

# Network Method of Moments

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MATLAB code: [github.com/yzhanghf](https://github.com/yzhanghf)

# Introduction and motivation

## Parametric network analysis:

- Parametric model  $\rightarrow$  point estimation  $\xrightarrow{?}$  inference

## Challenges:

- Inference may be difficult to derive
- Method is model-specific

## Non-parametric methods:

- **Less ambitious goal:** not learning **every detail** of network model, just **numerical features**
- **More flexibility:** model-free/applicable to many models; weak model assumptions
- **Computation efficiency:** easier/faster than fitting some models

# Introduction and motivation

## Network method of moments

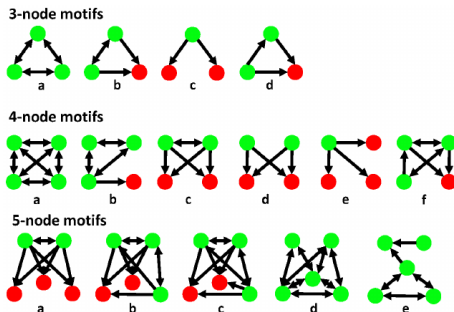
- **Network moments** extension of the classical *moments* in i.i.d. setting
- Fast and straightforward computation
- Model-free
- Universal and principled inference

**Question:** How to define *network moments*?

# Introduction and motivation

## Network moments (*Bickel et al. 2011*)

- Network moments are indexed by motifs
- Example: **edge**:  $\text{mean}_{ij}(A_{ij})$
- Example: **triangles**:  $\binom{n}{3}^{-1} \sum_{1 \leq i < j < k \leq n} A_{ij} A_{jk} A_{ki}$
- More examples (directed networks, *Jayavelu et al. (2014)*):



## Descriptive power of network moments

- Network comparison:  
Different network moments  $\Rightarrow$  different network models
- Knowing all moments determines exchangeable network model?  
“**Nearly yes**” (“Yes” for practitioners) (*Borgs et al, 2010*)
- May service some parametric models: ERGM:

$$\text{likelihood of } A \propto \exp \left\{ \sum_k \text{Motif}_k(A) \right\}$$

- Related topics:
  - One sample inference (*This paper*), (*Shao, Xia & Z., 2022+*)
  - Two sample inference (network comparison) (*Ghoshdastidar et al, 2017*), (*Shao et al, 2022+*)

**Major challenge:** distribution of network moments?

To better illustrate, we first describe the **base model**

# Problem formulation

## Data:

- Adjacency matrix:  $A \in \mathbb{R}^{n \times n}$

$$A_{ij} = A_{ji} = \begin{cases} 1 & i \leftrightarrow j \\ 0 & \text{otherwise} \end{cases}$$

- Symmetric edge probability matrix:  $W \in \mathbb{R}^{n \times n}$ :

$$A_{ij} | W \stackrel{\text{independent}}{\sim} \text{Bernoulli}(W_{ij})$$

# Problem formulation

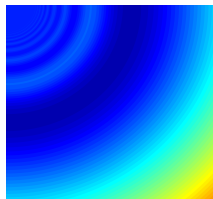
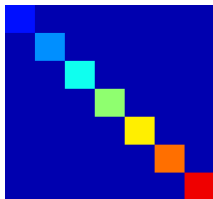
## Exchangeable networks (Aldous-Hoover representation):

- Latent **graphon function**  $f : [0, 1]^2 \rightarrow [0, 1]$
- Latent **node position**  $X_i \sim \text{Uniform}[0, 1]$ :

$$W_{ij} = \rho_n \cdot f(X_i, X_j)$$

$\rho_n$ : sparsity multiplier

- $f$  encodes structures;  $X_i$  encodes node's role; both **inestimable**





## Formulation of network motif:

- **Motif  $R$ :**  $r$  nodes and  $s$  edges
- Corresponding sample moment is the count statistic

$$\hat{U}_n := \binom{n}{r}^{-1} \sum_{1 \leq i_1 < \dots < i_r \leq n} h(A_{i_1, \dots, i_r}),$$

where

$$h(A_{i_1, \dots, i_r}) := \mathbb{1}_{[A_{i_1, \dots, i_r} \supseteq R]}$$

# Problem formulation

**Question:** How to characterize  $\hat{U}_n$ ?

- Design a proper variance estimator  $\hat{S}_n^2$ , and **studentize**:

$$\hat{T}_n := \frac{\hat{U}_n - \mathbb{E}[\hat{U}_n]}{\hat{S}_n}$$

**What's next?**

- 1 How to design  $\hat{S}_n$  =?
- 2 **Distribution of  $\hat{T}_n$ ?**

Before introducing our method, a quick literature review...

