

```

import pandas as pd
import matplotlib.pyplot as plt
import seaborn as sns
from statsmodels.graphics.factorplots import interaction_plot
import numpy as np
import statsmodels.api as sm
from statsmodels.stats.outliers_influence import
variance_inflation_factor as vif
from mlxtend.feature_selection import SequentialFeatureSelector as SFS
from sklearn.linear_model import LinearRegression
from mlxtend.plotting import plot_sequential_feature_selection as
plot_sfs
from sklearn.preprocessing import StandardScaler
from sklearn.linear_model import LassoCV
from sklearn.linear_model import Lasso
import scipy.stats as stats
from statsmodels.stats.anova import anova_lm

```

Data and Description

The goal of this analysis is to develop a regression model that predicts job profitability for a restoration company known as "Complete Restoration Utah". Each record in the dataset represents an individual restoration job completed by this company in Utah, with the response variable being "Gross Profit". The predictor variables will be job type, whether the company will be paying out of pocket or through insurance (referred to as "IsSelfPay"), year the house being restored built, and the estimator that worked on the job.

The data were obtained directly from the database of the software this company uses for scheduling and invoicing, "JobSight". A CSV file containing this dataset can be accessed [here](#).

Understanding the factors that most strongly influence job profitability can help the company identify key areas for improvement and strategic focus. Additionally, being able to predict profitability for individual jobs enables the business to forecast performance with greater confidence.

The central research question guiding this analysis is:

"What characteristics best predict job profitability, and how accurately can these variables be used to forecast future profit?"

By answering this question, the analysis aims to help business leaders for this restoration company make more informed, data-driven decisions that enhance financial outcomes and operational efficiency.

Summary Statistics

```

profit = pd.read_csv("JobProfitability.csv", delimiter = ',')

profit = profit.drop('JobId', axis=1)
profit.head()

```

	GrossProfit	JobType	IsSelfPay	YearBuilt	EstimatorUserId
\					
0	9400.31	Water	0	1974.0	26.0
1	3018.74	Water	1	1977.0	26.0
2	2176.64	Reconstruction	0	1995.0	26.0
3	2002.39	Reconstruction	0	1971.0	26.0
4	1017.27	Water	0	1998.0	26.0
EstimatorName					
0	Lucas Collier				
1	Lucas Collier				
2	Lucas Collier				
3	Lucas Collier				
4	Lucas Collier				
profit.describe()					
	GrossProfit	IsSelfPay	YearBuilt	EstimatorUserId	
count	5056.000000	5056.000000	4704.000000	5055.000000	
mean	4552.479502	0.366891	1997.985544	74.532938	
std	10965.357485	0.482004	21.109503	84.522028	
min	-395109.470000	0.000000	1850.000000	2.000000	
25%	1022.602500	0.000000	1987.000000	27.000000	
50%	2294.025000	0.000000	2001.000000	39.000000	
75%	5183.937500	1.000000	2015.000000	60.000000	
max	193312.870000	1.000000	2025.000000	283.000000	

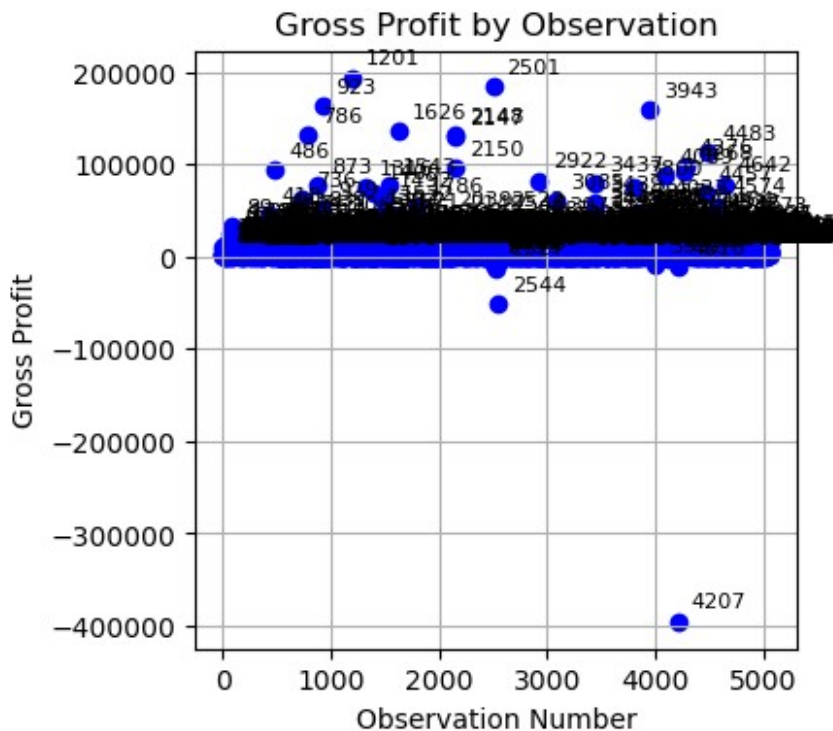
One thing that stands out here is the GrossProfit column. With an average of 4552.48 and a standard deviation of 10,965.36, it seems like there is a huge amount of variability in job profitability, which may make it hard to develop accurate predictions. However, the max value of 193,312.87 is over 17 standard deviations above the mean, and the min value of -395,109.47 is nearly 37 standard deviations away from the mean, so we definitely have some massive outliers here that are impacting the variability.

Response Variable Analysis

```
# Create scatter plot using the index
plt.figure(figsize=(4, 4))
plt.scatter(profit.index, profit['GrossProfit'], color='blue')

# Add labels to each point
for i in profit.index:
    plt.annotate(i, (i, profit.loc[i, 'GrossProfit']),
                 textcoords="offset points", xytext=(5,5), ha='left',
                 fontsize=8)
```

```
# Add titles and labels
plt.title('Gross Profit by Observation')
plt.xlabel('Observation Number')
plt.ylabel('Gross Profit')
plt.grid(True)
plt.show()
```



```
# Select specific rows by index
profit.loc[4207]
```

```
GrossProfit      -395109.47
JobType          Flood
IsSelfPay        0
YearBuilt        NaN
EstimatorUserId  2.0
EstimatorName    Jerry Pennock
Name: 4207, dtype: object
```

As noted earlier, there seems to be some extreme outliers in Gross Profit, so I decided to take a look at this right away. From this plot, it seems clear right away that the job 4207 is an extreme outlier. After looking into it a little more, this data point seems to be data entry error (the owner of the company entered the expenses for his pickleball court as a job with negative profit, so this doesn't reflect a real job). With that knowledge, I went ahead and removed this observation from the data set.

```

profit = profit[~profit.index.isin([4207])]

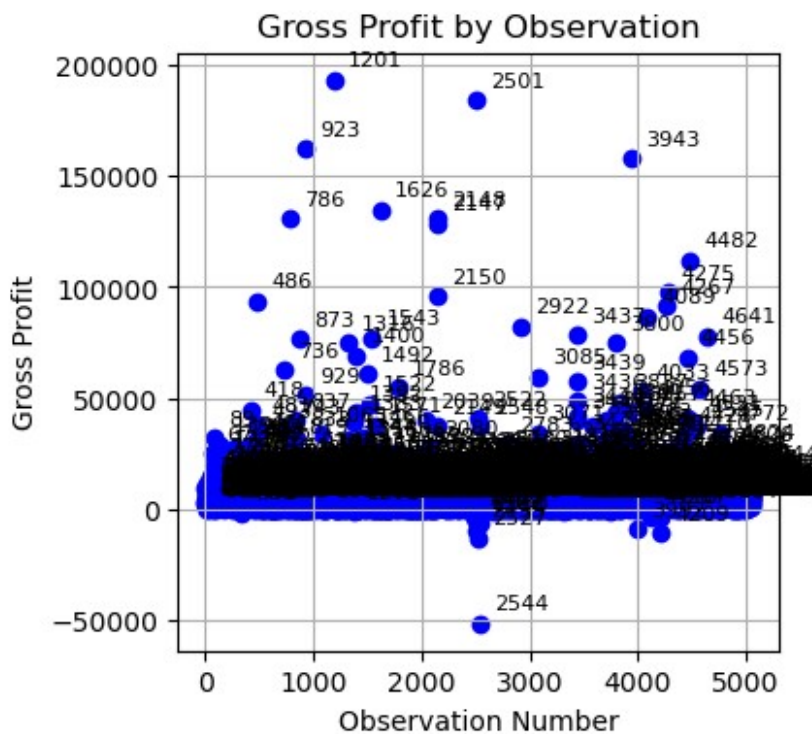
# Reset index just to be safe
profit = profit.reset_index(drop=True)

# Create scatter plot using the index
plt.figure(figsize=(4, 4))
plt.scatter(profit.index, profit['GrossProfit'], color='blue')

# Add labels to each point
for i in profit.index:
    plt.annotate(i, (i, profit.loc[i, 'GrossProfit']),
                 textcoords="offset points", xytext=(5,5), ha='left',
                 fontsize=8)

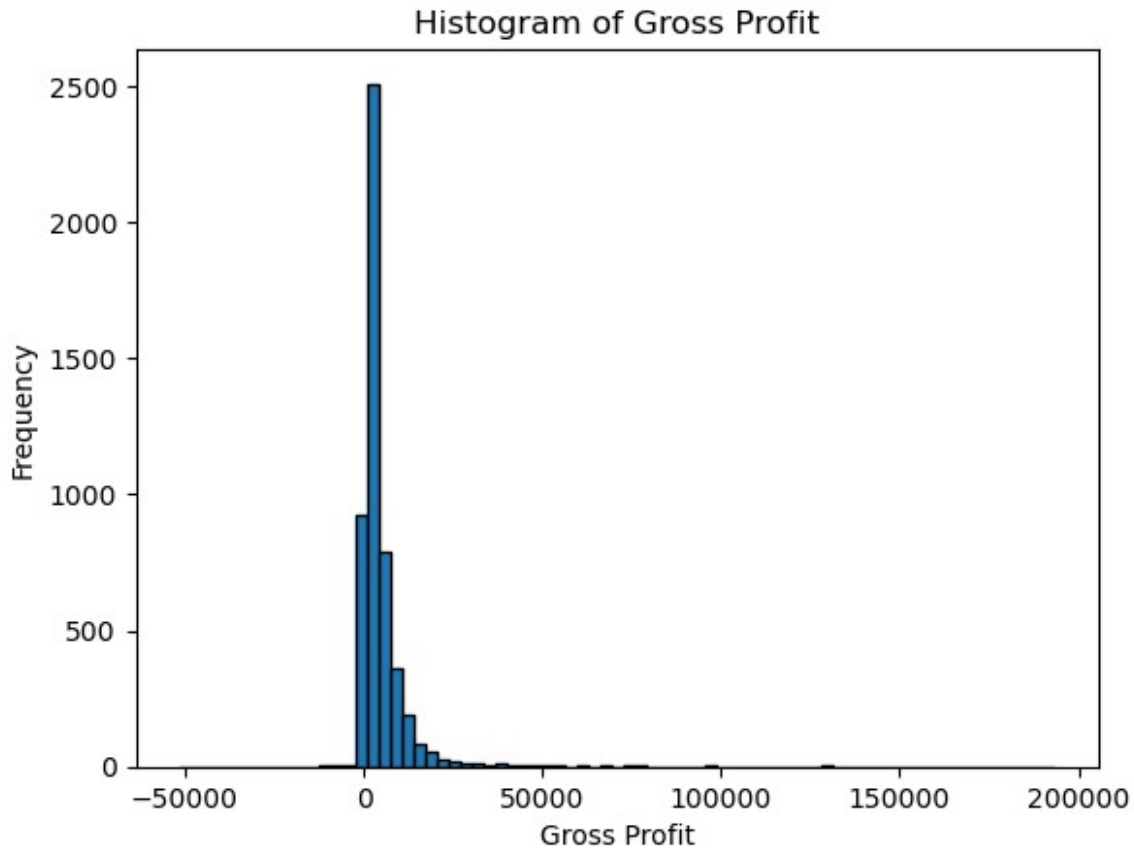
# Add titles and labels
plt.title('Gross Profit by Observation')
plt.xlabel('Observation Number')
plt.ylabel('Gross Profit')
plt.grid(True)
plt.show()

```



After removing observatoin 4207, there are still several observations far away from the mean, but not so drastically that they should be removed immediately. This is something we can revisit as we check each of the model assumptions.

```
plt.hist(profit['GrossProfit'], bins=75, edgecolor = 'black')
plt.xlabel("Gross Profit")
plt.ylabel("Frequency")
plt.title("Histogram of Gross Profit")
plt.show()
```



From this histogram alone there seems to be some right skew to the data, but in an effort to keep the model simple, we will hold off on a transformation at this point.

Continuous Predictor Analysis

```
profit[profit['YearBuilt'].isnull()]
```

	GrossProfit	JobType	IsSelfPay	YearBuilt
EstimatorUserId \				
53	1909.27	Water	0	NaN
26.0				
228	2559.94	Mitigation	0	NaN
26.0				
237	3096.67	Water	1	NaN
26.0				
373	6452.80	Reconstruction	0	NaN
26.0				

379	11623.19	Reconstruction	0	NaN
26.0				
...
..				
4712	1440.08	Sewer Loss	0	NaN
39.0				
4748	2610.94	Fire	1	NaN
39.0				
4781	1363.01	Fire	0	NaN
39.0				
4961	366.50	Mold	1	NaN
194.0				
4985	266.50	Mold	1	NaN
194.0				

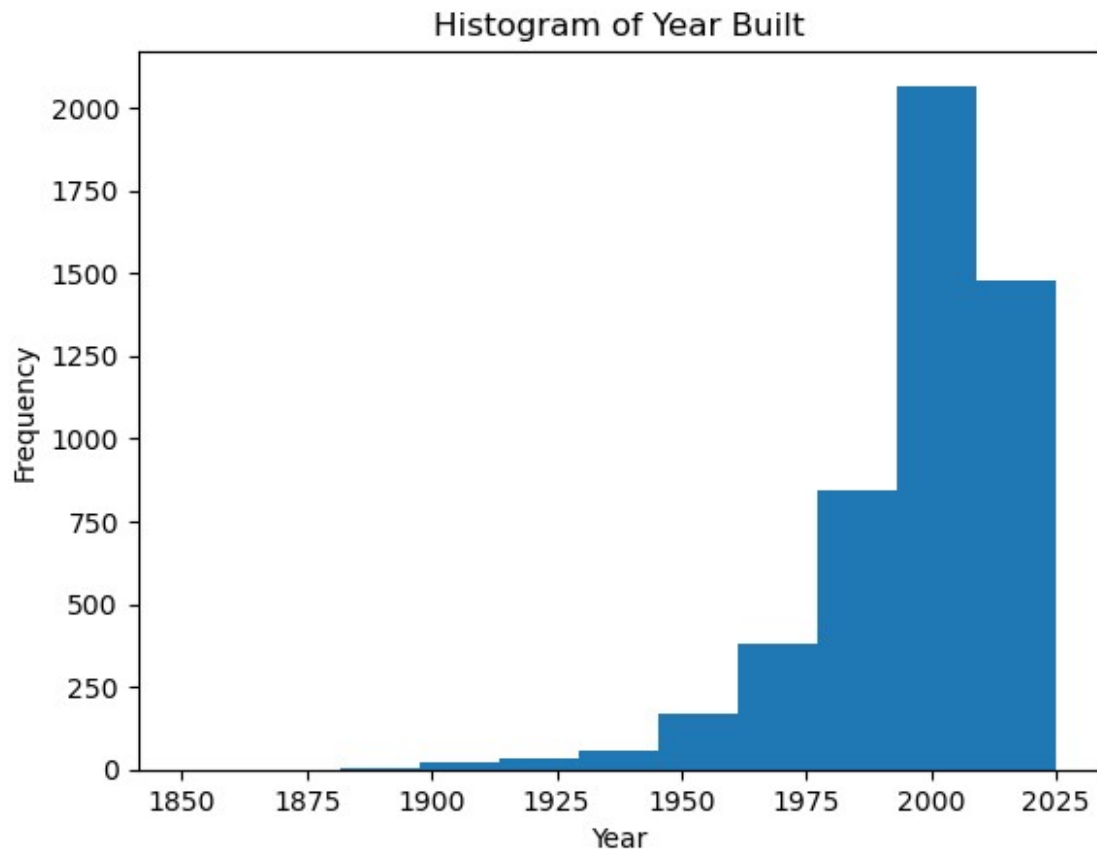
	EstimatorName
53	Lucas Collier
228	Lucas Collier
237	Lucas Collier
373	Lucas Collier
379	Lucas Collier
...	...
4712	Quade Bunker
4748	Quade Bunker
4781	Quade Bunker
4961	Julian Galarza
4985	Julian Galarza

[351 rows x 6 columns]

One thing to note with the "YearBuilt" column is that there are a lot of null values, either because the user didn't care to enter a year built or the home owner didn't know the year the house was built. According to the business, it's most likely the former. Because it seems totally random, I'll start by imputing these missing values with the overall median "YearBuilt".

