Introduction to statistical network analysis

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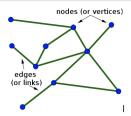
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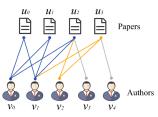
- Nodes (Vertices): participants in the network Network size: number of nodes
- Edges (Links): connections or relationships between nodes
- Examples [nodes (edges)]:
 - people (friendship)
 - webpages (hyperlinks)
 - paper (citation)
 - genes (regulatory actions)
 - brain regions (oxygen level correlation)
- Adjacency matrix: $A_{i,j}$ is the edge between nodes (i,j)

Classification of network data: (1) partite

• Unipartite network: any node may connect to anyone else



 Bipartite network: nodes partitioned into two groups, only between-group edges are possible



Classification of network data: (2) edge direction

- Undirected (Symmetric) network: $A_{i,j} = A_{j,i}$
- Directed (Asymmetric) network: $A_{i,j}$ and $A_{j,i}$ might not equal



Figure: Left: undirected network; right: directed network

Classification of network data: (2) edge weight

- Binary: $A_{i,j} \in \{0,1\}$ Example: friend (1) or non-friend (0)
- Signed: $A_{i,j} \in \{0,1,-1\}$ **Example:** friend (1) or foe (-1) or no interaction (0)
- Weighted:
 Example: trade surplus/deficit between countries

Research topics

Statistical research aims:

- · modeling network formation
- finding roles of individual nodes
 Examples: community detection, node embedding
- stochastic behavior of network features
 Examples: network moments
- link prediction

Research topics

Network data vs conventional data:

- Networks: no individual observation, only relational data
- Deriving "network analogues of classical techniques"
 Examples:
 - one-/two-sample test
 - cross-validation
 - method-of-moments
 - goodness-of-fit test
 - re/subsampling

Popular network models

Overview

Now we survey Network models: Erdos-Renyi model

- Simple → complex
- For simplicity, focus on undirected, binary networks
- Model probability matrix:

$$P_{i,j} := \mathbb{P}(A_{i,j} = 1)$$

Network models: Erdos-Renyi

Erdos-Renyi (ER) model (Newman 2018)

Model formula:

$$P_{i,j} := \mathbb{P}(A_{i,j} = 1) \equiv p$$

 Definition: network density is the (order of) average edge probability:

$$\rho_n \asymp \bar{P} := \binom{n}{2}^{-1} \sum_{1 \le i < j \le n} P_{i,j}$$

In ER model, $\rho_n \simeq p$

 Density is usually a crucial measure for assessing the problem difficulty of estimation/testing

Erdos-Renyi model

$$P_{i,j} := \mathbb{P}(A_{i,j} = 1) \equiv p \ltimes \rho_n$$

Model estimation:

Moment estimator:

$$\widehat{p} := \binom{n}{2}^{-1} \sum_{1 \le i < j \le n} A_{i,j}$$

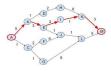
• Consistency (Chen et al. 2021): When $\binom{n}{2}p \to \infty$, $\widehat{p}/p \overset{p}{\to} 1$ Roughly speaking

$$\frac{|\widehat{p}-p|}{p} \overset{\text{approx.}}{\asymp} \frac{1}{\sqrt{n^2p}}$$

Erdos-Renyi model

ER model:

- Too simple to fit real-world data well
- But a good null model ("no structure/pattern") for testing (Gao & Lafferty 2017a)
- Neat for studying some network properties
 - Example: average shortest path (Katzav et al. 2018)



• Example: is the network connected? (Erdos & Rényi 1960)

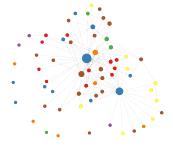


Answers to these questions are the same for: (i) ER model; and (ii) general model with all $P_{i,j} \approx \rho_n$

Definition: degree is the total number of edges of a node

$$d_i := \sum_{j:j \neq i} A_{i,j}$$

- Average degree: $\bar{d}:=\frac{1}{n}\sum_{i=1}^n d_i \asymp \rho_n n$
- Degree describes the node's popularity (hub, leaf)



Degree-driven network model:

- Network structure entirely determined by degrees
- Network → degrees: easy to compute
- Degree sequence $\stackrel{?}{\rightarrow}$ network model

Configuration model (Chung & Lu 2002):

- Not a probabilistic model for "network population"
- Given a degree sequence d_1, \ldots, d_n , generate network
 - **1** Set $S := \{1, \dots, 1, 2, \dots, 2, \dots, n, \dots, n\}$, each i repeats d_i times
 - Randomly select two entries from S, make an edge, delete them from S
 - Repeat step 2 until S is exhausted

β -model (Chatterjee et al. 2011):

• Formula:

$$P_{i,j} = \frac{e^{\beta_i + \beta_j}}{1 + e^{\beta_i + \beta_j}}$$

Parameters $\beta_i \in \mathbb{R}$

• Estimation: MLE, negative log-likelihood:

$$L(\beta) := \sum_{1 \le i < j \le n} \log \left(1 + e^{\beta_i + \beta_j} \right) - \sum_{i=1}^n \beta_i d_i$$

