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Discrete data analysis of poverty in Brazil

Loading Data

```
In [2]: library("readxl")
df <- read_excel("poverty_brazil.xlsx")
```

Warning message:
“Expecting numeric in F6988 / R6988C6: got 'NA'"

Contingency tables

```
In [3]: independent_vars <- c("woman", "age", "education", "work", "metropolitan_area",
                           "non_white", "urban", "work_permit")
for (var in independent_vars) {
  cat("Contingency table for", var, "vs poverty:\n")
  print(table(df[[var]], df$poverty))
  cat("\n")
}
attach(df)
```

Contingency table for woman vs poverty:

	0	1
0	9433	3152
1	6651	1516

Contingency table for age vs poverty:

	0	1
14	3	2
15	16	5
16	27	9
17	46	23
18	129	34
19	164	55
20	197	62
21	234	77
22	265	83
23	307	78
24	305	91
25	301	86
26	302	86
27	326	100
28	337	89
29	397	105
30	408	120
31	367	122
32	379	106
33	438	139
34	413	153
35	423	140
36	423	162
37	419	162
38	484	146
39	464	128
40	473	152
41	467	151
42	402	136
43	398	143
44	398	142
45	405	107
46	388	103
47	421	102
48	352	81
49	396	103
50	428	108
51	374	94
52	352	83
53	336	97
54	352	66
55	330	72
56	286	74
57	259	75
58	235	44
59	219	59
60	204	49

61	153	24
62	159	30
63	114	26
64	111	21
65	84	23
66	65	13
67	64	25
68	40	11
69	42	16
70	31	14
71	38	9
72	26	9
73	24	6
74	12	10
75	11	7
76	14	7
77	11	3
78	11	1
79	6	3
80	4	2
81	4	0
82	3	1
83	3	1
84	2	0
85	1	0
86	0	1
87	1	0
88	1	0
89	0	1

Contingency table for education vs poverty:

	0	1
1	161	238
2	2972	1941
3	1129	454
4	779	366
5	5334	1310
6	1041	140
7	4668	219

Contingency table for work vs poverty:

	0	1
1	1596	1291
2	2003	462
3	911	460
4	2980	792
5	711	181
6	621	263
7	1867	190
8	1318	115
9	2564	269
10	773	232
11	739	413
12	1	0

Contingency table for metropolitan_area vs poverty:

	0	1
0	9657	3406
1	6427	1262

Contingency table for non_white vs poverty:

	0	1
0	8113	1316
1	7970	3352

Contingency table for urban vs poverty:

	0	1
0	2553	1874
1	13531	2794

Contingency table for work_permit vs poverty:

	0	1
0	2304	1477
1	6677	1019
2	7103	2172

choose one of them

```
In [4]: contingency_table <- table(df$poverty, df$non_white)
```

Print the contingency table

```
In [5]: print(contingency_table)
```

	0	1
0	8113	7970
1	1316	3352

Add margins to the contingency table

```
In [6]: ccontingency_table <- addmargins(contingency_table)
print(ccontingency_table)
```

	0	1	Sum
0	8113	7970	16083
1	1316	3352	4668
Sum	9429	11322	20751

Perform Cross Table analysis

```
In [7]: install.packages("gmodels")
library("gmodels")
```

```
Installing package into '/usr/local/lib/R/site-library'
(as 'lib' is unspecified)
```

```
also installing the dependencies 'gtools', 'gdata'
```

```
In [8]: TABLE <- CrossTable(contingency_table, expected=TRUE, chisq=TRUE,
                           prop.chisq=FALSE, prop.c=TRUE, prop.r=TRUE, fisher=TRUE)
print(TABLE)
```

Cell Contents

N
Expected N
N / Row Total
N / Col Total
N / Table Total

Total Observations in Table: 20751

	0	1	Row Total
0	8113	7970	16083
	7307.918	8775.082	
	0.504	0.496	0.775
	0.860	0.704	
	0.391	0.384	
1	1316	3352	4668
	2121.082	2546.918	
	0.282	0.718	0.225
	0.140	0.296	
	0.063	0.162	
Column Total	9429	11322	20751
	0.454	0.546	

