- 1 Hoeffding's Inequality
- a. Use Chernoff bounds and Hoeffding's lemma to prove Hoeffding's inequality:

$$P_{r}(\frac{1}{n}\sum_{i=1}^{n}(X_{i}-\mu_{X_{i}}) \geq t) = P_{r}(\sum_{i=1}^{n}(X_{i}-\mu_{X_{i}}) \geq nt) = P_{r}(e^{\lambda\sum_{i=1}^{n}(X_{i}-\mu_{X_{i}})} \geq e^{\lambda nt})$$

$$\leq \min_{\lambda \geq 0} e^{-\lambda nt} \mathbb{E}[e^{\lambda\sum_{i=1}^{n}(X_{i}-\mu_{X_{i}})}] = \min_{\lambda \geq 0} e^{-\lambda nt} \prod_{i=1}^{n} \mathbb{E}[e^{\lambda(X_{i}-\mu_{X_{i}})}]$$

$$\leq \min_{\lambda \geq 0} e^{-\lambda nt} \prod_{i=1}^{n} e^{\frac{\lambda^{2}(b-a)^{2}}{8}} = \lim_{\lambda \geq 0} e^{-\lambda nt + \frac{n\lambda^{2}(b-a)^{2}}{8}}$$

Let  $g(\lambda) = -\lambda nt + \frac{n\lambda^2(b-a)^2}{8}$ . It's a quadratic function and achieves its minimum at

 $\frac{4t}{(b-a)^2}$ .

$$P_r(\frac{1}{n}\sum_{i=1}^n (X_i - \mu_{X_i}) \ge t) \le exp(-\frac{2nt^2}{(b-a)^2})$$

b. Give a simple distribution of  $X_i$  where the bound can be much sharper then Hoeffding's bound.

 $X_1, \ldots, X_n$  are i.i.d from a distribution with mean zero, bounded support [a, b], with variance  $\mathbb{E}[X^2] = \sigma^2$ . Then,

$$P_r(\frac{1}{n}\sum_{i=1}^n (X_i - \mu_{X_i}) \ge t) \le exp(-\frac{nt^2}{2(\sigma^2 + (b-a)t)}).$$

This equality is typically known as Bernstein's inequality. Since  $\sigma \leq b-a$ , when  $\sigma$  is small, this bound can be much sharper than Hoeffding's bound.

## 2 VC Dimension

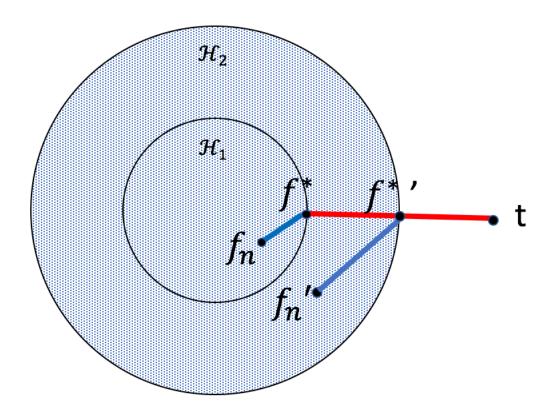
a. What is the VC-dimension of  $\mathcal{H}_1$  and  $\mathcal{H}_2$ .

The VC-dimension of  $\mathcal{H}_1$  is p. Because the VC-dimension of half-spaces in p dimensions is p+1.

The VC-dimension of  $\mathcal{H}_2$  is  $\binom{p+2}{2}$ . Because the feature space  $\phi(\mathbf{x})$  of the

second polynomial kernel is of degree 
$$\binom{p+2}{2} = \frac{(p+2)(p+1)}{2}$$
.

b. Draw a picture for the approximation and estimation error for  $\mathcal{H}_1$ ,  $\mathcal{H}_2$  and  $\hat{f}_1$ ,  $\hat{f}_2$  and write them down. Explain how the two errors change as n increases.



The blue lines are estimation error and the red lines are approximation error.

Approximation error for  $\mathcal{H}_1$  and  $\hat{f}_1$  is:

$$R^{true}(f^*) - R^*$$

Estimation error for  $\mathcal{H}_1$  and  $\hat{f}_1$  is:

$$R^{true}(f_n) - R^{true}(f^*) \le 2\sqrt{\frac{\log(N) + \log(\frac{2}{\delta})}{2n}}$$

Approximation error for  $\mathcal{H}_2$  and  $\hat{f}_2$  is:

$$R^{true}(f^{*\prime}) - R^{*}$$

Estimation error for  $\mathcal{H}_2$  and  $\hat{f}_2$  is:

$$R^{true}(f_n') - R^{true}(f^{*'}) \le 2\sqrt{\frac{log(N) + log(\frac{2}{\delta})}{2n}}$$

The estimation error decreases and the approximation error doesn't change as n increases.