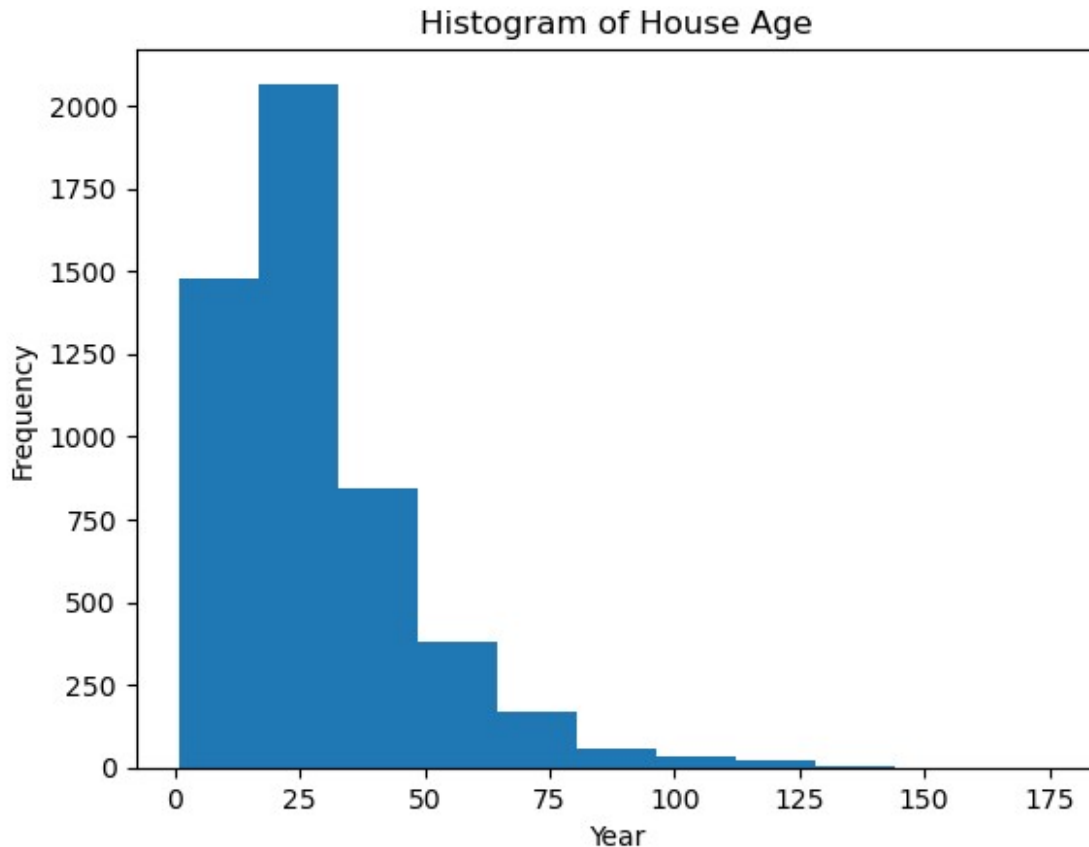
```
profit['YearBuilt'] =
profit['YearBuilt'].fillna(profit['YearBuilt'].median())

plt.hist(profit['YearBuilt'], bins=11)
plt.xlabel("Year")
plt.ylabel("Frequency")
plt.title("Histogram of Year Built")
plt.show()
```



There is definitely a left skew to the year of the house, which isn't surprising as we would expect there to be more newer houses than old. This data should be transformed now

```
profit['HomeAge'] = 2026 - profit['YearBuilt']  
  
plt.hist(profit['HomeAge'], bins=11)  
plt.xlabel("Age")  
plt.ylabel("Frequency")  
plt.title("Histogram of Home Age")  
plt.show()
```

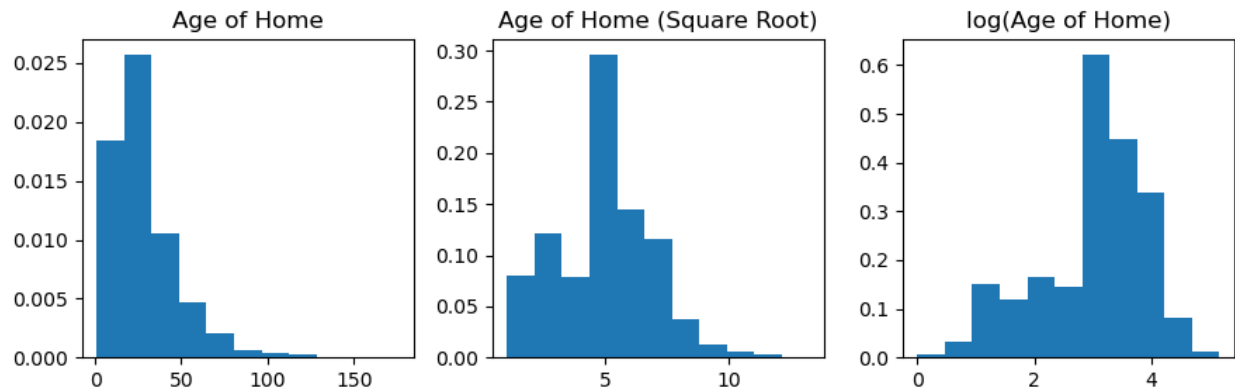


```
age_sqrt_trans = np.sqrt(profit['HomeAge'])
age_log_trans = np.log(profit['HomeAge'])

def hist_Y(variable, bin_size, ax, title):
    ax.hist(variable, density = True, bins = bin_size)
    ax.set_title(title)

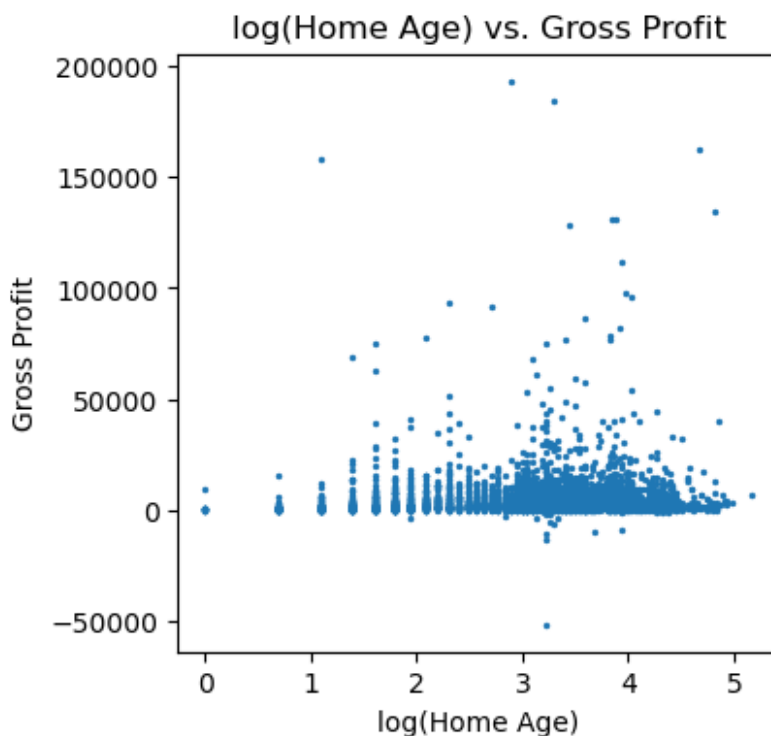
fig, axes = plt.subplots(1, 3, figsize = (9, 3))
hist_Y(profit['HomeAge'], 11, axes[0], 'Age of Home')
hist_Y(age_sqrt_trans, 11, axes[1], 'Age of Home (Square Root)')
hist_Y(age_log_trans, 11, axes[2], 'log(Age of Home)')
fig.tight_layout()
plt.show()

profit['HomeAgeTrans'] = age_log_trans
```

Rather than dealing with left skewed data, I decided to alter the data from the year the home was built to the age of the house to get right skewed data, which I found was easier to apply transformations to. I looked at both the square root of home age and the log of home age and found that a log transformation seemed to make a more normal distribution, so I decided to continue with that transformation.

```
fig = plt.figure(figsize = (4, 4))
plt.scatter(x = profit['HomeAgeTrans'],
            y = profit['GrossProfit'],
            s = 2)
plt.ylabel("Gross Profit")
plt.xlabel("log(Home Age)")
plt.title("log(Home Age) vs. Gross Profit")
plt.show()
```



```
print('Correlation Coefficient:',  
profit['HomeAgeTrans'].corr(profit['GrossProfit']))
```

Correlation Coefficient: 0.09953353407905821

Looking at the scatterplot of gross profit vs. log(home age), there doesn't seem to be any linear correlation here. The correlation coefficient supports this, showing a very weak positive linear correlation. Because home age is the only continuous predictor, there won't be any issue with multicollinearity between continuous predictors.

Categorical Predictors Analysis

```
profit['JobType'].value_counts()
```

JobType	
Mitigation	2015
Water	1438
Reconstruction	547
Mold	362
Flood	206
Sewer Loss	189
Fire	145
Other	137
Trauma	16

Name: count, dtype: int64

```
16 / 5056 * 100
```

0.31645569620253167

```
137 / 5056 * 100
```

2.709651898734177

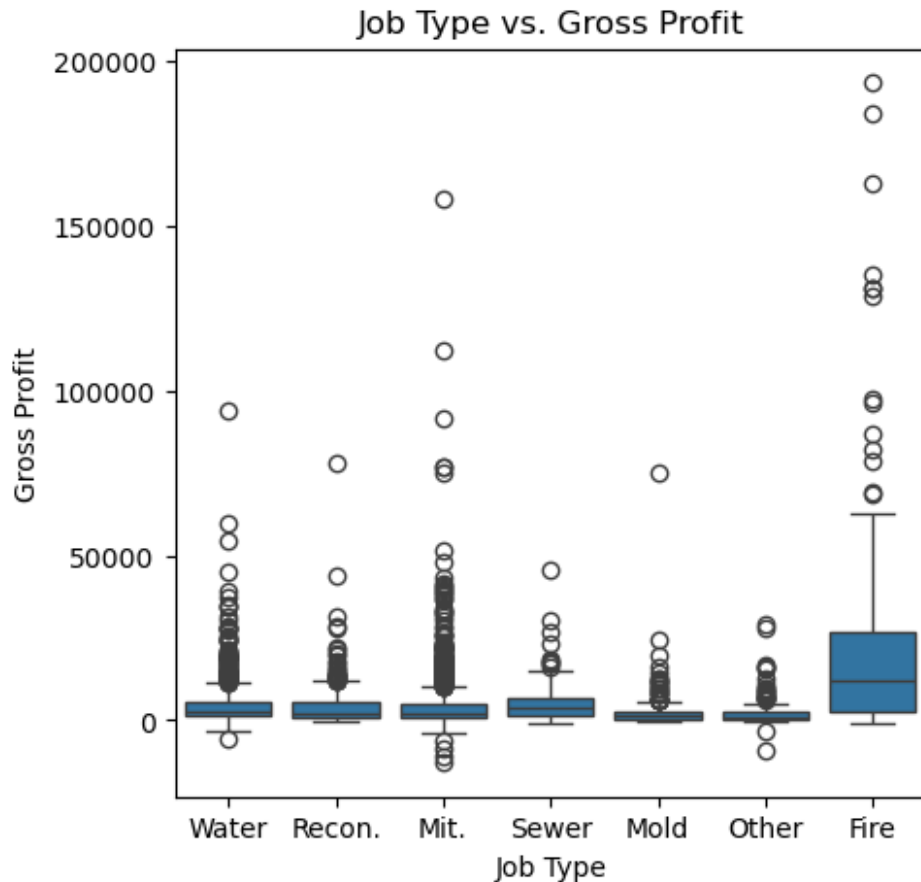
With the "Trauma" level for the JobType factor having very few observations (less than 1% of all rows are "Trauma"), any estimates from this level will have very wide standard errors and be generally unstable and unreliable. I think this group should be combined with the "Other" category to make for more stable estimates. The next lowest is "Other" which accounts for over 100 observations and about 3% of the data, so this and all other levels with more observations should be okay. Additionally, "Mitigation" and "Flood" represent the same job type, so those will need to be combined as well.

```
# Combine Trauma into Other  
profit['JobType'] = profit['JobType'].replace({'Trauma': 'Other'})  
profit['JobType'] = profit['JobType'].replace({'Flood': 'Mitigation'})  
  
# Optional: check counts again  
print(profit['JobType'].value_counts())
```

```
JobType
Mitigation      2221
Water           1438
Reconstruction   547
Mold             362
Sewer Loss       189
Other            153
Fire            145
Name: count, dtype: int64
```

```
profit['JobType'] = profit['JobType'].replace({
    'Reconstruction': 'Recon.',
    'Mitigation': 'Mit.',
    'Sewer Loss': 'Sewer'
})
```

```
plt.figure(figsize = (5,5))
sns.boxplot(x = 'JobType',
            y = 'GrossProfit',
            data = profit)
plt.xlabel('Job Type')
plt.ylabel('Gross Profit')
plt.title("Job Type vs. Gross Profit")
plt.show()
```



Looking at box plots for each level of Job Type, the median value of gross profit for each level is pretty close to the overall mean of gross profit, other than perhaps "Fire", which seems like it's a bit higher than the others. Most of the levels also seem like they are left skewed, so some transformations may be necessary here. There also seems to be some outliers within these levels, especially for the other level. "Mitigation" and "Fire" seem to have the most variability, which makes sense for mitigation given this level has the most observations. Overall, there may be some significance with the job type predictor due to "Fire" having a higher median than the others.

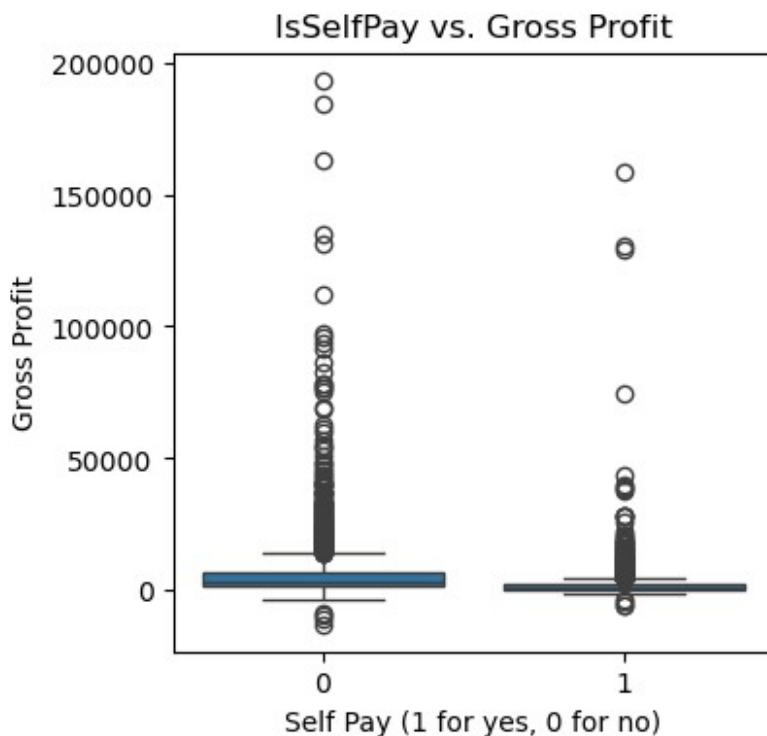
The median values for both Utah and other states seem to have roughly the same center with fairly normal shapes. Perhaps surprisingly, there is less variation in the other states group than there is in the Utah group. This may be due to the fact that there are more observations in the Utah group than the other states group. Overall, I'm not sure there is significance in this predictor based on the box plots.

```
profit['IsSelfPay'].value_counts()
```

```
IsSelfPay
0      3200
1      1855
Name: count, dtype: int64
```

There seems to be a sufficient number of rows of both jobs where the customer is paying with (0) and without insurance (1).

```
plt.figure(figsize = (4,4))
sns.boxplot(x = 'IsSelfPay',
            y = 'GrossProfit',
            data = profit)
plt.xlabel('Self Pay (1 for yes, 0 for no)')
plt.ylabel('Gross Profit')
plt.title("IsSelfPay vs. Gross Profit")
plt.show()
```



Both levels seem to be centered around the same point with relatively equal variability. Both seem to be left skewed, so a transformation may be needed here. Without any real difference in center between these two levels, I don't know how significant this predictor will be.

```
profit['EstimatorUserId'] = profit['EstimatorUserId'].astype('Int64')

est_summary = (
    profit.groupby(['EstimatorUserId', 'EstimatorName'], dropna=False)
        .size()
        .reset_index(name='NumJobs')
        .sort_values('NumJobs', ascending=False)
)

print(est_summary)
```

	EstimatorUserId	EstimatorName	NumJobs
3	27	Greg Prusak	1665
7	60	Jeffrey Peterson	859
4	39	Quade Bunker	668
11	283	Weston Clouse	625
2	26	Lucas Collier	496
6	53	Christian Martinez	305
10	194	Julian Galarza	177
9	78	Boyd Thatcher	126
1	20	Dan Goodwin	87
8	75	Todd Adams	21
5	46	Glenny P David	17
0	2	Jerry Pennock	8
12	<NA>	NaN	1

8 / 5056 * 100

0.15822784810126583

87 / 5056 * 100

1.720727848101266

The table above shows the list of estimators that have been on a job in this dataset. For the purposes of this first table, I have left both the Estimator Name and Estimator Id in for reference on which Id corresponds to which estimator, however for the rest of the analysis, I will be using the id as the levels for this predictor. The bottom 3 estimators have been on very few jobs (just over 1% combined), so I think it makes sense to combine these, as well as the one job with no estimator, into an "Other" category to reduce the variability of our estimates by having levels with so few observations. All other estimators have enough jobs that there shouldn't be a problem.

```
profit['EstimatorUserId'] = (
    profit['EstimatorUserId']
    .astype(str)
    .replace({'2': 'Other', '46': 'Other', '75': 'Other', '<NA>':
'Other'})
)
```

```
profit['EstimatorUserId'].value_counts()
```

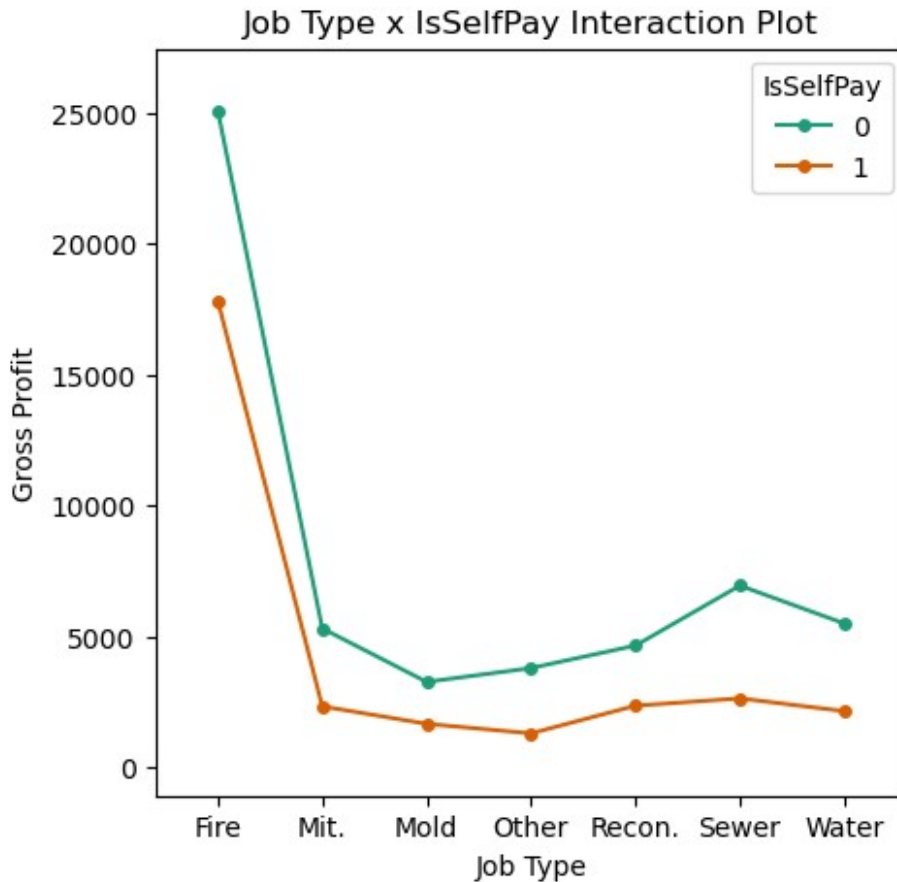
```
EstimatorUserId
27      1665
60      859
39      668
283     625
26      496
53      305
194     177
78      126
```



```

        colors = ["#1b9e77", "#d95f02"],
        ms = 8, # marker size
        ax = ax)
ax.set_xlabel('Job Type')
ax.set_ylabel('Gross Profit')
plt.title("Job Type x IsSelfPay Interaction Plot")
ax.legend(title = 'IsSelfPay')
<matplotlib.legend.Legend at 0x15076c03890>