# Distribution approximation

## Existing literature

- Asymptotic normality (*Bickel et al, 2011*)
- Network bootstraps:
  - Node sub-sampling: (*Bhattacharyya & Bickel, 2015*)
  - Node re-sampling: (*Green & Shalizi, 2017*)
  - Low-rank approximation then bootstrap estimated low-rank structure: (*Levin & Levina, 2019*)
- Limitations:
  - No finite sample accuracy guarantee, only consistency

$$\hat{T}_n \rightarrow N(0, 1)$$

in  $\xrightarrow{d}$ ,  $\xrightarrow{p}$ , etc

- Slow computation (bootstrap methods)

# Our method

Our paper:

- Analytical, higher-order accurate approximation to  $F_{\hat{T}_n}$
- Fast computation (eliminates bootstrap)
- Model-free & Versatility: applicable to non-smooth graphons  
(Choi, 2017)
- New theoretical insights
- Rate-optimal inference power + higher-order accurate risk control

# Key intuition

**Example:**  $R = \text{Edge}$ :

$$\begin{aligned}\hat{U}_n &= \binom{n}{2}^{-1} \sum_{1 \leq i < j \leq n} W_{ij} + \sum_{1 \leq i < j \leq n} (A_{ij} - W_{ij}) \\ &=: \binom{n}{2}^{-1} \sum_{1 \leq i < j \leq n} f(X_i, X_j) + \binom{n}{2}^{-1} \sum_{1 \leq i < j \leq n} \eta_{ij}\end{aligned}$$

## Decomposition of randomness

$$\hat{U}_n = U_n + (\hat{U}_n - U_n)$$

Here:

- $U_n = U_n(X_1, \dots, X_n)$ : variations of  $W$ , due to nodes' roles
- $\hat{U}_n - U_n$ : observational errors in  $A|W$
- Under mild conditions,  $\text{Var}(U_n) \gg \text{Var}(\hat{U}_n - U_n)$ , so we use  $U_n$  to design variance estimator

# Key intuition

**Example:**  $R = \text{Triangle}$ :

$$\begin{aligned}\hat{U}_n &= \binom{n}{3}^{-1} \sum_{1 \leq i < j < k \leq n} A_{ij} A_{jk} A_{ki} \\ &= \binom{n}{3}^{-1} \sum_{1 \leq i < j < k \leq n} w_{ij} w_{jk} w_{ki} \\ &\quad + \binom{n}{3}^{-1} \sum_{1 \leq i < j < k \leq n} \left\{ \eta_{ij} w_{jk} w_{ki} + w_{ij} \eta_{jk} w_{ki} + w_{ij} w_{jk} \eta_{ki} \right\} \\ &\quad + (\text{Product } \eta \text{ terms, remainder})\end{aligned}$$

# Noiseless version of the problem

- $U_n := \mathbb{E}[\hat{U}_n | W]$  is a **noiseless** U-statistic and admits a **Hoeffding's ANOVA decomposition**

$$U_n - \mathbb{E}[U_n] = \frac{r}{n} \sum_{i=1}^n g_1(X_i) + \frac{r(r-1)}{n(n-1)} \sum_{1 \leq i < j \leq n} g_2(X_i, X_j) + \dots$$

where the **uncorrelated**  $g_k$ 's are:

$$\begin{aligned} g_1(X_1) &:= \mathbb{E}[h(X_1, \dots, X_r) | X_1] - \mathbb{E}[h] \\ g_2(X_1, X_2) &:= \mathbb{E}[h(X_1, \dots, X_r) | X_1, X_2] - g_1(X_1) - g_1(X_2) - \mathbb{E}[h] \\ &\dots\dots \end{aligned}$$



