \frac{1}{N} \left( \sum_{i=1}^{N-1} X_i + X_N \right)$$

$$\bar{X}_{N-1} = \frac{1}{N-1} \sum_{i=1}^{N-1} X_i \Leftrightarrow (N-1) \bar{X}_{N-1} = \sum_{i=1}^{N-1} X_i$$

$$\bar{X}_N = \frac{1}{N} ((N-1) \bar{X}_{N-1} + X_N)$$

$$= \frac{N-1}{N} \bar{X}_{N-1} + \frac{1}{N} X_N$$

$$= \left(1 - \frac{1}{N}\right) \bar{X}_{N-1} + \frac{1}{N} X_N$$

$$= \bar{X}_{N-1} + \frac{1}{N} (X_N - \bar{X}_{N-1})$$

Optimistic Initial values 算術平均

estimation初期値を  $\bar{X}_{N-1}$  とする

→ arithmetic average は  $\bar{X}_{N-1}$  とする

→ 各  $i$  の bandit の estimated mean  $\mu_i$  max estimated mean  $\mu_j$  の方向に, 探索する

Role of Initial Value

→ High initial value : 初期値が大きいと  $\bar{X}_{N-1}$  が大きくなる



Low initial value : 初期値が小さいと  $\bar{X}_{N-1}$  が小さくなる

# UCB 1

- upper confidence bound

- is fair

Epsilon greedy : small prob of random exploration

Optimistic : naturally start at large value  
(each bandit will be chosen often)

- applying probability

$$P(\text{sample mean} - \text{true mean} \geq \text{error}) \leq f(\text{error})$$

$$\text{ex) } P(\text{sample mean} - \text{true mean} \geq \epsilon) \leq 1/\epsilon$$

$$\begin{aligned} & - \text{if } \epsilon \geq \frac{1}{\sqrt{N-1}} \text{ then } P(\text{sample mean} - \text{true mean} \geq \epsilon) \leq \frac{1}{\sqrt{N-1}} \\ & - \text{if } \epsilon < \frac{1}{\sqrt{N-1}} \text{ then } P(\text{sample mean} - \text{true mean} \geq \epsilon) \leq \frac{1}{\sqrt{N-1}} \end{aligned}$$

Markov inequality / Chebyshev inequality / Hoeffding

$$P(\bar{X}_n - E(X) \geq \epsilon) \leq e^{-2n\epsilon^2}$$

↓  
if  $\bar{X}_n - E(X) \geq \epsilon$  then  $P(\bar{X}_n - E(X) \geq \epsilon) \leq e^{-2n\epsilon^2}$

$$\hat{j} = \underset{j}{\operatorname{argmax}} \left( \bar{X}_{n_j} + \sqrt{2 \frac{\log N}{n_j}} \right)$$

↓  
sample mean

$N$  : total plays

$n_j$  : plays made on bandit  $j$

# Thompson Sampling

CI は 思い出す。

Small Dataset : Large confidence interval

Fat  $\rightarrow$  Explore more, skinny  $\rightarrow$  explore less

CLT: sum of  $R_{i,t}$   $R_{i,t} \sim$  normal Distribution

What we want?

— the distribution of mean (with rate)

$\rightarrow$  is a distribution of parameter  $\theta$   
(for a distribution is given)

$P(X|\theta)$ : 尤度  
参数  $\theta$  が与えられた時の,  $\theta$  の確率

if the evidence is 考えずに, 例外的に?  $\rightarrow$

evidence (が 与えられた), 計算  $p_i$  したい

(Monte Carlo は 計算  $p_i$  へて...)

事後分布  $\propto$  尤度  $\cdot$  事前分布

共役事前分布 について

ex) Gaussian  $\propto$  Gaussian  $\times$  Gaussian

例. 尤度を  $N(\mu, \sigma^2)$  分布

$\rightarrow$  共役事前分布 (conjugate Prior) は  
Beta 分布。

$$\begin{aligned} p(\theta | x) &\propto \left( \prod_{i=1}^n \theta^{x_i} (1-\theta)^{(1-x_i)} \right) \left( \frac{1}{B(\alpha, \beta)} \theta^{\alpha-1} (1-\theta)^{\beta-1} \right) \\ &\propto \left( \theta^{\sum_{i=1}^n x_i} (1-\theta)^{(n - \sum_{i=1}^n x_i)} \right) \left( \theta^{\alpha-1} (1-\theta)^{\beta-1} \right) \\ &= \left( \theta^{\sum_{i=1}^n x_i} (1-\theta)^{\sum_{i=1}^n (1-x_i)} \right) \left( \theta^{\alpha-1} (1-\theta)^{\beta-1} \right) \\ &= \left( \theta^{\sum_{i=1}^n x_i + \alpha - 1} (1-\theta)^{\sum_{i=1}^n (1-x_i) + \beta - 1} \right) \end{aligned}$$

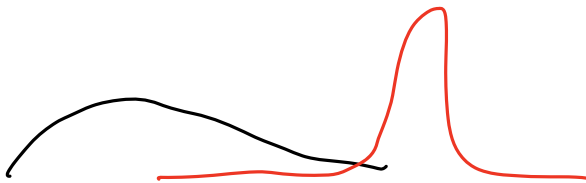
உயர்த்திப் பிறகு எழுதினார்களா?