```



The slope does not seem to change much for the two different levels of "IsSelfPay" between job types, so I don't believe an interaction is necessary here.

```

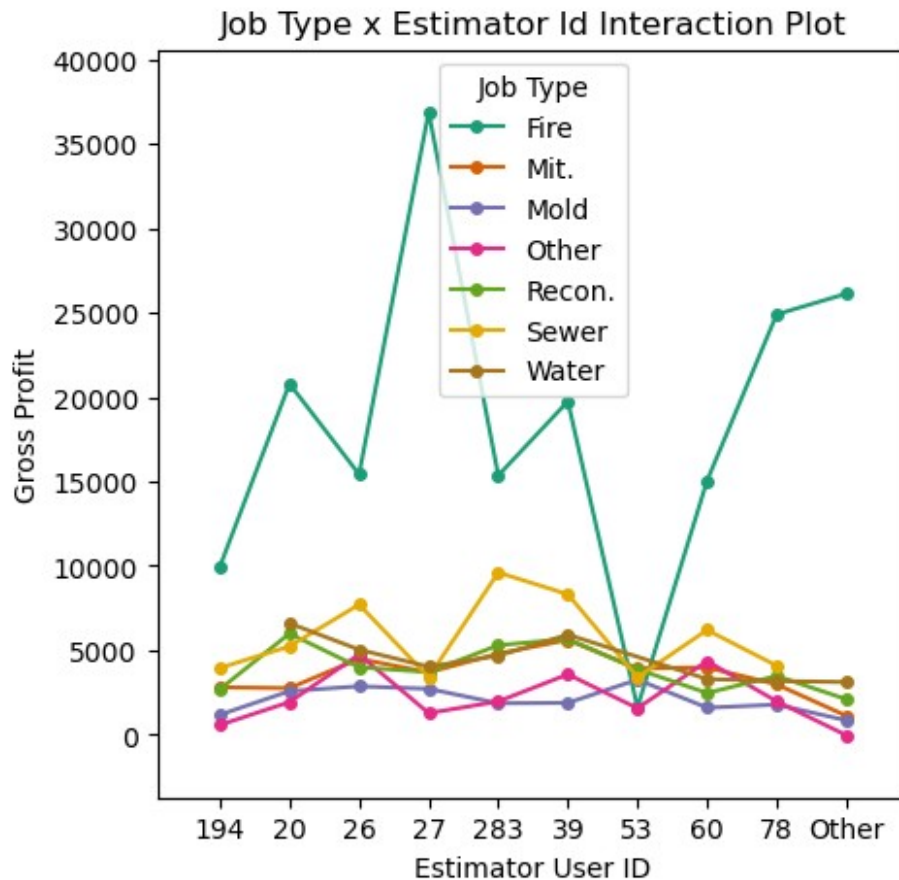
fig, ax = plt.subplots(figsize = (5, 5))
fig = interaction_plot(x = profit["EstimatorUserId"],
                      trace = profit["JobType"],
                      response = profit["GrossProfit"],
                      colors = ["#1b9e77", "#d95f02", "#7570b3",
                                "#e7298a", "#66a61e", "#e6ab02", "#a6761d"],
                      ms = 8, # marker size
                      ax = ax)
ax.set_xlabel('Estimator User ID')

```



```
ax.set_ylabel('Gross Profit')
plt.title("Job Type x Estimator Id Interaction Plot")
ax.legend(title = 'Job Type')
```

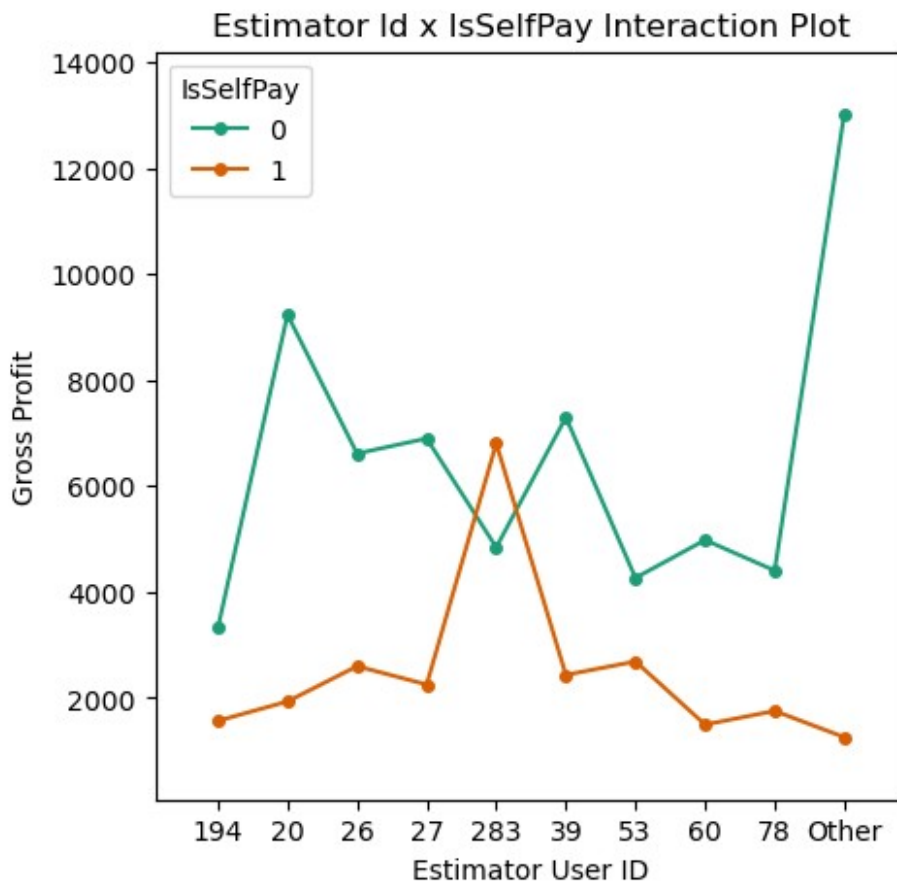
```
<matplotlib.legend.Legend at 0x1507441aad0>
```



The slopes for each job type across estimators seems to change pretty dramatically. I think the model will benefit from having an interaction between estimator and job type. This makes sense in a business context as well, as different estimators will likely have different evaluations for each job type.

```
fig, ax = plt.subplots(figsize = (5, 5))
fig = interaction_plot(x = profit["EstimatorUserId"],
                      trace = profit["IsSelfPay"],
                      response = profit["GrossProfit"],
                      colors = ["#1b9e77", "#d95f02"],
                      ms = 8, # marker size
                      ax = ax)
ax.set_xlabel('Estimator User ID')
ax.set_ylabel('Gross Profit')
plt.title("Estimator Id x IsSelfPay Interaction Plot")
ax.legend(title = 'IsSelfPay')
```

<matplotlib.legend.Legend at 0x1507766ed50>



The slope definitely seems to change for the levels of "IsSelfPay" among different estimators, so an interaction between estimator and "IsSelfPay" would likely improve our model.

```
# Bin HomeAgeTrans into 5 quantiles
profit['HomeAgeTrans_bin'] = pd.qcut(profit['HomeAgeTrans'], q=5,
duplicates='drop')

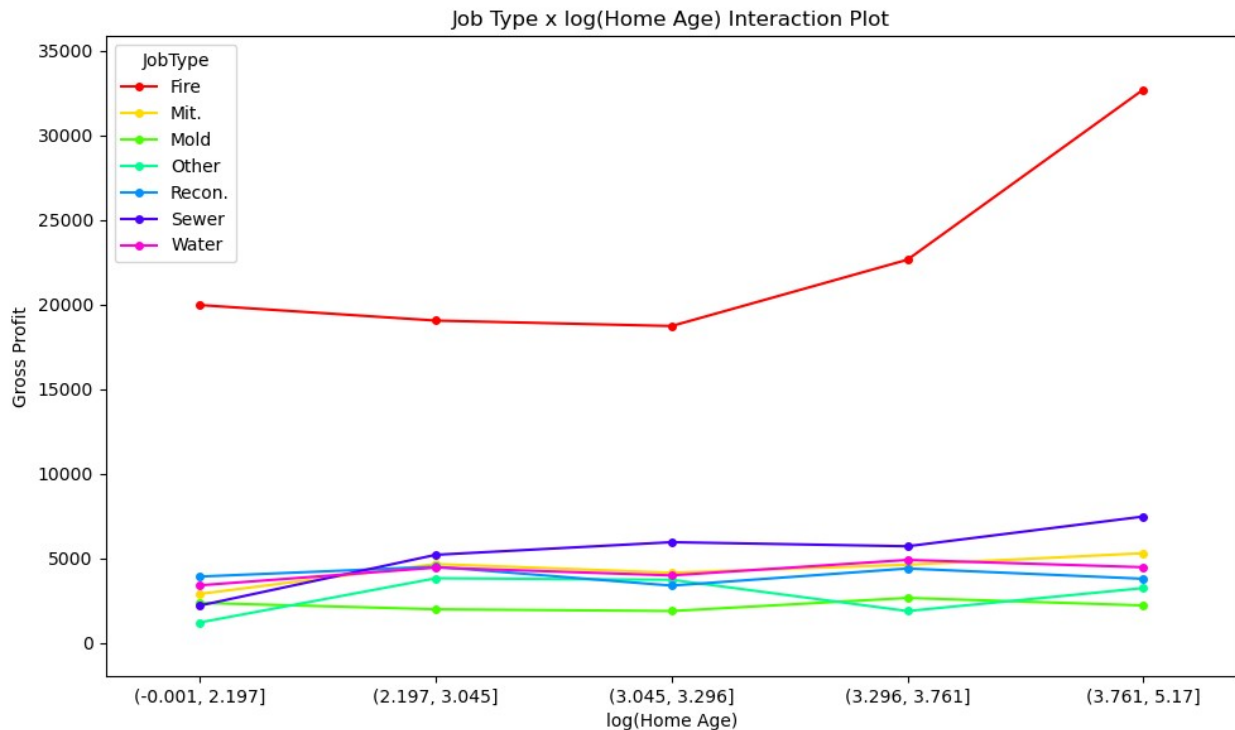
# Convert intervals to strings for plotting
profit['HomeAgeTrans_bin_str'] =
profit['HomeAgeTrans_bin'].astype(str)

fig, ax = plt.subplots(figsize=(10,6))
interaction_plot(
    x=profit['HomeAgeTrans_bin_str'],
    trace=profit['JobType'],
    response=profit['GrossProfit'],
    ms=8,
    ax=ax
)
```

```

ax.set_xlabel('log(Home Age)')
ax.set_ylabel('Gross Profit')
plt.title("Job Type x log(Home Age) Interaction Plot")
plt.tight_layout()
plt.show()

```



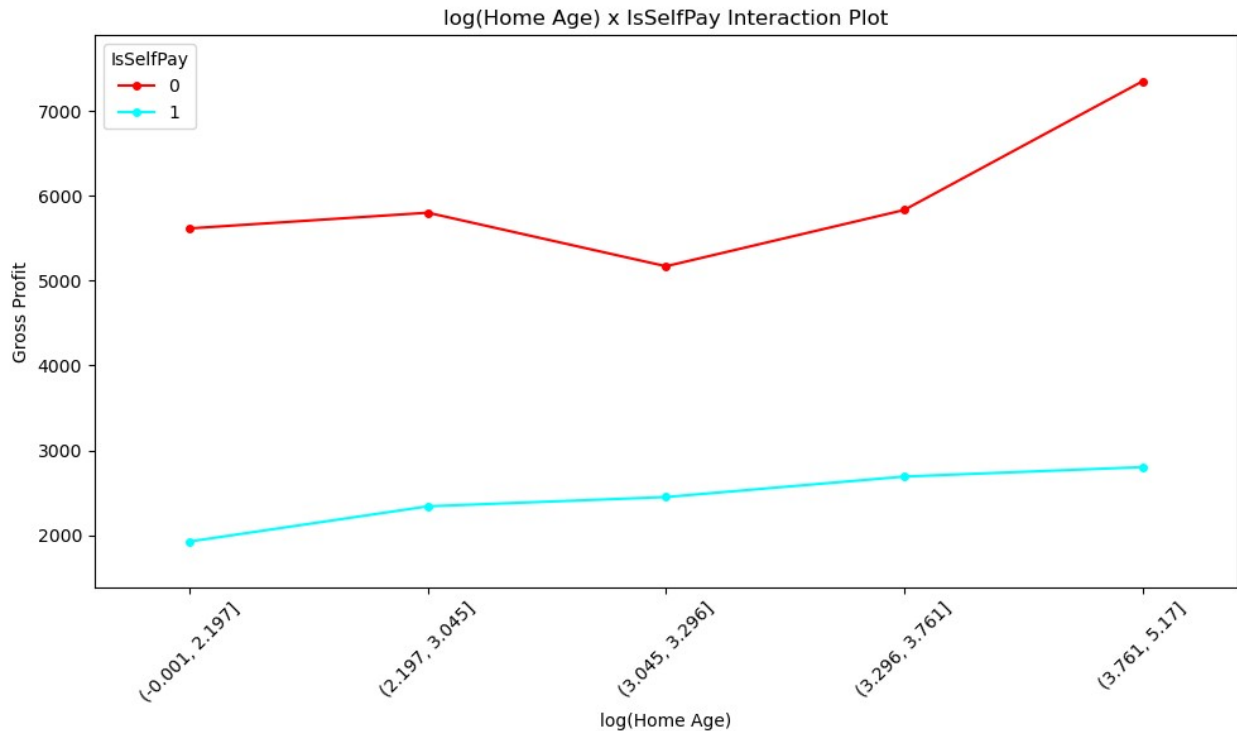
While most of the job types have relatively similar slopes across log(home age), the "Fire" and "Trauma" types differ significantly. I think an interaction would be beneficial here.

```

fig, ax = plt.subplots(figsize=(10,6))
interaction_plot(
    x=profit['HomeAgeTrans_bin_str'],
    trace=profit['IsSelfPay'],
    response=profit['GrossProfit'],
    ms=8,
    ax=ax
)

ax.set_xlabel('log(Home Age)')
ax.set_ylabel('Gross Profit')
plt.xticks(rotation=45)
plt.title("log(Home Age) x IsSelfPay Interaction Plot")
plt.tight_layout()
plt.show()

```



The interaction plot between log(home age) and "IsSelfPay" doesn't show too drastic of a difference between slopes, so I don't believe an interaction is necessary here.

Overall, it seems that our model could benefit from interactions between Job Type and Estimator Id, IsSelfPay and Estimator Id, and Job Type and log(Home Age).

```
profit_dummy = pd.get_dummies(profit, columns = ['JobType',
'IsSelfPay', 'EstimatorUserId'])

y = profit_dummy['GrossProfit']

# Make sure for each of the categorical predictors, you leave one of
the levels out of the
# model. This will be your baseline level
X = sm.add_constant(profit_dummy[['HomeAgeTrans',
'JobType_Fire', 'JobType_Mit.',
'JobType_Mold', 'JobType_Recon.', 'JobType_Sewer', 'JobType_Water',
'IsSelfPay_0',
'EstimatorUserId_194',
'EstimatorUserId_20', 'EstimatorUserId_26', 'EstimatorUserId_27',
'EstimatorUserId_283',
'EstimatorUserId_39',
'EstimatorUserId_53', 'EstimatorUserId_60', 'EstimatorUserId_78']])
X.apply(lambda col: col.astype(int) if col.dtype ==
bool else col))

mod = sm.OLS(y, X)
res = mod.fit()
```

```
profit['residuals'] = res.resid
profit['fittedvalues'] = res.fittedvalues
res.summary()
```

```
<class 'statsmodels.iolib.summary.Summary'>
"""
```

OLS Regression Results						
=====						
=====						
Dep. Variable:		GrossProfit	R-squared:			
0.161						
Model:		OLS	Adj. R-squared:			
0.158						
Method:		Least Squares	F-statistic:			
56.95						
Date:		Thu, 23 Oct 2025	Prob (F-statistic):			
1.37e-177						
Time:		14:55:20	Log-Likelihood:			
-52954.						
No. Observations:		5054	AIC:			
1.059e+05						
Df Residuals:		5036	BIC:			
1.061e+05						
Df Model:		17				
Covariance Type:		nonrobust				
=====						
=====						
		coef	std err	t	P> t	
[0.025 0.975]						

const		-1134.1911	1506.300	-0.753	0.452	-
4087.194	1818.812					
HomeAgeTrans		509.1094	152.694	3.334	0.001	
209.763	808.456					
JobType_Fire		2.028e+04	1010.072	20.074	0.000	
1.83e+04	2.23e+04					
JobType_Mit.		1404.3825	733.218	1.915	0.056	-
33.045	2841.810					
JobType_Mold		131.4083	837.867	0.157	0.875	-
1511.175	1773.992					
JobType_Recon.		863.6177	798.464	1.082	0.279	-
701.719	2428.955					
JobType_Sewer		2397.6839	941.809	2.546	0.011	
551.329	4244.039					
JobType_Water		1188.2156	740.482	1.605	0.109	-

263.451	2639.882					
IsSelfPay_0		2917.4410	275.949	10.572	0.000	
2376.461	3458.421					
EstimatorUserId_194		-905.4266	1437.141	-0.630	0.529	-
3722.847	1911.994					
EstimatorUserId_20		497.4698	1565.415	0.318	0.751	-
2571.425	3566.365					
EstimatorUserId_26		1036.3494	1330.559	0.779	0.436	-
1572.125	3644.823					
EstimatorUserId_27		1225.7433	1294.266	0.947	0.344	-
1311.580	3763.067					
EstimatorUserId_283		179.3769	1325.422	0.135	0.892	-
2419.027	2777.781					
EstimatorUserId_39		1598.4369	1316.618	1.214	0.225	-
982.708	4179.581					
EstimatorUserId_53		-117.0320	1376.588	-0.085	0.932	-
2815.744	2581.680					
EstimatorUserId_60		-110.3159	1307.554	-0.084	0.933	-
2673.691	2453.059					
EstimatorUserId_78		-841.0535	1486.325	-0.566	0.572	-
3754.896	2072.789					

```

=====
=====
Omnibus:                7358.089    Durbin-Watson:
1.907
Prob(Omnibus):          0.000    Jarque-Bera (JB):
3327195.935
Skew:                   8.650    Prob(JB):
0.00
Kurtosis:               127.501    Cond. No.
114.
=====
=====

```

Notes:

[1] Standard Errors assume that the covariance matrix of the errors is correctly specified.