# Design of variance estimator

## Design

$$\hat{S}_n^2 = \frac{r^2}{n^2} \cdot \sum_{i=1}^n \left\{ \underbrace{\frac{1}{\binom{n-1}{r-1}} \sum_{\substack{1 \leq i_1 < \dots < i_{r-1} \leq n \\ i_1, \dots, i_{r-1} \neq i}} h(A_{i, i_1, \dots, i_{r-1}}) - \hat{U}_n}_{\text{Estimates } g_1(X_1)} \right\}^2$$

- **Theorem 3.3 Z. & Xia, (2022)**  $\hat{S}_n^2$  is equivalent to **network jackknife** (Maesono, 1997), but computes faster and more convenient for analysis

**Next: Distribution of  $\hat{T}_n$ ?**

# Distribution of $\hat{T}_n$ ?

If we simplify it (we can't, but just for this moment)...

$$U_n - \mathbb{E}[U_n] = \frac{r}{n} \sum_{i=1}^n g_1(X_i) + \frac{r(r-1)}{n(n-1)} \sum_{1 \leq i < j \leq n} g_2(X_i, X_j) + \dots$$

- **Edgeworth expansion** for i.i.d. data  
(need **Cramer's condition**):

$$\text{CDF}\left(\frac{\bar{g}_1(X) - \mu}{\sigma_{g_1(X_1)}/\sqrt{n}}; u\right) = \Phi(u) - \varphi(u) \frac{\mathbb{E}[g_1(X_1)^3](u^2 - 1)}{6\sqrt{n} \cdot \sigma_{g_1(X_1)}^3} + O(n^{-1})$$

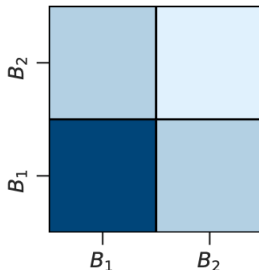
# Edgeworth expansion for noiseless U-statistics

## Cramer's condition:

- For noiseless U-statistic,  $g_1(X_1)$  satisfies:

$$\limsup_{t \rightarrow \infty} \left| \mathbb{E}[e^{itg_1(X_1)}] \right| < 1$$

- Cramer's condition  $\approx g_1(X_1)$  is **continuous**
- Violation:** block model (think  $R = \text{Edge}$ , then  $n(g_1(X_i) + \mu) = \text{expected degree of node } i$ )



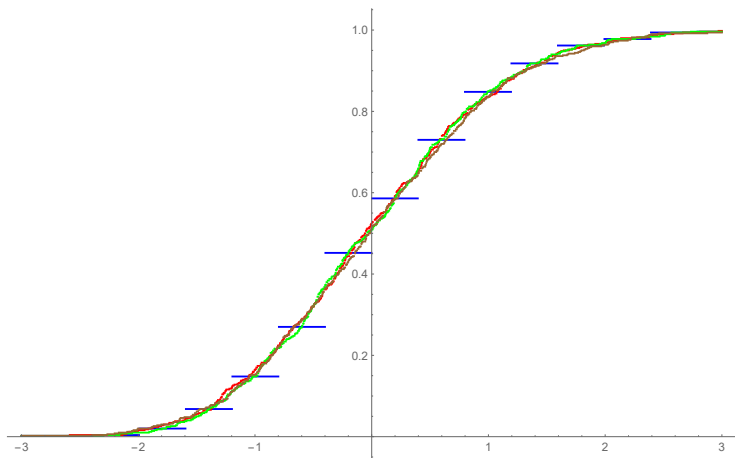
# Key insight

$$\hat{U}_n = \underbrace{U_n}_{\text{Randomness from } X_1, \dots, X_n} + \underbrace{(\hat{U}_n - U_n)}_{\text{Randomness from } A|W}$$

Key findings:  $\hat{U}_n - U_n$  provides a self-smoothing effect

- **Observational error:** behaves like a Gaussian smoother
- **Sparsity:** amplifies the smoothing effect to eliminate potential discontinuity in  $U_n$

# Observational noise smooths CDF



IID case illustration: data  $X_1, \dots, X_{10}$  are Bernoulli. Blue: original  $\bar{X}$ , Red: plus uniform noise, bandwidth  $n^{-1/2}$ , Green: mixed with normal noise, bandwidth  $(\log n/n)^{1/2}$ , Brown, normal noise bandwidth  $n^{-1/2}$

