## 基科 32 曾柯又 2013012266

**7.53** 考虑 
$$\phi_a = W + aU$$

则
$$E_{\theta}\phi_a = E_{\theta}W + aE_{\theta}U$$

$$= \tau(\theta)$$

故  $\forall a, \phi_a$  为  $\tau(\theta)$  的无偏估计量。

$$\overrightarrow{\mathbf{m}} Var_{\theta} \phi_a = Var(W + aU)$$

$$= Var_{\theta}W + aCov_{\theta}(W, U) + a^{2}Var_{\theta}U$$

如果存在  $\theta_0$  使得  $Cov_{\theta_0}(W,U) \neq 0$  则可以取 a 使得  $aCov_{\theta_0}(W,U) +$  $a^2 Var_{\theta_0}(U) < 0$ 

即若  $Cov_{\theta_0}(W, U) > 0$ 

$$\Re \ a \in \left(-\frac{Cov_{\theta_0}(W, U)}{Var_{\theta_0}W}, 0\right)$$

若 
$$Cov_{\theta_0}(W,U) < 0$$

$$\mathbb{R} \ a \in (0, -\frac{Cov_{\theta_0}(W, U)}{Var_{\theta_0}W})$$

此时  $Var_{\theta_0}\phi_a < Var_{\theta_0}W$ , 即 W 不为最佳无偏估计

## 7.57

(a) 
$$ET = \times P(\sum_{i=1}^{n} X_i > X_{n+1}|p) = h(p)$$

**(b)**  $X_1, ... X_{n+1}$  的联合分布为:

$$f(x_1, \dots, x_{n+1}) = p^{\sum_{i=1}^{n+1} x_i} (1-p)^{n+1-\sum_{i=1}^{n+1} x_i}$$

由因子分解定理  $\sum_{i=1}^{n+1} X_i$  为一个充分统计量,且由于伯努力分布为指 数分布族且参数空间包含了  $\mathbb{R}$  的开集,故  $T = \sum_{i=1}^{n+1} X_i$  为完全充分统计 量,因此  $\phi(\sum_{i=1}^{n+1} X_i) = E(T|\sum_{i=1}^{n+1} X_i)$  为最佳无偏估计。 记  $T_2 = \sum_{i=1}^{n+1} X_i$ 

记 
$$T_2 = \sum_{i=1}^{n+1} X_i$$

$$P(\sum_{i=1}^{n} X_{i} > X_{n+1} | T_{2} = t_{2}) = \frac{P(\sum_{i=1}^{n} X_{i} > X_{n+1}, T_{2} = t_{2})}{P(T_{2} = t_{2})}$$

$$T_{2} = 0 \text{ 时, } 显然 P(\sum_{i=1}^{n} X_{i} > X_{n+1}, T_{2} = 0) = 0$$

$$\Rightarrow P(\sum_{i=1}^{n} X_{i} > X_{n+1} | T_{2} = 0) = 0$$

$$T_{2} = 1 \text{ 时: } P(\sum_{i=1}^{n} X_{i} > X_{n+1}, T_{2} = 1) = (1 - p) \binom{n}{1} p (1 - p)^{n-1}$$

$$\overrightarrow{m} P(T_{2} = 1) = \binom{n+1}{1} p (1 - p)^{n}$$

$$\Rightarrow P(\sum_{i=1}^{n} X_{i} > X_{n+1} | T_{2} = 1) = \frac{n}{n+1}$$

$$T_{2} = 2 \text{ If } P(\sum_{i=1}^{n} X_{i} > X_{n+1}, T_{2} = 2) = (1 - p) \binom{n}{2} p^{2} (1 - p)^{n-2}$$

$$\overrightarrow{m} P(T_{2} = 2) = \binom{n+1}{2} p^{2} (1 - p)^{n-1}$$

$$\Rightarrow P(\sum_{i=1}^{n} X_{i} > X_{n+1} | T_{2} = 2) = \frac{n-1}{n+1}$$

$$T_{2} > 2 \text{ If }, \quad \overrightarrow{a} \text{ $\beta$ } \vec{f} \text{ If } P(\sum_{i=1}^{n} X_{i} > X_{n+1} | T_{2}) \equiv 1$$

$$\text{ But } \phi(T_{2}) = \begin{cases} 0 & T_{2} = 0 \\ \frac{n}{n+1} & T_{2} = 1 \\ \frac{n-1}{n+1} & T_{2} = 2 \end{cases}$$

$$\text{ But } \phi(T_{2}) = \begin{cases} 0 & T_{2} = 0 \\ \frac{n}{n+1} & T_{2} = 2 \\ 1 & T_{2} > 2 \end{cases}$$

7.58

(a) 
$$L(\theta|x) = (\frac{\theta}{2})^{|x|} (1-\theta)^{1-|x|}$$
  
 $\log L = |x| \log(\frac{\theta}{2}) + (1-|x|) \log(1-\theta)$   
由  $\frac{\partial}{\partial \theta} (\log L) = \frac{|x|}{\theta} - \frac{1-|x|}{1-\theta} = 0$   
 $\Rightarrow \theta = |x|$   
且  $\frac{|x|}{\theta} - \frac{1-|x|}{1-\theta}$  单调递减,故极值为极大值  
因此 $\hat{\theta} = |x|$  为极大似然估计

**(b)** 
$$ET(X) = 2P(X = 1) = \theta$$

(c) 由因子分解,
$$|x|$$
 是  $\theta$  的充分统计量, 考虑  $\phi(|x|) = E(T\big||x|)$ 

$$\phi(0) = E(T|x = 0) = 0$$

$$\phi(1) \ = \ E(T\big||x| \ = \ 1) \ = \ 2P(x \ = \ 1\big||x| \ = \ 1) \ = \ 2\frac{P(x=1)}{|x|=1} \ = \ 2\frac{(\frac{\theta}{2})}{2\frac{\theta}{2}} \ = \ 1 \ \ \mathbb{I}$$

$$\phi(|x|) = |x|$$
 是一个更好的  $\theta$  的无偏估计量