

Mathematical Statistics Project

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2024 July 20 , Saturday

Part I

MLE for Weibull Distribution

Given a sample x_1, x_2, \dots, x_n from a Weibull distribution with probability density function.

$$X_1, \dots, X_n \sim W(\alpha, \beta)$$

$$f_{\alpha, \beta}(x) = \frac{\alpha}{\beta} \left(\frac{x}{\beta} \right)^{\alpha-1} e^{-\left(\frac{x}{\beta} \right)^\alpha}, \quad \alpha, \beta, x > 0$$

a) Define $L(\alpha, \beta) := \mathbf{f} \alpha, \beta(\mathbf{x})$

$$L(\alpha, \beta) := f_{\alpha, \beta}(x) = \prod_{i=1}^n f_{\alpha, \beta}(x_i) = \prod_{i=1}^n \frac{\alpha}{\beta} \left(\frac{x_i}{\beta} \right)^{\alpha-1} e^{-\left(\frac{x_i}{\beta} \right)^\alpha} = \frac{\alpha^n}{\beta^{n\alpha}} \prod_{i=1}^n x_i^{\alpha-1} e^{-\sum_{i=1}^n \left(\frac{x_i}{\beta} \right)^\alpha}$$

$$\ln L(\alpha, \beta) = \ln \left(\frac{\alpha^n}{\beta^{n\alpha}} \right) + \ln \left(\prod_{i=1}^n x_i^{\alpha-1} \right) + \ln \left(e^{-\sum_{i=1}^n \left(\frac{x_i}{\beta} \right)^\alpha} \right)$$

$$\ln L(\alpha, \beta) = n \ln \alpha - n\alpha \ln \beta + \sum_{i=1}^n \ln x_i^{\alpha-1} - \sum_{i=1}^n \left(\frac{x_i}{\beta} \right)^\alpha$$

$$\ln L(\alpha, \beta) = n \ln \alpha - n\alpha \ln \beta + (\alpha - 1) \sum_{i=1}^n \ln x_i - \sum_{i=1}^n \left(\frac{x_i}{\beta} \right)^\alpha \quad (1)$$

According to equation (1), we were able to obtain the likelihood function.

b) MLE (α, β) based on $L(\alpha, \beta)$

$$\ln L(\alpha, \beta) = n \ln \alpha - n\alpha \ln \beta + (\alpha - 1) \sum_{i=1}^n \ln x_i - \sum_{i=1}^n \left(\frac{x_i}{\beta} \right)^\alpha$$

$$\hat{\beta} := \frac{\partial \ln L(\alpha, \beta)}{\partial \beta} = 0 \Rightarrow -n\alpha\beta^{-1} + \alpha\beta^{-\alpha-1} \sum_{i=1}^n x_i^\alpha = 0$$

$$n\alpha\beta^{-1} = \alpha\beta^{-\alpha-1} \sum_{i=1}^n x_i^\alpha \Rightarrow n\alpha = \alpha\beta^{-\alpha} \sum_{i=1}^n x_i^\alpha$$

$$\beta^\alpha = \frac{\sum_{i=1}^n x_i^\alpha}{n}$$

$$\hat{\beta} = \left(\frac{\sum_{i=1}^n x_i^\alpha}{n} \right)^{\frac{1}{\alpha}} \quad (2)$$

The above likelihood equation (2) where β is MLE depends on .

- $\ln L(\alpha, \beta) = n \ln(\alpha) - n\alpha \ln(\beta) + (\alpha - 1) \sum_{i=1}^n \ln(x_i) - \sum_{i=1}^n \left(\frac{x_i}{\beta} \right)^\alpha$

$$\hat{\alpha} := \frac{\partial \ln L(\alpha, \beta)}{\partial \alpha} = 0 \Rightarrow \frac{n}{\alpha} - n \ln(\beta) + \sum_{i=1}^n \ln(x_i) - \sum_{i=1}^n \left(\frac{x_i}{\beta} \right)^\alpha \ln \left(\frac{x_i}{\beta} \right) = 0$$

$$\frac{n}{\alpha} - \sum_{i=1}^n \ln(\beta) + \sum_{i=1}^n \ln(x_i) - \sum_{i=1}^n \left(\frac{x_i}{\beta} \right)^\alpha \ln \left(\frac{x_i}{\beta} \right) = 0$$