Then

$$\widehat{\beta} := \arg\min_{\beta} L(\beta)$$

• Consistency: if true β_i 's are O(1), then $\|\widehat{\beta} - \beta\|_{\infty} \stackrel{p}{\to} 0$

β -model:

$$P_{i,j} = \frac{e^{\beta_i + \beta_j}}{1 + e^{\beta_i + \beta_j}}$$

$$L(\beta) := \sum_{1 \le i < j \le n} \log \left(1 + e^{\beta_i + \beta_j} \right) - \sum_{i=1}^n \beta_i d_i$$

Pros and cons:

- (+) Scalable to very large and sparse networks (Shao et al. 2023)
- (+) Privacy protection (only publicize degrees) (Karwa & Slavković 2016)
- (-) Limited expressivity

- Nodes are partitioned into K communities (blocks, groups)
- Formula (Holland et al. 1983)

$$P_{i,j} := \rho_n P_{c_i,c_j}$$

 c_i : which community node i belongs to

 ρ_n : network sparsity rescaler

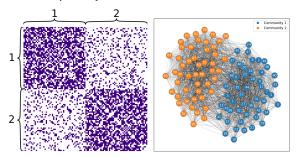


Figure: Left: heatmap of adjacency *A*; right: network plot (nodes have been sorted)

SBM:

$$P_{i,j} := \rho_n B_{c_i,c_j}$$

where B: community-level edge probability matrix

• Estimation: MLE:

$$\begin{split} L(c,B) &:= \sum_{1 \leq i < j \leq n} \left\{ A_{i,j} \log B_{c_i,c_j} + (1 - A_{i,j}) \log (1 - B_{c_i,c_j}) \right\} \\ &(\widehat{c},\widehat{B}) := \underset{c_i \in [1:K], B \in [0,1]^{K \times K}}{\arg \max} L(c,B) \end{split}$$

- Estimating c_{true} is community detection (Abbe 2018)
- If $\rho_n \gg n^{-1} \log n$, then MLE is consistent: as $n \to \infty$,
 - $\hat{c} = c_{\text{true}}$ a.s. (Bickel & Chen 2009)
 - $\widehat{B} \xrightarrow{p} B_{\text{true}}$

Question: How to understand this requirement on ρ_n ?

$$L(c,B) := \sum_{1 \le i < j \le n} \left\{ A_{i,j} \log B_{c_i,c_j} + (1 - A_{i,j}) \log (1 - B_{c_i,c_j}) \right\}$$

- Exact MLE infeasible (combinatorial optimization for c)
- In practice: Tabu search (Zhao et al. 2012)
 - 1 Initialize c and B
 - 2 Iterate until convergence:
 - For $i \in [1:n]$: update c_i
 - Update B

SBM:

$$P_{i,j} := \rho_n P_{c_i,c_j}$$

MLE is too slow, can we do better?

Alternative formula:

$$P = \rho_n Z B Z^T$$

where membership matrix Z may look like:

$$\begin{pmatrix} 1 & 0 & 0 & 0 \\ 1 & 0 & 0 & 0 \\ 0 & 0 & 0 & 1 \\ \vdots & \vdots & \vdots & \vdots \\ 0 & 1 & 0 & 0 \end{pmatrix}$$

corresponding to $c=(1,1,4,\ldots,2)$, i.e., each row of Z is a one-hot vector indicating membership

• Verify: $P_{i,j} = \rho_n Z_{i,\cdot} B Z_{j,\cdot}^T = \rho_n B_{c_i,c_j}$

SBM (matrix form):

$$P_{n\times n} = \rho_n Z_{n\times K} B_{K\times K} (Z^T)_{K\times n}$$

Therefore:

- P is rank-K
- **SVD** of *P*:

$$P = USU^T$$

where $S: K \times K$ diagonal; $U: n \times K$ orthonormal: $U^TU = I$

- U only has K different rows, one for each community (Qin & Rohe 2013)
- cluster rows of $U \Rightarrow$ true community labels

SBM (matrix form):

$$P = \rho_n ZBZ^T$$

In practice, we only observe A ...

- Suppose we know K
- Leading-K SVD of A:

$$A \approx \widehat{U}\widehat{S}\widehat{U}^T$$

(network sparsity absorbed into \widehat{S})

- Cluster the rows of $\widehat{U}\Rightarrow$ community detection
- The above method is called spectral clustering (Lei & Rinaldo 2015)
- Consistency: $\rho_n \gg n^{-1} \log n \Rightarrow$ misclassification rate $\stackrel{p}{\rightarrow} 0$

SBM:

$$P_{i,j} := \rho_n P_{c_i,c_j}$$

- We only explained "block model"
- What about "stochastic"?
- Two-stage generation (Bickel & Chen 2009):
 - 1 "Stochastic": each node chooses its community:

$$c_1, \ldots, c_n \overset{\text{i.i.d.}}{\sim} \mathsf{Multinomial}(q_1, \ldots, q_K)$$