- c. The VC-dimension of the set of sin functions with arbitrarily large or small frequency is infinite. But the number of parameters of sin(ax) is only one.
- 3 Ridge Regression
- a. Derive the closed form solution of  $\hat{eta}^{ridge}$

$$\begin{split} \mathbf{F}(\vec{\lambda}) &= \left\| \vec{y} - \bar{\bar{X}} \vec{\beta} \right\|_{2}^{2} + \lambda \left\| \vec{\beta} \right\|_{2}^{2} \\ \mathbf{F}(\vec{\lambda}) &= -2\bar{\bar{X}}^{\mathrm{T}} (\vec{y} - \bar{\bar{X}} \vec{\beta}) + 2\lambda \vec{\beta} = 2(-2\bar{\bar{X}}^{\mathrm{T}} \vec{y} + \bar{\bar{X}}^{\mathrm{T}} \bar{\bar{X}} \vec{\beta} + \lambda \vec{\beta}) = 0 \\ \bar{\bar{X}}^{\mathrm{T}} \vec{y} &= (\bar{\bar{X}}^{\mathrm{T}} \bar{\bar{X}} + \lambda \bar{\bar{\mathbf{I}}}) \vec{\beta}^{*} \\ \bar{\beta}^{*} &= (\bar{\bar{X}}^{\mathrm{T}} \bar{\bar{X}} + \lambda \bar{\bar{\mathbf{I}}})^{-1} \bar{\bar{X}}^{\mathrm{T}} \vec{y} \end{split}$$

b.

Since we know

$$\begin{pmatrix} \mathbf{0} \\ \mathbf{L}\boldsymbol{\gamma} \\ \mathbf{0} \end{pmatrix} = \begin{pmatrix} \mathbf{1}_n^T \\ \mathbf{U}_p^T \\ \mathbf{U}_{n-p-1}^T \end{pmatrix} \mathbf{U}_p \mathbf{L} \mathbf{V}^T \boldsymbol{\beta}$$

we have

$$L\gamma = U_p^T U_p L V^T \beta$$

So,

$$\boldsymbol{\gamma} = \boldsymbol{V}^T \boldsymbol{\beta}$$

In the original ridge regression, we have

$$\boldsymbol{\beta}^* = (\boldsymbol{X}^T \boldsymbol{X} + \lambda \boldsymbol{I})^{-1} \boldsymbol{X}^T \boldsymbol{Y}$$

If we plug this equation into the equation above, we have

$$\hat{\mathbf{y}} = \mathbf{V}^{T} (\mathbf{X}^{T} \mathbf{X} + \lambda \mathbf{I})^{-1} \mathbf{X}^{T} \mathbf{Y}$$

$$= \mathbf{V}^{T} \left( \left( \mathbf{U}_{p} \mathbf{L} \mathbf{V}^{T} \right)^{T} \mathbf{U}_{p} \mathbf{L} \mathbf{V}^{T} + \lambda \mathbf{I} \right)^{-1} \mathbf{X}^{T} \mathbf{Y}$$

$$= \mathbf{V}^{T} \left( \mathbf{V} \mathbf{L} \mathbf{U}_{p}^{T} \mathbf{U}_{p} \mathbf{L} \mathbf{V}^{T} + \lambda \mathbf{I} \right)^{-1} \mathbf{X}^{T} \mathbf{Y}$$

$$= \mathbf{V}^{T} (\mathbf{V} \mathbf{L}^{2} \mathbf{V}^{T} + \lambda \mathbf{I})^{-1} \left( \mathbf{U}_{p} \mathbf{L} \mathbf{V}^{T} \right)^{T} \mathbf{Y}$$

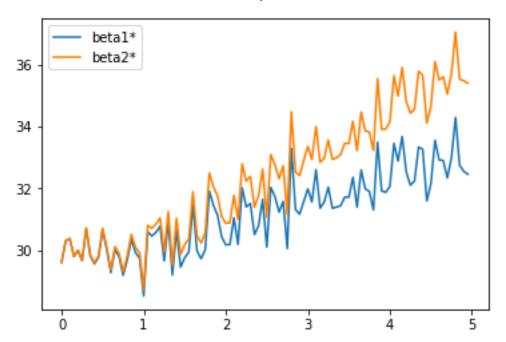
$$= \mathbf{V}^{T} (\mathbf{V} \mathbf{L}^{2} \mathbf{V}^{T} + \lambda \mathbf{V} \mathbf{V}^{T})^{-1} \mathbf{V} \mathbf{L} \mathbf{U}_{p}^{T} \mathbf{Y}$$

$$= \mathbf{V}^{T} (\mathbf{V} \mathbf{L}^{2} \mathbf{V}^{T} + \mathbf{V} \lambda \mathbf{V}^{T})^{-1} \mathbf{V} \mathbf{L} \mathbf{U}_{p}^{T} \mathbf{Y}$$

$$= \mathbf{V}^{T} \mathbf{V} (\mathbf{L}^{2} + \lambda \mathbf{I})^{-1} \mathbf{V}^{T} \mathbf{V} \mathbf{L} \mathbf{U}_{p}^{T} \mathbf{Y}$$

$$= (\mathbf{L}^{2} + \lambda \mathbf{I})^{-1} \mathbf{L} \mathbf{U}_{p}^{T} \mathbf{Y}$$

So, the closed form of  $\hat{\gamma}$  is  $(L^2 + \lambda I)^{-1}LU_p^T Y$ .



For large  $\lambda$ , penalty dominates the loss function. If  $\lambda$  is big, the sum of squares of the coefficients must be small. So, the MSE of both  $\beta_1$  and  $\beta_2$  increases as  $\lambda$  increases.

The sum of squares of  $\beta_1$  is 5.64 and the sum of squares of  $\beta_2$  is 10.72. So, ridge regression forces  $\beta_2$  to change more than  $\beta_2$ , and that's why  $\beta_2$  becomes more different from the original  ${\beta_2}^*$ . So, the difference between the MSE of  $\beta_2$  and  $\beta_1$ 

C.