$$\frac{n}{\alpha} + \sum_{i=1}^n \ln \left(\frac{x_i}{\beta} \right) - \sum_{i=1}^n \left(\frac{x_i}{\beta} \right)^\alpha \ln \left(\frac{x_i}{\beta} \right) = 0$$

$$\Rightarrow \frac{n}{\alpha} + \sum_{i=1}^n \ln \left(\frac{x_i}{\beta} \right) - \sum_{i=1}^n \left(\frac{x_i}{\beta} \right)^\alpha \ln \left(\frac{x_i}{\beta} \right) = 0$$

$$\frac{n}{\alpha} = \sum_{i=1}^n \ln \left(\frac{x_i}{\beta} \right) \left(\left(\frac{x_i}{\beta} \right)^\alpha - 1 \right) \Rightarrow \hat{\alpha} = \frac{n}{\sum_{i=1}^n \ln \left(\frac{x_i}{\beta} \right) \left(\left(\frac{x_i}{\beta} \right)^\alpha - 1 \right)}$$

$$\hat{\alpha} = \left(\frac{\sum_{i=1}^n \ln \left(\frac{x_i}{\beta} \right) \left(\left(\frac{x_i}{\beta} \right)^\alpha - 1 \right)}{n} \right)^{-1}$$

$$\hat{\alpha} = \left(\frac{\sum_{i=1}^n \left(\frac{x_i}{\beta} \right)^\alpha \ln(x_i)}{n} - \frac{\sum_{i=1}^n \left(\frac{x_i}{\beta} \right)^\alpha \ln(\beta)}{n} - \frac{\sum_{i=1}^n \ln(x_i)}{n} + \frac{\sum_{i=1}^n (\ln \beta)}{n} \right)^{-1}$$

$$\hat{\beta} = \left(\frac{\sum_{i=1}^n (x_i)^\alpha}{n} \right)^{\frac{1}{\alpha}} \Rightarrow \hat{\alpha} = \left(\frac{\sum_{i=1}^n (x_i)^\alpha \ln(x_i)}{\sum_{i=1}^n (x_i)^\alpha} - \frac{\sum_{i=1}^n \ln(x_i)}{n} \right)^{-1} \quad (3)$$

c) Let β be known

$$L_\beta(\alpha) := L(\alpha, \beta_0) := f_{\alpha, \beta_0}(\mathbf{x}) = \prod_{i=1}^n f_{\alpha, \beta_0}(x_i) = \prod_{i=1}^n \frac{\alpha}{\beta_0} \left(\frac{x_i}{\beta_0} \right)^{\alpha-1} e^{-\left(\frac{x_i}{\beta_0} \right)^\alpha}$$

$$\frac{\alpha^n}{\beta_0^{n\alpha}} \cdot \prod_{i=1}^n x_i^{\alpha-1} \cdot e^{-\sum_{i=1}^n \left(\frac{x_i}{\beta_0} \right)^\alpha}$$

$$\hat{\alpha}(\beta) := \frac{\partial \ln L(\alpha, \beta_0)}{\partial \alpha} = 0 \Rightarrow n\alpha^{-1} - n \ln \beta_0 + \sum_{i=1}^n \ln x_i - \sum_{i=1}^n \left(\frac{x_i}{\beta_0} \right)^\alpha \ln \left(\frac{x_i}{\beta_0} \right) = 0$$

$$\frac{n}{\alpha} - \sum_{i=1}^n \ln(\beta_0) + \sum_{i=1}^n \ln(x_i) - \sum_{i=1}^n \left(\frac{x_i}{\beta_0} \right)^\alpha \ln \left(\frac{x_i}{\beta_0} \right) = 0$$

$$\frac{n}{\alpha} + \sum_{i=1}^n \ln \left(\frac{x_i}{\beta_0} \right) - \sum_{i=1}^n \left(\frac{x_i}{\beta_0} \right)^\alpha \ln \left(\frac{x_i}{\beta_0} \right) = 0$$

$$\Rightarrow \frac{n}{\alpha} + \sum_{i=1}^n \ln \left(\frac{x_i}{\beta_0} \right) - \sum_{i=1}^n \left(\frac{x_i}{\beta_0} \right)^\alpha \ln \left(\frac{x_i}{\beta_0} \right) = 0$$

$$\frac{n}{\alpha} = \sum_{i=1}^n \ln \left(\frac{x_i}{\beta_0} \right) \left(\left(\frac{x_i}{\beta_0} \right)^\alpha - 1 \right) \Rightarrow \hat{\alpha} = \frac{n}{\sum_{i=1}^n \ln \left(\frac{x_i}{\beta_0} \right) \left(\left(\frac{x_i}{\beta_0} \right)^\alpha - 1 \right)}$$