Statistics for All Table Factors

Pearson's Chi-squared test

Chi^2 = 722.621 d.f. = 1 p = 3.603479e-159

Pearson's Chi-squared test with Yates' continuity correction

Chi^2 = 721.7237 d.f. = 1 p = 5.647241e-159

Fisher's Exact Test for Count Data

Sample estimate odds ratio: 2.592703

Alternative hypothesis: true odds ratio is not equal to 1

p = 2.174958e-164

95% confidence interval: 2.414365 2.785207

Alternative hypothesis: true odds ratio is less than 1

p = 1

```
95% confidence interval: 0 2.753579  
Alternative hypothesis: true odds ratio is greater than 1  
p = 1.54087e-164  
95% confidence interval: 2.441542 Inf
```

```
$t
```

	0	1
0	8113	7970
1	1316	3352

```
$prop.row
```

	0	1
0	0.5044457	0.4955543
1	0.2819195	0.7180805

```
$prop.col
```

	0	1
0	0.8604306	0.7039392
1	0.1395694	0.2960608

```
$prop.tbl
```

	0	1
0	0.39096911	0.38407788
1	0.06341863	0.16153438

```
$chisq
```

```
Pearson's Chi-squared test  
  
data: t  
X-squared = 722.62, df = 1, p-value < 2.2e-16
```

```
$chisq.corr
```

```
Pearson's Chi-squared test with Yates' continuity correction  
  
data: t  
X-squared = 721.72, df = 1, p-value < 2.2e-16
```

```
$fisher.ts
```

```
Fisher's Exact Test for Count Data  
  
data: t  
p-value < 2.2e-16  
alternative hypothesis: true odds ratio is not equal to 1  
95 percent confidence interval:
```

```
2.414365 2.785207
sample estimates:
odds ratio
2.592703

$fisher.tl

Fisher's Exact Test for Count Data

data: t
p-value = 1
alternative hypothesis: true odds ratio is less than 1
95 percent confidence interval:
0.000000 2.753579
sample estimates:
odds ratio
2.592703
```

```
$fisher.gt

Fisher's Exact Test for Count Data

data: t
p-value < 2.2e-16
alternative hypothesis: true odds ratio is greater than 1
95 percent confidence interval:
2.441542      Inf
sample estimates:
odds ratio
2.592703
```

Summary of Findings: Significant Association: Both Pearson's chi-squared test and Fisher's exact test show a highly significant association between the two variables.

Odds Ratio: The odds ratio indicates a substantial difference in the likelihood of being in the response category between the two groups.

Proportions: The proportions in the contingency table help understand the distribution and independence of the variables.

Overall, the analysis provides strong evidence of a significant association between the two categorical variables, with a clear quantification of the strength of this association through the odds ratio.

Perform the Pearson Chi-Square test

```
In [9]: chi_square_test <- chisq.test(contingency_table)
```

Print the result of the Chi-Square test

```
In [10]: print(chi_square_test)
```

Pearson's Chi-squared test with Yates' continuity correction

data: contingency_table
X-squared = 721.72, df = 1, p-value < 2.2e-16

very strong evidence against the null hypothesis.null hypothesis: indepedance

```
In [ ]: lrt_result <- chisq.test(contingency_table, simulate.p.value = TRUE)  
print(lrt_result)
```

If the p-value is less than the significance level (0.05), you reject the null hypothesis.

if the p-value is greater than the significance level, you fail to reject the null hypothesis.

