

On A Mallows-type Model For (Ranked) Choices



Yifan Feng¹², and Yuxuan Tang²

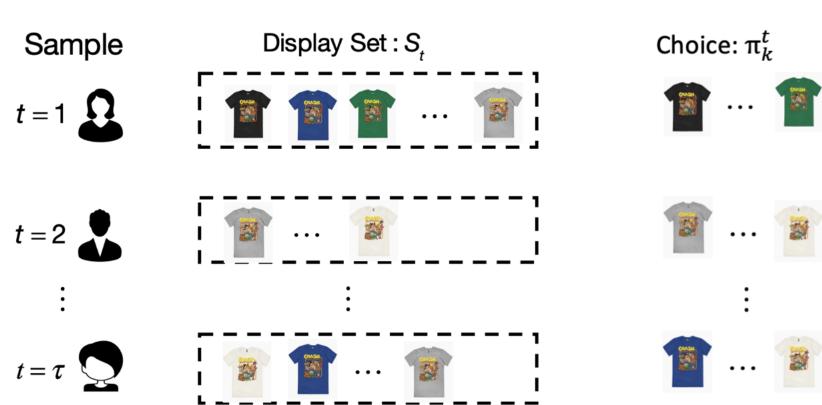
¹Department of Analytics and Operations, and ²Institute of Operations Research and Analytics, National University of Singapore, Singapore

Summary of Results

- We identify a new distance-based (Mallows-type) ranking model.
- It aggregates into a simple closed-form probability of any top-k subranking π_k among an arbitrary subset S, i.e., $\mathbb{P}(\pi_k|S)$.
- We develop effective parameter learning with theoretically proven consistency.

A Motivating Example

To learn the preference over n product prototypes, a company displays different subsets to customers and ask for top-k feedback.



How to aggregate the population's preferences from their ranked choices?

What are Mallows-type Models?

A collection of **Probability distribution of rankings**.

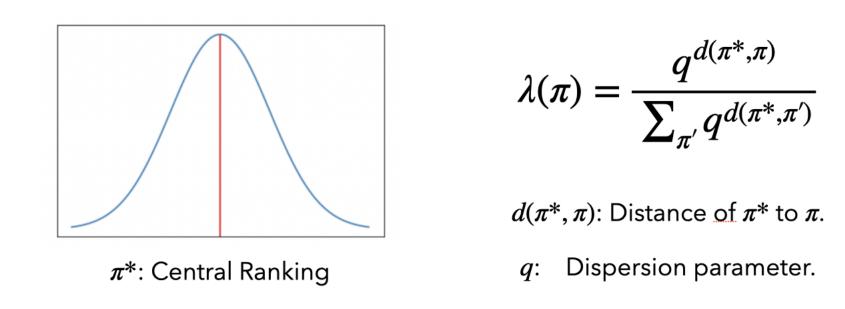


Figure 1: The probability of a ranking $\lambda(\pi)$ exponentially decays as its distance to the central ranking increases.

Different distance functions lead to different distributions.

Challenge

Under any existing Mallows-type model, it's difficult to obtain ranked choice probabilities even when k = 1.

A New Distance Function: Reverse Major Index (RMJ)

Assume the central ranking is the identity ranking, i.e, $\pi^* = (1, 2, ..., n)$.

$$\mathbf{d_R}(\pi) = \sum_{\mathbf{i}=\mathbf{1}}^{\mathbf{n}-\mathbf{1}} \mathbb{I}\{\pi(\mathbf{i}) > \pi(\mathbf{i}+\mathbf{1})\} \cdot (\mathbf{n}-\mathbf{i})$$
 $\mathbf{k}=\mathbf{1}$