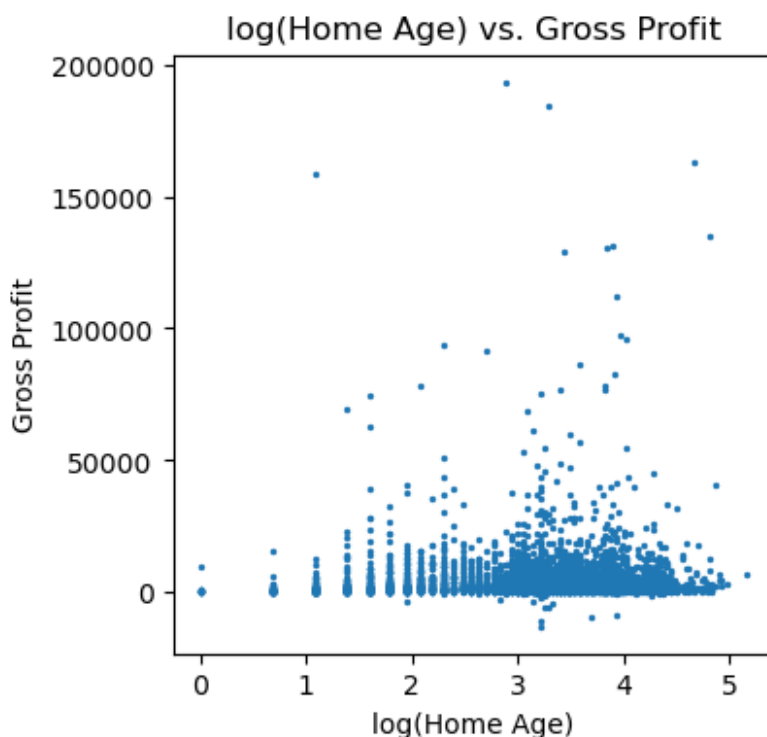
"""

From an initial look at this model, it seems like estimator is not a significant predictor. I'm guessing this will be left out of the model in the variable selection process. It also seems like the intervals for each job type are very wide, sometimes even a \$3000 range, so while these seem significant, I'm not sure how much value there will be in the predictions of this model. But we will see!

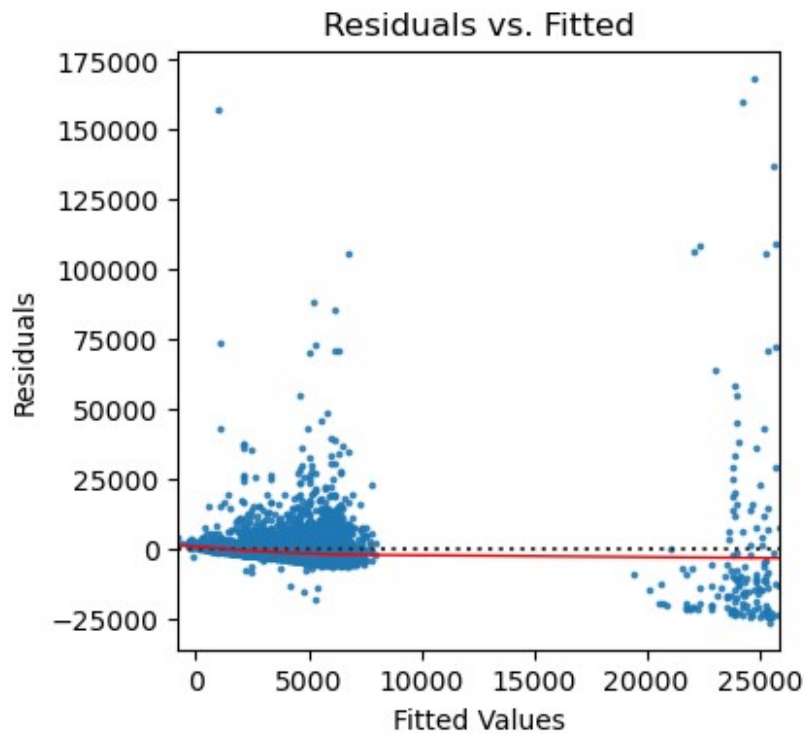
Assumptions Analysis (Cursory Check)

Xs vs. Y Are Linear

```
fig = plt.figure(figsize = (4, 4))
plt.scatter(x = profit['HomeAgeTrans'],
            y = profit['GrossProfit'],
            s = 2)
plt.ylabel("Gross Profit")
plt.xlabel("log(Home Age)")
plt.title("log(Home Age) vs. Gross Profit")
plt.show()
```



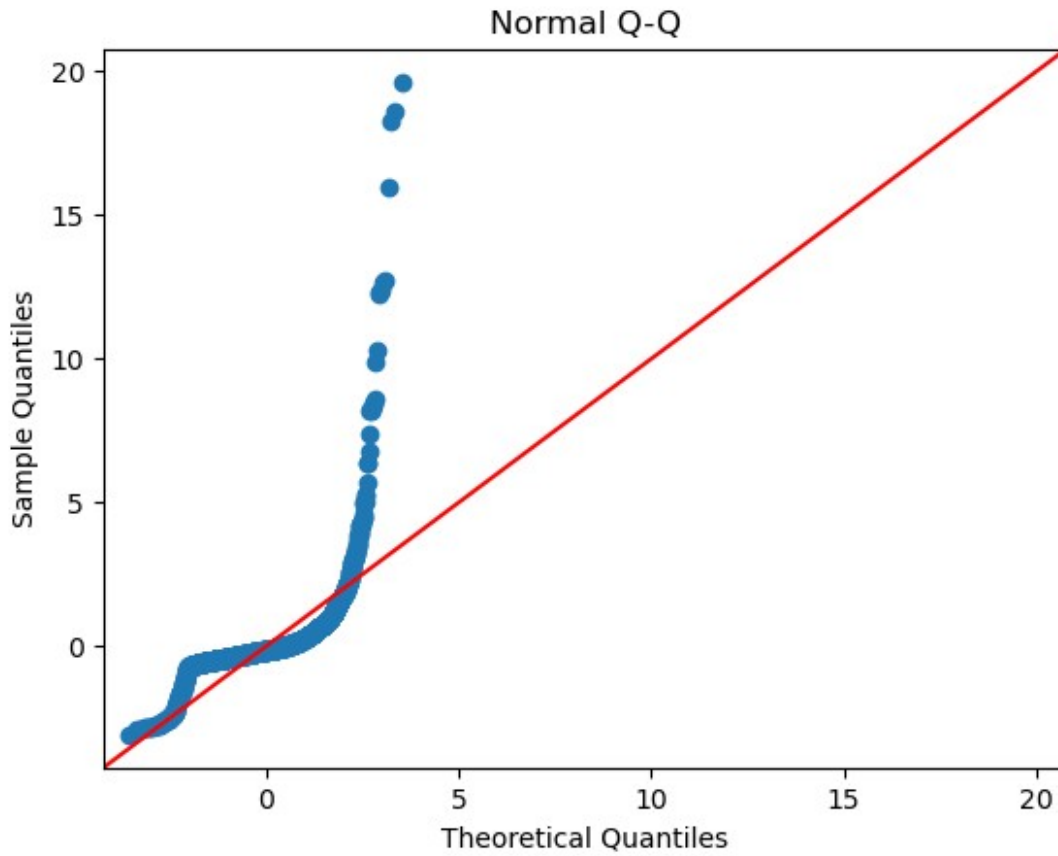
```
fig = plt.figure(figsize = (4, 4))
sns.residplot(x = profit['fittedvalues'],
              y = profit['residuals'],
              lowess = True,
              scatter_kws = {'s': 3},
              line_kws = {'color': 'red', 'lw': 1})
plt.title("Residuals vs. Fitted")
plt.ylabel("Residuals")
plt.xlabel("Fitted Values")
plt.show()
```



It doesn't seem that any line would better fit the scatterplot data than a linear trend line, and the line in the residuals vs. fitted plot is very straight. This assumption is met.

Residuals are normally distributed and centered at zero

```
sm.qqplot(profit['residuals'],  
           line = '45',  
           fit = True)  
plt.title("Normal Q-Q")  
plt.show()
```

From an initial look at this assumption, it's clear that the residuals are not normally distributed. A transformation will have to be applied here. After exploring some possible transformations, the best one I found was a log transformation of the response variable.

```
inv_trans = np.log(profit['GrossProfit'] +
np.abs(profit['GrossProfit'].min()) + 1)
profit['GrossProfitTrans'] = inv_trans

profit_dummy = pd.get_dummies(profit, columns = ['JobType',
'IsSelfPay', 'EstimatorUserId'])

y = profit_dummy['GrossProfitTrans']

# Make sure for each of the categorical predictors, you leave one of
the levels out of the
# model. This will be your baseline level
X = sm.add_constant(profit_dummy[['HomeAgeTrans',
'JobType_Fire', 'JobType_Mit.',
'JobType_Mold', 'JobType_Recon.', 'JobType_Sewer', 'JobType_Water',
'IsSelfPay_0',
'EstimatorUserId_194',
'EstimatorUserId_20', 'EstimatorUserId_26', 'EstimatorUserId_27',
'EstimatorUserId_283',
```

```
        'EstimatorUserId_39',  
'EstimatorUserId_53', 'EstimatorUserId_60', 'EstimatorUserId_78']]  
        .apply(lambda col: col.astype(int) if col.dtype ==  
bool else col))
```

```
mod = sm.OLS(y, X)  
res = mod.fit()
```

```
profit['residuals_trans'] = res.resid  
profit['fittedvalues_trans'] = res.fittedvalues  
res.summary()
```

```
<class 'statsmodels.iolib.summary.Summary'>  
"""
```

OLS Regression Results

```
=====
```

```
Dep. Variable:          GrossProfitTrans    R-squared:  
0.170  
Model:                      OLS    Adj. R-squared:  
0.167  
Method:                   Least Squares    F-statistic:  
60.56  
Date:                Thu, 23 Oct 2025    Prob (F-statistic):  
1.54e-188  
Time:                14:55:56    Log-Likelihood:  
-860.79  
No. Observations:          5054    AIC:  
1758.  
Df Residuals:              5036    BIC:  
1875.  
Df Model:                  17
```

```
Covariance Type:          nonrobust
```

```
=====
```

```
=====
```

		coef	std err	t	P> t
[0.025	0.975]				

const		9.4137	0.050	187.177	0.000
9.315	9.512				
HomeAgeTrans		0.0226	0.005	4.428	0.000
0.013	0.033				
JobType_Fire		0.5706	0.034	16.920	0.000
0.505	0.637				
JobType_Mit.		0.0647	0.024	2.641	0.008
0.017	0.113				

JobType_Mold	0.0045	0.028	0.160	0.873	-
0.050 0.059					
JobType_Recon.	0.0494	0.027	1.852	0.064	-
0.003 0.102					
JobType_Sewer	0.1336	0.031	4.247	0.000	
0.072 0.195					
JobType_Water	0.0689	0.025	2.786	0.005	
0.020 0.117					
IsSelfPay_0	0.1561	0.009	16.939	0.000	
0.138 0.174					
EstimatorUserId_194	-0.0075	0.048	-0.157	0.875	-
0.102 0.087					
EstimatorUserId_20	-0.0859	0.052	-1.643	0.100	-
0.188 0.017					
EstimatorUserId_26	0.0856	0.044	1.927	0.054	-
0.001 0.173					
EstimatorUserId_27	0.0707	0.043	1.636	0.102	-
0.014 0.155					
EstimatorUserId_283	0.0275	0.044	0.621	0.535	-
0.059 0.114					
EstimatorUserId_39	0.1073	0.044	2.440	0.015	
0.021 0.193					
EstimatorUserId_53	0.0399	0.046	0.867	0.386	-
0.050 0.130					
EstimatorUserId_60	0.0272	0.044	0.623	0.534	-
0.058 0.113					
EstimatorUserId_78	-0.0169	0.050	-0.340	0.734	-
0.114 0.080					

```

=====
=====
Omnibus:                6400.446    Durbin-Watson:
1.934
Prob(Omnibus):          0.000    Jarque-Bera (JB):
13961468.348
Skew:                   -5.931    Prob(JB):
0.00
Kurtosis:               260.213    Cond. No.
114.
=====
=====

```

Notes:

[1] Standard Errors assume that the covariance matrix of the errors is correctly specified.

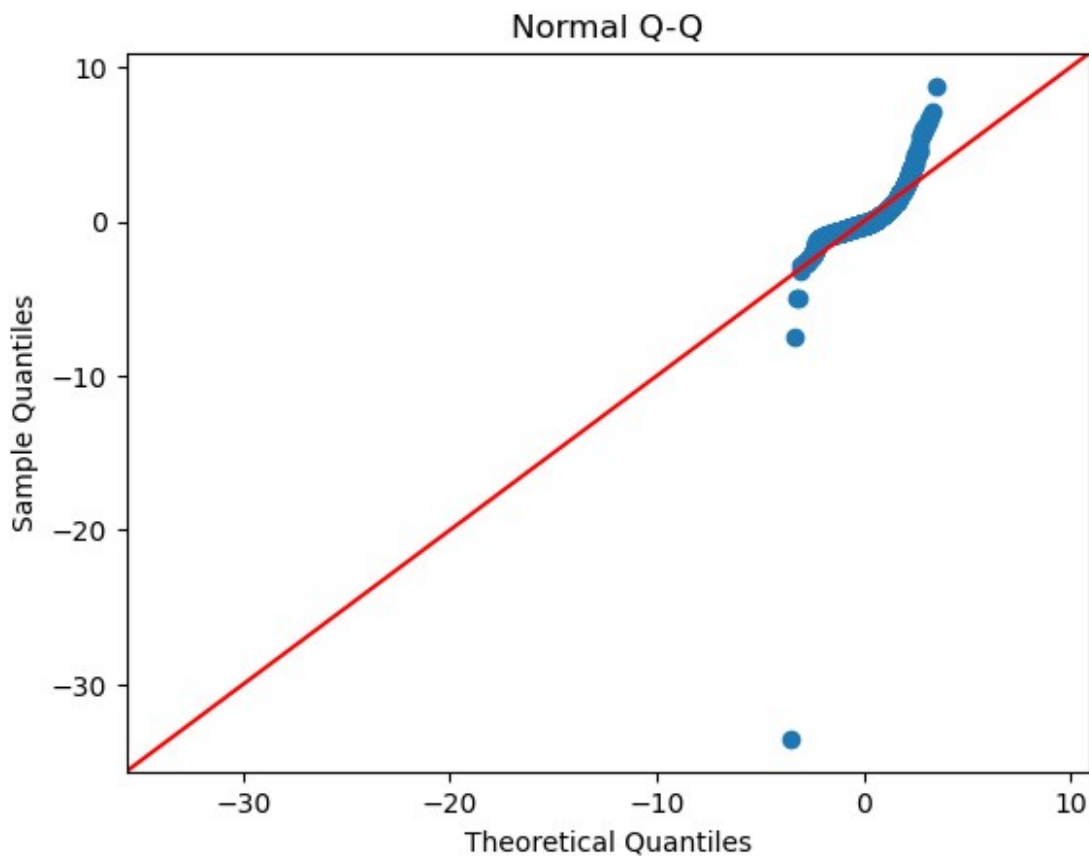
"""

```

sm.qqplot(profit['residuals_trans'],
           line = '45',
           fit = True)
plt.title("Normal Q-Q")

```

```
plt.show()
```



```
min_idx = profit['residuals_trans'].idxmin()  
profit.loc[[min_idx]]
```

	GrossProfit	JobType	IsSelfPay	YearBuilt	EstimatorUserId
EstimatorName \					
2527	-12808.53	Mit.	0	2001.0	20 Dan Goodwin

	HomeAge	HomeAgeTrans	HomeAgeTrans_bin	HomeAgeTrans_bin_str \
2527	25.0	3.218876	(3.045, 3.296]	(3.045, 3.296]

	residuals	fittedvalues	GrossProfitTrans	residuals_trans \
2527	-18132.392228	5323.862228	0.0	-9.621234

	fittedvalues_trans
2527	9.621234

After applying the transformation, it looks like there is one residual that is very far away from all the other residuals. I identified this point as observation 2544. After speaking with the company, this was a job that was created before invoicing was put in place, so gross profit was calculated

differently than it is now. The company advised that I remove this data point, as it no longer reflects accurately.

```
profit = profit[~profit.index.isin([2544])]

profit = profit.drop('GrossProfitTrans', axis=1)
profit = profit.drop('residuals_trans', axis=1)
profit = profit.drop('fittedvalues_trans', axis=1)

inv_trans = np.log(profit['GrossProfit'] +
np.abs(profit['GrossProfit'].min()) + 1)
profit['GrossProfitTrans'] = inv_trans

profit_dummy = pd.get_dummies(profit, columns = ['JobType',
'IsSelfPay', 'EstimatorUserId'])

y = profit_dummy['GrossProfitTrans']

# Make sure for each of the categorical predictors, you leave one of
the levels out of the
# model. This will be your baseline level
X = sm.add_constant(profit_dummy[['HomeAgeTrans',
                                'JobType_Fire', 'JobType_Mit.',
'JobType_Mold', 'JobType_Recon.', 'JobType_Sewer', 'JobType_Water',
                                'IsSelfPay_0',
                                'EstimatorUserId_194',
'EstimatorUserId_20', 'EstimatorUserId_26', 'EstimatorUserId_27',
'EstimatorUserId_283',
                                'EstimatorUserId_39',
'EstimatorUserId_53', 'EstimatorUserId_60', 'EstimatorUserId_78']])
    .apply(lambda col: col.astype(int) if col.dtype ==
bool else col))

mod = sm.OLS(y, X)
res = mod.fit()

profit['residuals_trans'] = res.resid
profit['fittedvalues_trans'] = res.fittedvalues
res.summary()

<class 'statsmodels.iolib.summary.Summary'>
"""
                                OLS Regression Results
=====
=====
Dep. Variable:          GrossProfitTrans    R-squared:
0.170
Model:                                OLS    Adj. R-squared:
0.167
Method:                Least Squares    F-statistic:
```

60.56
Date: Thu, 23 Oct 2025 Prob (F-statistic):
1.54e-188
Time: 14:56:03 Log-Likelihood:
-860.79
No. Observations: 5054 AIC:
1758.
Df Residuals: 5036 BIC:
1875.
Df Model: 17

