

MATH 526: Discrete State Stochastic Processes  
Lecture 23

**Chapter 5: Martingales**

Prakash Chakraborty

University of Michigan

## 5.5: Convergence

Martingales, submartingales, and supermartingales have a tendency to converge. Why?

From analysis: a decreasing sequence, bounded from below, always has a limit. Supermartingales can be viewed as the stochastic analogue of decreasing sequences (note that a supermartingale, itself, **does NOT have to be decreasing in time**, but its expectation is). Hence, it is natural to expect that, if a supermartingale is bounded from below, it has a limit.

### Theorem (Martingale Convergence)

Let  $X_n \geq C$  be a **supermartingale**. Then, there exists a random variable  $X_\infty$ , satisfying:

$$X_\infty = \lim_{n \rightarrow \infty} X_n, \quad \text{with probability one,} \quad \text{convergence.}$$

and

$$\mathbb{E}X_\infty \leq \mathbb{E}X_0$$

$$\mathbb{E}(M_{n+1} \mid \mathcal{F}_n^X) \leq M_n$$

### Theorem (Martingale Convergence 2)

Let  $X_n$  be a *supermartingale*, such that  $\mathbb{E}|X_n| \leq C$ . Then, there exists a random variable  $X_\infty$ , satisfying:

$$X_\infty = \lim_{n \rightarrow \infty} X_n, \quad \text{with probability one,}$$

and

$$\mathbb{E}X_\infty \leq \mathbb{E}X_0$$

### Remark

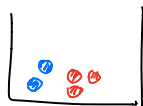
- ▶ any *martingale* is a *supermartingale*, hence, the *above theorems hold for martingales* as well;
- ▶ if  *$X_n$  is a submartingale*, then  *$-X_n$  is a supermartingale*, hence, we can reformulate the above convergence theorems for submartingales:

## 5.5: Convergence

The martingale convergence theorems, often, make it easy to prove that a given stochastic process has a limit. However, it tells us nothing about the distribution of the limiting random variable (compare to the case of Markov processes, where the limiting distribution is given by the stationary distribution of the process, and it can be computed via its generator).

For martingales, we, typically, first, apply the convergence theorem to conclude that the limit exists, and, then find the limit.

### Example (Polya's urn)



Initially  $K=2$   
 $X_n =$  fraction of red balls  
 at time  $n$ .  
 $P(X_1 = \frac{2}{3}) = \frac{1}{2}$

- $X_n$  adapts to itself.  $|X_n| < \infty$

$$\mathbb{E}(X_{n+1} | \mathcal{F}_n^X) = X_n$$

$$X_n \cdot \frac{(k+n)X_n + 1}{k+n+1} + (1-X_n) \cdot \frac{(k+n) \cdot X_n}{k+n+1} = X_n$$

- Initially  $K = r + b$   
 $\sim \text{Beta}(r, b)$

Apply the Thm:

$$X_n \rightarrow X_\infty$$

$$P(X_\infty = \frac{j}{n+2})$$

$$= \binom{n}{j} \frac{1}{2} \times \dots \times \frac{j-1}{j} \times \frac{1}{j+1} \times \dots \times \frac{n-j+1}{n+1} = \frac{1}{n+1}$$

$$x^{r-1} (1-x)^{b-1}$$

$$\text{Beta}(r, b) \rightarrow \frac{\Gamma(r+b)}{\Gamma(r)\Gamma(b)}$$

# Applications

**Example (Asymmetric simple random walk)** Let  $\{X_n\}$  be i.i.d. random variables, taking values 1 and  $-1$ , with probabilities  $p$  and  $q = 1 - p$ , respectively. Introduce

$$S_0 = 0, \quad S_n = \sum_{i=1}^n X_i$$

$\rightarrow 0$

$p \neq q$  positive  
 $\Rightarrow M_n \geq 0$

Assume that  $p \neq q$ . Then  $(M_n = (q/p)^{S_n})$  is a martingale.

Since  $M_n \geq 0$ , we can apply the first Martingale Convergence theorem, to conclude that there is a random variable  $M_\infty$ , such that:

$$(q/p)^{S_n} \rightarrow M_\infty,$$

$$\begin{cases} \frac{q}{p} < 1. & S_n \rightarrow +\infty \\ \frac{q}{p} > 1. & S_n \rightarrow -\infty \end{cases} \quad \text{Intuition}$$

as  $n \rightarrow \infty$ .

Let us show that  $M_\infty = 0$ , with probability one. *WTS. 如果  $M_\infty \neq 0$ , 那么*

Pick any random outcome  $\omega$ , such that  $M_\infty(\omega) \neq 0$ . Let us show that, for such  $\omega$ ,

$$\lim_{n \rightarrow \infty} M_n(\omega) \neq M_\infty(\omega),$$

$\omega \in A \quad \lim_{n \rightarrow \infty} M_n(\omega) \neq M_\infty(\omega).$   
那么  $P(A) = 0$

and, hence, the probability of all such  $\omega$  is zero. We will prove it by contradiction.