2 "Block model":

$$P_{i,j} := \rho_n P_{c_i,c_j}$$

Stochastic or not doesn't matter so much in practice

Arguably the "most famous" network model

- Simple, yet expressive
- Communities are very interpretable
 Examples:
 - social circles (friendship network) (Yang et al. 2013)
 - functional regions (brain image network) (Wu et al. 2022)
 - research areas (citation network) (Ji et al. 2022)
- Towards general models: SBM with growing K

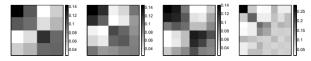
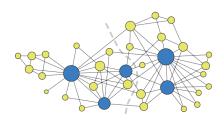


Figure: Choi et al. (2012): as *K* increases, SBM captures more structural details



- What may go wrong with SBM?
- Why? (Recall spectral clustering ...)
- In community detection, we want to adjust for degree heterogeneity

• Formula:

$$P_{i,j} = \rho_n \theta_i \theta_j B_{c_i,c_j}$$

where θ_i : degree correction

Matrix form:

$$P = \rho_n \Theta ZBZ^T \Theta$$

where $\Theta = \operatorname{diag}(\theta_1, \dots, \theta_n)$

• **Identifiability:** require $\sum_{i:c_i=k} \theta_i = 1$ for each community k



Figure: Karrer & Newman (2011): DCSBM adjusts for degrees

- Spectral clustering (Von Luxburg 2007):
 - 1 Compute Laplacian:

$$L := D^{-1/2}AD^{-1/2}$$

where $D = \operatorname{diag}(d_1, \ldots, d_n)$

Interpretation: reweighted edges

$$L_{i,j} = \frac{A_{i,j}}{\sqrt{d_i d_j}}$$

(how much attention do we pay to our relationship?)

2 Leading-K SVD of L:

$$L \approx \widehat{U}\widehat{S}\widehat{U}^T$$

3 Cluster rows of \widehat{U}

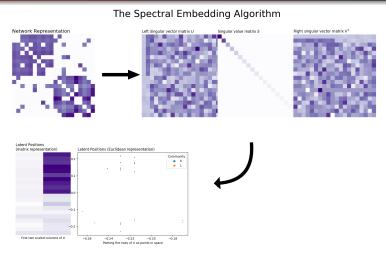


Figure: Illustration of spectral community detection for SBM (Bridgeford et al. 2022)

Quick recap

Which methods have we seen so far?

- Moment estimator for ER model
- MLE for β-model and SBM
- Spectral method for SBM

Overlapping communities:

- Airoldi et al. (2008) Mixed-Membership Stochastic Blockmodel (MMSB):
 - 1 Prior step:

$$ec{\pi}_i \overset{ ext{i.i.d.}}{\sim} ext{Dirichlet}$$

2 For each (i, j), draw $c_i \sim \text{Multinom.}(\vec{\pi}_i)$, do the same for j. Then generate $P_{i,j}$:

$$P_{i,j} = B_{c_i,c_j}$$

- * Parameter estimation: EM algorithm
- Zhang et al. (2020) Overlapping Continuous Community Assignment Model (OCCAM):
 - Matrix form:

$$P = \rho_n \Theta Z B Z^T \Theta$$

but now Z may have continuous rows, satisfying $\|Z_{i,\cdot}\|_2 = 1$ Similarly, (Jin et al. 2024): $\|Z_{i,\cdot}\|_1 = 1$

Parameter estimation: spectral clustering

Other estimation methods:

Modularity (Girvan & Newman 2002, Newman 2006)

$$\frac{1}{2m}\sum_{i< j}\left(A_{i,j}-\frac{d_id_j}{2m}\right)\mathbb{1}_{[c_i=c_j]}$$

m: total number of edges

- Analytical approximate EM algorithm (Ball et al. 2011)
- Semi-definite programming (SDP)

$$\arg\max_{X}\langle A,X\rangle,\quad \text{s.t., } X_{ii}\equiv 1,X\succeq 0$$

Then spectral cluster *X* (Cai & Li 2015, Amini & Levina 2018)

- "SCORE" (Jin 2015)
- Spectral clustering + majority vote refinement (Gao et al. 2018)

Extension: popularity-adjusted SBM (Sengupta & Chen 2018):

- Motivation: overall sparse but locally dense networks DCSBM cannot describe this scenario:
 - $i, j \in \text{same community } k$
 - *i* is globally more popular than *j*
 - ullet ... but not among members of some community ℓ
- **Formulation:** for $i \in \text{community } k, j \in \text{community } \ell$:

$$P_{i,j} = \Lambda_i^{(\ell)} \Lambda_j^{(k)}$$

where Λ encodes:

- · community-level edge probabilities
- node's popularity for each (k, ℓ)

You can think:
$$\Lambda_i^{(\ell)} = \theta_i^{(\ell)} \sqrt{\rho_n B_{k,\ell}}$$
 (recall $c_i = k$)

Reduce to DCSBM:

$$\Lambda_i^{(\ell)} = \theta_i \sqrt{\rho_n B_{k,\ell}} \quad \Rightarrow \quad P_{i,j} = \rho_n \theta_i \theta_j B_{k,\ell}$$