Covariance Type: nonrobust

		coef	std err	t	P> t	
[0.025 0.975]						

const		9.4137	0.050	187.177	0.000	
9.315	9.512					
HomeAgeTrans		0.0226	0.005	4.428	0.000	
0.013	0.033					
JobType_Fire		0.5706	0.034	16.920	0.000	
0.505	0.637					
JobType_Mit.		0.0647	0.024	2.641	0.008	
0.017	0.113					
JobType_Mold		0.0045	0.028	0.160	0.873	-
0.050	0.059					
JobType_Recon.		0.0494	0.027	1.852	0.064	-
0.003	0.102					
JobType_Sewer		0.1336	0.031	4.247	0.000	
0.072	0.195					
JobType_Water		0.0689	0.025	2.786	0.005	
0.020	0.117					
IsSelfPay_0		0.1561	0.009	16.939	0.000	
0.138	0.174					
EstimatorUserId_194		-0.0075	0.048	-0.157	0.875	-
0.102	0.087					
EstimatorUserId_20		-0.0859	0.052	-1.643	0.100	-
0.188	0.017					
EstimatorUserId_26		0.0856	0.044	1.927	0.054	-
0.001	0.173					
EstimatorUserId_27		0.0707	0.043	1.636	0.102	-
0.014	0.155					
EstimatorUserId_283		0.0275	0.044	0.621	0.535	-
0.059	0.114					
EstimatorUserId_39		0.1073	0.044	2.440	0.015	
0.021	0.193					

EstimatorUserId_53	0.0399	0.046	0.867	0.386	-
0.050	0.130				
EstimatorUserId_60	0.0272	0.044	0.623	0.534	-
0.058	0.113				
EstimatorUserId_78	-0.0169	0.050	-0.340	0.734	-
0.114	0.080				

```
=====
=====
Omnibus:                6400.446    Durbin-Watson:
1.934
Prob(Omnibus):          0.000    Jarque-Bera (JB):
13961468.348
Skew:                   -5.931    Prob(JB):
0.00
Kurtosis:               260.213    Cond. No.
114.
=====
=====
```

Notes:

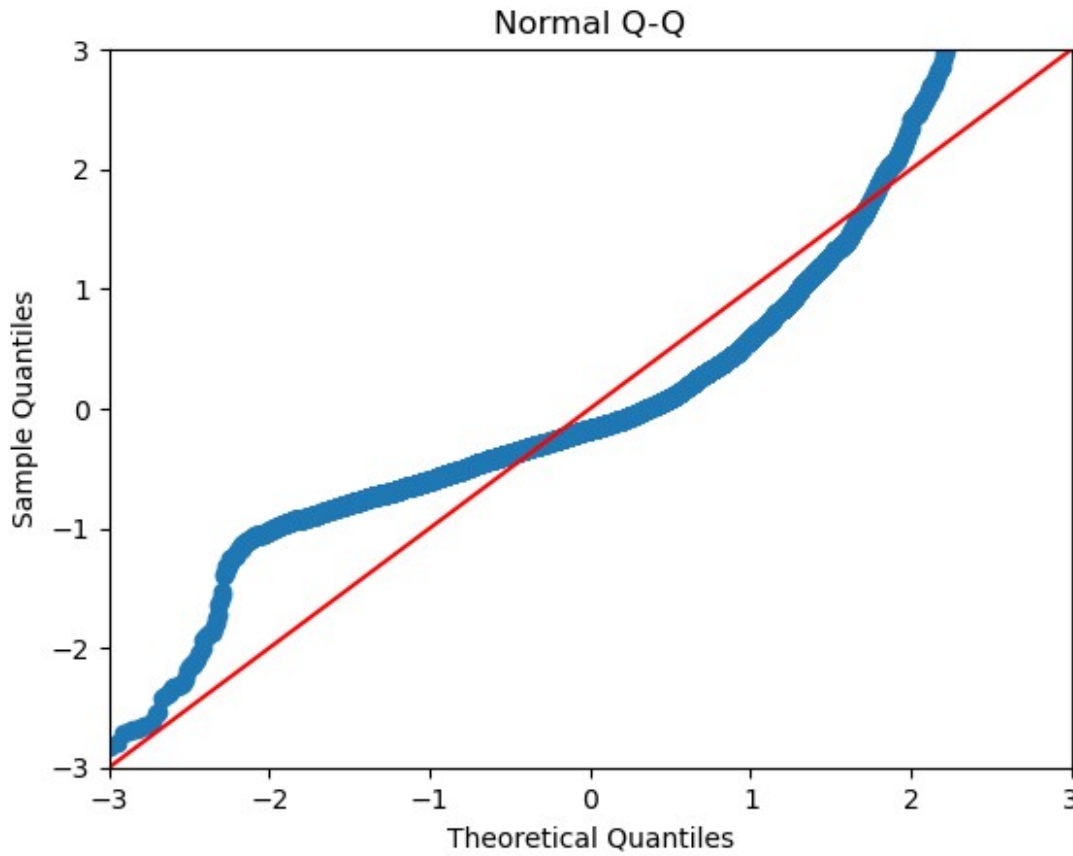
[1] Standard Errors assume that the covariance matrix of the errors is correctly specified.

"""

```
sm.qqplot(profit['residuals_trans'],
           line = '45',
           fit = True)
plt.title("Normal Q-Q")

plt.xlim(-3, 3)
plt.ylim(-3, 3)

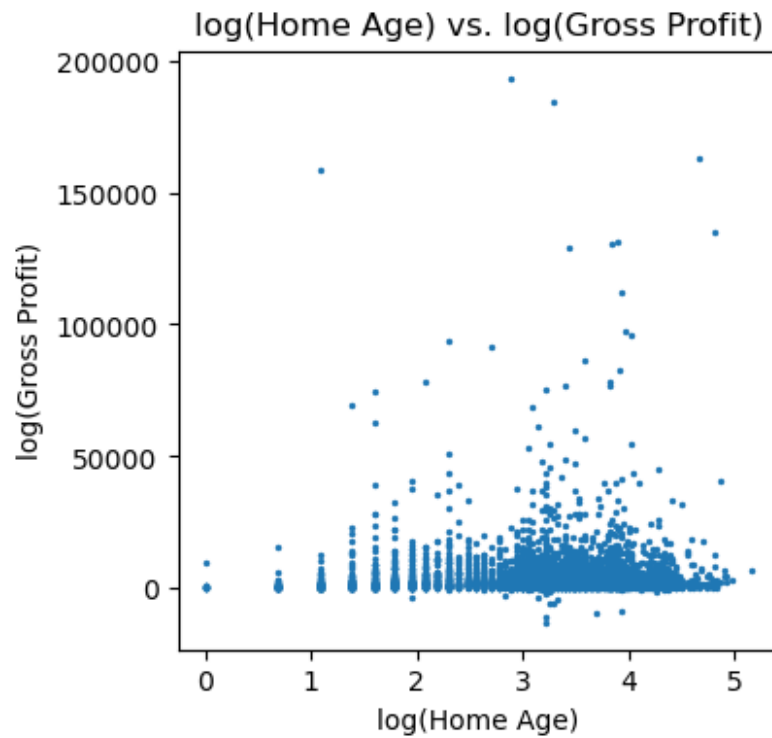
plt.show()
```



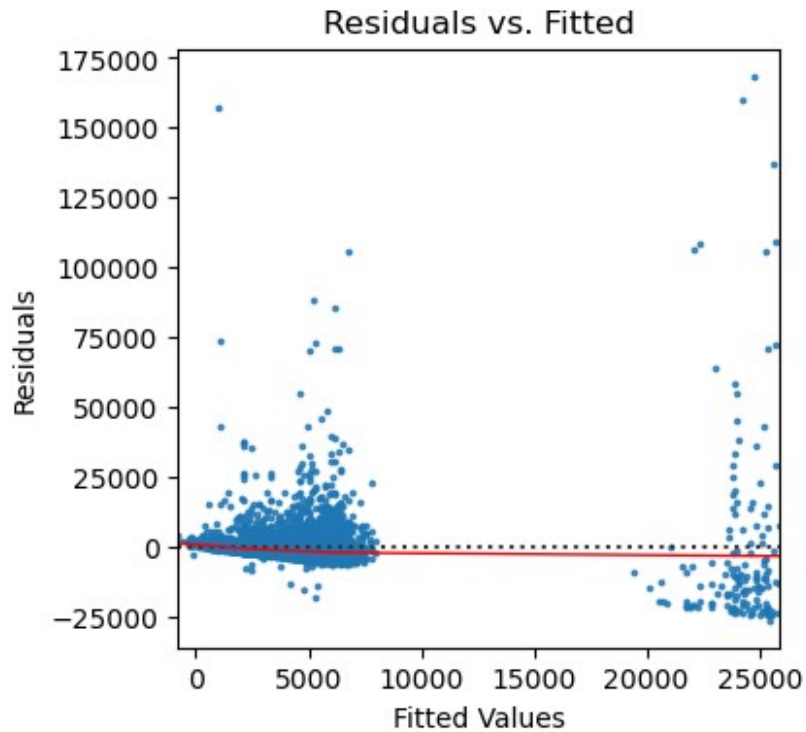
While this has improved the assumption, it's still clear that the assumption has not been met. I recall from Simple Linear Regression that transformations can be applied to the predictor variable as well, I'm not sure if that applies in Multiple Linear Regression too or how you would go about that, but that would be the next thing I try.

Xs vs. Y Are Linear (After Transformation)

```
fig = plt.figure(figsize = (4, 4))
plt.scatter(x = profit['HomeAgeTrans'],
            y = profit['GrossProfit'],
            s = 2)
plt.ylabel("log(Gross Profit)")
plt.xlabel("log(Home Age)")
plt.title("log(Home Age) vs. log(Gross Profit)")
plt.show()
```

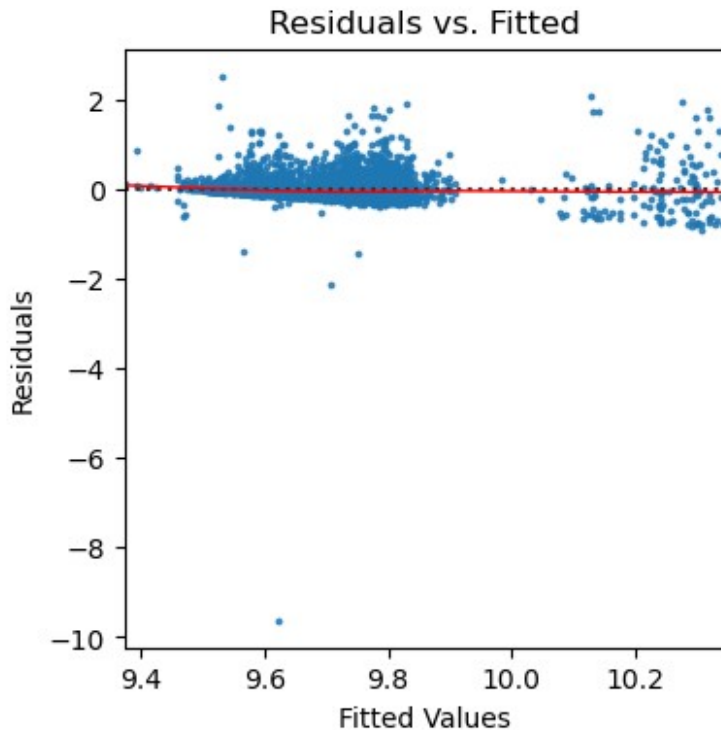
```
fig = plt.figure(figsize = (4, 4))
sns.residplot(x = profit['fittedvalues'],
              y = profit['residuals'],
              lowess = True,
              scatter_kws = {'s': 3},
              line_kws = {'color': 'red', 'lw': 1})
plt.title("Residuals vs. Fitted")
plt.ylabel("Residuals")
plt.xlabel("Fitted Values")
plt.show()
```



The linearity assumption seems to hold true after the transformation.

The residuals have constant variance across all values of x

```
fig = plt.figure(figsize = (4, 4))
sns.residplot(x = profit['fittedvalues_trans'],
              y = profit['residuals_trans'],
              lowess = True,
              scatter_kws = {'s': 3},
              line_kws = {'color': 'red', 'lw': 1})
plt.title("Residuals vs. Fitted")
plt.ylabel("Residuals")
plt.xlabel("Fitted Values")
plt.show()
```



There doesn't appear to be any funneling shape with the Residuals vs. Fitted plot. I think this assumption has been met.

No Influential Points

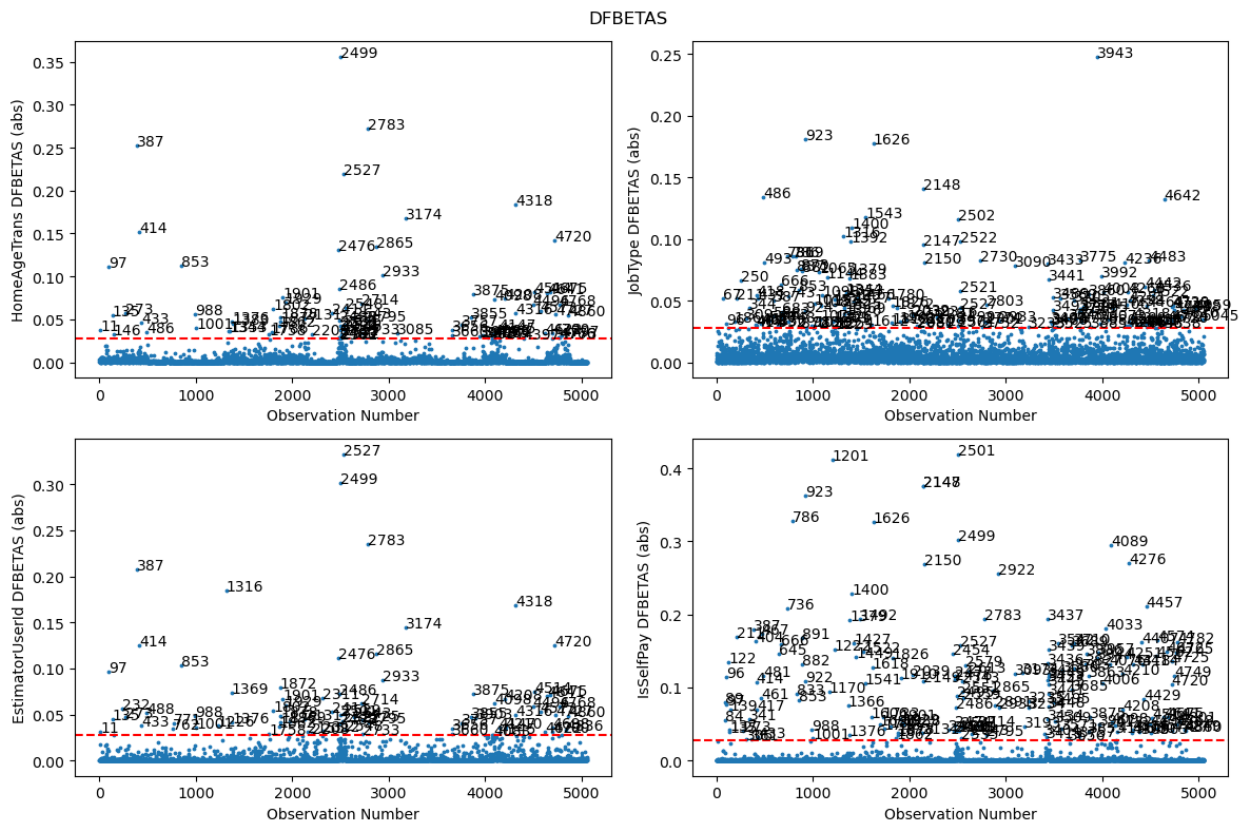
```
def plot_dfbetas(column, ax, label_level = 2 / np.sqrt(len(profit))):
    idx = profit.columns.get_loc(column)
    profit['dfbetas_' + column] = res.get_influence().dfbetas[:, idx]

    ax.set_ylabel(column + " DFBETAS (abs)")
    ax.set_xlabel("Observation Number")
    ax.scatter(profit.index, np.abs(profit['dfbetas_' + column]), s =
3)
    ax.axhline(y = label_level, color = 'r', linestyle = 'dashed')

    # optional: label outliers
    outliers = profit[np.abs(profit['dfbetas_' + column]) >
label_level]
    for i in list(outliers.index):
        ax.annotate(str(i), (i, np.abs(profit['dfbetas_' + column]
[i])))

fig, axes = plt.subplots(2, 2, figsize = (12, 8))
plt.suptitle("DFBETAS")
plot_dfbetas("HomeAgeTrans", axes[0, 0])
plot_dfbetas("JobType", axes[0, 1])
plot_dfbetas("EstimatorUserId", axes[1, 0])
```

```
plot_dfbetas("IsSelfPay", axes[1, 1])
fig.tight_layout()
plt.show()
```



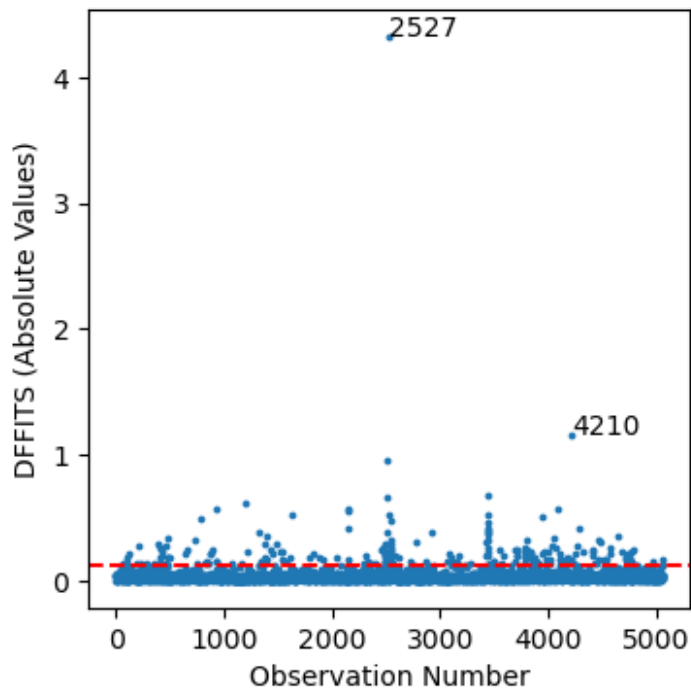
```
profit['dffits'] = res.get_influence().dffits[0]

fig, ax = plt.subplots(figsize = (4, 4))
ax.set_ylabel("DFFITS (Absolute Values)")
ax.set_xlabel("Observation Number")

# scatter + cutoff line
ax.scatter(profit.index,
           np.abs(profit['dffits']),
           s = 3)
ax.axhline(y = 2 * np.sqrt(len(res.params) / len(profit)),
           color = 'r',
           linestyle = 'dashed')

# Change the y_cutoff value to whatever number you want - all the
# points above that value will be labeled with their row number
y_cutoff_dffits = .9
outliers_dffits = profit[np.abs(profit['dffits']) > y_cutoff_dffits]
for i in outliers_dffits.index:
    ax.annotate(str(i), (i, np.abs(profit['dffits'])[i])))
```

```
plt.show()
```



```
profit.loc[2527]
```

GrossProfit	-12808.53
JobType	Mit.
IsSelfPay	0
YearBuilt	2001.0
EstimatorUserId	20
EstimatorName	Dan Goodwin
HomeAge	25.0
HomeAgeTrans	3.218876
HomeAgeTrans_bin	(3.045, 3.296]
HomeAgeTrans_bin_str	(3.045, 3.296]
residuals	-17564.649972
fittedvalues	4756.119972
GrossProfitTrans	0.0
residuals_trans	-9.621234
fittedvalues_trans	9.621234
dfbetas_HomeAgeTrans	-0.219132
dfbetas_JobType	-0.044258
dfbetas_EstimatorUserId	-0.332248
dfbetas_IsSelfPay	0.157758
dffits	-4.329471

Name: 2527, dtype: object

From the DFFITS plot, it seems our only real potential influential point is 2527, which does show up in our log(home age) plot and estimator id plot. After speaking to the company, this was a legitimate job with a particularly bad gross profit, so the observation should be kept in the model. This assumption has been met.

