

MATH 526: Discrete State Stochastic Processes  
Lecture 20

**Chapter 5: Martingales**

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## 5.2: Martingales, Basics

Random Walk:  $(X_n)_{n=0}^{\infty}$  are i.i.d. r.v.'s, with mean  $\mu$ , and  $M_n = \sum_{i=0}^n X_i$ . Notice that  $M_n$  is completely determined by  $X_0, \dots, X_n$ .

### Definition (Adapttness)

*discrete time Stochastic Process*

We say that the sequence of r.v.'s  $(M_n)$  is **adapted** to the sequence of r.v.'s  $(X_n)$  if

*Model of randomness*  $M_n = f_n(X_0, \dots, X_n)$ , for all  $n$ ,

where each  $f_n$  is a **deterministic function**.

- ▶ Any sequence is always adapted to itself.
- ▶ Assume  $(M_n)$  is adapted to  $(X_n)$ .  $f_n = X_n$ 
  - ▶ The r.v.'s  $X_0, \dots, X_n$  play the role of **information available at time  $n$** : their values are observed at time  $n$  and they determine the value of  $M_n$ .
  - ▶  $X_0, \dots, X_n$  **contain equal, or greater, amount of information** than  $M_0, \dots, M_n$ .  *$M_n$  是一个函数, 减少信息量*
  - ▶  $M_0, \dots, M_n$  may contain **strictly less information** than  $X_0, \dots, X_n$ : when it is impossible to recover  $X_n$  from  $M_1, \dots, M_n$ . For example,  $M_n = X_n^2$ .
  - ▶ If  $X_n$  can be recovered from  $M_0, \dots, M_n$ , then,  $M_0, \dots, M_n$  contain **the same amount of information** as  $X_0, \dots, X_n$ . In the random walk example, we have:  $X_n = M_n - M_{n-1}$ .

## 5.2: Martingales, Basics

Let us go back to the example of random walk. Because the random walk  $(M_n)$  is adapted to  $(X_n)$ , we have:

$$\Rightarrow \mathbb{E}(M_n | X_0, \dots, X_n) = M_n.$$

And

$$\mathbb{E}(M_{n+1} | X_0, \dots, X_n) = M_n + \mu.$$

**Notation:** for brevity, we will denote the family of variables on which we condition,  $(X_0, \dots, X_n)$ , by  $\mathcal{F}_n^X$ .

Thus,

$$\mathbb{E}(M_{n+1} | \mathcal{F}_n^X) = M_n + \mu.$$

If  $\mu = 0$

$$\mathbb{E}(M_{n+1} | \mathcal{F}_n^X) = M_n,$$

and  $(M_n)$  becomes a **martingale**.

### Definition (Martingale)

A discrete time stochastic process  $(M_n)_{n=0}^{\infty}$  is a **martingale** with respect to  $(\mathcal{F}_n)_{n=0}^{\infty}$ , if

1.  $(M_n)_{n=0}^{\infty}$  is **adapted** to  $(\mathcal{F}_n)_{n=0}^{\infty}$ ,
2.  $\mathbb{E}|M_n| < \infty$ , for all  $n \geq 0$ ,
3. and  $\mathbb{E}(M_{n+1} | \mathcal{F}_n^X) = M_n$ , for all  $n \geq 0$ .

The condition  $\mathbb{E}|M_n| < \infty$  is needed to ensure that the conditional expectation is well defined.

Martingale property means that the conditional expectation of the next value of the process, given all currently available information, is simply the current value of the process. Statistically speaking, **the increments of a martingale process are pure noise: the best prediction for the next value of a martingale is simply its current value.**

### Remark

If it is not specified with respect to what the process is a martingale, then, either the available information is clear from the context, or the specific structure of available information is not important, or the process is a martingale with respect to itself:

$$\mathbb{E}(M_{n+1} \mid \mathcal{F}_n^M) = M_n.$$

Martingale should be understood as something that **doesn't change on average**. It is the stochastic analogue of a constant process. Analogously, one can introduce processes that increase or decrease on average.