# Applications

$$M_n = \left(\frac{q}{p}\right)^{S_n} \geq 0$$

Assume that  $\lim_{n \rightarrow \infty} M_n(\omega) = M_\infty(\omega)$ ,

Since  $M_\infty(\omega)$  is a limit of nonnegative numbers, it has to be nonnegative itself.

Since, in addition, it is not equal to zero, we have:  $M_\infty(\omega) > 0$ . This means that  $M_n(\omega)$  has a positive limit, and, therefore,

Contradiction:  $M_\infty(\omega) > 0$

$$\log(M_n(\omega))$$

has a finite limit, as  $n \rightarrow \infty$ .

Then, as a difference of two sequences with the same limit,

如果  $M_\infty(\omega) > 0$

$$\Rightarrow \lim_{n \rightarrow \infty} \log(M_n(\omega)) < \infty$$

$$|S_{n+1}(\omega) \log(q/p) - S_n(\omega) \log(q/p)| \rightarrow 0,$$

as  $n \rightarrow \infty$ .

On the other hand, we notice that

$$|S_{n+1}(\omega) - S_n(\omega)| = 1$$

$$|S_{n+1}(\omega) \log(q/p) - S_n(\omega) \log(q/p)| = \log(q/p) = \text{const} \neq 0,$$

which is a contradiction.

Hence, when  $M_\infty(\omega) \neq 0$ ,  $M_\infty(\omega)$  cannot coincide with the limit of  $M_n(\omega)$ , and, due to the Martingale Convergence theorem, the set of all such random outcomes  $\omega$  has probability zero. In other words,  $M_\infty = 0$ , with probability one.

MCT  
 $\Rightarrow \mathbb{P} = 0$

The previous example provides an alternative proof of the fact that a simple random walk with  $p < q$ , hits any lower level, and, similarly, a simple random walk with  $p > q$  hits any upper level, with probability one. Indeed, if  $p < q$ , then,  $q/p > 1$  and

$$(q/p)^{S_n} \rightarrow 0$$

is equivalent to

$$S_n \rightarrow -\infty$$

Similar argument applies in the case  $p > q$ .

## Remark

The convergence of  $S_n$  to  $+\infty$  or  $-\infty$ , in turn, implies that

$$\mathbb{P}_x(\tau < \infty) = 1,$$

where  $\tau = V_a \wedge V_b$ , with some integers  $a \leq x \leq b$ . Recall that the fact that  $\tau$  is finite is needed for the computation of exit probabilities and expected exit times, when we apply the Optional Sampling Theorem. We used to argue for the finiteness of  $\tau$  using the results of Markov chain theory. As the above examples shows, the same conclusion can be obtained by using martingale convergence instead.

*recurrence.  
transient*

**Example** Consider a sequence  $\{X_i\}_{i=1}^{\infty}$  of non-constant independent identically distributed random variables with zero mean, such that, with probability one,  $|X_i| \leq C$  for any  $i \geq 1$ , with a fixed constant  $C$ . Consider

$$S_n = \sum_{i=1}^n X_i, \quad S_0 = 0, \quad \rightarrow m.G.$$

and the stopping time

$$T_a = \min \{n \geq 0 : S_n \geq a\}, \quad \text{for } a > 0.$$

$$\mu=0, \quad |X_i| \leq C$$

$$\Rightarrow T_a < \infty \quad w/p=1$$

Q: Let us show that  $\mathbb{P}(T_a < \infty) = 1$ .

Due to Optional Sampling Theorem,  $(S_{n \wedge T_a})$  is a martingale. In addition,  $S_{n \wedge T_a} \leq a + C$ . Therefore, due to the Martingale Convergence Theorem, with probability one,

$$\lim_{n \rightarrow \infty} S_{n \wedge T_a} = S_{\infty},$$

for some random variable  $S_{\infty}$ .



$T_a = \infty \Rightarrow S_n(\omega) \nexists \text{ converge}$

Let us show that the convergence is only possible if  $T_a < \infty$ .  $\Leftrightarrow$  converge

Pick  $\omega$  such that  $T_a(\omega) = \infty$ . Then,  $S_{n \wedge T_a}(\omega) = S_n(\omega)$ , and, hence, the convergence is possible only if  $S_n(\omega)$  converges to a finite limit, as  $n \rightarrow \infty$ .

Assume that there is a positive probability that  $T_a = \infty$ . Then, there is a positive probability that  $(S_n)$  converges to a finite limit as  $n \rightarrow \infty$ . Therefore, for  $\varepsilon > 0$ , there is a positive probability that at least for one  $n$ , the event

$$A_n = \{|X_n| \leq \varepsilon, |X_{n+1}| \leq \varepsilon, \dots\}$$

$$\lim_{n \rightarrow \infty} |S_{n+1} - S_n| = 0$$

occurs. In other words

$$\mathbb{P}\left(\bigcup_{n=1}^{\infty} A_n\right) > 0$$

$$\Rightarrow X_n = 0 \cdot \forall n > N.$$

Let us choose  $\varepsilon > 0$  so that  $\mathbb{P}(|X_n| \leq \varepsilon) = \delta < 1$ . Then,  $\mathbb{P}(A_n) = \delta^\infty = 0$ , and, since there are countably many  $A_n$ 's, the probability of their union is also zero. Thus, we obtain a contradiction. Therefore,  $\mathbb{P}(T_a = \infty) = 0$ .