No additional predictors required

This assumption is certainly not met. There are several predictors that would be very beneficial to have, the biggest one being the square footage of the home. However, because this data is not required to be input by the users and most users don't bother, less than 1% of the jobs have square footage reported, so it wasn't possible to use this predictor in the model. There are likely other predictors, too, that would be beneficial that we don't have data for (i.e., equipment used, technician hours, etc.).

No Multicollinearity

Because I only have one continuous predictor in the model, I won't have to worry about multicollinearity. This assumption is met.

Variable Selection

Sequential Replacement

```
y = profit_dummy['GrossProfitTrans']
x = (profit_dummy[['HomeAgeTrans',
                  'JobType_Fire', 'JobType_Mit.', 'JobType_Mold',
                  'JobType_Recon.', 'JobType_Sewer', 'JobType_Water',
                  'IsSelfPay_0',
                  'EstimatorUserId_194', 'EstimatorUserId_20',
                  'EstimatorUserId_26', 'EstimatorUserId_27', 'EstimatorUserId_283',
                  'EstimatorUserId_39', 'EstimatorUserId_53',
                  'EstimatorUserId_60', 'EstimatorUserId_78']]
     .apply(lambda col: col.astype(int) if col.dtype ==
bool else col))

# Tell python which dummy variables belong to the same categorical
variable:
feature_groups = [[0], [1, 2, 3, 4, 5, 6], [7], [8, 9, 10, 11, 12, 13,
14, 15, 16]]

seqrep_selection = SFS(LinearRegression(fit_intercept = True),
                        k_features = (1, len(feature_groups)), #
defined above
                        forward = True,
                        floating = True,
                        scoring = 'neg_mean_squared_error',
                        feature_groups = feature_groups, # defined
above
                        cv = 5)
```

```

seqrep = seqrep_selection.fit(x, y)

print('Sequential Replacement Stepwise Selection:',
      seqrep.k_feature_names_)

Sequential Replacement Stepwise Selection: ('HomeAgeTrans',
      'JobType_Fire', 'JobType_Mit.', 'JobType_Mold', 'JobType_Recon.',
      'JobType_Sewer', 'JobType_Water', 'IsSelfPay_0')

metric_dict = seqrep.get_metric_dict()
x_axis = sorted(metric_dict.keys())

metric_dict_sorted = dict(sorted(metric_dict.items(), key = lambda x:
x[1]['avg_score']))

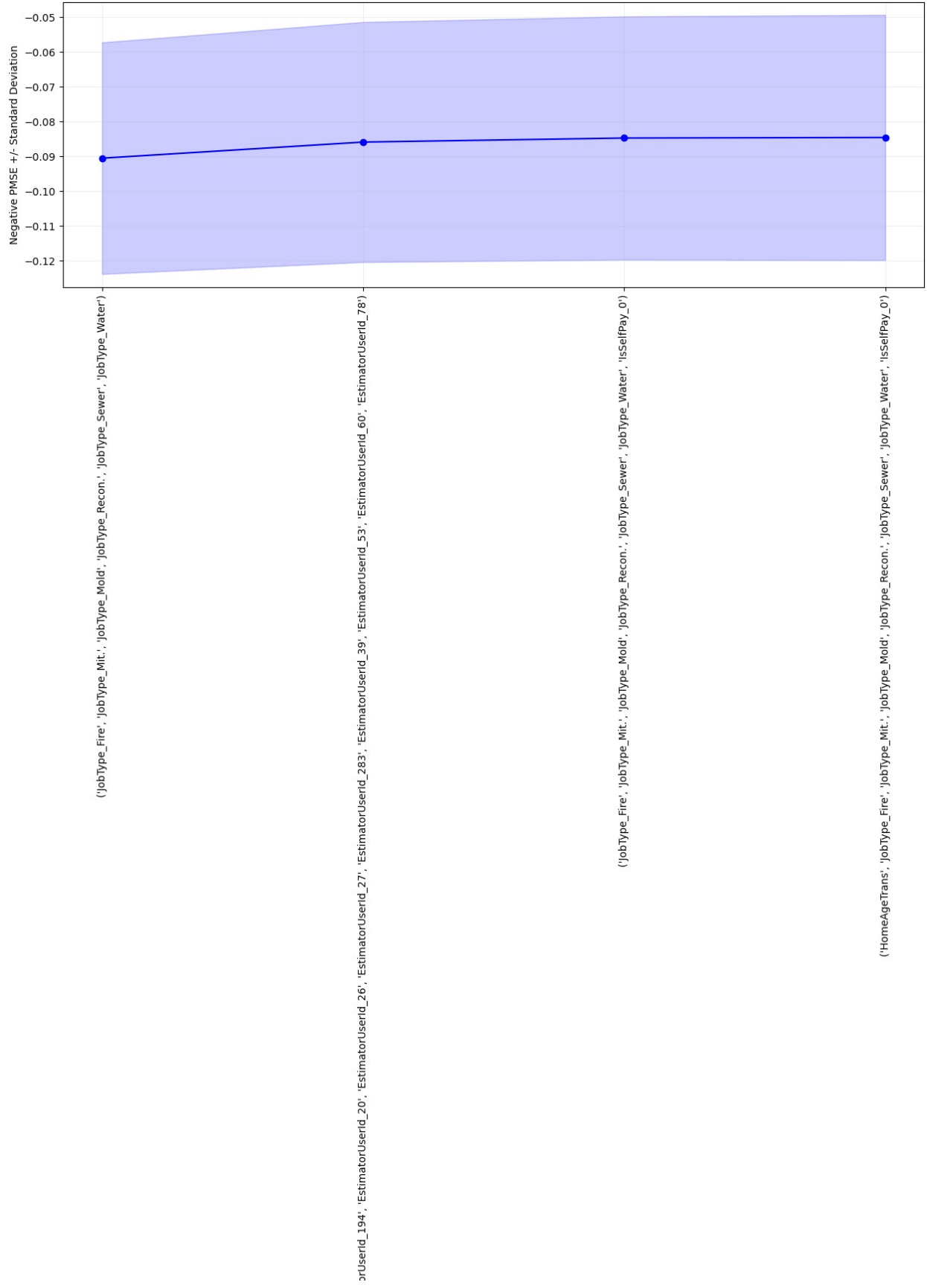
fig = plt.figure(figsize = (15, 5))
k_feat = metric_dict_sorted.keys()
k_feat_list = list(k_feat)
avg = [metric_dict_sorted[k]['avg_score'] for k in k_feat]

upper, lower = [], []
for k in k_feat:
    upper.append(metric_dict_sorted[k]['avg_score'] +
                  metric_dict_sorted[k]['std_dev'])
    lower.append(metric_dict_sorted[k]['avg_score'] -
                  metric_dict_sorted[k]['std_dev'])

plt.fill_between(x_axis,
                 upper,
                 lower,
                 alpha = 0.2,
                 color = 'blue',
                 lw = 1)

plt.plot(x_axis, avg, color = 'blue', marker = 'o')
plt.ylabel('Negative PMSE +/- Standard Deviation')
plt.xlabel('Variables Included')
plt.xticks(x_axis,
           [str(metric_dict_sorted[k]['feature_names']) for k in
            k_feat_list],
           rotation = 90)
plt.grid(alpha = 0.2)
plt.show()

```




```
# Code to see the next best models
```

```
seqrep_results = pd.DataFrame.from_dict(seqrep.get_metric_dict()).T
seqrep_results_sorted = seqrep_results.sort_values(by = 'avg_score',
ascending = False)
seqrep_results_sorted.head(20)
```

	feature_idx \
3	(0, 1, 2, 3, 4, 5, 6, 7)
2	(1, 2, 3, 4, 5, 6, 7)
4	(0, 1, 2, 3, 4, 5, 6, 7, 8, 9, 10, 11, 12, 13, ...)
1	(1, 2, 3, 4, 5, 6)

	cv_scores	avg_score \
3	[-0.06138460182048177, -0.06145348787493667, -...	-0.084532
2	[-0.0621493477100802, -0.06209155828807331, -0...	-0.084687
4	[-0.06390586336468634, -0.06213517157308179, -...	-0.085847
1	[-0.07267836627286879, -0.07056815954972857, -...	-0.090475

	feature_names	ci_bound
std_dev \		
3	(HomeAgeTrans, JobType_Fire, JobType_Mit., Job...	0.045302
0.035246		
2	(JobType_Fire, JobType_Mit., JobType_Mold, Job...	0.044925
0.034953		
4	(HomeAgeTrans, JobType_Fire, JobType_Mit., Job...	0.044342
0.0345		
1	(JobType_Fire, JobType_Mit., JobType_Mold, Job...	0.042775
0.03328		

	std_err
3	0.017623
2	0.017477
4	0.01725
1	0.01664

From the sequential replacement method, it seems that every number of predictors was roughly the same. To make the simplest model possible, we could select job type as our only predictor, but for the purposes of this project, based on this method, I would select the next best model that contains two predictors, which would be the model with job type and "IsSelfPay".

LASSO

```
X = (profit_dummy[['HomeAgeTrans',
'JobType_Fire', 'JobType_Mit.',
'JobType_Mold', 'JobType_Recon.', 'JobType_Sewer', 'JobType_Water',
'IsSelfPay_0',
'EstimatorUserId_194',
'EstimatorUserId_20', 'EstimatorUserId_26', 'EstimatorUserId_27',
'EstimatorUserId_283',
'EstimatorUserId_39',
```

```

'EstimatorUserId_53', 'EstimatorUserId_60', 'EstimatorUserId_78']]
        .apply(lambda col: col.astype(int) if col.dtype ==
bool else col))

#Standardize continuous predictors ONLY
y = profit_dummy['GrossProfitTrans'].ravel()
variable_names = ['HomeAgeTrans']
scaler = StandardScaler().fit(X[variable_names])
X[variable_names] = scaler.transform(X[variable_names])

# Create a list of possible alphas
potential_alphas = np.logspace(-4, 2, 500)
LASSOCV_model = LassoCV(alphas = potential_alphas,
                        cv = 5,
                        random_state = 12345,
                        max_iter = 10000,
                        fit_intercept = True)

# Fit the model
LASSOCV_model.fit(X, y)

# Get the list of alphas and corresponding MSEs
alphas = LASSOCV_model.alphas_
pmse_means = np.mean(LASSOCV_model.mse_path_, axis = 1)
pmse_std_error = np.std(LASSOCV_model.mse_path_,
                        axis = 1,
                        ddof = 1) / np.sqrt(5) # 5-fold CV

# Find the alpha that minimizes MSE
alpha_index_min = np.argmin(pmse_means)
alpha_min = alphas[alpha_index_min]

# Find the MSE that is one standard error away from the minimum MSE
one_se_above_min = min(pmse_means) + pmse_std_error[alpha_index_min]

# Find the largest alpha with MSE less than or equal to
one_se_above_min
alpha_index_lse = np.where(pmse_means <= one_se_above_min)[0][0]
alpha_lse = alphas[alpha_index_lse]

print("Minimum alpha:", alpha_min)
print("One SE alpha:", alpha_lse)

C:\Users\zacha\AppData\Local\Temp\ipykernel_17144\2773064357.py:9:
FutureWarning: Series.ravel is deprecated. The underlying array is
already 1D, so ravel is not necessary. Use `to_numpy()` for
conversion to a numpy array instead.
    y = profit_dummy['GrossProfitTrans'].ravel()

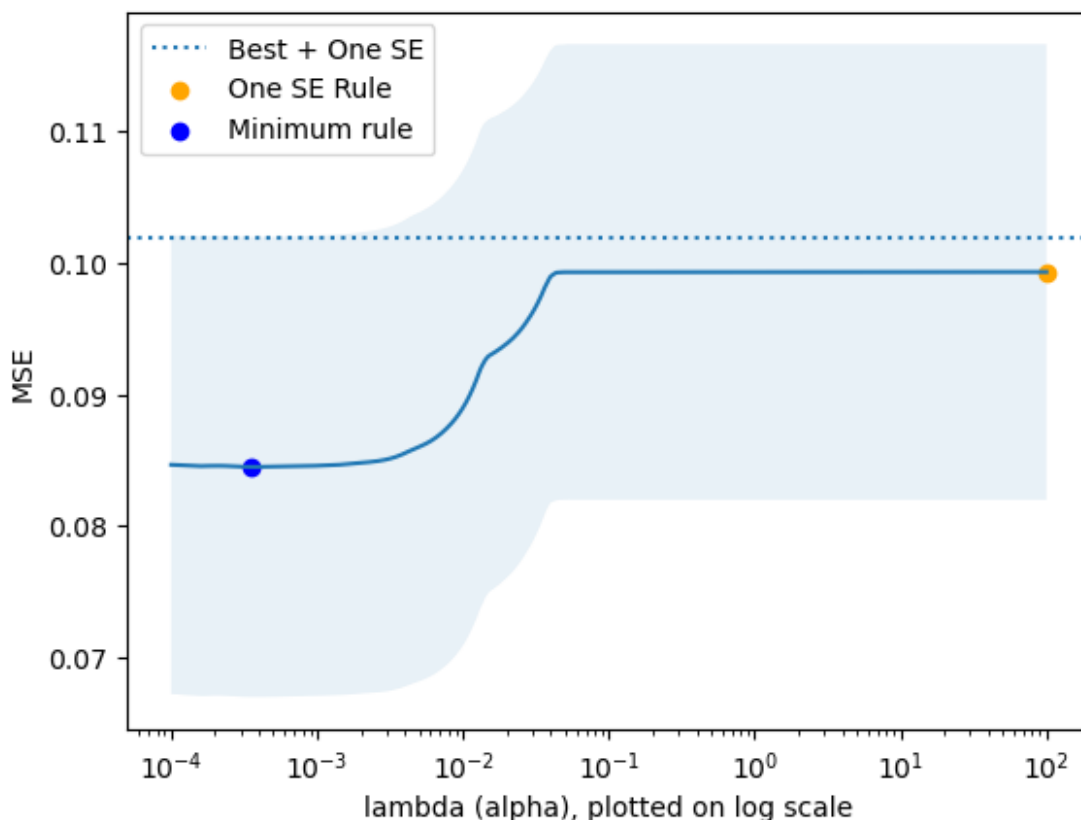
```

Minimum alpha: 0.0003573602246065788

One SE alpha: 100.0

Optional plot

```
plt.plot(alphas,
         pmse_means)
plt.fill_between(alphas,
                 pmse_means + pmse_std_error,
                 pmse_means - pmse_std_error,
                 alpha = 0.1)
plt.axhline(one_se_above_min,
            linestyle = 'dotted',
            label = 'Best + One SE')
plt.scatter([alpha_1se],
            [pmse_means[alpha_index_1se]],
            marker = 'o',
            color = 'orange',
            label = 'One SE Rule')
plt.scatter([alpha_min],
            [pmse_means[alpha_index_min]],
            marker = 'o',
            color = 'blue',
            label = 'Minimum rule')
plt.legend()
plt.xscale('log')
plt.xlabel('lambda (alpha), plotted on log scale')
plt.ylabel('MSE')
plt.show()
```



```
# LASSO results with lambda within one SE of the mean
LASSO_lse = Lasso(alpha = alpha_lse,
                  fit_intercept = True)
LASSO_lse.fit(X, y)

# Get coefficients
print(list(zip(LASSO_lse.coef_, X)))

[(np.float64(0.0), 'HomeAgeTrans'), (np.float64(0.0), 'JobType_Fire'),
 (np.float64(-0.0), 'JobType_Mit.'), (np.float64(-0.0),
 'JobType_Mold'), (np.float64(-0.0), 'JobType_Recon.'),
 (np.float64(0.0), 'JobType_Sewer'), (np.float64(0.0),
 'JobType_Water'), (np.float64(0.0), 'IsSelfPay_0'), (np.float64(-0.0),
 'EstimatorUserId_194'), (np.float64(-0.0), 'EstimatorUserId_20'),
 (np.float64(0.0), 'EstimatorUserId_26'), (np.float64(-0.0),
 'EstimatorUserId_27'), (np.float64(0.0), 'EstimatorUserId_283'),
 (np.float64(0.0), 'EstimatorUserId_39'), (np.float64(-0.0),
 'EstimatorUserId_53'), (np.float64(-0.0), 'EstimatorUserId_60'),
 (np.float64(-0.0), 'EstimatorUserId_78')]

# LASSO results with min(lambda)
LASSO_min = Lasso(alpha = alpha_min,
                  fit_intercept = True)
LASSO_min.fit(X, y)
```

```
# Get coefficients
print(list(zip(LASSO_min.coef_, X)))

[(np.float64(0.01981858917109515), 'HomeAgeTrans'),
 (np.float64(0.5084826199381963), 'JobType_Fire'),
 (np.float64(0.01866709057987206), 'JobType_Mit.'), (np.float64(-
0.036304985627142024), 'JobType_Mold'), (np.float64(0.0),
'JobType_Recon.'), (np.float64(0.07850674845783903), 'JobType_Sewer'),
 (np.float64(0.021878140154526663), 'JobType_Water'),
 (np.float64(0.15595531054638995), 'IsSelfPay_0'), (np.float64(-
0.02432633499232881), 'EstimatorUserId_194'), (np.float64(-
0.09470278807167248), 'EstimatorUserId_20'),
 (np.float64(0.05473146977657855), 'EstimatorUserId_26'),
 (np.float64(0.04150226942405857), 'EstimatorUserId_27'), (np.float64(-
0.0), 'EstimatorUserId_283'), (np.float64(0.07634652968487178),
'EstimatorUserId_39'), (np.float64(0.0063195989013819075),
'EstimatorUserId_53'), (np.float64(-0.0), 'EstimatorUserId_60'),
 (np.float64(-0.03317133360597296), 'EstimatorUserId_78')]
```

It seems like there's something wrong with my code for the LASSO method. When I look at the results within one SE of the mean, it suggests that I remove all variables from my model. When looking at the results with min(lambda), though, the model suggests keeping all of the predictors.