Simple Closed-form Choice Probability

Let a display set $S = \{x_1, x_2, \dots, x_M\}$ be such that $x_1 < x_2 < \dots < x_M$. Under the RMJ-based ranking model

$$\mathbb{P}(x_i|S) = \frac{q^{i-1}}{1+q+\ldots+q^{M-1}}$$

Parameter Learning: Maximum Likelihood Estimation

Given historical data $H_T = (S_1, x_1, \dots, S_T, x_T)$. The MLE formulation is

$$\sum_{t=1}^{T} \log \left(\frac{1-q}{1-q^{|S_t|}} \right) + \log q \underbrace{\sum_{\substack{(i,j):i\neq j}} \mathbb{I} \left\{ j >_{\pi} i \right\} \cdot w_{ij}}_{\dagger} \tag{1}$$

- $w_{ij} := \sum_{t=1}^{T} \mathbb{I} \{ \{i, j\} \subseteq S_t \text{ and } x_t = i \}.$
- Intuitively, a positive wij is an indication that item i should be preferred to item j.
- in (1) has an integer programming **reformulation**, and it's a **well-studied** weighted feedback arc set problem on tournaments.

Given π and set $\alpha = -\log q$, MLE is a **concave** function of α .

Under some mild conditions, the estimator $(\hat{\pi}, \hat{q})$ is consistent.

$$k \geq 1$$

Ranked Choice Probability and Its Learning

Given a display set S and a π_k such that $R(\pi_k) \subseteq S$, we have

$$\mathbb{P}(\pi_{\mathsf{k}}|\mathsf{S}) = \mathsf{q}^{\mathsf{d}_{\mathsf{S}}(\pi_{\mathsf{k}}) + \mathsf{L}_{\mathsf{S}}(\pi_{\mathsf{k}})} \cdot rac{\psi(|\mathsf{S}| - \mathsf{k}, \mathsf{q})}{\psi(|\mathsf{S}|, \mathsf{q})},$$

where
$$d_S(\pi_k) := \sum_{i=1}^{k-1} \mathbb{I}\{\pi_k(i) > \pi_k(i+1)\} \cdot (|S|-i), L_S(\pi_k) := |\{x \in R^c(\pi_k) \cap S : x < \pi_k(k)\}|, \psi(n,q) := \prod_{i=1}^n \left(1+q+\dots+q^{(i-1)}\right).$$

Given historical data $H_T = (S_1, \pi_k^1, \dots, S_T, \pi_k^T)$, where $\pi_k^t = (x_1^t, \dots, x_k^t)$.

The MLE for the central ranking can be obtained from the same integer programming with a generalized definition of wij below

$$w_{ij} = \sum_{t=1}^{T} \left[\sum_{h=1}^{k-1} (|S_t| - h) \cdot \mathbb{I}\{x_h^t = i, x_{h+1}^t = j\} + \mathbb{I}\{x_k^t = i\} \cdot \mathbb{I}\{j \in S_t \setminus \{x_1^t, \dots, x_{k-1}^t\} \right]$$

Sampling

Probability Distribution of Top-k

$$\lambda(\pi_{\mathsf{k}}) = \mathsf{q}^{\mathsf{d}(\pi_{\mathsf{k}}) + \mathsf{L}(\pi_{\mathsf{k}})} \cdot rac{\psi(\mathsf{n} - \mathsf{k}, \mathsf{q})}{\psi(\mathsf{n}, \mathsf{q})},$$

$$\begin{split} \lambda(\pi_k) &= \mathsf{q}^{\mathsf{d}(\pi_k) + \mathsf{L}(\pi_k)} \cdot \frac{\psi(\mathsf{n} - \mathsf{k}, \mathsf{q})}{\psi(\mathsf{n}, \mathsf{q})}, \\ \text{where } \mathsf{d}(\pi_k) \; := \; \sum_{i=1}^{k-1} \mathbb{I}\left\{\pi_k(\mathsf{i}) > \pi_k(\mathsf{i}+1)\right\} \, \cdot \, (\mathsf{n} - \mathsf{i}), \\ \mathsf{L}(\pi_k) := |\{\mathsf{x} \in \mathsf{R}^c\left(\pi_k\right) : \mathsf{x} < \pi_k(\mathsf{k})\}|. \end{split}$$