MLE for Λ's ⇒ community detection

popularity-adjusted SBM (Cont'd)

$$P_{i,j} = \Lambda_i^{(\ell)} \Lambda_j^{(k)}$$

This is subtle: MLE for Λ 's \Rightarrow community detection

- MLE gives you all $\Lambda_i^{(\ell)}$'s ..
- ... but that's insufficient for determining $P_{i,j}$
- For each (*i*, *j*):

$$P_{i,j} = \Lambda_i^{(?)} \Lambda_j^{(?)}$$

Random dot-product model (RDPG)

- Generalization of SBM/DCSBM (Young & Scheinerman 2007)
- Formula: recall DCSBM (no ρ_n):

$$P = \Theta Z B Z^T \Theta$$

Suppose $B \succ 0$ (diagonal all positive), then

$$P = (\Theta Z B^{1/2}) \cdot (\Theta Z B^{1/2}) =: X X^T$$

where $X \in \mathbb{R}^{n \times K}$: latent space matrix

- In other words:
 - Each node has an embedding X_i,.
 - Edge probability: $P_{i,j} = \langle X_{i,\cdot}, X_{j,\cdot} \rangle$
- **Estimation:** can estimate X up to a rotation/reflection O: Under mild conditions, \widehat{X} by SVD: $\|\widehat{X} XO\|_{2 \to \infty} \stackrel{p}{\to} 0$ (Cape et al. 2019); CLT (Athreya et al. 2022)

Random dot-product model (RDPG)

Approximately low-rank structure is ubiquitous

• Udell & Townsend (2019): for any A

$$\inf_{\operatorname{rank}(\widehat{A}) \leq C \log n/\varepsilon^2} \|A - \widehat{A}\|_{\max} \leq \varepsilon \|A\|$$

 Xu (2018): Holder smoothness in graphon (will explain later) ⇒ polynomial eigenvalue decay

$$\sum_{i>r} \lambda_i^2 = O(r^{-\beta} + n^{-1})$$

Random dot-product model (RDPG)

$$P = XX^T$$
, $A \approx \widehat{X}\widehat{X}^T$, $\widehat{X} \approx XO$

Assortativity constraint?

Generalized RDPG (Rubin-Delanchy et al. 2022)

$$P = X \begin{pmatrix} I_p & 0 \\ 0 & -I_q \end{pmatrix} X^T$$

Estimation: still SVD A:

$$A = \widehat{U}\widehat{S}\widehat{V}^T$$

compare the signs of columns: $\widehat{U}_{\cdot,i}$ vs $\widehat{V}_{\cdot,i}$ to know sign

Random dot-product model (RDPG)

$$P = XX^T$$
, $A \approx \widehat{X}\widehat{X}^T$, $\widehat{X} \approx XO$

Interpretability

• Rows i of \widehat{X} : embedding of node i

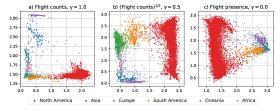


Figure: Gallagher et al. (2024): embedding of airports in flight network

 However, this embedding is limited to that particular network, due to the unknown O
 Cannot be directly used for comparing nodes' roles across networks

Latent space models

Distance model (Hoff et al. 2002)

• Formulation:

$$Logit(P_{i,j}) := \alpha + \beta x_{i,j} + ||z_i - z_j||$$

where

- x_{i,j}: observable edge covariate (two node covariates → edge covariate)
- z_i: latent space position
- Identifiability: $\sum_{i=1}^{n} z_i = \vec{0}$
- Estimation: Variational EM

Multiplicative Z term (Ma et al. 2020)

• Formulation:

$$Logit(P_{i,j}) := \alpha + \beta x_{i,j} + \langle z_i, z_j \rangle$$

Identifiability condition omitted here

Estimation: (regularized) MLE by (projected) gradient

Exponential random graph models (ERGM) (Hunter et al. 2008)

- Directly model the distributions of subgraphs (motifs)
- Dependent edge generation
- Not a convenient generative model (model $\stackrel{?}{\rightarrow}$ data)

Formula:

- Set of motifs (subgraphs) \$\mathscr{R}_1, \ldots, \mathscr{R}_m\$
 Examples: edge, triangle, three-star
- Given an adjacency matrix a, let R(a) count the frequency of R in a
- Network likelihood (Hunter & Handcock 2006):

$$\mathbb{P}(A=a) \propto \exp \left\{ \sum_{i=1}^{m} \theta_i R_i(a) \right\}$$

To fit an ERGM to data:

- Choose the set of motifs, compute $R_i(a)$'s
- Estimate θ_i 's

$$\mathbb{P}(A=a) \propto \exp\left\{\sum_{i=1}^{m} \theta_{i} R_{i}(a)\right\}$$

Given θ , how to generate data?

- Start with an initial adjacency
- Each time, decide whether to flip an edge (i, j), resulting A (before flip) $\rightarrow A'$ (after flip)
- Metropolis-Hastings (Liu & Liu 2001): accept the flip w.p.

$$\min (1, \exp \left[\sum_{i} \theta_{i} \{R_{i}(A') - R_{i}(A)\}\right])$$

- Iterate many rounds (with burn-in and thinning)
 - Burn-in: discard the first few rounds (e.g., first 1000 rounds)
 - Thinning: take data from every few rounds, as if they were i.i.d.

$$\mathbb{P}(A=a) \propto \exp\left\{\sum_{i=1}^{m} \theta_{i} R_{i}(a)\right\}$$

Given A = a, how to estimate θ ?