Based on these two methods, with one suggesting job type and "IsSelfPay" and the other suggesting all four predictors, I would end up choosing the model with only job type and "IsSelfPay", one reason being that I don't trust the results of the LASSO method, but the other being that the initial look at the model didn't support estimator id being included either, and the model is much simpler with log(home age) removed. Therefore I think that the best, simplest model would be the one using "Job Type" and "IsSelfPay".

```
profit_dummy = pd.get_dummies(profit, columns = ['JobType',
'IsSelfPay'])

y = profit_dummy['GrossProfitTrans']

# Make sure for each of the categorical predictors, you leave one of
the levels out of the
# model. This will be your baseline level
X = sm.add_constant(profit_dummy[['JobType_Fire', 'JobType_Mit.',
'JobType_Mold', 'JobType_Recon.', 'JobType_Sewer', 'JobType_Water',
'IsSelfPay_0']])
                    .apply(lambda col: col.astype(int) if col.dtype ==
bool else col))

mod = sm.OLS(y, X)
res = mod.fit()

profit['residuals_simple'] = res.resid
```

```
profit['fittedvalues_simple'] = res.fittedvalues
res.summary()

<class 'statsmodels.iolib.summary.Summary'>
"""
                                OLS Regression Results
=====
=====
Dep. Variable:      GrossProfitTrans    R-squared:
0.153
Model:              OLS    Adj. R-squared:
0.152
Method:             Least Squares    F-statistic:
130.7
Date:               Thu, 23 Oct 2025    Prob (F-statistic):
2.48e-177
Time:              14:58:30    Log-Likelihood:
-909.80
No. Observations:   5054    AIC:
1836.
Df Residuals:       5046    BIC:
1888.
Df Model:           7

Covariance Type:    nonrobust

=====
=====
                                coef    std err          t      P>|t|      [0.025
0.975]
-----
-----
const              9.5248      0.024    397.701      0.000      9.478
9.572
JobType_Fire        0.5711      0.034    16.906      0.000      0.505
0.637
JobType_Mit.        0.0683      0.024     2.808      0.005      0.021
0.116
JobType_Mold        0.0135      0.028     0.482      0.630     -0.041
0.069
JobType_Recon.      0.0568      0.027     2.131      0.033      0.005
0.109
JobType_Sewer       0.1439      0.032     4.554      0.000      0.082
0.206
JobType_Water       0.0841      0.025     3.397      0.001      0.036
0.133
IsSelfPay_0         0.1611      0.009    18.687      0.000      0.144
0.178
=====
=====
```

```

=====
Omnibus:                                6470.088    Durbin-Watson:
1.906
Prob(Omnibus):                          0.000    Jarque-Bera (JB):
14381002.472
Skew:                                    -6.060    Prob(JB):
0.00
Kurtosis:                               264.045    Cond. No.
20.9
=====
=====

Notes:
[1] Standard Errors assume that the covariance matrix of the errors is
correctly specified.
"""

```

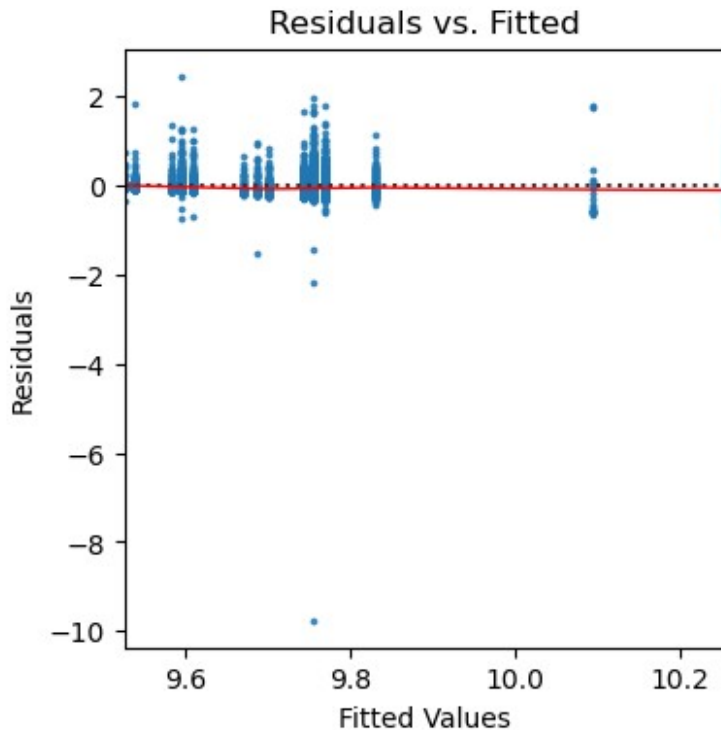
Assumption Checks (In-Depth)

Xs vs Y are linear

```

fig = plt.figure(figsize = (4, 4))
sns.residplot(x = profit['fittedvalues_simple'],
              y = profit['residuals_simple'],
              lowess = True,
              scatter_kws = {'s': 3},
              line_kws = {'color': 'red', 'lw': 1})
plt.title("Residuals vs. Fitted")
plt.ylabel("Residuals")
plt.xlabel("Fitted Values")
plt.show()

```



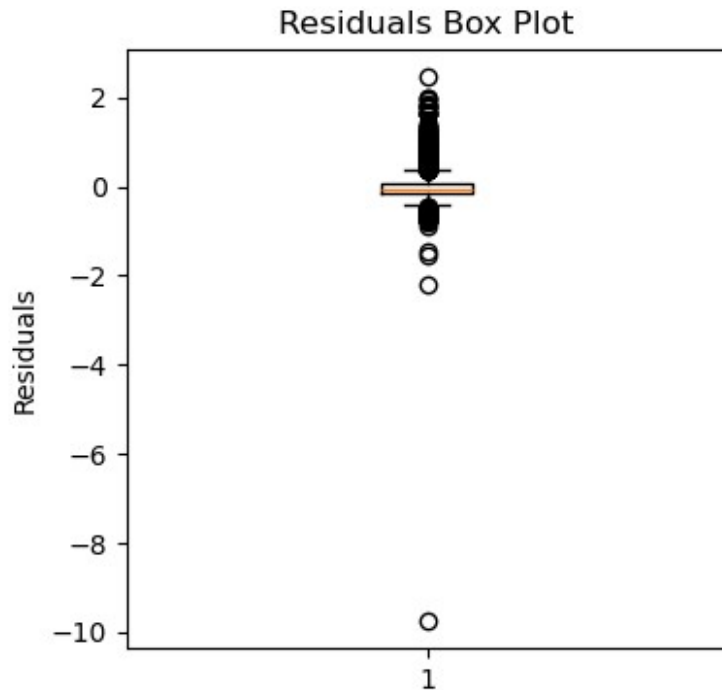
Now that we are left with only categorical predictors, a scatter plot or partial regression plot doesn't make sense in this context. The residuals vs fitted values plot shows a very straight line, however, so this assumption is met.

Residuals are independent

There is some potential for dependence here. One possibility could be if a single home had multiple restoration jobs done, those observations would be more closely related than restoration jobs from other homes, however if this is the case I would guess it's pretty rare. Another could be timing, if certain times of year produce more restoration jobs than others (maybe rainy seasons cause more floods, or summer has more fires than winter, etc.). I'm not sure if this is the case, but in a future model it may make sense to speak with an industry expert about both of these issues to see if this should be considered or not. Other than that, there isn't anything that jumps out as obviously causing dependence. I would say this assumption is met for now, but should be investigated further.

Residuals are normally distributed and centered at 0

```
fig = plt.figure(figsize = (4, 4))
plt.boxplot(profit['residuals_simple'])
plt.ylabel("Residuals")
plt.title("Residuals Box Plot")
plt.show()
```

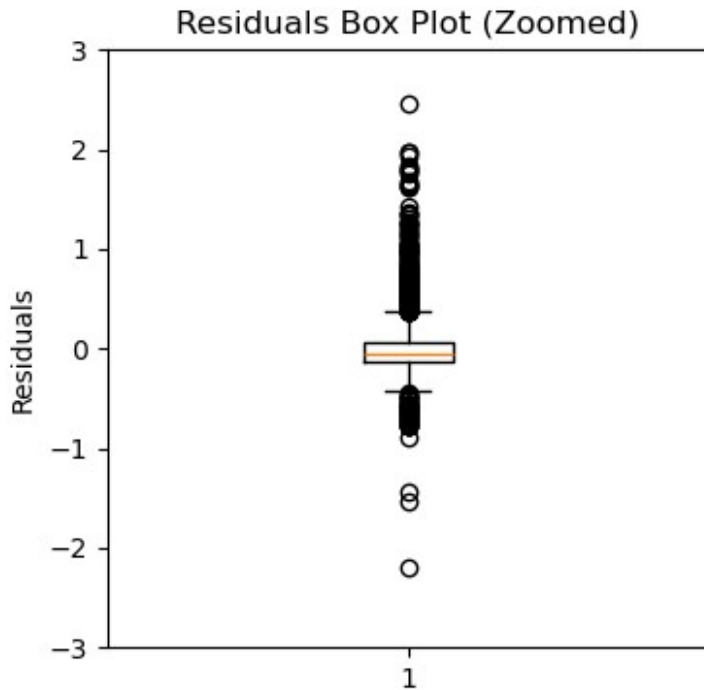
```
lowest_idx = profit['residuals_simple'].idxmin()
profit.loc[lowest_idx]
```

```
GrossProfit      -12808.53
JobType          Mit.
IsSelfPay        0
YearBuilt        2001.0
EstimatorUserId  20
EstimatorName    Dan Goodwin
HomeAge          25.0
HomeAgeTrans     3.218876
HomeAgeTrans_bin (3.045, 3.296]
HomeAgeTrans_bin_str (3.045, 3.296]
residuals        -18132.392228
fittedvalues     5323.862228
GrossProfitTrans 0.0
residuals_trans  -9.621234
fittedvalues_trans 9.621234
residuals_simple -9.754213
fittedvalues_simple 9.754213
Name: 2527, dtype: object
```

There seems to be one residual that is very far away from all the other residuals. It seems that this residual is the same observation mentioned earlier, so it should be left in the model.

```
fig = plt.figure(figsize = (4, 4))
plt.boxplot(profit['residuals_simple'])
```

```
plt.ylabel("Residuals")
plt.title("Residuals Box Plot (Zoomed)")
plt.ylim(-3, 3)
plt.show()
```



```
fig = plt.figure(figsize = (4, 4))

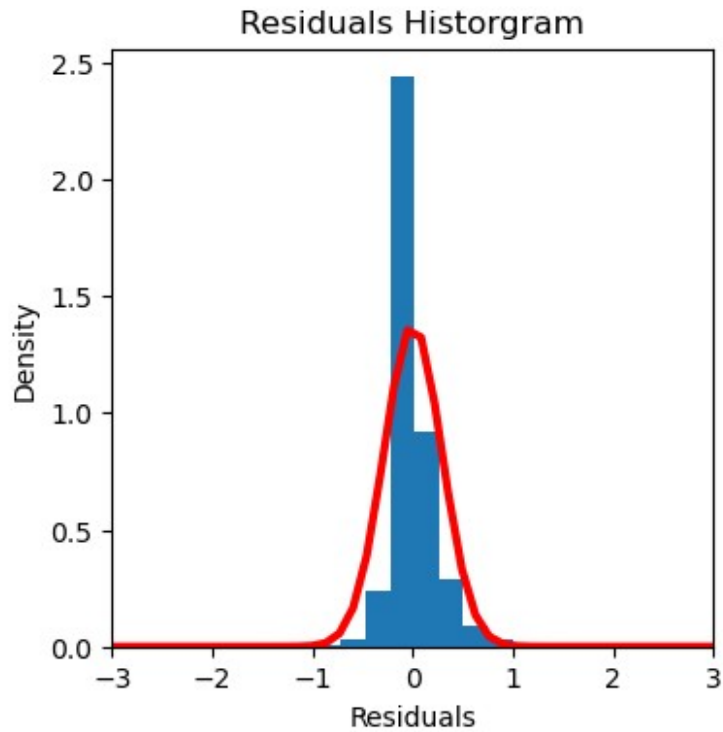
# plot histogram (density = True so that it's on the same scale as the
# normal distribution)
plt.hist(profit['residuals_simple'],
         density = True,
         bins = 50)
plt.xlabel("Residuals")
plt.ylabel("Density")

# calculate mean and standard deviation
mean = np.mean(profit['residuals_simple'])
sd = np.std(profit['residuals_simple'])

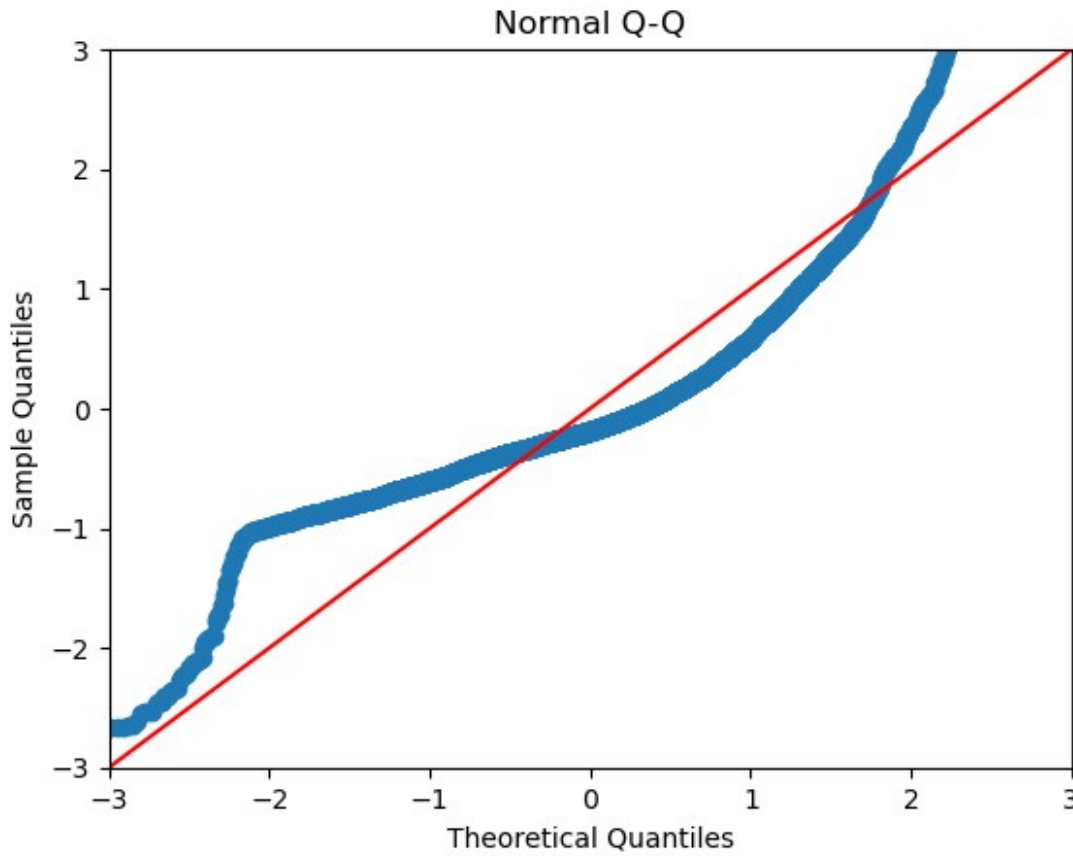
# generate x values to plot
xmin, xmax = plt.xlim()
x = np.linspace(xmin, xmax, 100)

# plot normal distribution curve
plt.plot(x,
         stats.norm.pdf(x, mean, sd),
         color = 'r',
```

```
        lw = 3)
plt.xlim(-3, 3)
plt.title("Residuals Histogram")
plt.show()
```



```
sm.qqplot(profit['residuals_simple'],
           line = '45',
           fit = True)
plt.title("Normal Q-Q")
plt.xlim(-3, 3)
plt.ylim(-3, 3)
plt.show()
```



```
stats.shapiro(profit['residuals_simple'])
```

C:\Users\zacha\anaconda3\Lib\site-packages\scipy\stats_axis_nan_policy.py:586: UserWarning: scipy.stats.shapiro: For N > 5000, computed p-value may not be accurate. Current N is 5054.
 res = hypotest_fun_out(*samples, **kws)

```
ShapiroResult(statistic=np.float64(0.6595934254077522),
pvalue=np.float64(2.6738786314381885e-72))
```

The Q-Q plot still shows a very non-normal distribution for the residuals, and the shapiro-wilk test confirms this, although the box plot and histogram look pretty normal when ignoring the one residual that is far below the others, though you can still see some right skew here as well. I wonder how much this one residual is messing up the normality of the rest of the data, and if this point should be removed even if it is valid data. For now, I will consider this assumption not met.

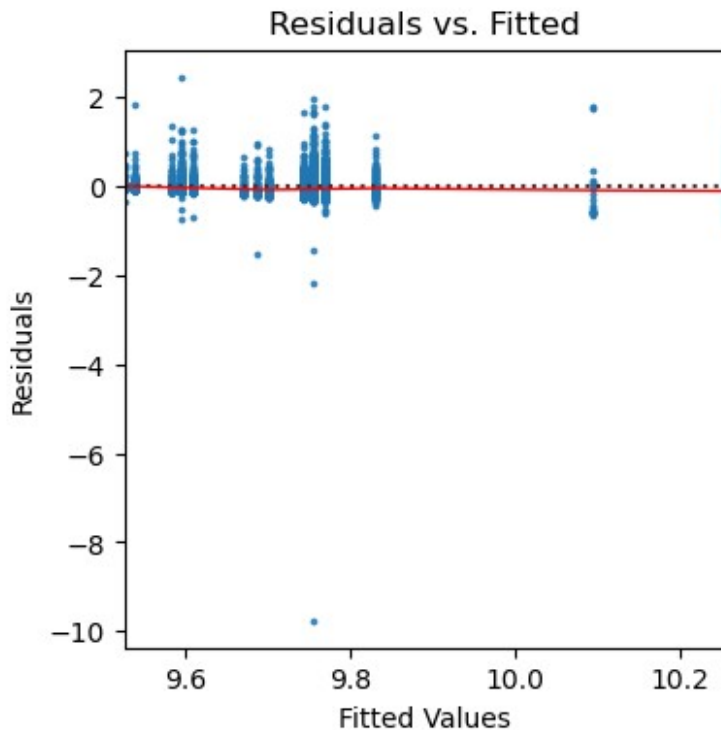
Constant variance among residuals

```
fig = plt.figure(figsize = (4, 4))
sns.residplot(x = profit['fittedvalues_simple'],
              y = profit['residuals_simple'],
              lowess = True,
```

```

        scatter_kws = {'s': 3},
        line_kws = {'color': 'red', 'lw': 1})
plt.title("Residuals vs. Fitted")
plt.ylabel("Residuals")
plt.xlabel("Fitted Values")
plt.show()

```



There is no funneling shape to the residuals vs fitted values plot. This assumption is met.