### Definition (Submartingale)

$(M_n)$  is a **submartingale** with respect to  $(X_n)$ , if

1.  $(M_n)$  is adapted to  $(X_n)$ ,
2.  $\mathbb{E}|M_n| < \infty$ , for all  $n \geq 0$ ,
3. and  $\mathbb{E}(M_{n+1} \mid \mathcal{F}_n^X) \geq M_n$ , for all  $n \geq 0$ .

R.W w/ positive mean,

non-increasing

### Definition (Supermartingale)

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2.  $\mathbb{E}|M_n| < \infty$ , for all  $n \geq 0$ ,
3. and  $\mathbb{E}(M_{n+1} | \mathcal{F}_n^X) \leq M_n$ , for all  $n \geq 0$ .

In the random walk example,  $(M_n)$  is a submartingale if  $\mu \geq 0$ :

$$\mathbb{E}(M_{n+1} | \mathcal{F}_n^X) = M_n + \mu \geq M_n.$$

Similarly,  $(M_n)$  is a supermartingale if  $\mu \leq 0$ :

$$\mathbb{E}(M_{n+1} | \mathcal{F}_n^X) = M_n + \mu \leq M_n.$$

### Remark

Notice that martingale can be defined as a process which is a submartingale and a supermartingale at the same time.

## 5.2: Martingales, Basics

We can, sometimes, modify (*compensate*) a sub- or supermartingale, to make it a martingale.

**Example (Asymmetric simple random walk).** Assume that  $\{X_i\}$  are i.i.d. random variables, taking value 1 with probability  $p$ , and  $-1$  with probability  $q = 1 - p$ .

$$S_n = \sum_{i=1}^n X_i.$$

And

$$\mathbb{E}(S_{n+1} | \mathcal{F}_n^X) = S_n + \underbrace{p - q}_{\mu} \rightarrow \text{非 } mG.$$

Thus,  $(S_n)$  is a **submartingale** if  $p \geq q$ , and it is a **supermartingale** if  $p \leq q$ .

① Define  $M_n := S_n - (p - q)n$ . Then,  $M_n \sim \{X_1, \dots, X_n\}$ .

$$\begin{aligned} \mathbb{E}(M_{n+1} | \mathcal{F}_n^X) &= \mathbb{E}(S_{n+1} | \mathcal{F}_n^X) - (p - q)(n + 1) \\ &= S_n + (p - q) - (p - q)(n + 1) \\ &= S_n - (p - q)n = M_n. \end{aligned}$$

Hence,  $(M_n)$  is a **martingale**.

## 5.2: Martingales, Basics

**Martingales do not necessarily have to arise as sums of independent r.v.'s (random walks).** In particular, there is another way to construct a martingale from  $(S_n)$ :

②

$$M_n := (q/p)^{S_n}.$$

Let us check that  $(M_n)$  is a martingale with respect to  $(X_n)$ :

► Adaptedness with respect to  $(X_n)$  is immediate.

►

$$\mathbb{E}|M_n| \text{ ☀ } < \max \left\{ \left( \frac{q}{p} \right)^{X_0+n}, \left( \frac{p}{q} \right)^{X_0+n} \right\} < \infty.$$

►

$$\begin{aligned} \mathbb{E}(M_{n+1} | \mathcal{F}_n^X) &= \text{☀ } M_n \cdot \mathbb{E} \left[ \left( \frac{q}{p} \right)^{X_{n+1}} \right] \\ &= M_n \cdot \left( \frac{q}{p} \cdot p + \frac{p}{q} \cdot q \right) \\ &= M_n \cdot 1 \quad c \cdot p + \frac{1}{c} q = 1 \\ &= M_n. \end{aligned}$$

In the above, we used the following **property of conditional expectation**: if  $X$  is a deterministic function of  $(Y_1, \dots, Y_n)$ , then

$$\mathbb{E}(XZ | Y_1, \dots, Y_n) = X \mathbb{E}(Z | Y_1, \dots, Y_n),$$

for any r.v.  $Z$ , for which the conditional expectation is well defined.

$$\begin{array}{c} 1 \\ p \end{array} \times \begin{array}{c} -1 \\ p-1 \end{array}$$



## 5.2: Martingales, Basics

### Basic properties of martingales.

Assume that  $(M_n)$  is a **martingale** with respect to  $(X_n)$ .