- MLE is hard, don't know normalizing constant
- Score equation:

$$\nabla \ell(\theta) = 0$$
,

that is, for each i,

$$R_i(a) = \mathbb{E}[R_i(A)]$$

Robbins–Monro update (Snijders 2002):

$$\theta_i^{(n+1)} = \theta_i^{(n)} + \alpha_n \{R_i(a) - \widehat{\mathbb{E}}^{\theta^{(n)}}[R_i(A)]\},$$

where: $\widehat{\mathbb{E}}[R_i(A)]$ evaluated by MH-MC step size α_i : $\sum_i \alpha_i = \infty$, $\sum_i \alpha_i^2 < \infty$

Overview:

- · A very general framework
- Contains the aforementioned models as special cases

Two-stage network generation:

- **1** Latent node positions: $X_1, \ldots, X_n \stackrel{\text{i.i.d.}}{\sim} \text{Uniform}[0, 1]$
- 2 Latent graphon function:

$$P_{i,j} = f(X_i, X_j) \tag{1}$$

Understandings:

- X_i describes "what kind of node"
- f captures all structural information in the network

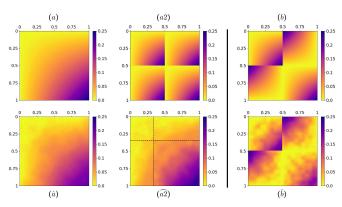


Figure: Illustration of some graphons (Sischka & Kauermann 2025)

Examples:

- ER model: $f(x,y) \equiv p$
- *β*-model:

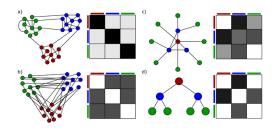
$$\beta_i = g(X_i)$$

$$f(x, y) = \text{Logit}^{-1}(g(X_i) + g(X_j))$$

for proper $g(\cdot)$

Examples: SBM

- K communities \Leftrightarrow partition [0,1] into K sub-intervals
- $X_i \in k$ th interval $\Leftrightarrow i \in$ community k
- Graphon *f*: piece-wise constant



Aldous-Hoover theorem (Zhao 2023)

• Exchangeable network: for any permutation $\pi : \mathbb{N} \to \mathbb{N}$,

$$(A_{i,j}) \stackrel{d}{=} (A_{\pi(i),\pi(j)})$$

Aldous-Hoover: Exchangeable network has the representation

$$A_{i,j} \sim \mathsf{Bernoulli}(f(X_i, X_j)),$$

for $X_i \stackrel{\text{i.i.d.}}{\sim} \mathsf{Uniform}[0,1]$

- Exchangeability ⇒ graphon model
- All nodes i.i.d. choose their roles ⇒ exchangeability

Examples:

- RDPG with i.i.d. latent positions ⇒ graphon model
- Latent space model with i.i.d. latent positions ⇒ graphon model
- Recall ERGM:

$$\mathbb{P}(A=a) \propto \exp\left\{\sum_{i=1}^{m} \theta_{i} R_{i}(a)\right\}$$

where each $R_i(a)$: some motif count Chatterjee & Diaconis (2013): ERGM \approx graphon model (as $n \to \infty$)

$$X_1, \dots, X_n \overset{\text{i.i.d.}}{\sim} \mathsf{Uniform}[0,1]$$

 $P_{i,j} = f(X_i, X_j)$

Graphon estimation:

- Cannot estimate *f* or *X*_i's (not identifiable)
- May estimate P_{i,j}'s
- Smoothness assumption on *f* [in equiv. class]
- How to estimate $P_{i,j}$'s?

$$X_1, \dots, X_n \overset{\text{i.i.d.}}{\sim} \text{Uniform}[0, 1]$$

 $P_{i,j} = f(X_i, X_j)$

Graphon estimation:

- Method 1: universal singular value thresholding (USVT) (Chatterjee 2015)
 - **1** SVD *A*:

$$A = \widehat{U}\widehat{S}\widehat{V}^T$$
, (S diag. sorted)

- **2** Keep singular values above $C\sqrt{n}$ for C>2Say, keep the first $K\Rightarrow (\widetilde{U},\widetilde{S},\widetilde{V})=(\widehat{U}_{\cdot,1:K},\widehat{S}_{1:K,1:K},\widehat{V}_{\cdot,1:K})$
- 3 Estimator:

$$\widehat{P} := \widetilde{U}\widetilde{S}\widetilde{V}^T$$

Achieves SOTA MSE (Xu 2018, Luo & Gao 2024)

$$X_1, \dots, X_n \overset{\text{i.i.d.}}{\sim} \text{Uniform}[0, 1]$$

 $P_{i,j} = f(X_i, X_j)$

Graphon estimation:

