$$E[(Y - \hat{f}(x))^2 | X = x] = [f(x) - \hat{f}(x)]^2 + Var(\epsilon)$$

MSE =
$$1/n \sum (y_i - \hat{f}(x))^2 = E[(Y - \hat{f}(x))^2]$$

$$E[(y_0 - \hat{f}(x_0))^2] = Var(\hat{f}(x_0)) + (f(x_0) + E(\hat{f}(x_0)))^2 + Var(\epsilon)$$

Classification Error (Loss) = $1/n \sum I[y_i \neq \widehat{C(x_i)}]$

$$RSS = \sum (y_i - \widehat{\mathbb{Z}}_0 - \widehat{\mathbb{Z}}_1 x_i)^2 - Matrix form (Y - x\beta)^T (Y - x\beta) = \epsilon^T \epsilon$$

TSS =
$$\sum (y_i - \overline{y})^2$$
; R² = $\frac{TSS - RSS}{TSS}$; RSE = $\sqrt{\frac{RSS}{n-n}}$

$$\beta_0 = \overline{y} - \beta_1 \overline{x}; \quad \beta_1 = \frac{\sum (x_i - \overline{x}) (y_i - \overline{y})}{\sum (x_i - \overline{x})^2} = \frac{\sum x_i (y_i - \overline{y})}{\sum x_i (x_i - \overline{x})}$$

$$SE(\widehat{\beta_0})^2 = \sigma^2 \left[\frac{1}{n} + \frac{\overline{x^2}}{\sum (x_i - \overline{x})^2} \right], SE(\widehat{\beta_1})^2 = \sigma^2 \left[\frac{1}{n} + \frac{\sigma^2}{\sum (x_i - \overline{x})^2} \right], \text{ where } \sigma^2 = Var(\epsilon);$$

95% Confidence Interval for $\beta_1 = [\widehat{\mathbb{Q}}_1 - 2 * SE(\widehat{\mathbb{Q}}_1), \widehat{\mathbb{Q}}_1 + 2 * SE(\widehat{\mathbb{Q}}_1)]$

$$\widehat{\sigma^2} = \frac{RSS}{n-2}$$
; $\square = (x^T x)^{-1} x^T y \leftarrow OLS$ Estimate

$$Var(\mathbb{Z}) = (x^{\mathsf{T}}x)^{-1}\sigma^2 \leftarrow \text{Matrix Form} \mid \text{Non-Matrix Form} \rightarrow \text{SE}(\widehat{\beta_1})^2 = \frac{\sigma^2}{\sum (x_i - \overline{x})^2}$$

$$\beta_{Px1} = \operatorname{argmin}_{\beta}(y_i - x\beta)^{\mathsf{T}}(y_i - x\beta); \quad \text{T-statistic } (\beta_1) = \frac{\widehat{\beta_1}}{SE(\widehat{\beta_1})}$$

Correlation Coefficient:
$$r = \frac{\sum (x_i - \overline{x}) (y_i - \overline{y})}{\sqrt{\sum (x_i - \overline{x})^2} \sqrt{\sum (y_i - \overline{y})^2}}$$
; F-statistic = $\frac{(TSS - RSS)/p}{RSS/(n - p - 1)}$

Logistic Regression p(x) =
$$\frac{e^{(\beta_0 + \beta_1 X)}}{1 + e^{(\beta_0 + \beta_1 X)}}$$

If
$$Y_i = \beta_0 + \beta_1 x_i + \epsilon_i$$
 and $\epsilon_i \sim N(0, \sigma^2)$

$$\ell = \prod_{i=1}^{p} \frac{1}{\sqrt{2\pi\sigma^{2}}} e^{\left(-\frac{(\epsilon_{i}-0)^{2}}{2\sigma^{2}}\right)} = \prod_{i=1}^{p} \frac{1}{\sqrt{2\pi\sigma^{2}}} e^{\left(-\frac{(y_{i}-\beta_{0}-\beta_{1}x_{i})^{2}}{2\sigma^{2}}\right)}$$