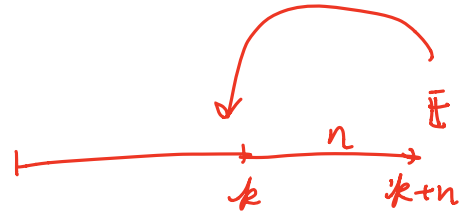
- ▶  $\mathbb{E}(M_{k+n} | \mathcal{F}_k^X) = M_k$ , for all  $n, k \geq 0$ .

Proof. ☀  $\mathbb{E}(M_{k+2} | \mathcal{F}_k^X)$

$$\begin{aligned} \text{Tower prop} &= \mathbb{E} \left( \mathbb{E}(M_{k+2} | \mathcal{F}_{k+1}^X) | \mathcal{F}_k^X \right) \\ &= \mathbb{E}(M_{k+1} | \mathcal{F}_k^X) = M_k \end{aligned}$$

- ▶  $\mathbb{E}M_n = \mathbb{E}M_0$ , for all  $n \geq 0$ .

Proof. ☀  $\mathbb{E}(\mathbb{E}(M_n | \mathcal{F}_0^X)) = \mathbb{E}(M_0)$  <sup>Prop 1</sup>  
 $\mathbb{E}(M_n) =$



Assume  $(M_n)$  is a **submartingale** (**supermartingale**) with respect to  $(X_n)$ :

- ▶  $\mathbb{E}[M_{k+n} | \mathcal{F}_k^X] \geq (\leq) M_k$ , for all  $n, k \geq 0$ .
- ▶  $\mathbb{E}[M_n] \geq (\leq) \mathbb{E}M_0$ , for all  $n \geq 0$ .

## 5.2: Martingales, Basics

Functions of stochastic processes as martingales (submartingales, supermartingales).

Let us make a connection between martingales and Markov chains.

### Question

How to determine if a given function of a Markov chain is a martingale?

### Theorem

Let  $(X_n)_{n=0}^{\infty}$  be a **Markov chain**, with state space  $\mathbb{S}$  and transition matrix  $p$ , and assume that function  $f$  satisfies:  $\mathbb{E}|M_n| < \infty$   $\mathbb{E}|M_n| = \sum$

- ▶  $\sum_{y \in \mathbb{S}} p^n(x, y) |f(y, n)| < \infty$ , for all  $x \in \mathbb{S}$  and  $n \geq 0$ ;  $\rightarrow \mathbb{E}|M_n|$
- ▶  $\sum_y p(x, y) f(y, n+1) = f(x, n)$ , for all  $x \in \mathbb{S}$  and  $n \geq 0$ .  $\rightarrow \mathbb{E}(M_{n+1} | \mathcal{F}_n^X)$ .

Then,  $(M_n = f(X_n, n))$  is a **martingale** with respect to  $(X_n)$ .

### Remark

If  $f(x)$  is independent of  $n$ , then the second condition in the theorem becomes:  $\sum_y p(x, y) f(y) = f(x)$ . This equation is something we have seen before, when computing the exit probabilities. This is not a coincidence: we will see that, in fact, the martingale theory allows for a much simpler computation of the exit probabilities.

## 5.2: Martingales, Basics

Proof. ☀

$$\mathbb{E}|M_n| = \sum_{y \in S} p^n(x, y) \cdot |f(y, n)| < \infty \quad \text{for all } x \in S$$

$$\mathbb{E}(M_{n+1} \mid \mathcal{F}_n^x) = \sum p(x, y) \cdot f(y, n+1) = f(x, n)$$

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HAS BEEN GIVEN

### Remark

In fact, we have shown that, for a Markov chain  $(X_n)$ , with  $\mathbb{E}|X_n| < \infty$ , **the following useful (and intuitively obvious) identity holds:**


$$\mathbb{E}\left[f(X_{n+1}, n+1) \mid \mathcal{F}_n^X\right] = \sum_{y \in \mathbb{S}} f(y, n+1) p(X_n, y).$$

## 5.2: Martingales, Basics

**Example (Branching process).** Denote by  $X_n$  the number of members in the  $n$ -th generation of a population. Each member gives birth to a random number of children. These random numbers  $\{Y_i\}$  are independent across members and across generations, and they are identically distributed with mean  $\mu$ .