$$\alpha_{\hat{\beta}} = \left(\frac{\sum_{i=1}^n \left(\frac{x_i}{\beta_0} \right)^\alpha \ln(x_i)}{n} - \frac{\sum_{i=1}^n \left(\frac{x_i}{\beta_0} \right)^\alpha \ln(\beta_0)}{n} - \frac{\sum_{i=1}^n \ln(x_i)}{n} + \frac{\sum_{i=1}^n \ln(\beta_0)}{n} \right)^{-1}$$

$$\frac{\partial^2 \ln L(\alpha, \beta_0)}{\partial \alpha^2} = \frac{-n}{\alpha^2} - \sum_{i=1}^n \ln \left(\frac{x_i}{\beta_0} \right)^2 \left(\frac{x_i}{\beta_0} \right)^\alpha < 0 \Rightarrow \alpha < \sqrt{\frac{\sum_{i=1}^n \ln \left(\frac{x_i}{\beta_0} \right)^2 \left(\frac{x_i}{\beta_0} \right)^\alpha}{n}} \quad (4)$$

The above statement (4) must be true so that α can take its maximum value.

d) Define $L(\beta) := L(\hat{\alpha}(\beta), \beta)$ Find $\hat{\beta}$ based on $L(\beta)$

$$L(\beta) = L(\alpha_{\hat{\beta}}, \beta) = \frac{\alpha_{\hat{\beta}}^n}{\beta^{n\alpha_{\hat{\beta}}}} \cdot \prod_{i=1}^n x_i^{\alpha_{\hat{\beta}}-1} \cdot e^{-\sum_{i=1}^n \left(\frac{x_i}{\beta} \right)^{\alpha_{\hat{\beta}}}}$$

$$\ln L(\alpha_{\hat{\beta}}, \beta) = n \ln(\alpha_{\hat{\beta}}) - n\alpha_{\hat{\beta}} \ln(\beta) + (\alpha_{\hat{\beta}} - 1) \sum_{i=1}^n \ln(x_i) - \sum_{i=1}^n \left(\frac{x_i}{\beta} \right)^{\alpha_{\hat{\beta}}}$$

$$\hat{\beta} := \frac{\partial \ln L(\alpha_{(\hat{\beta})}, \beta)}{\partial \beta} = 0 \Rightarrow -\frac{n\alpha_{(\hat{\beta})}}{\beta} - \frac{-\alpha_{(\hat{\beta})}\beta^{\alpha_{(\hat{\beta})}-1}}{\beta^{2\alpha_{(\hat{\beta})}}} \sum_{i=1}^n (x_i)^\alpha = 0$$

$$\frac{n\alpha_{(\hat{\beta})}}{\beta} = \alpha_{(\hat{\beta})}\beta^{-1-\alpha_{(\hat{\beta})}} \sum_{i=1}^n (x_i)^{\alpha_{(\hat{\beta})}} \Rightarrow \beta^{\alpha_{(\hat{\beta})}} = \frac{\sum_{i=1}^n (x_i)^{\alpha_{(\hat{\beta})}}}{n}$$

$$\hat{\beta} = \left(\frac{\sum_{i=1}^n (x_i)^{\alpha_{(\hat{\beta})}}}{n} \right)^{\frac{1}{\alpha_{(\hat{\beta})}}}$$

$$\alpha_{(\hat{\beta})} = \frac{n}{\sum_{i=1}^n \ln \left(\frac{x_i}{\beta_0} \right) \left(\left(\frac{x_i}{\beta_0} \right)^\alpha - 1 \right)} \neq 0$$

$$\hat{\beta} = \left(\frac{\sum_{i=1}^n (x_i)^{\frac{1}{\sum_{i=1}^n \ln \left(\frac{x_i}{\beta_0} \right) \left(\left(\frac{x_i}{\beta_0} \right)^\alpha - 1 \right)}}}{n} \right)^{\frac{1}{\sum_{i=1}^n \ln \left(\frac{x_i}{\beta_0} \right) \left(\left(\frac{x_i}{\beta_0} \right)^\alpha - 1 \right)}} = \left(\frac{\sum_{i=1}^n (x_i)^{\frac{1}{\sum_{i=1}^n \ln \left(\frac{x_i}{\beta_0} \right) \left(\left(\frac{x_i}{\beta_0} \right)^\alpha - 1 \right)}}}{n} \right)^{\frac{1}{\sum_{i=1}^n \ln \left(\frac{x_i}{\beta_0} \right) \left(\left(\frac{x_i}{\beta_0} \right)^\alpha - 1 \right)}}$$