Maximum Likelihood Estimate = Min RSS

Bayes Theorem: $Pr(Y=k|X=x) = (Pr(X=x|Y=k) \cdot Pr(Y=k)) / Pr(X=x)$

Bayes Classifier:
$$p_k(x) = \frac{\pi_k f_k(x)}{\sum\limits_{l=1}^K \pi_l f_l(x)}$$
 and $log(p_k(x)) = log(\pi_k) + log(f_k(x)) + const$

$$\widehat{\delta \mathbb{Z}}(x) = x \cdot \frac{\widehat{\mu \mathbb{Z}}}{\widehat{\sigma^2}} - \frac{\widehat{\mu \mathbb{Z}}}{\widehat{2\sigma^2}} + log(\widehat{\pi}\mathbb{Z}) \leftarrow LDA$$
 Single-variate Discriminant

$$\begin{split} &\delta_k = x^T \boldsymbol{\Sigma}^{-1} \boldsymbol{\mu}_k - \boldsymbol{\mu}_k^T \boldsymbol{\Sigma}^{-1} \boldsymbol{\mu}_k + \log \pi_k < \text{- LDA Multivariate Discriminant} \\ &\delta_k = -\frac{1}{2} x^T \; \boldsymbol{\Sigma}_k^{-1} x \; + x^T \boldsymbol{\Sigma}_k^{-1} \boldsymbol{\mu}_k - \frac{1}{2} \boldsymbol{\mu}_k^T \boldsymbol{\Sigma}^{-1} \boldsymbol{\mu}_k - \frac{1}{2} \log \left| \boldsymbol{\Sigma}_k \right| + \log \pi_k < -\frac{1}{2} \boldsymbol{\mu}_k^T \boldsymbol{\Sigma}^{-1} \boldsymbol{\mu}_k - \frac{1}{2} \log \left| \boldsymbol{\Sigma}_k \right| + \log \pi_k < -\frac{1}{2} \boldsymbol{\mu}_k^T \boldsymbol{\Sigma}^{-1} \boldsymbol{\mu}_k - \frac{1}{2} \log \left| \boldsymbol{\Sigma}_k \right| + \log \pi_k < -\frac{1}{2} \boldsymbol{\mu}_k^T \boldsymbol{\Sigma}^{-1} \boldsymbol{\mu}_k - \frac{1}{2} \log \left| \boldsymbol{\Sigma}_k \right| + \log \pi_k < -\frac{1}{2} \boldsymbol{\mu}_k^T \boldsymbol{\Sigma}^{-1} \boldsymbol{\mu}_k - \frac{1}{2} \log \left| \boldsymbol{\Sigma}_k \right| + \log \pi_k < -\frac{1}{2} \boldsymbol{\mu}_k^T \boldsymbol{\Sigma}^{-1} \boldsymbol{\mu}_k - \frac{1}{2} \log \left| \boldsymbol{\Sigma}_k \right| + \log \pi_k < -\frac{1}{2} \boldsymbol{\mu}_k^T \boldsymbol{\Sigma}^{-1} \boldsymbol{\mu}_k - \frac{1}{2} \log \left| \boldsymbol{\Sigma}_k \right| + \log \pi_k < -\frac{1}{2} \boldsymbol{\mu}_k^T \boldsymbol{\Sigma}^{-1} \boldsymbol{\mu}_k - \frac{1}{2} \log \left| \boldsymbol{\Sigma}_k \right| + \log \pi_k < -\frac{1}{2} \boldsymbol{\mu}_k^T \boldsymbol{\Sigma}^{-1} \boldsymbol{\mu}_k - \frac{1}{2} \log \left| \boldsymbol{\Sigma}_k \right| + \log \pi_k < -\frac{1}{2} \log \left| \boldsymbol{\Sigma}_k \right| + \log \pi_k < -\frac{1}{2} \log \left| \boldsymbol{\Sigma}_k \right| + \log \pi_k < -\frac{1}{2} \log \left| \boldsymbol{\Sigma}_k \right| + \log \pi_k < -\frac{1}{2} \log \left| \boldsymbol{\Sigma}_k \right| + \log \pi_k < -\frac{1}{2} \log \left| \boldsymbol{\Sigma}_k \right| + \log \pi_k < -\frac{1}{2} \log \left| \boldsymbol{\Sigma}_k \right| + \log \pi_k < -\frac{1}{2} \log \left| \boldsymbol{\Sigma}_k \right| + \log \pi_k < -\frac{1}{2} \log \left| \boldsymbol{\Sigma}_k \right| + \log \pi_k < -\frac{1}{2} \log \left| \boldsymbol{\Sigma}_k \right| + \log \pi_k < -\frac{1}{2} \log \left| \boldsymbol{\Sigma}_k \right| + \log \pi_k < -\frac{1}{2} \log \left| \boldsymbol{\Sigma}_k \right| + \log \pi_k < -\frac{1}{2} \log \left| \boldsymbol{\Sigma}_k \right| + \log \pi_k < -\frac{1}{2} \log \left| \boldsymbol{\Sigma}_k \right| + \log \pi_k < -\frac{1}{2} \log \left| \boldsymbol{\Sigma}_k \right| + \log \pi_k < -\frac{1}{2} \log \left| \boldsymbol{\Sigma}_k \right| + \log \pi_k < -\frac{1}{2} \log \left| \boldsymbol{\Sigma}_k \right| + \log \pi_k < -\frac{1}{2} \log \left| \boldsymbol{\Sigma}_k \right| + \log \pi_k < -\frac{1}{2} \log \left| \boldsymbol{\Sigma}_k \right| + \log \pi_k < -\frac{1}{2} \log \left| \boldsymbol{\Sigma}_k \right| + \log \pi_k < -\frac{1}{2} \log \left| \boldsymbol{\Sigma}_k \right| + \log \pi_k < -\frac{1}{2} \log \left| \boldsymbol{\Sigma}_k \right| + \log \pi_k < -\frac{1}{2} \log \left| \boldsymbol{\Sigma}_k \right| + \log \pi_k < -\frac{1}{2} \log \left| \boldsymbol{\Sigma}_k \right| + \log \pi_k < -\frac{1}{2} \log \left| \boldsymbol{\Sigma}_k \right| + \log \pi_k < -\frac{1}{2} \log \left| \boldsymbol{\Sigma}_k \right| + \log \pi_k < -\frac{1}{2} \log \left| \boldsymbol{\Sigma}_k \right| + \log \pi_k < -\frac{1}{2} \log \left| \boldsymbol{\Sigma}_k \right| + \log \pi_k < -\frac{1}{2} \log \left| \boldsymbol{\Sigma}_k \right| + \log \pi_k < -\frac{1}{2} \log \left| \boldsymbol{\Sigma}_k \right| + \log \pi_k < -\frac{1}{2} \log \left| \boldsymbol{\Sigma}_k \right| + \log \pi_k < -\frac{1}{2} \log \left| \boldsymbol{\Sigma}_k \right| + \log \pi_k < -\frac{1}{2} \log \left| \boldsymbol{\Sigma}_k \right| + \log \pi_k < -\frac{1}{2} \log \left| \boldsymbol{\Sigma}_k \right| + \log \pi_k < -\frac{1}{2} \log \left| \boldsymbol{\Sigma}_k \right| + \log \pi_k < -\frac{1}{2} \log \left| \boldsymbol{\Sigma}_k \right| + \log \pi_k < -\frac{1}{2} \log \left| \boldsymbol{\Sigma}_k \right| + \log \pi_k < -\frac{1}{2} \log \left| \boldsymbol{\Sigma}_k$$