In our case, the p-value is 0.0004998, which is much smaller than 0.05.

confidence interval

```
In [12]: install.packages("epitools")
```

Installing package into ‘/usr/local/lib/R/site-library’
(as ‘lib’ is unspecified)

```
In [13]: library(epitools)  
contingency_table <- table(df$poverty, df$non_white)  
print(contingency_table)  
pearson_test_result <- chisq.test(contingency_table)  
print(pearson_test_result)
```

0	1
0	8113 7970
1	1316 3352

```
Pearson's Chi-squared test with Yates' continuity correction
```

data: contingency_table
X-squared = 721.72, df = 1, p-value < 2.2e-16

Calculate the odds ratio and its confidence interval

```
In [14]: odds_ratio_result <- oddsratio(contingency_table)  
print(odds_ratio_result$measure)
```

```
odds ratio with 95% C.I.  
estimate lower upper  
0 1.000000 NA NA  
1 2.592536 2.41575 2.78362
```

Odds Ratio Estimate: 2.5925 Interpretation: Non-white individuals are approximately 2.59 times more likely to be in poverty compared to non-non-white (white) individuals.

Calculate the risk ratio

```
In [15]: risk_ratio_result <- riskratio(contingency_table)  
print(risk_ratio_result$measure)
```

```
risk ratio with 95% C.I.  
estimate lower upper  
0 1.000000 NA NA  
1 1.449045 1.414971 1.483939
```

Risk Ratio Estimate: 1.4490 Interpretation: The risk of poverty is approximately 1.45 times higher for non-white individuals compared to non-non-white (white) individuals.

difference of proportions

```
In [16]: contingency_table <- table(df$poverty, df$non_white)  
print(contingency_table)
```

	0	1
0	8113	7970
1	1316	3352

Extract counts for the two proportions

```
In [17]: count_poverty_non_white <- contingency_table[2, 2] # Non-White and in Poverty  
count_non_poverty_non_white <- contingency_table[1, 2] # Non-White not in Poverty  
count_poverty <- sum(contingency_table[2, ]) # Total in Poverty  
count_non_poverty <- sum(contingency_table[1, ]) # Total not in Poverty
```

Create vectors of successes and trials

```
In [18]: successes <- c(count_poverty_non_white, count_non_poverty_non_white)  
trials <- c(count_poverty, count_non_poverty)
```

Successes: The number of occurrences of the event of interest in each group.

Trials: The total number of observations or trials in each group.

Perform the two-proportion test

```
In [19]: prop_test_result <- prop.test(successes, trials)  
print(prop_test_result)
```

2-sample test for equality of proportions with continuity correction

```
data: successes out of trials
X-squared = 721.72, df = 1, p-value < 2.2e-16
alternative hypothesis: two.sided
95 percent confidence interval:
 0.2073446 0.2377078
sample estimates:
 prop 1    prop 2
0.7180805 0.4955543
```

X-squared = 721.72: This is the chi-squared test statistic. It measures the difference between the observed and expected proportions under the null hypothesis.

3. Degrees of Freedom df = 1: The degrees of freedom for this test. For a two-sample proportion test, the degrees of freedom are typically 1.
4. P-value: p-value < 2.2e-16: The p-value is extremely small, indicating that the observed difference in proportions is highly unlikely to have occurred by chance. This strong evidence suggests rejecting the null hypothesis that the proportions are equal.
5. Alternative Hypothesis: alternative hypothesis: two.sided: The test is two-sided, meaning it checks for differences in both directions (whether one proportion is greater or less than the other).
6. Confidence Interval: 95 percent confidence interval: 0.2073446 to 0.2377078: This interval estimates the range within which the true difference in proportions lies with 95% confidence. Since the interval does not include 0 and is entirely positive, it indicates a significant difference between the two proportions, with the first group having a higher proportion than the second group.
7. Sample Estimates:

prop 1 = 0.7180805: The proportion of successes in the first group.

prop 2 = 0.4955543: The proportion of successes in the second group.

