1 Kingman's Coalescent

 \mathbf{a}

The only possible transition at time t is to go from N_t to $N_t - 1$, so the rate as a function of the number of particles k is

$$g(k, k-1) = \begin{cases} \binom{k}{2} & \text{for } 1 < k \le L \\ 0 & \text{otherwise.} \end{cases}$$

The generator is given by

$$\begin{split} (\mathcal{L}f)(k) &= \sum_{\substack{j \in S \\ j \neq k}} g(k,j) [f(j) - f(k)] \\ &= g(k,k-1) [f(k-1) - f(k)] \\ &= \binom{k}{2} \bigg[f(k-1) - f(k) \bigg], \end{split}$$

and the master equation

$$\frac{d}{dt}\pi_{t}(k) = \sum_{j \neq k} \pi_{t}(j)g(j,k) - \sum_{j \neq k} \pi_{t}(k)g(k,j)
= \pi_{t}(k+1)g(k+1,k)
= \pi_{t}(k+1)\binom{k+1}{2} - \pi_{t}(k)\binom{k}{2}.$$

The process has one absorbing state at k=1 since it needs two particle to coalesce, so once it reaches the one particle state nothing will happen. This also describes the stationary distribution given by $(1,0,0,\ldots)$, and since the process converges to this stationary distribution from any starting state, the process is ergodic.

b

Since the expected staying time in state k is 1/g(k, k-1), the total time to go from k=L to k=1 is the sum of all the expected times, i.e.,

$$\mathbb{E}(T) = \sum_{k=2}^{L} \left[\binom{k}{2} \right]^{-1} = \sum_{k=2}^{L} \frac{2! (k-2)!}{k!}$$

$$= 2 \sum_{k=2}^{L} \frac{1}{k(k-1)} = 2 \sum_{k=2}^{L} \left(\frac{1}{k-1} - \frac{1}{k} \right)$$

$$= 2 \left[\left(\frac{1}{1} - \frac{1}{2} \right) + \left(\frac{1}{2} - \frac{1}{3} \right) + \dots + \left(\frac{1}{1-L} - \frac{1}{L} \right) \right]$$

$$= 2 \left(1 - \frac{1}{L} \right)$$

 \mathbf{c}

The generator for the rescaled process N_t/L is given by

$$(\mathcal{L}f)(x) = {xL \choose 2} \left[f\left(x - \frac{1}{L}\right) - f(x) \right],$$

where x = k/L. Using Taylor expansion for f(x-1/L) around f(x) gives

$$f\left(x - \frac{1}{L}\right) = f(x) - \frac{1}{L}f'(x) + \frac{1}{2L^2}f''(x) + \dots, \text{ so}$$
$$(\mathcal{L}f)(x) = {xL \choose 2} \left[\frac{1}{2L^2}f''(x) - \frac{1}{L}f'(x)\right]$$

Slowing the process down just applies a factor of 1/L to the rates. The generator for $X_t^L = \frac{1}{L} N_{t/L}$ is then

$$(\mathcal{L}f)(x) = \frac{1}{L} \cdot \frac{(Lx)!}{2! (Lx - 2)!} \left[\frac{1}{2L^2} f''(x) - \frac{1}{L} f'(x) \right]$$

$$= \frac{(Xx)(Lx - 1)}{2X} \left[\frac{1}{2L^2} f''(x) - \frac{1}{L} f'(x) \right]$$

$$= \frac{1}{2} \left[\frac{x^2}{2L} f''(x) - x^2 f'(x) - \frac{x^2}{2L} f''(x) + \frac{x}{L} f'(x) \right]$$

$$= -\frac{x^2}{2} f'(x) \text{ as } L \to \infty.$$

Since the process X_t^L starts at x = L/L = 1 and always terminates at $x = 1/L \neq 0$, the state space (0,1] is open at 0 and $X_0 = 1$.

We see that this corresponds to the generator of a diffusion process with only the drift term but no diffusion which specifies the randomness. Therefore the process is deterministic because there is no randomness.

In order to compute X_t explicitly, we start with an SDE with no diffusion term and integrate both sides

$$dX_{t} = a(X_{t}, t)dt$$

$$\int_{0}^{t} dX_{s} = \int_{0}^{t} \frac{-X_{s}^{2}}{2} ds$$

$$\int_{0}^{t} X_{s}^{-2} dX_{s} = -\int_{0}^{t} \frac{1}{2} ds$$

$$X_{s}^{-1} \Big|_{0}^{t} = \frac{1}{2} s \Big|_{0}^{t}$$

By evaluating this integral and substituting $X_0 = 1$, we arrive at the explicit solution

$$X_t = \frac{2}{2+t} \, .$$

This is compatible with the result from part b) because the results are for differently scaled processes. i.e., 2/(2+t) describes the path of X_t^L which we have defined to be $\frac{1}{L}N_{t/L}$, whereas the mean time to absorption calculated in part b) is for the unscaled process N_t . By applying the appropriate scaling, we have

$$N_t = L \cdot \frac{2}{2 + tL} \,,$$

from which we can deduce that

$$N_t = 1 \implies t = \mathbb{E}(T) = 2\left(1 - \frac{1}{L}\right),$$

which is in line with our results from part b).