- Method 2: Neighborhood smoothing (NS) (Zhang et al. 2017)
 - **1** Dissimilarity measure of graphon slices $f(\cdot,X_i)$ and $f(\cdot,X_j)$:

$$d(i,j) := \max_{i' \neq i,j} |\langle A_{i,\cdot} - A_{j,\cdot}, A_{i',\cdot} \rangle| / n$$

- 2 For each node i, set neighborhood $\mathcal{N}_i := \{j : \text{smallest } h = \log n/n \text{ proportion of } d(i, j) \text{'s} \}$
- 3 Estimator:

$$\widehat{P}_{i,j} := \frac{1}{|\mathcal{N}_i|} \sum_{i' \in \mathcal{N}_i} A_{i',j}$$

Achieves state-of-the-art " $2 \rightarrow \infty$ error"; explicit formula

Selected research topics

Overview

We will cover these selected topics

- Method-of-moments
- Goodness-of-fit tests
- Cross-validation
- Graph matching
- Extensions to weighted, directed, bipartite graphs
- Other topics (briefly mention)

Method-of-moments

Network moments

Definition: a network moment indexed by a motif R is

$$\widehat{U}_n := \binom{n}{r}^{-1} \sum_{i_1 < \dots < i_r} \mathbb{1}_{[A_{i_1, \dots, i_r} \cong R]}$$

• Examples: triangle count

$$\widehat{U}_n = \binom{n}{3}^{-1} \sum_{i < j < k} A_{i,j} A_{j,k} A_{k,i}$$

• CLT:

$$\frac{\widehat{U}_n - \mathbb{E}[U_n]}{\widehat{S}_n} \xrightarrow{p} N(0,1),$$

where S_n is some variance estimator (Bickel et al. 2011, Zhang & Xia 2022)

 Moments can be used to identify network models Application: network two-sample test

Is the network "structureless"? (Gao & Lafferty 2017b)

- $H_0: P_{i,j} \equiv p$ (ER as null model)
- *H_a*: non-constant graphon model
- Method-of-moment test: null distribution:

$$n^{3/2}g(p)\cdot (T_2,T_3)^T \stackrel{d}{\to} N(0,I_2),$$
 (2)

where $T_2 = E^3 - T$ and $T_2 = 3E^2(1 - E) - V$, here:

- E counts edges
- T counts triangles
- V counts V-shapes

 $g(\cdot)$ is known function; easy to estimate p

GoF test for SBM (Lei 2016)

- $H_0: K = K_0 \text{ vs } H_a: K > K_0$
- Based on observed A, fit SBM with $K = K_0$, get \widehat{P}
- Idea: is the residual network "pure noise"?
- Test statistic:

$$\widetilde{A}_{i,j} := \frac{A_{i,j} - \widehat{P}_{i,j}}{\sqrt{(n-1)\widehat{P}_{i,j}(1-\widehat{P}_{i,j})}}, \quad i \neq j$$

Then under H_0 ,

$$n^{2/3} \cdot \big\{ \pm \lambda_1(\widetilde{A}) - 2 \big\} \overset{d}{\to} \mathsf{Tracy\text{-}Widom}(1),$$

where $\lambda_1 \geq \cdots \geq \lambda_n$ are sorted eigenvalues

Application: decide K

Key to build valid GoF method:

- Number of continuous parameters is small under H₀
- Example: ER model has 1 parameter
- **Example:** SBM has $O(n+K^2)$, but the n memberships can be perfectly recovered for not-very-sparse networks, so we just deal with $O(K^2)$ parameters
- Example: DCSBM (Zhang & Amini 2023, Jin et al. 2025)
 H₀: model is an SBM (no degree corrections)

No known method to test GoF for:

- β-model
- Latent space models, including RDPG
- · General graphon model

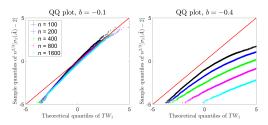


Figure: Shao et al. (2023): exactly mimicking Lei (2016) for β -model won't work (each curve)

Cross-validation

CV for network data?

- Motivation: select tuning parameter(s)
 Example: number of communities K in SBM
 Example: bandwidth h in neighborhood smoothing
- Challenge: we may observe only one A, how to split data?

Method 1 (node-split CV) (Chen & Lei 2018)

Write

$$A = \begin{pmatrix} A^{(11)} & A^{(12)} \\ A^{(21)} & A^{(22)} \end{pmatrix}$$

- 2 Use $[A^{(11)}, A^{(12)}]$ to train model (say, SBM)
- **3 Validation:** look at how well fitted model predicts $A^{(22)}$ (beware $A^{(21)} = \{A^{(12)}\}^T$ already used for training)

Cross-validation

Method 2 (edge-splitting CV): (Li et al. 2020)

- 1 Randomly sample $M_{i,j} \stackrel{\text{i.i.d.}}{\sim}$ Bernoulli(q); set $M_{j,i} = M_{i,j}$ Training data: $\{(i,j): M_{i,j} = 1\}$ Validation data: $\{(i,j): M_{i,j} = 0\}$
- 2 Fit a model to training data using, say, low-rank method May fill in zeros to validation data spots This outputs \widetilde{A} , then $\widehat{A} = \widetilde{A}/q$ is link predictor
- 3 See how well \widehat{A} predicts A entries on validation data

- **Data:** two adjacency matrices $A^{(1)}, A^{(2)} \in \{0,1\}^{n \times n}$
- Problem: estimate a permutation matrix Π minimizing

$$||A^{(1)} - \Pi A^{(2)} \Pi^T||_F^2 \tag{3}$$

- Applications (Conte et al. 2004, Lyzinski et al. 2015):
 - Account matching in different social networks
 - Connectome alignment in brain imaging
 - Structurally similar chemical particles
 - Network two-sample test
- Complexity: (3) is NP (not known to be P)
- Statistical question: if A's are structured, would the problem be more solvable?