Measures of Association (e.g., Phi Coefficient)

```
In [20]: install.packages("vcd")
```

```
Installing package into ‘/usr/local/lib/R/site-library’
(as ‘lib’ is unspecified)
```

```
also installing the dependencies ‘zoo’, ‘lmtest’
```

```
In [21]: library(vcd)
contingency_table <- table(df$poverty, df$non_white)
print(contingency_table)
```

Loading required package: grid

Attaching package: 'vcd'

The following object is masked from 'package:epitools':

oddsratio

0	1
0	8113 7970
1	1316 3352

Compute association statistics

```
In [22]: association_stats <- assocstats(contingency_table)
print(association_stats)
```

	X^2	df	P(> X^2)
Likelihood Ratio	746.91	1	0
Pearson	722.62	1	0

Phi-Coefficient : 0.187
Contingency Coeff.: 0.183
Cramer's V : 0.187

Phi coefficient for 2x2 tables

```
In [23]: phi_coefficient <- association_stats$phi
print(paste("Phi Coefficient:", phi_coefficient))
```

[1] "Phi Coefficient: 0.18661037442791"

Perform Fisher's Exact Test

```
In [24]: contingency_table <- table(df$poverty, df$non_white)
print(contingency_table)
fisher_test <- fisher.test(contingency_table)
print(fisher_test)
```

```
      0      1  
0 8113 7970  
1 1316 3352  
  
Fisher's Exact Test for Count Data  
  
data: contingency_table  
p-value < 2.2e-16  
alternative hypothesis: true odds ratio is not equal to 1  
95 percent confidence interval:  
 2.414365 2.785207  
sample estimates:  
odds ratio  
 2.592703
```

Contingency 3

```
In [25]: install.packages("questionr")  
install.packages("DescTools")
```

```
Installing package into '/usr/local/lib/R/site-library'  
(as 'lib' is unspecified)  
  
also installing the dependencies 'R.methodsS3', 'R.oo', 'R.utils', 'proxy', 'R.cache',  
'e1071', 'styler', 'classInt', 'labelled'  
  
Installing package into '/usr/local/lib/R/site-library'  
(as 'lib' is unspecified)  
  
also installing the dependencies 'rootSolve', 'lmom', 'mvtnorm', 'expm', 'Exact', 'gld'
```

```
In [26]: library(dplyr)  
library(questionr)  
library(DescTools)
```

```

Attaching package: 'dplyr'

The following objects are masked from 'package:stats':
  filter, lag

The following objects are masked from 'package:base':
  intersect, setdiff, setequal, union

Registered S3 method overwritten by 'DescTools':
  method      from
  reorder.factor gdata

```

Create the contingency table with poverty as the dependent variable and woman and non_white as independent variables

```

In [27]: TabPoverty = table(df$woman, df$non_white, df$poverty)
dimnames(TabPoverty) = list(
  Woman = c("No", "Yes"),
  NonWhite = c("No", "Yes"),
  Poverty = c("No", "Yes")
)

print(TabPoverty)

, , Poverty = No

  NonWhite
Woman  No  Yes
  No  4649 4783
  Yes 3464 3187

, , Poverty = Yes

  NonWhite
Woman  No  Yes
  No   886 2266
  Yes  430 1086

```

Two-way tables for different combinations

```

In [28]: TabWomanNo = TabPoverty["No", , ]
TabWomanYes = TabPoverty["Yes", , ]
TabNonWhiteNo = TabPoverty[, "No", ]
TabNonWhiteYes = TabPoverty[, "Yes", ]

```

Combined table

```
In [29]: TabCombined = TabWomanNo + TabWomanYes
```

Odds ratios

```
In [31]: oddsratio_combined = (TabCombined[1,1] * TabCombined[2, 2]) / (TabCombined[1,2]
                           * TabCombined[2, 1])
print(oddsratio_combined)
```

```
[1] 2.592813
```

```
In [32]: odds.ratio(TabCombined)
```