$$Pr(y = k|x = x) = \frac{e^{\delta_k(x)}}{\sum_{\ell}^{K} e^{\delta_{\ell}(x)}}$$

Maximum Log Likelihood QDA = $\max \sum log(\frac{1}{\sqrt{2\pi\sigma^2}}e^{(-\frac{(x_i-\mu_{vi})^2}{2\sigma^2})}) = \max$

$$\sum_{i=1}^{N} (-\log \sigma - \frac{1}{2\sigma^2} (x_i - \mu_{y_i})^2 + consts)$$

$$X = [a b, c d]; det(X) = ad - bc; X^{-1} = (1/detX)[d - b, -c a]$$

↑Flexibility → ↑Var + ↓Bias

Confusion Matrix True

Pred Correct Type 1(False +)

Type 2(False -) Correct

FPR: FP / (TN + FP); (type 1)

TPR: TP / (TP + FN)

QDA

$$C(x) : j$$
, if $p_i(x) = max\{p_1(x),....,p_k(x)\}$, where $p_k(x) = Pr(Y=k \mid X=x)$

Multivariate Gaussian PDF:

$$f\left(oldsymbol{x}
ight) = rac{1}{\left(2\pi
ight)^{p/2} |oldsymbol{\Sigma}|^{1/2}} \exp\left\{-rac{1}{2} \left(oldsymbol{x} - oldsymbol{\mu}
ight)^T oldsymbol{\Sigma}^{-1} \left(oldsymbol{x} - oldsymbol{\mu}
ight)
ight\}$$