$$\frac{\partial^2 \ln L(\alpha_{(\hat{\beta})}, \beta)}{\partial \beta^2} = n\alpha_{(\hat{\beta})}\beta^{\alpha_{(\hat{\beta})}-1} \neq 0$$

e) Compare $\hat{\beta}$ in section (b) and (d).

In part (b), we see that the likelihood function depends on the parameter α . But this is not the case in part (d) and it is independent of the parameter.

Part II

a) Is $L(\beta)$ at Part I in section (d) Likelihood?

No, it is basically Profile Likelihood. Because we fix a parameter and replace it with the relevant function and maximize the resulting probability.

In Likelihood Function we consider all the parameters.

But in Function Likelihood Profile, we remove disturbing parameters by expressing them as a function of desired parameters.

b) Let $(\alpha = 2, \beta = 1)$ generate 10000 and Compare $L(\beta)$, $L(\alpha, \beta)$

$$\text{likelihood: } L(\alpha, \beta) = \frac{\alpha^n}{\beta^{n\alpha}} \cdot \prod_{i=1}^n x_i^{\alpha-1} \cdot e^{-\sum_{i=1}^n \left(\frac{x_i}{\beta} \right)^\alpha} = 2^{10000} \cdot \prod_{i=1}^{10000} x_i \cdot e^{-\sum_{i=1}^{10000} x_i^2}$$

$$\text{Profile likelihood: } L(\beta) = L(\hat{\alpha}_{(\beta)}, \beta) = \frac{\alpha_{(\hat{\beta})}^n}{\beta^{n\alpha_{(\hat{\beta})}}} \cdot \prod_{i=1}^n x_i^{\alpha_{(\hat{\beta})}-1} \cdot e^{-\sum_{i=1}^n \left(\frac{x_i}{\beta} \right)^{\alpha_{(\hat{\beta})}}}$$

$$L(\beta) = \alpha_{\hat{\beta}} \cdot \prod_{i=1}^{10000} x_i^{\alpha_{\hat{\beta}} - 1} \cdot e^{-\sum_{i=1}^{10000} x_i^{\alpha_{\hat{\beta}}}}$$

$$\alpha_{\hat{\beta}} = \frac{n}{\sum_{i=1}^n \ln \left(\frac{x_i}{\beta_0} \right) \left(\left(\frac{x_i}{\beta_0} \right)^{\alpha} - 1 \right)} = \frac{10000}{\sum_{i=1}^{10000} \ln(x_i) \left((x_i)^2 - 1 \right)}$$

```
In [7]: #B)
# Load necessary library
library(ggplot2)

# Set parameters
alpha <- 2
beta <- 1
n <- 10000

# Generate samples
set.seed(123)
samples <- rweibull(n, shape = alpha, scale = beta)
```

```
In [8]: # Define the log-likelihood functions
logL_alpha_beta <- function(alpha, beta, x) {
  n <- length(x)
  logL <- n * log(alpha) - n * alpha * log(beta) + (alpha - 1) * sum(log(x)) - sum((x
  return(logL)
}

logL_beta <- function(beta, alpha, x) {
  n <- length(x)
  logL <- n * log(alpha) - n * alpha * log(beta) + (alpha - 1) * sum(log(x)) - sum((x
  return(logL)
}
```