# Expansion expansion

Network Edgeworth expansion:

$$G_n(x) := \Phi(x) - \frac{\varphi(x)}{\sqrt{n} \cdot \xi_1^3} \cdot \left\{ \frac{2x^2 + 1}{6} \cdot \mathbb{E}[g_1^3(X_1)] \right. \\ \left. + \frac{r-1}{2} \cdot (x^2 + 1) \mathbb{E}[g_1(X_1)g_1(X_2)g_2(X_1, X_2)] \right\}$$

where  $\xi_1^2 := \mathbb{E}[g_1^2(X_1)]$ ,  $\Phi(x)$  and  $\varphi$ :  $N(0, 1)$  CDF and PDF

- $\hat{G}_n(x) :=$  empirical version

# Main theorems

- **Theorems 3.1 & 3.2 (Z. & Xia, (2022))** Assume

- 1  $\rho_n^{-2s} \cdot \text{Var}(g_1(X_1)) \geq \text{Const} > 0$
- 2 **Dense regime:** Either  $R$  is **acyclic** and  $\rho_n = \omega(n^{-1/2})$ , or  $R$  is **cyclic** and  $\rho_n = \omega(n^{-1/r})$
- 3 Either  $\rho_n = O((\log n)^{-1})$  or Cramer's condition holds

We have

$$\sup_{u \in \mathbb{R}} \left| F_{\hat{T}_n}(u) - G_n(u) \right| = O(\mathcal{M}(\rho_n, n; R)) \ll n^{-1/2}$$

where

$$\mathcal{M}(\rho_n, n; R) := \begin{cases} (\rho_n \cdot n)^{-1} \log^{1/2} n + n^{-1} \log^{3/2} n & \text{for acyclic } R \\ \rho_n^{-r/2} \cdot n^{-1} \log^{1/2} n + n^{-1} \log^{3/2} n & \text{for cyclic } R \end{cases}$$

Result also holds:  $G_n(u)$  replaced by  $\hat{G}_n(u)$  ( $O$  replaced by  $O_p$  with  $n^{-1}$  tail probability)

# Main theorems (continued)

- **Theorem 3.4 (Z. & Xia, (2022))** Assume

- ① (same as before)
- ② **Sparse regime:** Either  $R$  is **acyclic** and  $n^{-1} \prec \rho_n \preceq n^{-1/2}$ , or  $R$  is **cyclic** and  $n^{-2/r} \prec \rho_n \preceq n^{-1/r}$
- ③ (same as before)

We have

$$\begin{aligned} \sup_{u \in \mathbb{R}} \left| F_{\hat{T}_n}(u) - G_n(u) \right| &\asymp \sup_{u \in \mathbb{R}} \left| F_{\hat{T}_n}(u) - \Phi(u) \right| \\ &= O(\mathcal{M}(\rho_n, n; R)) \wedge o_p(1) \gg n^{-1/2} \end{aligned}$$

- **Sparse regime:** Berry-Esseen bound dominates  $n^{-1/2}$ ; using  $N(0, 1)$  approximation is good enough



# Applications

## One-sample inference:

- **Hypothesis testing:**

- Type I error =  $\alpha + O(\mathcal{M}(\rho_n, n; R))$ :

$$\text{Estimated p-value} = 2 \min \left\{ \widehat{G}_n(t^{(\text{obs})}), 1 - \widehat{G}_n(t^{(\text{obs})}) \right\}$$

- **Optimal separation condition** under  $H_a$ .