Sampling of Next Position

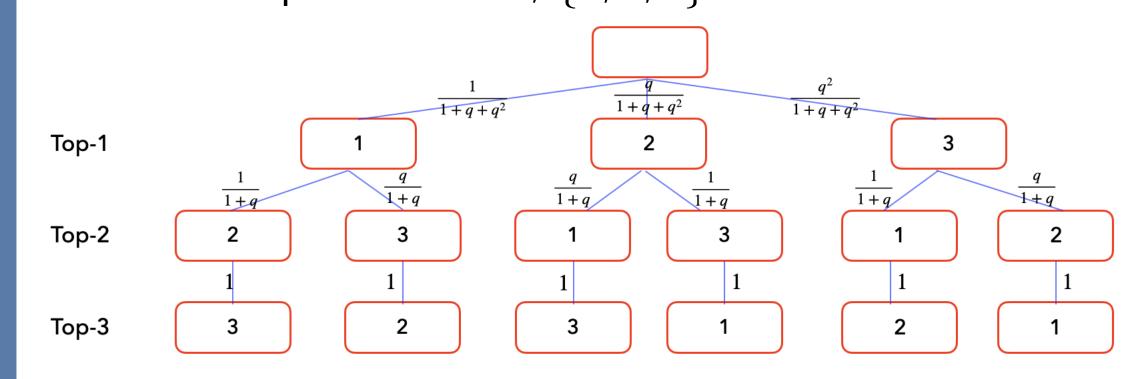
Given π_k such that $\pi_k(k) = z$, the conditional probability for the (k+1)-positioned item is

$$\mathbb{P}\left(\pi_{\mathsf{k}+1} = \pi_{\mathsf{k}} \oplus \mathsf{y} \mid \pi_{\mathsf{k}}\right) = \frac{\mathsf{q}^{\mathsf{h}(\mathsf{y}|\mathsf{z})-1}}{1 + \mathsf{q} + \cdots + \mathsf{q}^{\mathsf{n}-\mathsf{k}-1}},$$

$$h(y|z) = \begin{cases} \sum_{x \in R^c(\pi_k)} \mathbb{I}\{z < x \le y\} & \text{if } y > z, \\ n - k - \sum_{x \in R^c(\pi_k)} \mathbb{I}\{y < x < z\} & \text{if } y < z. \end{cases}$$

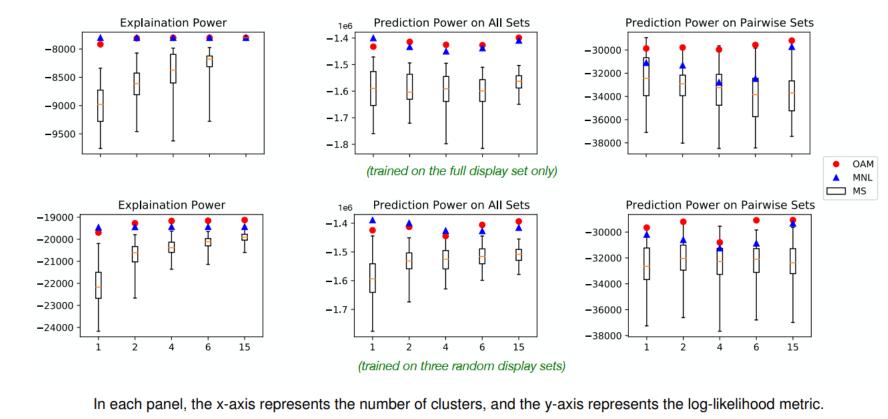
O(nk)

An example of 3 items, $\{1, 2, 3\}$.



Experiments

Real data. We test top-1 prediction power on Sushi preference and E-commerce data. Here are the performances:



We also conduct robustness check on top-k choice and test our estimation method when n is large.

Full Paper is Available at:



https://arxiv.org/abs/2207.01783