No Influential Points

```

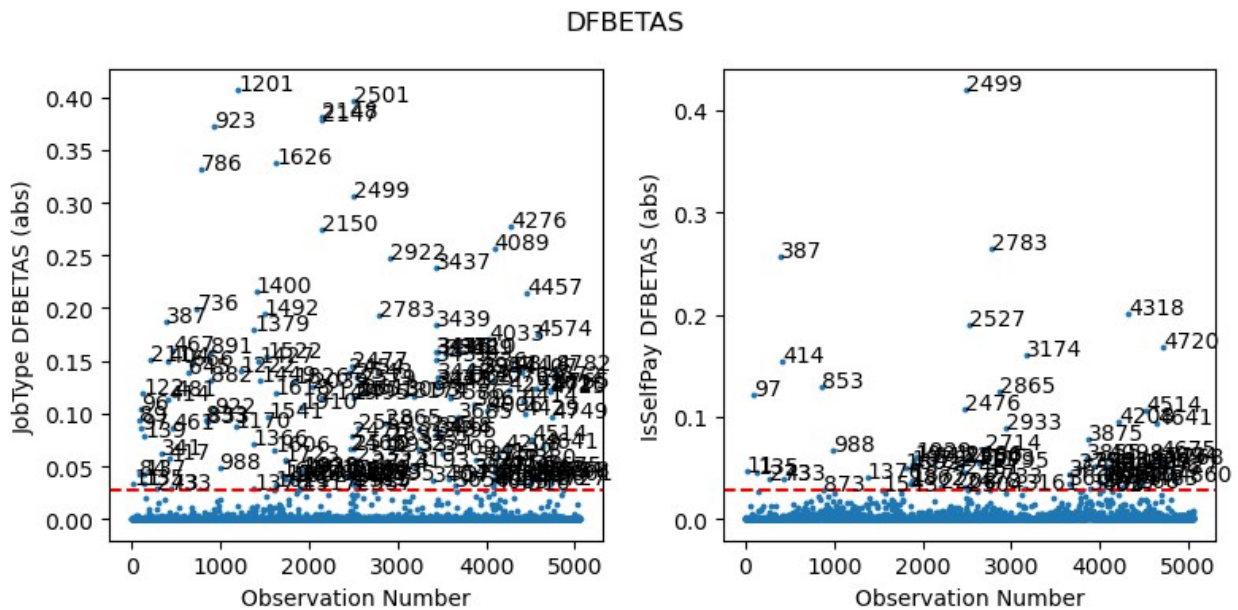
def plot_dfbetas(column, ax, label_level = 2 / np.sqrt(len(profit))):
    idx = profit.columns.get_loc(column)
    profit['dfbetas_' + column] = res.get_influence().dfbetas[:, idx]

    ax.set_ylabel(column + " DFBETAS (abs)")
    ax.set_xlabel("Observation Number")
    ax.scatter(profit.index, np.abs(profit['dfbetas_' + column]), s =
3)
    ax.axhline(y = label_level, color = 'r', linestyle = 'dashed')

    # optional: label outliers
    outliers = profit[np.abs(profit['dfbetas_' + column]) >
label_level]
    for i in list(outliers.index):
        ax.annotate(str(i), (i, np.abs(profit['dfbetas_' + column]
[i])))

```

```
fig, axes = plt.subplots(1, 2, figsize = (8, 4))
plt.suptitle("DFBETAS")
plot_dfbetas("JobType", axes[0])
plot_dfbetas("IsSelfPay", axes[1])
fig.tight_layout()
plt.show()
```



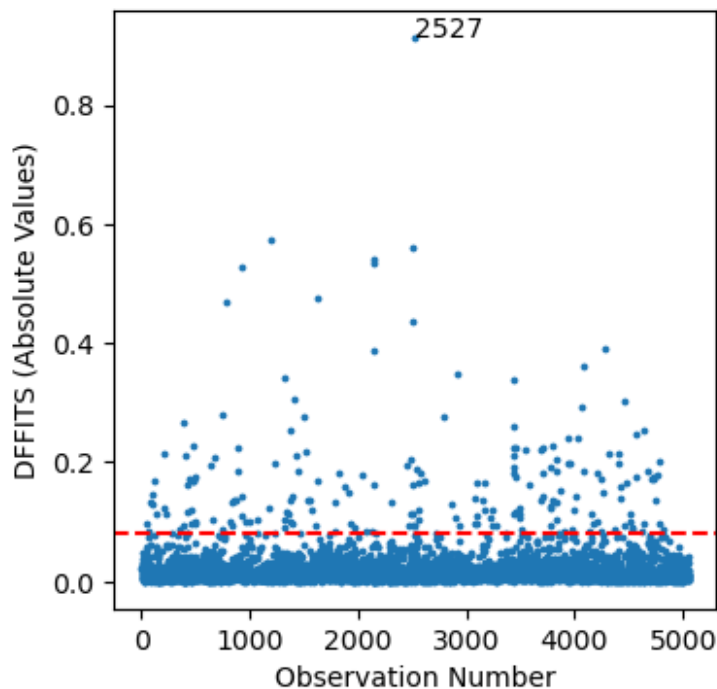
```
profit['dffits'] = res.get_influence().dffits[0]

fig, ax = plt.subplots(figsize = (4, 4))
ax.set_ylabel("DFFITS (Absolute Values)")
ax.set_xlabel("Observation Number")

# scatter + cutoff line
ax.scatter(profit.index,
           np.abs(profit['dffits']),
           s = 3)
ax.axhline(y = 2 * np.sqrt(len(res.params) / len(profit)),
           color = 'r',
           linestyle = 'dashed')

# Change the y_cutoff value to whatever number you want - all the
# points above that value will be labeled with their row number
y_cutoff_dffits = .9
outliers_dffits = profit[np.abs(profit['dffits']) > y_cutoff_dffits]
for i in outliers_dffits.index:
    ax.annotate(str(i), (i, np.abs(profit['dffits'])[i])))
```

```
plt.show()
```



Again observation 2527 shows up. I would consider this assumption to be met with the knowledge that observation 2527 shouldn't be removed from the model.

No additional predictors required

This assumption is certainly not met. There are several predictors that would be very beneficial to have, the biggest one being the square footage of the home. However, because this data is not required to be input by the users and most users don't bother, less than 1% of the jobs have square footage reported, so it wasn't possible to use this predictor in the model. There are likely other predictors, too, that would be beneficial that we don't have data for (equipment used, technician hours, etc.).

No multicollinearity

Because we have no continuous predictors, this assumption is not a concern for this model.

Interaction Terms

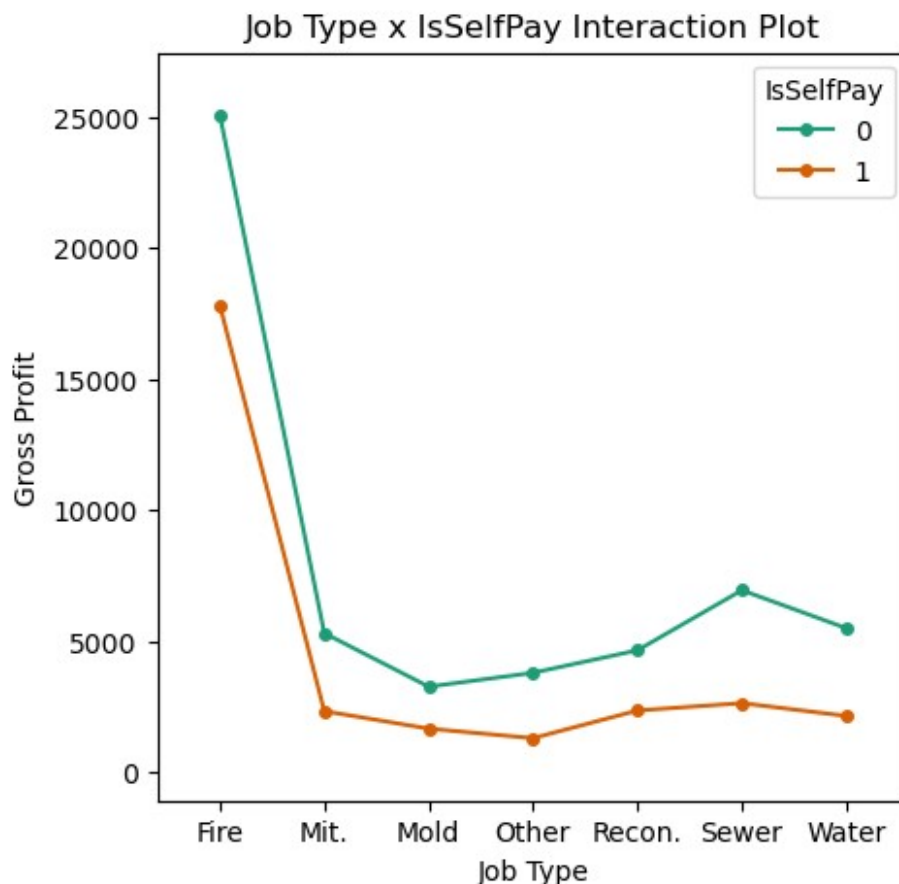
```
fig, ax = plt.subplots(figsize = (5, 5))
fig = interaction_plot(x = profit["JobType"],
                      trace = profit["IsSelfPay"],
                      response = profit["GrossProfit"],
                      colors = ["#1b9e77", "#d95f02"],
                      ms = 8, # marker size)
```

```

        ax = ax)
ax.set_xlabel('Job Type')
ax.set_ylabel('Gross Profit')
plt.title("Job Type x IsSelfPay Interaction Plot")
ax.legend(title = 'IsSelfPay')

<matplotlib.legend.Legend at 0x15076a5ec10>

```



From our EDA, we determined that an interaction may make sense between job type and estimator id, "IsSelfPay" and estimator id, and job type and log(home age), but because we removed estimator id and log(home age) from the model, I'll only check if an interaction makes sense between job type and "IsSelfPay", however I don't expect a significant interaction because the interaction plot between job type and "IsSelfPay" showed perhaps the least evidence of an interaction of any of my interaction plots.

```

profit_dummy = pd.get_dummies(profit, columns = ['JobType',
'IsSelfPay'])

bool_cols = profit_dummy.select_dtypes(include='bool').columns
profit_dummy[bool_cols] = profit_dummy[bool_cols].astype(int)

job_types = profit['JobType'].unique()

```



```

int_list = []
for job_type in job_types:
    if job_type != 'Other':
        profit_dummy[job_type + '_IsSelfPay_1'] =
profit_dummy['JobType_' + job_type] * profit_dummy['IsSelfPay_1']
        int_list.append(str(job_type + '_IsSelfPay_1'))

y = profit_dummy['GrossProfit']

baselines = ['JobType_Other', 'IsSelfPay_0']
X = sm.add_constant(
    profit_dummy[['JobType_Fire', 'JobType_Mit.', 'JobType_Mold',
'JobType_Recon.', 'JobType_Sewer', 'JobType_Water',
'IsSelfPay_0']
        + int_list]
    .apply(lambda col: col.astype(int) if col.dtype == bool else col)
)

mod_int = sm.OLS(y, X)
res_int = mod_int.fit()
profit['residuals_int'] = res_int.resid
profit['fittedvalues_int'] = res_int.fittedvalues
res_int.summary()

```

```

<class 'statsmodels.iolib.summary.Summary'>
"""

```

OLS Regression Results

```

=====
=====
Dep. Variable:          GrossProfit    R-squared:
0.156
Model:                  OLS    Adj. R-squared:
0.154
Method:                 Least Squares    F-statistic:
71.52
Date:                   Thu, 23 Oct 2025    Prob (F-statistic):
3.35e-174
Time:                   15:00:20    Log-Likelihood:
-52970.
No. Observations:      5054    AIC:
1.060e+05
Df Residuals:          5040    BIC:
1.061e+05
Df Model:              13

Covariance Type:      nonrobust

=====
=====

```

=====		coef	std err	t	P> t	
[0.025	0.975]					

const		1287.2703	1017.391	1.265	0.206	-
707.258	3281.798					
JobType_Fire		2.128e+04	1236.035	17.218	0.000	
1.89e+04	2.37e+04					
JobType_Mit.		1513.1787	992.413	1.525	0.127	-
432.383	3458.740					
JobType_Mold		-528.9889	1226.724	-0.431	0.666	-
2933.902	1875.924					
JobType_Recon.		859.1508	1057.303	0.813	0.416	-
1213.623	2931.925					
JobType_Sewer		3150.2024	1239.951	2.541	0.011	
719.360	5581.045					
JobType_Water		1703.1595	1005.247	1.694	0.090	-
267.562	3673.881					
IsSelfPay_0		2495.0592	1402.376	1.779	0.075	-
254.208	5244.326					
Water_IsSelfPay_1		-853.8692	1482.049	-0.576	0.565	-
3759.329	2051.591					
Recon._IsSelfPay_1		197.8739	1630.936	0.121	0.903	-
2999.471	3395.219					
Mit._IsSelfPay_1		-480.6550	1452.764	-0.331	0.741	-
3328.703	2367.393					
Sewer_IsSelfPay_1		-1813.0119	1923.992	-0.942	0.346	-
5584.872	1958.848					
Mold_IsSelfPay_1		893.8533	1691.499	0.528	0.597	-
2422.220	4209.927					
Fire_IsSelfPay_1		-4765.4007	2507.821	-1.900	0.057	-
9681.821	151.020					
=====						
=====						
Omnibus:		7354.211	Durbin-Watson:			
1.892						
Prob(Omnibus):		0.000	Jarque-Bera (JB):			
3289463.061						
Skew:		8.646	Prob(JB):			
0.00						
Kurtosis:		126.781	Cond. No.			
47.7						
=====						
=====						

Notes:

[1] Standard Errors assume that the covariance matrix of the errors is

```
correctly specified.  
"""
```

```
anova_lm(res, res_int)
```

	df_resid	ssr	df_diff	ss_diff	F	Pr(>F)
0	5046.0	4.241516e+02	0.0	NaN	NaN	NaN
1	5040.0	3.756112e+11	6.0	-3.756112e+11	-839.999999	1.0

The F test confirms what I hypothesised, the interaction is not significant (p-value = 1.0). We will continue with the model that does not include any interactions.

Model Analysis

```
print('R-squared adjusted:', round(res.rsquared_adj, 4))
```

```
R-squared adjusted: 0.1523
```

The R-squared adjusted value indicates that this model explains about 15% of the variability in log(gross profit), which is not very much.

```
print("The test statistic for the F-test is",  
      round(res.fvalue, 2),  
      "and the corresponding p-value is",  
      res.f_pvalue)
```

```
The test statistic for the F-test is 130.69 and the corresponding p-  
value is 2.476424400733913e-177
```

The F-test confirms that the model is significant (, which means that there is evidence of a statistically significant relationship between log(gross profit) and job type and "IsSelfPay".

```
rmse = np.sqrt(np.sum((profit['GrossProfitTrans'] -  
profit['fittedvalues_simple']) ** 2) / (len(profit) - 2))  
print('RMSE (log-scale):', rmse)
```

```
RMSE (log-scale): 0.28975362606927735
```

```
mae = np.sum(np.abs(profit['GrossProfitTrans'] -  
profit['fittedvalues_simple'])) / (len(profit) - 2)  
print('MAE (log-scale):', mae)
```

```
MAE (log-scale): 0.16671367814489826
```

```
print('log(gross profit) min:', profit['GrossProfitTrans'].min())  
print('log(gross profit) max:', profit['GrossProfitTrans'].max())
```

```
log(gross profit) min: 0.0  
log(gross profit) max: 12.236225446077071
```

On average, our model predictions for log(gross profit) are off by between 0.17 and 0.29 log-units. For a log(gross profit) range of \$12.24, this seems like an acceptable level of error.

```
res.conf_int()

              0          1
const      9.477877  9.571781
JobType_Fire  0.504857  0.637304
JobType_Mit.  0.020612  0.115982
JobType_Mold -0.041494  0.068521
JobType_Recon. 0.004546  0.109006
JobType_Sewer  0.081969  0.205890
JobType_Water  0.035564  0.132618
IsSelfPay_0    0.144188  0.177987

print(0.504857 - 0.081969, 0.637304 - 0.205890)

0.422888 0.43141399999999996
```

From our confidence intervals, we can see the effect of each of our levels relative to the baseline level. For job type, the baseline level was "Other". This means that job types of "Fire" make on average between 0.50 and 0.64 log-units more than "Other" job types. Mitigation jobs, reconstruction, sewer and water jobs all also make more than "Other" jobs, while mold jobs made on average between 0.04 log-units less than "Other" jobs, and 0.07 log-units more than "Other" jobs. Overall, it seems that fire jobs made more than any other job, with the next highest being sewer jobs. Fire jobs made on average between 0.42 and 0.43 log-units more than sewer jobs. Jobs in which the customer used insurance when paying also made on average between 0.14 and 0.18 log-units more than jobs where the customer paid themselves.

```
print('Test Statistics')
print(res.tvalues)
print('P-Values')
print(res.pvalues)

Test Statistics
const      397.700788
JobType_Fire  16.905887
JobType_Mit.   2.807822
JobType_Mold   0.481619
JobType_Recon.  2.131066
JobType_Sewer  4.553958
JobType_Water  3.397182
IsSelfPay_0   18.687438
dtype: float64
P-Values
const      0.000000e+00
JobType_Fire  2.068544e-62
JobType_Mit.   5.006930e-03
JobType_Mold   6.300977e-01
JobType_Recon.  3.313188e-02
```

```
JobType_Sewer      5.388516e-06
JobType_Water      6.861012e-04
IsSelfPay_0        2.091188e-75
dtype: float64
```

From the test statistics and p-values we can see that the only level that was not statistically significant was the mold job type. This means that both job type and "IsSelfPay" are significant predictors of log(gross profit).

```
#Average log(gross profit) when the customer is paying with insurance
for a mitigation job
pred = res.get_prediction([1, 0, 1, 0, 0, 0, 0, 1])
pred.summary_frame(alpha = 0.05).iloc[:, [0, 2, 3]]
```

	mean	mean_ci_lower	mean_ci_upper
0	9.754213	9.740624	9.767803

The average log(gross profit) for a mitigation job when a customer is not paying themselves is between 9.74 and 9.77 log units.

```
#Average log(gross profit) when the customer is paying with insurance
for a mitigation job
pred = res.get_prediction([1, 1, 0, 0, 0, 0, 0, 0])
pred.summary_frame(alpha = 0.05).iloc[:, [0, 4, 5]]
```

	mean	obs_ci_lower	obs_ci_upper
0	10.09591	9.525387	10.666433

For a new fire job when a customer is paying themselves, the company can expect to make between 9.53 and 10.67 log units.

Summary

The goal with this analysis was to use data that is available before a job has been completed to predict the gross profit of the job. Ultimately, it was determined that the best model to predict gross profit used the job type and whether the customer was paying themselves or through insurance. Overall, the model was statistically significant, meaning that using job type and "IsSelfPay" leads to better predictions than simply referring to the mean of gross profit, however the model explains only ~15% of the variability in gross profit. This means that while there is a statistically significant relationship between gross profit and job type and "IsSelfPay", there may not be any real world application for this model without expanding to include other predictors, such as square footage or equipment used, for example.