Approach 1: low-rank approximation

- Assume model is low-rank: $A \sim P = XX^T$
- Leading-k SVD:

$$A \approx \widehat{X}\widehat{X}^T$$

Davis-Kahan theorem:

$$\hat{X} \approx XO$$

for (inestimable) orthonormal $O: O^TO = I$

Minimize:

$$\|\widehat{X}^{(1)} - \Pi \widehat{X}^{(2)} O\|_F$$

over Π and O

$$\|\widehat{X}^{(1)} - \Pi \widehat{X}^{(2)}O\|_F$$
 (4)

Approach 1: low-rank approximation (cont'd)

- joint optimization of (4) is still difficult
- Zhang (2018): first estimate O, then Π (no iteration!)
- **Key idea:** match row distributions, get \widehat{O} ; then easy to estimate Π
- Alternatively, Tang et al. (2014) uses a rotation-invariant kernel to run MMD (Gretton et al. 2012) on X-rows

Approach 2: local network moments

Correlated ER model:

$$\mathrm{Cor}\big(A_{i,j}^{(1)},A_{\pi^*(i),\pi^*(j)}^{(2)}\big) = \rho$$

for some true permutation π^*

- Method 1: degree profile (Ding et al. 2021):
 - **1** For each node i, compute degree distribution of its neighbors: F_i
 - 2 Similarity matrix:

$$S_{i,j} := \operatorname{distance}(F_i, F_j)$$

3 Estimate π by Hungarian algorithm

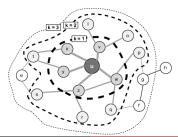
Approach 2: local network moments (cont'd)

- Method 2: large neighborhood statistics (Mossel & Xu 2020):
 - \bullet Seed I_0 : a small set of correctly matched nodes
 - **2** For each k and each $j_1 \in \text{net 1}$, $j_2 \in \text{net 2}$, compute

```
w_{k;j_1,j_2} := \{i : i \in I_0, i \in \{k\text{-hop neighborhoods of both } i \text{ and } j\}\}
```

k: tuning parameter; if $|w_{k;j_1,j_2}|$ is large, match (j_1,j_2)

 Many recent (strong) results (Ding & Du 2023, Mao et al. 2024) and more ...



Extensions to weighted, directed, bipartite graphs

Weighted networks:

- $A_{i,j} \sim \mathsf{Bernoulli} \Rightarrow A_{i,j} \sim \mathscr{F}_{\theta}$
- Example: Poisson β -model (Yan et al. 2016)

$$\mathbb{P}(A_{i,j} = x) = \frac{e^{(\beta_i + \beta_j)x}}{x!} e^{-e^{\beta_i + \beta_j}}$$

Example: weighted SBM (Xu et al. 2020)

$$A_{i,j} \sim \mathscr{F}_{c_i,c_j}$$

There are $\binom{K}{2}$ different \mathscr{F} 's

 Typically the same methods (MLE, spectral, ...) work for the weighted version

Extensions to weighted, directed, bipartite graphs

Directed networks:

- Each node may have two roles: sender and receiver
- Example: directed SBM (Wang & Wong 1987)

$$\mathbb{P}(A_{i,j}=1)=P_{c_i,d_j},$$

where

- c_i: community label as sender
- d_i: community label as receiver
- Spectral clustering:

$$A \approx \widehat{U} S \widehat{V}^T$$
,

then:

- Cluster rows of \widehat{U} to estimate \vec{c}
- Cluster rows of \hat{V} to estimate \vec{d}

Extensions to weighted, directed, bipartite graphs

Bipartite networks:

- Two node sets of distinct meanings
 Example: recommender system: buyers
 ⇔ goods
 Example: drug-target network
- A is $m \times n$
- Method and theory: mostly similar to directed networks Some differences:
 - Unlike directed networks, no self-loops
 - Asymptotics: m,n may grow at different rates

Other topics

Resampling networks

Problem description:

- Observe: one A
- Goal: sampling distribution of some model functional (e.g., triangle count)

Recall i.i.d. case: observe X_1, \ldots, X_n , goal: sample mean

- Parametric bootstrap: use X to fit a model $F_{\hat{\theta}}$, resample from F, repeatedly get X_1^*, \ldots, X_n^* and bootstrapped \bar{X}^*
- Non-parametric bootstrap: resample from X, repeatedly get X_1^*, \ldots, X_n^* and bootstrapped \bar{X}^*
- These two methods produce similar \bar{X}^* distributions

Resampling networks

Network? ... observe *A*, goal: triangle count

- **Method 1:** use A to fit \widehat{P} , then repeatedly generate A^* from \widehat{P} and compute triangle count
- **Method 2** (Green & Shalizi 2017): repeatedly resample nodes, each time induce A^* , count triangles Specifically, k_1^*, \ldots, k_n^* resampled from [n], then

$$A_{i,j}^* := A_{k_i,k_j}$$

 Methods 1 and 2 may give you very different triangle count sampling distributions!