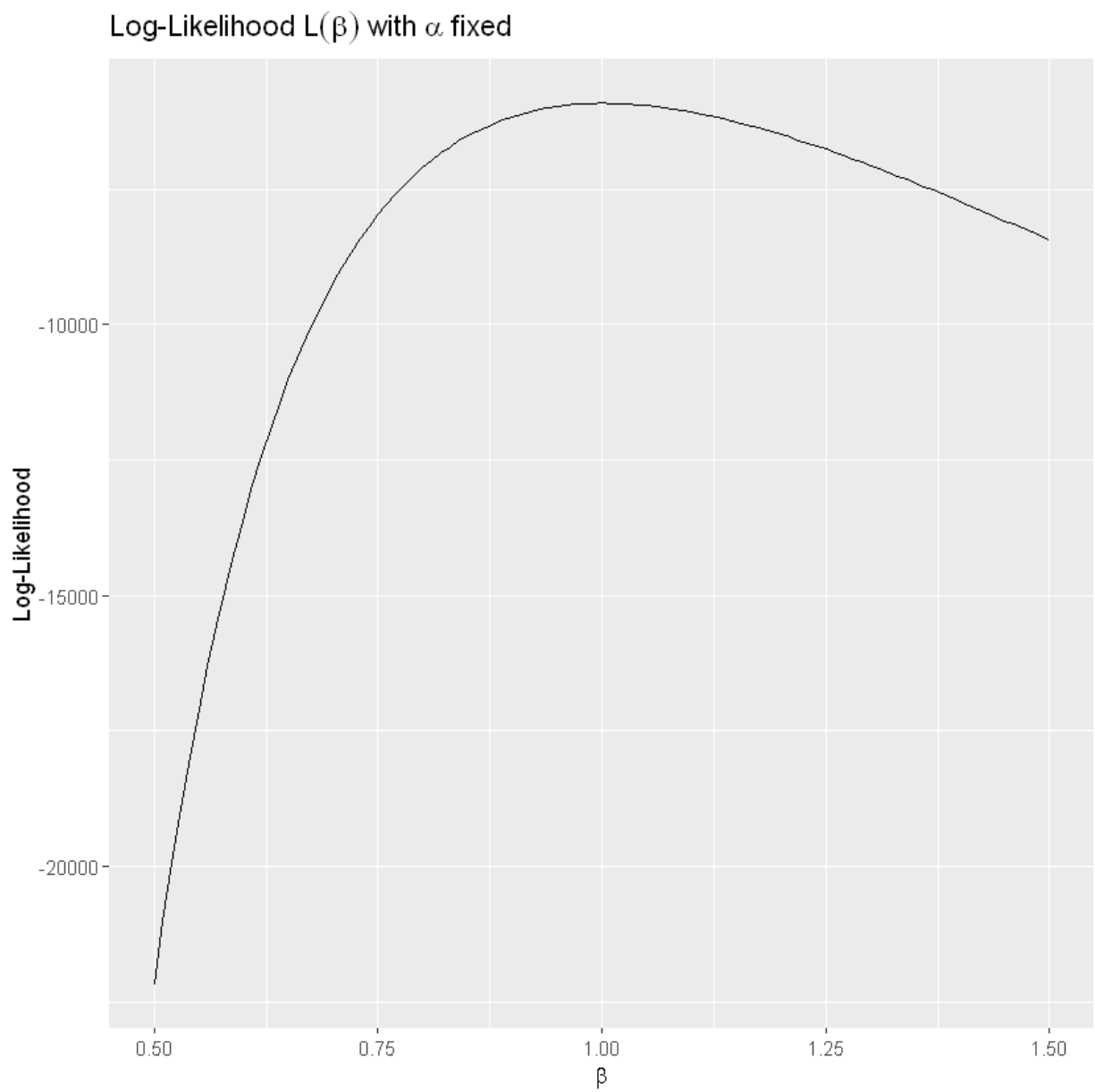
```
In [9]: # Define a range for beta
beta_range <- seq(0.5, 1.5, by = 0.01)

# Compute log-likelihood values for L(beta) with alpha fixed
logL_beta_values <- sapply(beta_range, function(b) logL_beta(b, alpha, samples))

# Compute log-likelihood values for L(alpha, beta)
logL_alpha_beta_values <- outer(beta_range, beta_range, Vectorize(function(a, b) logL_

# Convert to data frame for plotting
logL_beta_df <- data.frame(beta = beta_range, logL = logL_beta_values)
```

```
In [10]: # Plot log-likelihood for L(beta)
ggplot(logL_beta_df, aes(x = beta, y = logL)) +
  geom_line() +
  ggtitle(expression(paste("Log-Likelihood ", L(beta), " with ", alpha, " fixed"))) +
  xlab(expression(beta)) +
  ylab("Log-Likelihood")
```



```
In [15]: # Plot Log-Likelihood for  $L(\alpha, \beta)$ 
contour(beta_range, beta_range, logL_alpha_beta_values, xlab = expression(alpha), ylab = expression(beta),
        main = expression(paste("Log-Likelihood ", L(alpha, beta))),
        levels = pretty(range(logL_alpha_beta_values), 10))
```

Log-Likelihood $L(\alpha, \beta)$