2 Orstein-Uhlenbeck process

 \mathbf{a}

Let $f(X_t) = X_t$, then $f'(X_t) = 1$ and $f''(X_t) = 0$, and using the evolution equation,

$$\frac{d}{dt}\mathbb{E}[X_t] = \mathbb{E}[-\alpha X_t] = -\alpha \mathbb{E}[X_t],$$

SO

$$\frac{d}{dt}\mathbb{E}[X_t] + \alpha \mathbb{E}[X_t] = 0$$

is the required ODE. Solving this gives $m(t) := \mathbb{E}[X_t] = Ae^{-\alpha t}$, where A is an arbitrary constant.

Similarly, let $f(X_t) = X_t^2$, then $f'(X_t) = 2X_t$ and $f''(X_t) = 2$, and

$$\frac{d}{dt}\mathbb{E}[X_t^2] = \mathbb{E}[-\alpha X_t \cdot 2X_t + \frac{1}{2}\sigma^2 \cdot 2] = -\alpha \mathbb{E}[X_t^2] + \sigma^2,$$

SO

$$\frac{d}{dt}\mathbb{E}[X_t^2] + 2\alpha\mathbb{E}[X_t^2] = \sigma^2,$$

which can be solved to give $\mathbb{E}[X_t^2] = \sigma^2/2\alpha + Be^{-2\alpha t}$, where B is an arbitrary constant. Therefore,

$$v(t) = \mathbb{E}[X_t^2] - (\mathbb{E}[X_t])^2 = \frac{\sigma^2}{2\alpha} + Be^{-2\alpha t} - (Ae^{-\alpha t})^2$$
$$= \frac{\sigma^2}{2\alpha} + Ce^{-2\alpha t}$$

3 Moran model and Wright-Fisher diffusion

 \mathbf{a}

The state space of $(X_t : t \ge 0)$ is $\{(x_1, x_2, \dots, x_L) : x_i \in \{1, 2, \dots, L\}\}$. The process is not reducible since once a type has been completely killed off, the process can no longer return to a state where one of the individuals is of that type.

The stationary distributions are $\{(x, x, \dots, x) : x \in \{1, 2, \dots, L\}\}$ because the process is only stationary once only a single type survives.

b

 $(N_t: t \ge 0)$ is a Markov process and its state space is $\{0, 1, \dots, L\}$.

The rates of this process for $x \neq 0$ or L is

$$\begin{cases} g(x, x+1) &= x \cdot \frac{L-x}{L-1} \\ g(x, x-1) &= (L-x) \cdot \frac{x}{L-1} \end{cases} ... 28?????$$

The reasoning is that in order to gain an individual, one of the x current individuals must impose its type on someone of a different type. Since it does so randomly, the probability of killing someone of a different type is $\frac{L-x}{L-1}$. We also get g(x, x-1) by applying similar logic.

The generator is

$$(\mathcal{L}f)(x) = \frac{x(L-x)}{L-1}[f(x+1) - 2f(x) + f(x-1)].$$

The process is, once again, not irreducible since x=0 and x=L are absorbing states. This also means that there exist infinitely many stationary distributions of the form

$$(\alpha, 0, \dots, 1 - \alpha)$$
, $\alpha \in [0, 1]$.

The limiting distribution as $t \to \infty$ is

$$\left(\frac{L-1}{L},0,\ldots,0,\frac{1}{L}\right)$$

since the probability of any one individual dominating is exactly the same due to symmetry, and there was only one individual of type k at t = 0.

 \mathbf{c}

By setting $f(N_t) = N_t$ and using the evolution equation, we obtain

$$\frac{d}{dt}\mathbb{E}[N_t] = \mathbb{E}\left[\frac{N_t(L - N_t)}{L - 1}\left((N_t + 1) - 2(N_t) + (N_t - 1)\right)\right] = 0,$$

which means that N_t is constant in time. Given that $N_0 = n$, we may deduce that $m_1(t) \equiv n$.

Similarly, by setting $f(N_t) = N_t^2$,

$$\frac{d}{dt}\mathbb{E}[N_t^2] = \mathbb{E}\left[\frac{N_t(L-N_t)}{L-1}\left((N_t^2 + 2N_t + 1) - 2(N_t^2) + (N_t^2 - 2N_t + 1)\right)\right]
= \mathbb{E}\left[\frac{2N_t(L-N_t)}{L-1}\right]
= \frac{2L}{L-1}\mathbb{E}[N_t] - \frac{2}{L-1}\mathbb{E}[N_t^2]$$

i.e.,

$$\frac{d}{dt}\mathbb{E}[N_t^2] + \frac{2}{L-1}\mathbb{E}[N_t^2] = \frac{2Ln}{L-1}.$$

Solving this gives

$$m_2(t) = \mathbb{E}[N_t^2] = Ln + Ae^{\frac{2t}{1-L}}$$

In the scaling limit $t \to \infty$, the exponential term vanishes so we just have $m_2(t) \to Ln$.