→ Beta(1, 1) は - 標準分布

もしドメイン知識がなければこれを使用する。

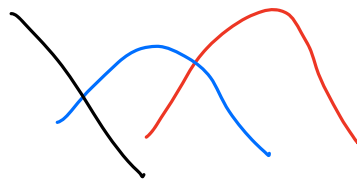
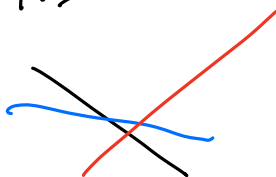
$\chi$  は陽子数に比例する新粒子

ex) Prior =  $\text{Beta}(1,1)$ , collect 1, post  $(1+1, 1+1) = \text{Beta}(2,1)$   
 $(2,1)$  1  $\text{Beta}(3,1)$   
 $(3,1)$  0  $\text{Beta}(3,2)$



→ is a suboptimal policy, 黒い方より緑の方の方が  
suboptimal なの方が 2 倍ある

有盤



1. 何をどういってやるか、  
どうやるか

Reynolds coming from normal distribution

- precision = 1 / variance

$$p(x|\mu, \tau) = \prod_{i=1}^N \sqrt{\frac{\tau}{2\pi}} e^{-\frac{\tau}{2} (x_i - \mu)^2}$$

$$X \sim \mathcal{N}(\mu, \tau^{-1}), \quad \mu|X \sim \mathcal{N}(m, \lambda^{-1})$$

$$p(\mu|X) = \frac{p(\mu, X)}{p(X)} = \frac{p(X|\mu) p(\mu)}{p(X)}$$

$$X \sim \mathcal{N}(\mu, \tau^{-1}), \quad \mu \sim \mathcal{N}(m_0, \lambda_0^{-1}), \quad \mu|X \sim \mathcal{N}(m, \lambda^{-1})$$

$$p(\mu|X) \propto p(X|\mu) p(\mu)$$

$$= \left( \prod_{i=1}^N \sqrt{\frac{\tau}{2\pi}} e^{-\frac{\tau}{2} (x_i - \mu)^2} \right) \left( \sqrt{\frac{\lambda_0}{2\pi}} e^{-\frac{\lambda_0}{2} (\mu - m_0)^2} \right)$$

$$= \left( \left[ \frac{\tau}{2\pi} \right]^N e^{-\frac{\tau}{2} \sum_{i=1}^N (x_i - \mu)^2} \right) \left( \dots \right) \quad \text{const } \tau \downarrow \sum_{i=1}^N x_i \sim 9$$

$$\propto \left( e^{-\frac{\tau}{2} \sum_{i=1}^N (x_i - \mu)^2} \right) \left( e^{-\frac{\lambda_0}{2} (\mu - m_0)^2} \right)$$

$$= e^{-\frac{\tau}{2} \sum_{i=1}^N (x_i - \mu)^2 - \frac{\lambda_0}{2} (\mu - m_0)^2}$$

$$= e^{-\frac{\tau}{2} \sum_{i=1}^N (\mu^2 - 2\mu x_i + x_i^2) - \frac{\lambda_0}{2} (\mu^2 - 2\mu m_0 + m_0^2)}$$

$$= \exp \left( -\frac{\tau}{2} \sum_{i=1}^N (\mu^2 - 2\mu x_i + x_i^2) - \frac{\lambda_0}{2} (\mu^2 - 2\mu m_0 + m_0^2) \right)$$

$$= \exp \left( -\frac{\tau}{2} (N\mu^2 - 2\mu \sum_{i=1}^N x_i + \sum_{i=1}^N x_i^2) - \frac{\lambda_0}{2} \mu^2 - \frac{\lambda_0}{2} 2\mu m_0 - \frac{\lambda_0}{2} m_0^2 \right)$$

$$\propto \exp \left( -\frac{\tau}{2} (N\mu^2 - 2\mu \sum_{i=1}^N x_i) - \frac{\lambda_0}{2} (\mu^2 - 2\mu m_0) \right)$$

$$= \exp \left( -\frac{\tau N + \lambda_0}{2} \mu^2 + \left( \tau \sum_{i=1}^N x_i + \lambda_0 m_0 \right) \mu \right)$$

$$\begin{aligned}
 p(\mu|x) &= \sqrt{\frac{\lambda}{2\pi}} \exp\left(-\frac{\lambda}{2}(\mu-m)^2\right) \\
 &= \sqrt{\frac{\lambda}{2\pi}} \exp\left(-\frac{\lambda}{2}(\mu^2 - 2m\mu + m^2)\right) \\
 \mathcal{L} &= \exp\left(-\frac{\lambda}{2}(\mu^2 - 2m\mu)\right) \\
 &= \exp\left(-\frac{\lambda}{2}\mu^2 + m\lambda\mu\right)
 \end{aligned}$$

$$\exp\left(-\frac{\tau N + \lambda_0}{2}\mu^2 + \left(\tau \sum_{i=1}^N x_i + \lambda_0 m_0\right)\mu\right)$$

$$\exp\left(-\frac{\lambda}{2}\mu^2 + m\lambda\mu\right)$$

$$\lambda = \tau N + \lambda_0$$

$$m\lambda = \tau \sum_{i=1}^N x_i + \lambda_0 m_0 \Leftrightarrow m = \frac{1}{\lambda} \left( \tau \sum_{i=1}^N x_i + \lambda_0 m_0 \right)$$

$$m = \frac{1}{\tau N + \lambda_0} \left( \tau \sum_{i=1}^N x_i + \lambda_0 m_0 \right)$$

$$\frac{x - \mu}{\sigma} = z$$

$$x = \sigma z + \mu$$