$(X_n)$  is a MC with the state space  $\mathbb{S} = \{0, 1, 2, \dots\}$  and a transition matrix  $p$ .

Notice that  $\mathbb{E}(X_{n+1} | X_n = x) = \sum_{y \in \mathbb{S}} p(x, y)y = \mathbb{E} \sum_{i=1}^x Y_i = \boxed{\mu x}$ .

Introduce  $Z_n = f(X_n, n) = X_n / \mu^n$ . We show that  $(Z_n)$  is a martingale:  *a fn of MC that is MG*

①  $Z_n$  adapted to  $\mathcal{F}_n^X$ . ✓

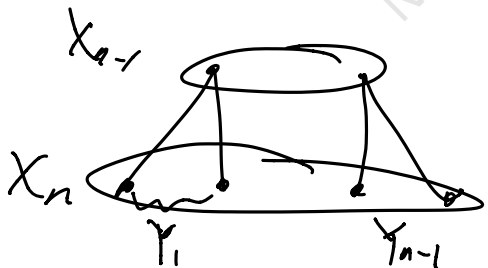
$$\textcircled{2} \mathbb{E}(|Z_n|) = \mathbb{E}(X_n / \mu^n) = \frac{1}{\mu^n} \mathbb{E}(X_n)$$

$$= \frac{1}{\mu^n} \mathbb{E}(\mathbb{E}(X_n | X_{n-1}))$$

$$= \frac{1}{\mu^n} \cdot \mu \cdot \mathbb{E}(X_{n-1})$$

$$= \frac{\mathbb{E}(X_{n-1})}{\mu^{n-1}} = \mathbb{E}(|Z_{n-1}|)$$

$$= \mathbb{E}(|Z_0|)$$



## 5.2: Martingales, Basics

$$\textcircled{2} \mathbb{E}|Z_n| = \sum_{y=0}^{\infty} p^n(x, y) |f(y, n)|$$

$$= \sum_{y=0}^{\infty} p^n(x, y) \frac{|y|}{\mu^n}$$

$$= \frac{1}{\mu^n} \sum_{y=0}^{\infty} \sum_{z=0}^{\infty} p^{n-1}(x, z) p(z, y) |y|$$

$$= \frac{1}{\mu^n} \sum_{z=0}^{\infty} p^{n-1}(x, z) \sum_{y=0}^{\infty} p(z, y) |y|$$

$$\mathbb{E}(X_1 | X_0 = z)$$

$$= \mu \cdot |z|$$

$$= \sum_{z=0}^{\infty} p^{n-1}(x, z) \frac{|z|}{\mu^{n-1}} = \dots < \infty$$

$$\textcircled{3} \mathbb{E}(Z_{n+1} | \mathcal{F}_n^X) = \frac{1}{\mu^{n+1}} \cdot \mu X_n = \frac{X_n}{\mu^n} = Z_n$$

### Theorem

- ▶ If  $(M_n)$  is a martingale and  $f$  is a convex function, such that  $\mathbb{E}|f(M_n)| < \infty$ , then,  $(f(M_n))$  is a submartingale. In particular, if  $(M_n)$  is a martingale, and  $\mathbb{E}M_n^2 < \infty$ , then,  $(M_n^2)$  is a submartingale.
- ▶ If  $(M_n)$  is a submartingale and  $f$  is a nondecreasing convex function, such that  $\mathbb{E}|f(M_n)| < \infty$ , then,  $(f(M_n))$  is a submartingale.

**Proof.** Using Jensen's inequality, we obtain:

$$\mathbb{E}(f(M_{n+1}) | \mathcal{F}_n) \geq f(\mathbb{E}(M_{n+1} | \mathcal{F}_n)) \geq f(M_n).$$

↓

Janson

$$p_1 f(x_1) + p_2 f(x_2) \geq f(p_1 x_1 + p_2 x_2)$$

$$mG + \text{convex} \Rightarrow \text{sub } mG.$$

## 5.2: Martingales, Basics

The increments of a martingale are orthogonal.