- **Length-optimal CI**; nominal level =  $1 - \alpha + O(\mathcal{M}(\rho_n, n; R))$ :

$$\left( \widehat{U}_n \pm \widehat{q}_{\widehat{T}_n; \alpha/2} \cdot \widehat{S}_n \right)$$

where (**Cornish-Fisher expansion**):

$$\begin{aligned} \widehat{q}_{\widehat{T}_n; \alpha} := & z_\alpha - \frac{1}{\sqrt{n} \cdot \widehat{\xi}_1^3} \left\{ \frac{2z_\alpha^2 + 1}{6} \cdot \widehat{E}[g_1^3(X_1)] \right. \\ & \left. + \frac{r-1}{2} (z_\alpha^2 + 1) \widehat{\mathbb{E}}[g_1(X_1)g_1(X_2)g_2(X_1, X_2)] \right\} \end{aligned}$$

- Computes much faster than **bootstrap iteration** (*Beran, 1987, 1988*)

Understanding the accuracy of network bootstraps:

- Old theory (*Bhattacharya & Bickel, 2015*), (*Green & Shalizi, 2017*), (*Levin & Levina, 2019*):  
 $o(1)$  (only consistency, no finite-sample error rate)
- Our theory implies:  $o(n^{-1/2})$  (when  $\rho_n \gg n^{-1/2}$ )

Our method vs network bootstraps:

- Our error bound  $\ll$  network bootstraps
- Our computation time  $\ll$  network bootstraps

# Simulations

Set up:

- Graphons: SBM, smooth graphon, “non-smooth graphon”
- Motifs: edge, triangle, V-shape

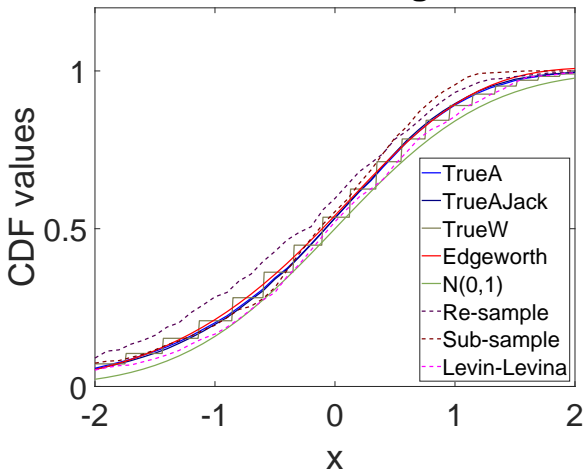
Benchmarks:

- $N(0, 1)$  approximation
- Node re-sampling (*Green & Shalizi, 2017*), Alg. 1
- Node sub-sampling (*Bhattacharya & Bickel, 2015*), Alg. 1

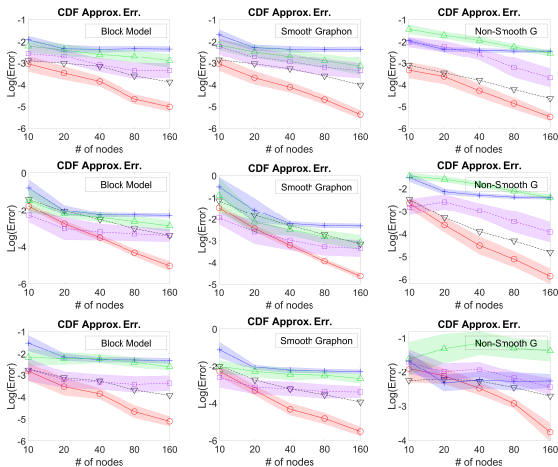
Performance measured by:

- $\|\hat{F}_{\hat{T}_n} - F_{\hat{T}_n}\|_{\infty}$  on  $(-2, 2)$
- 500 experiments, within each iteration:  $10^5$  Monte Carlo repeats to evaluate true CDF

## BlockModel, Triangle, n=80



**Figure:** CDF curves,  $n = 80$ , SBM, triangle, bootstrap sample: 500. TrueA is  $F_{\hat{T}_n}$ ; TrueAJack is  $F_{\hat{T}_{n,jackknife}}$ ; TrueW is  $F_{T_n}$ ; Edgeworth is our EEE; Re-sample is *Green & Shalizi, (2017)*; Sub-sample is *Bhattacharyya & Bickel, (2015)*; Levin-Levina is *Levin & Levina, (2019)*.