What happened?

- **Method 1.5:** use A to fit \widehat{P} , then still resample nodes, each time induce \widehat{P}^* , then sample A^* from \widehat{P}^* , count triangles
- Method 1.5 \approx Method 2 \neq Method 1
- Reading: Zhang & Xia (2022)

Dynamic networks

What do data look like?

- Scenario 1: network snapshots: $\{A^{(t)}\}_{t=1}^T$ (Sewell & Chen 2015, Pensky 2019)
 - **Examples:** friendship, international trade
- Scenario 2: stream of edges: (i, j,t), where t: time stamp (Less studied) (Perry & Wolfe 2013, Fang et al. 2024)
 Examples: emails, contact (epidemiology)

Network snapshots:

- Build model for each layer
- Parameters evolve over time / across layers
- Different layers may have similar/shared parameters
- Tensor analysis (Dr. Xia's tutorial tomorrow)
- Some works: convert edge stream to snapshots

Nodal covariates

Sometimes you observe more than a network ...

- Data: $A \in \mathbb{R}^{n \times n}$ and $X \in \mathbb{R}^{n \times p}$
- Examples: individual covariates (gender, age, job, ...)
- Structure of X may relate to structure of A

Use X to help modeling A

Example: covariate-assisted community detection (Binkiewicz et al. 2017)

• Method:

Apply spectral clustering to: $A + \lambda XX^T$

- λ : tuning parameter
- Edge covariates can be utilized similarly

Nodal covariates

Use A to model/predict X

Example: network regression (Le & Li 2022)

- **Data:** on each node, observe (X_i, Y_i)
- Goal: use X and A to predict Y
- **Model:** $C_X := \text{ColumnSpace}(X), C_P := \text{ColumnSpace}(P),$

$$\mathbb{E}[Y|X] = \underbrace{X\beta}_{\in C_X} + \underbrace{X\theta}_{\in C_X \cap C_P} + \underbrace{\alpha}_{\in C_P}$$

and $(X\beta) \perp (X\theta)$ and $(X\theta) \perp \alpha$

• Estimation strategy: first estimate θ

Nodal covariates

Some understandings

- Beware of differences in X and A structures
 (considering X might or might not help A's model fitting)
 (Zhang et al. 2016, Deshpande et al. 2018)
- Be careful when building joint models
 Example: shall we worry about this model for joint community detection?

$$P_{i,j} \propto \exp\left\{B_{c_i,c_j} + C_{X_i,X_j}\right\}$$

where suppose each $X_i \in [K]$, here

- c_i: network community label
- *X_i*: covariate community label

Sampling network

What if we have to explore the network?

- We might not be able to sample nodes i.i.d. from {all kinds of nodes}
- Example (snowball): start from an initial set of people, ask them to refer friends to our study, and friends ask their further friends, ...
- Study strategy:
 - Model the true (complete) network
 - 2 Specify a sampling scheme
 - 3 Analyze what the sampling scheme will produce

Sampling network

Example: random walk (Athreya & Röllin 2016)

- Mechanism. Start from i_0 ; randomly pick a neighbor, walk to i_1 ; randomly pick a neighbor of i_1 , walk to i_2 , ... Obtain the induced network observation
- **Theorem.** Network \sim graphon f(u,v), then the network surveyed by random walk is approximately generated by

degree-weighted graphon:
$$f(X_i, X_j)$$
,

but now X_i is **not** Uniform[0,1], instead, has pdf:

$$p_X(u) := \frac{\int_0^1 f(u, v) dv}{\int_0^1 \int_0^1 f(u', v) du' dv} \propto \mathbb{E}[\mathsf{degree}] \text{ at } u$$

 Many more other sampling schemes (non-backtracking, snowball, ...) (Rohe 2019)

Core periphery structure

- Aforementioned models attempt to model all nodes
- But often there are seclusive nodes with few connections (Rombach et al. 2014)
- Core: dense and structured subnetwork
- Periphery: the rest of network

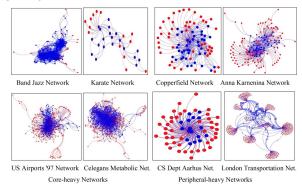


Figure: Blue: core; red: periphery (Meghanathan 2024)

Core periphery structure

Miao & Li (2023)

- Informative model for core part
- Non/Less-informative model for periphery part

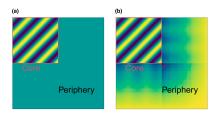


Figure: Two periphery models: left: ER; right: configuration (Miao & Li 2023)

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