Logistic regression model: $log(\frac{p(x)}{1-p(x)}) = \beta_0 + \beta_1 X$

Log reg. PDF:
$$p(x) = \frac{e^{\beta_0 + \beta_1 x}}{1 + e^{\beta_0 + \beta_1 x}}$$

Log Likelihood Log reg.: $log \ell(x) = log(\prod_{i=1}^{N} (p(x_i)^{y_i} (1 - p(x_i))^{1-y_i}))$

 $Var(X+Y) = Var(X)+Var(Y) + 2Cov(X,Y); Var(aX) = a^2Var(X); Cov(aX,bY) = abCov(X,Y);$

TSS = $\sqrt{RSE/(n-2)}/(1-R^2)$; LOOCV: CV = 1/n \sum MSE

Speeding up LOOCV: CV = $1/n \sum ((y_i - \widehat{y_i})/(1 - h_i))^2$

$$h_i = 1/n + ((x_i - \overline{x})^2 / \Sigma (x_i - \overline{x})^2)$$

K-fold cv: ∑ n_k/n MSE

classification cv: ∑ n_k/n Err

Err =
$$\sum I(y_i \neq \widehat{y_i})/n_k$$

$$C_p = \frac{1}{n} \left(\text{RSS} + 2d\hat{\sigma}^2 \right), \text{ AIC} = -2 \log L + 2 \cdot d$$

L is likelihood function

$$BIC = \frac{1}{n} \left(RSS + \log(n) d\hat{\sigma}^2 \right).$$

Maximum Log Likelihood = $\max \sum log(\frac{1}{\sqrt{2\pi\sigma^2}}e^{(-\frac{(x_i-\mu_{\gamma i})^2}{2\sigma^2})})$

$$R^2 = \frac{TSS - RSS}{TSS} = 1 - \frac{RSS}{TSS}$$

Adjusted
$$R^2 = 1 - \frac{\text{RSS}/(n-d-1)}{\text{TSS}/(n-1)}$$
.

Ridge regression (parameter λ), ℓ_2 penalty

Constrained Ridge

$$\min_{\beta} \text{RSS}(\beta) + \lambda \sum_{j} \beta_{j}^{2} =$$

$$\min_{\beta} \sum_{i=1}^{n} \left(y_{i} - \beta_{0} - \sum_{i=1}^{p} \beta_{j} x_{ij} \right)^{2} + \lambda \sum_{i} \beta_{j}^{2}$$

$$\hat{\beta}^{\text{ridge}} = \mathop{\arg\min}_{\beta} \| Y - X^T \beta \|^2 + \lambda \|\beta\|^2,$$

Another way to formulate the ridge regression is

$$\begin{split} \hat{\beta}^{\text{ridge}} &= \mathop{\arg\min}_{\beta} \|Y - X^T \beta\|^2 \\ &\text{subject to } \|\beta\|^2 \leq t, \end{split}$$

For ridge regression: $\hat{\beta} = (X^T X + \lambda I)^{-1} X^T Y$

$$\tilde{x}_{ij} = \frac{x_{ij}}{\sqrt{\frac{1}{n}\sum_{i=1}^{n}(x_{ij}-\overline{x}_{j})^{2}}}$$
 <= Standardized predictors

$$\min_{\beta} \mathrm{RSS}(\beta) + \lambda \sum_{j} |\beta_{j}| =$$
 <= Lasso
$$\min_{\beta} \sum_{i=1}^{n} \left(y_{i} - \beta_{0} - \sum_{j=1}^{p} \beta_{j} x_{ij} \right)^{2} + \lambda \sum_{j} |\beta_{j}|$$

$$\underset{\beta}{\text{minimize}} \sum_{i=1}^{n} \left(y_i - \beta_0 - \sum_{j=1}^{p} \beta_j x_{ij} \right)^2 \quad \text{subject to} \quad \sum_{j=1}^{p} |\beta_j| \le s$$

Constrained Lasso =>

Polynomial =>
$$Y = \beta_0 + \beta_1 X + \beta_2 X^2 + \dots + \beta_p X^p$$

Step function => $Y = \beta_0 + \beta_1 I(X < c_1) + \beta_2 I(c_1 < X < c_2) + \dots + \beta_p I(c_p < X)$
Cubic Spline => $Y_i = \beta_0 + \beta_1 X_i + \beta_2 X_i^2 + \beta_3 X_i^2 + \beta_4 (X_i - \xi_1)_+ + \dots + \beta_{k+3} (X_i - \xi_k)_+$

$$\underset{g \in \mathcal{S}}{\text{minimize}} \sum_{i=1}^{n} (y_i - g(x_i))^2 + \lambda \int g''(t)^2 dt$$