Countable Intersection of  $P=1$  is  $P=1$

# Applications

**Example.** Suppose a sequence of independent events  $\{A_n\}_{n=1}^{\infty}$  satisfies

$$\sum_{i=1}^n \mathbb{P}(A_i) \rightarrow \infty, \quad \text{as } n \rightarrow \infty.$$

Notice that this fact implies

$$\prod_{i=1}^n (1 - \mathbb{P}(A_i)/2) \rightarrow 0, \quad \text{as } n \rightarrow \infty.$$

$|A_i| > 0$

Consider the random variables  $X_i = \mathbf{1}_{A_i}$  (each  $X_i$  takes value 1, if the event  $A_i$  occurs, and zero otherwise).

**Q:** Let us show that, with probability one,  $Y_n = \prod_{i=1}^n (1 + X_i) \rightarrow \infty$ , as  $n \rightarrow \infty$ . Define  $X_i = \mathbf{1}_{A_i}$ , for  $i = 1, 2, \dots$ , and

$$M_n = \prod_{i=1}^n \frac{1}{(1 - \mathbb{P}(A_i)/2)(1 + X_i)}, \quad M_0 = 1.$$

$$M_{n+1} = M_n \cdot \frac{1}{(1 - \mathbb{P}(A_{n+1})/2)(1 + X_{n+1})}$$

$$(1 - \mathbb{P}(A_i))$$

Notice that  $(M_n)$  is a martingale:

$M_n$  positive

$$\mathbb{E}|M_n| = \prod_{i=1}^n \frac{\mathbb{E}(1/(1 + X_i))}{1 - \mathbb{P}(A_i)/2} = \prod_{i=1}^n \frac{\mathbb{P}(A_i)/2 + 1 - \mathbb{P}(A_i)}{1 - \mathbb{P}(A_i)/2} = 1.$$

same = 1

$$\mathbb{E}(M_{n+1} | \mathcal{F}_n^X) = \left( \prod_{i=1}^n \frac{1}{(1 - \mathbb{P}(A_i)/2)(1 + X_i)} \right) \frac{\mathbb{E}(1/(1 + X_{n+1}))}{(1 - \mathbb{P}(A_{n+1})/2)} = M_n.$$

Since  $M_n \geq 0$ , we can apply the Martingale Convergence Theorem to conclude that, with probability one,  $(M_n)$  converges to a finite limit  $M_\infty$ , as  $n \rightarrow \infty$ .

Recall that

$$\prod_{i=1}^n (1 - \mathbb{P}(A_i)/2) \rightarrow 0, \quad \text{as } n \rightarrow \infty.$$

Therefore,

$$\prod_{i=1}^n (1 + X_i) = \frac{1}{M_n \prod_{i=1}^n (1 - \mathbb{P}(A_i)/2)} \rightarrow \frac{1}{M_\infty \cdot 0} = \infty,$$

as  $n \rightarrow \infty$ .

## Remark

This example proves the Borel-Cantelli lemma, which states that, under the conditions of this exercise, with probability one, an infinite number of events  $A_i$  occur.

说明有无限个  $X_i = 1$

$$X \rightarrow X+1 \quad P_X$$

$$X \rightarrow X \quad 1 - P_X - q_X$$

$$X \rightarrow X-1 \quad q_X$$

$$\frac{P_1}{q_1} + \frac{P_1}{q_1} \cdot \frac{P_2}{q_2} + \dots + \frac{P_1 \dots P_x}{q_1 \dots q_x}$$

$$\phi(x) = \sum_{y=1}^x \prod_{z=1}^{y-1} \frac{P_z q_z}{q_z P_z} \quad ?$$

$$\textcircled{1} M_n = \phi(X_n \wedge V_0) \text{ is } mG.$$

$$M_n = \frac{P_1}{q_1} + \dots + \frac{P_1 \dots P_{x-1}}{q_1 \dots q_{x-1}} \cdot \frac{P_x}{q_x}$$

$$\textcircled{1} \cdot \phi(X_n \wedge V_0) \text{ adapted to itself}$$

$$\cdot \mathbb{E} |M_n| < \infty$$

$$\cdot \mathbb{E}(M_{n+1} | \mathcal{F}_n^X) =$$

$$\frac{M_n - M_{n-1}}{M_{n-1} - M_n}$$

$$\textcircled{2} P_x(V_a < V_b) \quad 1 - P_X - q_X$$

$$= \frac{\phi(b) - \phi(x)}{\phi(b) - \phi(a)}$$

$$M_{n-1}$$

$$\frac{P_x^2}{q_x} + (1 - q_x)$$

$$\frac{q_x}{P_x} \cdot P_x$$

$$+ \cancel{\frac{q_x}{P_x}} (-q_x)$$