**Figure: Motifs:** row 1: Edge; row 2: Triangle; row 3: Vshape. CDF approximation errors. **Our method**,  $N(0,1)$ , **Green & Shalizi (2017)**, **Bhattacharyya & Bickel (2015)**, **Levina & Levina (2019)**

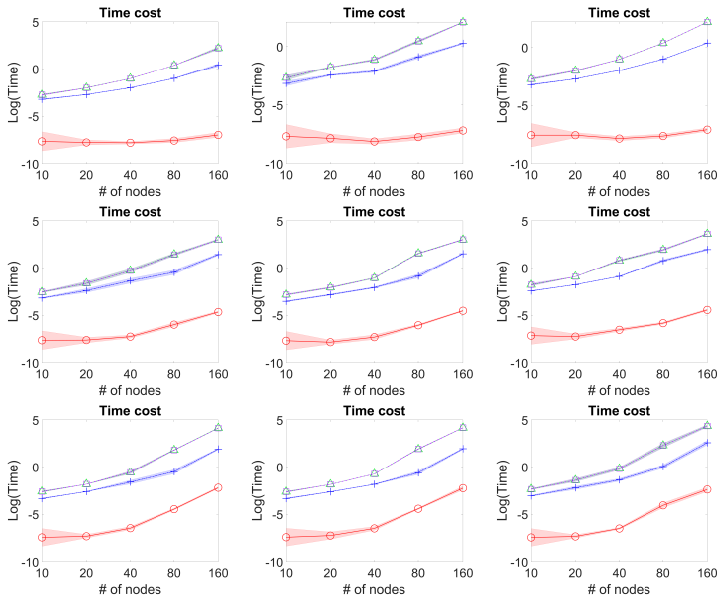
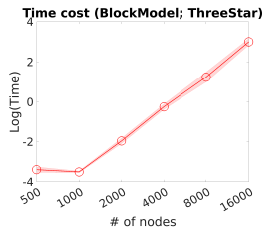
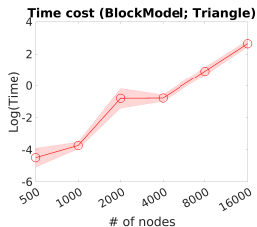
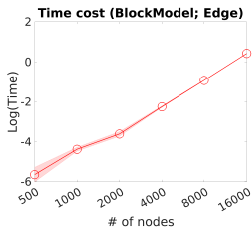


Figure: Log-time, Our method, Green & Shalizi (2017), Bhattacharrya & Bickel (2015), Levina & Levina (2019)

**Table:** Performance measures of 95% confidence intervals  
 $n = 80$ ,  $\rho_n \asymp 1$ , graphon: block model

Method	Edge	Triangle	V-shape	Three star
Our method	Coverage = 0.957(0.202)	0.953(0.211)	0.956(0.205)	0.952(0.213)
	Length = 0.097(0.010)	0.040(0.008)	0.200(0.033)	0.145(0.033)
	LogTime = -8.448(0.110)	-7.214(0.083)	-7.165(0.082)	-7.180(0.353)
Norm. Approx.	0.950(0.218)	0.934(0.248)	0.942(0.235)	0.932(0.251)
	0.097(0.010)	0.040(0.008)	0.200(0.033)	0.145(0.033)
	No time cost*	No time cost	No time cost	No time cost
Bhattacharyya & Bickel (2015)	0.842(0.365)	0.870(0.337)	0.852(0.355)	0.852(0.355)
	0.068(0.009)	0.031(0.007)	0.147(0.026)	0.113(0.025)
	-2.591(0.008)	-2.160(0.026)	-2.127(0.024)	-0.992(0.006)
Green & Shalizi (2017)	0.938(0.241)	0.944(0.230)	0.934(0.249)	0.938(0.241)
	0.096(0.013)	0.044(0.010)	0.204(0.038)	0.150(0.037)
	-1.198(0.007)	0.499(0.032)	0.142(0.035)	0.383(0.010)
Levina & Levina (2019)	0.942(0.234)	0.942(0.234)	0.942(0.234)	0.942(0.234)
	0.099(0.013)	0.043(0.010)	0.209(0.039)	0.155(0.038)
	-1.188(0.004)	0.507(0.028)	0.142(0.027)	0.489(0.004)

\*  $N(0, 1)$  costs the same time as ours in evaluating the studentization



**Figure:** Scalability of our method on large networks.



# Thank you!

Edgeworth expansions for network moments  
Z. and Xia, *Annals of Statistics* (2022)

Thank you! Any questions?