

Math 733 - Fall 2020

Homework 2

Due: 09/27, 10pm

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1. *Proof.* Assume X and Y are random variables from (Ω, \mathcal{F}) to (S, \mathcal{S}) , then

$$\{\omega : f(X(\omega)) \in B\} = \{\omega : X(\omega) \in f^{-1}(B)\} \in \mathcal{F}$$

If and only if $f^{-1}(B) \in \sigma(X)$, $Y = f(X)$ is measurable w.r.t $\sigma(X)$. □

2. *Proof.*

$$\begin{aligned} E[X^p] &= \int_0^1 y^p \cdot \mathbb{P}(X = y) dy \\ &= \int_0^{1-\varepsilon} y^p \cdot \mathbb{P}(X = y) dy + \int_{1-\varepsilon}^1 y^p \cdot \mathbb{P}(X = y) dy \quad \forall \varepsilon \in [0, 1] \end{aligned}$$

When $p \rightarrow \infty$,

$$\int_0^{1-\varepsilon} y^p \cdot \mathbb{P}(X = y) dy \rightarrow 0$$

$$\begin{aligned} \int_{1-\varepsilon}^1 y^p \cdot \mathbb{P}(X = y) dy &\leq \int_{1-\varepsilon}^1 y^p dy \\ &= \frac{1}{p+1} \cdot (1 - (1-\varepsilon)^{p+1}) \\ &\leq \frac{1}{1+p} \rightarrow 0 \quad (\text{as } p \rightarrow \infty) \end{aligned}$$

So, $E[X^p] = 0$ as $p \rightarrow \infty$. □

3. *Proof.*

(a) Consider *Derangement formula*

$$\mathbb{P}(X_n = 0) = \frac{D(n)}{n!}$$

where

$$D(n) = n! \cdot \sum_{k=2}^n \frac{(-1)^k}{k!}$$

Then

$$\mathbb{P}(X_n = 0) = \sum_{k=2}^n \frac{(-1)^k}{k!}$$

When $n \rightarrow \infty$,

$$\mathbb{P}(X_n = 0) \rightarrow \sum_{k=2}^{\infty} \frac{(-1)^k}{k!}$$

This is one of the expression of $\frac{1}{e}$.

$$\lim_{n \rightarrow \infty} \mathbb{P}(X_n = 0) = \frac{1}{e}$$

(b) Noticed that

$$\begin{aligned}\mathbb{P}(X_n = 1) &= \frac{\binom{n}{1} D(n-1)}{n!} = \frac{D(n-1)}{(n-1)!} = \mathbb{P}(X_{n-1} = 0) \\ \mathbb{P}(X_n = 2) &= \frac{\binom{n}{2} D(n-2)}{n!} = \frac{1}{2} \cdot \frac{D(n-2)}{(n-2)!} = \frac{1}{2} \mathbb{P}(X_n = 0) \\ \mathbb{P}(X_n = k) &= \frac{\binom{n}{k} D(n-k)}{n!} = \frac{1}{k} \mathbb{P}(X_n = k-1)\end{aligned}$$

So, we have

$$\mathbb{P}(X_n = k) = \frac{1}{k!} \cdot \mathbb{P}(X_n = 0)$$

$$\begin{aligned}E[X_n] &= \sum_{k=0}^n k \cdot \mathbb{P}(X_n = k) \\ &= \sum_{k=0}^n k \cdot \frac{1}{k!} \mathbb{P}(X_n = 0) \\ &= \mathbb{P}(X_n = 0) \cdot \sum_{k=1}^n \frac{1}{(k-1)!} \\ &= \sum_{k=2}^n \frac{(-1)^k}{k!} \cdot \sum_{k=1}^n \frac{1}{(k-1)!}\end{aligned}$$

By the way, when n is large enough, $E[X_n] \rightarrow 1$.

□

4. *Proof.* Consider the Integral form of *Cauchy-Schwarz inequality*.

Let $f = y \cdot \sqrt{\mathbb{P}(Y = y)}$, $g = \sqrt{\mathbb{P}(Y = y)}$.

By the non-negativity of Y

$$\left(\int \left(y \cdot \sqrt{\mathbb{P}(Y = y)} \right)^2 dy \right) \cdot \left(\int \left(\sqrt{\mathbb{P}(Y = y)} \right)^2 dy \right) \geq \left(\int y \cdot \mathbb{P}(Y = y) dy \right)^2$$

That is

$$\mathbb{P}(Y > 0) \geq \frac{(E[Y])^2}{E[Y^2]}$$

□

5. *Proof.* When $X_1 = X_2 = \dots = X_n = S$, $\prod_{j=1}^n g_j = 1$

Consider

$$\mathbb{P}(X_1 = x_1, X_2 = S, \dots, X_n = S) = g_1(x_1) \cdot C \quad C \text{ is a constant}$$

So we have

$$\mathbb{P}(X_2 = x_2, \dots, X_n = x_n | X_1 = x_1) = \frac{\prod_{j=2}^n g_j(x_j)}{C}$$

Thus, X_1 is independent. So does X_2, \dots, X_n . On the other hand, if they are independent, we have

$$\mathbb{P}(X_1 = x_1, \dots, X_n = x_n) = \prod_{j=1}^n \mathbb{P}(X_j = x_j)$$

Here, let $g_j(x) = \mathbb{P}(X_j = x)$ for all $1 \leq j \leq n$, they are non-negative functions from S to \mathbb{R} .

□

6. *Proof.* Consider the expression of ω in binary. The probability of $X_n = 1$ means the probability that the n th decimal place is 1. By the definition and arbitrariness of ω , we know

$$\mathbb{P}(X_n = 0) = \mathbb{P}(X_n = 1) = \frac{1}{2}$$

That means,

$$X_n \sim \text{Bernoulli}(1/2), \forall n \in \mathbb{N}$$

And their independence is obvious. □