### Theorem

If  $(M_n)$  is a martingale, such that  $\mathbb{E}M_n^2 < \infty$ , for all  $n \geq 0$ , then, for any time indices  $0 \leq i \leq j \leq k \leq n$ , we have:

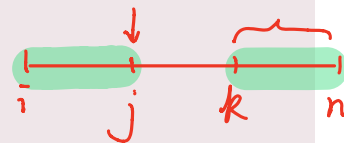
$$\mathbb{E}((M_n - M_k)M_j) = 0,$$

$$\mathbb{E}((M_n - M_k)(M_j - M_i)) = 0,$$

provided the above expectations are well defined.

$$\mathbb{E}XY \leq \sqrt{\mathbb{E}X^2 \cdot \mathbb{E}Y^2}$$

$\Rightarrow \mathbb{E}XY$  exists.



Proof.

$$\mathbb{E}((M_n - M_k)M_j) = \mathbb{E}(\mathbb{E}((M_n - M_k)M_j | \mathcal{F}_k)) = \mathbb{E}(M_j \mathbb{E}(M_n - M_k | \mathcal{F}_k)) = 0.$$

$$\mathbb{E}(M_n - M_k)$$

$$= \mathbb{E}(\mathbb{E}(M_n - M_k | \mathcal{F}_k))$$

$$= \mathbb{E}(\mathbb{E}(M_n | \mathcal{F}_k) - M_k)$$

$$(\mathbb{E}(M_n | \mathcal{F}_k) - M_k) = 0$$



## 5.2: Martingales, Basics

We also have this useful result.

### Theorem

If  $(M_n)$  is a martingale, such that  $\mathbb{E}M_n^2 < \infty$ , for all  $n \geq 0$ , then  $(Y_n)$  is a martingale, where

$$M_{n+1}^2 = (M_{n+1} - M_n + M_n)^2 = (M_{n+1} - M_n)^2 + M_n^2 + 2(M_{n+1} - M_n) \cdot M_n$$
$$Y_n = M_n^2 - \sum_{k=1}^n \mathbb{E}((M_k - M_{k-1})^2 | \mathcal{F}_{k-1})$$

$\downarrow$   
0

convex Jensen

The process  $(Y_n)$  is called the **compensated square of a martingale**.

$$S_n^2 - \sum_{k=1}^n \mathbb{E}((S_k - S_{k-1})^2 | \mathcal{F}_{k-1}) = S_n^2 - \sum_{k=1}^n \mathbb{E}(X_k^2) = S_n^2 - \sigma^2 \cdot n$$

**Example.** In particular, for a sequence of i.i.d. r.v.'s  $(X_n)$ , with  $\mathbb{E}X_n = 0$  and  $\mathbb{E}X_n^2 = \sigma^2$ , we have:  $S_n = \sum_{i=1}^n X_i$  and  $S_n^2 - \sigma^2 n$  are both martingales.

Suggested Exercises (Durrett, 3rd ed.), not for submission

Exercises 5.1, 5.2.

**5.1. Brother-sister mating.** Consider the six state chain defined in Exercise 1.66. Show that the total number of A's is a martingale and use this to compute the probability of getting absorbed into the 2,2 (i.e., all A's state) starting from each initial state.

**5.2. Lognormal stock prices.** Consider the special case of Example 5.4 in which  $X_i = e^{\eta_i}$  where  $\eta_i \sim \text{normal}(\mu, \sigma^2)$ . For what values of  $\mu$  and  $\sigma$  is  $M_n = M_0 \cdot X_1 \cdots X_n$  a martingale?

$$\begin{aligned} \mathbb{E}(M_{n+1} | \mathcal{F}_n^X) &= \mathbb{E}(M_0 \cdots X_{n+1} | \mathcal{F}_n) \\ &= M_0 \cdot X_1 \cdots X_n \cdot \mathbb{E}(X_{n+1}) \\ &= \mu \cdot M_n \Rightarrow \mu = 1 \end{aligned}$$

$$\mathbb{E} |M_n| < \infty$$

$$M_0 \cdot \underbrace{\mathbb{E}(X_i)^n}_{=1}$$