Smoothing Spline =>

Gini Index =>
$$G = \sum_{k=1}^{K} \hat{P}_{mk} (1 - \hat{P}_{mk})$$

Cross-Entropy =>
$$D = -\sum_{k=1}^{K} \hat{P}_{mk} log_2(\hat{P}_{mk})$$

If $Var(\alpha X + (1 - \alpha)Y)$ then minimizing this is $\alpha = \frac{\sigma_Y^2 - \sigma_{XY}}{\sigma_X^2 + \sigma_Y^2 - 2\sigma_{XY}}$

SVM (Hard-Margin)

SVM (Soft-Margin)

$$\max_{\beta_0,\beta_1,\dots,\beta_p,M} \beta_0,\beta_1,\dots,\beta_p,\epsilon_1,\dots,\epsilon_n,M$$

$$\max_{\beta_0,\beta_1,\dots,\beta_p,M} \text{subject to } \sum_{j=1}^p \beta_j^2 = 1,$$

$$\sup_{j=1} \sum_{j=1}^p \beta_j^2 = 1,$$

$$y_i(\beta_0 + \beta_1 x_{i1} + \beta_2 x_{i2} + \dots + \beta_p x_{ip}) \geq M(1 - \epsilon_i),$$

$$y_i(\beta_0 + \beta_1 x_{i1} + \beta_2 x_{i2} + \dots + \beta_p x_{ip}) \geq M \ \forall i = 1,\dots,n.$$

$$\epsilon_i \geq 0, \ \sum_{i=1}^n \epsilon_i \leq C,$$

$$\epsilon_1,\dots,\epsilon_n \ \text{are } slack \ variables$$

$$f(x) = \hat{\beta}_0 + \hat{\beta}_1 x_1 + \dots + \hat{\beta}_p x_p = \hat{\beta}_0 + \sum_{i=1}^n \hat{\alpha}_i \langle x, x_i \rangle$$

Linear support vector classifier =>

$$\text{Kernel version $^{\Lambda\Lambda} = >$} f(x) = \hat{\beta}_0 + \sum_{i=1}^n \hat{\alpha}_i k(x,x_i)$$

Thei version
$$K = >$$
 $E = 1$ $K(x_i, x_{i'}) = \left(1 + \sum_{i=1}^p x_{ij} x_{i'j}\right)^{r}$

Polynomial Kernel =>

$$K(x_i, x_{i'}) = \exp(-\gamma \sum_{j=1}^{p} (x_{ij} - x_{i'j})^2).$$

Radial Kernel =>

PCA and Eigen: $maximize \ \varphi \ with \ \ \varphi_1^{\ T} \Sigma \varphi_1 \ subject \ to \ \ \varphi_1^{\ T} \varphi_1 = 1$ which is equivalent to $\ \ \varphi_1^{\ T} \Sigma \varphi_1 - \lambda (\varphi_1^{\ T} \varphi_1 - 1)$ and then taking gradient and setting to zero ends up being λ

PCA: $Z_1 = \Phi_{11}X_1 + \Phi_{21}X_2 + \dots + \Phi_{p1}X_p$ for single component

$$\underset{\phi_{11},...,\phi_{p1}}{\text{maximize}} \frac{1}{n} \sum_{i=1}^{n} \left(\sum_{j=1}^{p} \phi_{j1} x_{ij} \right)^{2} \text{ subject to } \sum_{j=1}^{p} \phi_{j1}^{2} = 1.$$

$$Z_{nx1}\Phi_1^T = x_{nx2}^* = denoised data$$

$$\hat{\Sigma} = \frac{X^T X}{n}$$

$$\sum_{j=1}^{p} Var(x_j) = \sum_{m=1}^{M} Var(Z_m)$$

K-means:

$$\min C_1 ... C_k \sum_{k=1}^K WCV(C_k) \ where \ WCV(C_k) \ = \frac{1}{|C_k|} \sum_{i,i' \in C_k} \sum_{j=1} (x_{ij} - x_{i'j})^2 = \frac{1}{|C_k|} \sum_{i \in C_k} \sum_{j=1} (x_{ij} - \bar{x}_{kj})^2$$

As K increases, cost function decreases