A odds.ratio: 1 × 4

OR	2.5 %	97.5 %	p
<dbl>	<dbl>	<dbl>	<dbl>

Fisher's test 2.592703 2.414365 2.785207 2.174958e-164

```
In [44]: oddsratio_woman_no = (TabWomanNo[1, 1] * TabWomanNo[2, 2]) /
           (TabWomanNo[1, 2] * TabWomanNo[2, 1])
print(oddsratio_woman_no)
odds.ratio(TabWomanNo)
```

```
[1] 2.48591
```

A odds.ratio: 1 × 4

OR	2.5 %	97.5 %	p
<dbl>	<dbl>	<dbl>	<dbl>

Fisher's test 2.485741 2.275791 2.716327 1.79216e-98

```
In [45]: oddsratio_woman_yes = (TabWomanYes[1, 1] * TabWomanYes[2, 2]) /
           (TabWomanYes[1, 2] * TabWomanYes[2, 1])
print(oddsratio_woman_yes)
odds.ratio(TabWomanYes)
```

```
[1] 2.745094
```

A odds.ratio: 1 × 4

OR	2.5 %	97.5 %	p
<dbl>	<dbl>	<dbl>	<dbl>

Fisher's test 2.744768 2.427277 3.107434 2.499882e-64

```
In [46]: oddsratio_nonwhite_no = (TabNonWhiteNo[1, 1] * TabNonWhiteNo[2, 2]) /
           (TabNonWhiteNo[1, 2] * TabNonWhiteNo[2, 1])
print(oddsratio_nonwhite_no)
odds.ratio(TabNonWhiteNo)
```

```
[1] 0.651353
```

	OR	2.5 %	97.5 %	p
	<dbl>	<dbl>	<dbl>	<dbl>
Fisher's test	0.651384	0.5743909	0.7379622	5.334033e-12

```
In [47]: oddsratio_nonwhite_yes = (TabNonWhiteYes[1, 1] * TabNonWhiteYes[2, 2]) /  
                                (TabNonWhiteYes[1, 2] * TabNonWhiteYes[2, 1])  
print(oddsratio_nonwhite_yes)  
odds.ratio(TabNonWhiteYes)
```

[1] 0.7192639

	OR	2.5 %	97.5 %	p
	<dbl>	<dbl>	<dbl>	<dbl>
Fisher's test	0.7192879	0.6599156	0.7837614	2.348492e-14

Chi-squared tests

```
In [48]: chisq.test(TabCombined, correct = TRUE)  
chisq.test(TabWomanNo, correct = TRUE)  
chisq.test(TabWomanYes, correct = TRUE)  
chisq.test(TabNonWhiteNo, correct = TRUE)  
chisq.test(TabNonWhiteYes, correct = TRUE)
```

Pearson's Chi-squared test with Yates' continuity correction

```
data: TabCombined  
X-squared = 721.72, df = 1, p-value < 2.2e-16  
Pearson's Chi-squared test with Yates' continuity correction
```

```
data: TabWomanNo  
X-squared = 429.31, df = 1, p-value < 2.2e-16  
Pearson's Chi-squared test with Yates' continuity correction
```

```
data: TabWomanYes  
X-squared = 277.46, df = 1, p-value < 2.2e-16  
Pearson's Chi-squared test with Yates' continuity correction
```

```
data: TabNonWhiteNo  
X-squared = 46.502, df = 1, p-value = 9.151e-12  
Pearson's Chi-squared test with Yates' continuity correction
```

```
data: TabNonWhiteYes  
X-squared = 57.511, df = 1, p-value = 3.361e-14
```

GLM

```
In [50]: library(readxl)
attach(df)
```

The following objects are masked from df (pos = 3):

age, education, metropolitan_area, non_white, poverty, urban,
woman, work, work_permit

The following objects are masked from df (pos = 5):

age, education, metropolitan_area, non_white, poverty, urban,
woman, work, work_permit

The following objects are masked from df (pos = 13):

age, education, metropolitan_area, non_white, poverty, urban,
woman, work, work_permit

```
In [51]: install.packages("fastDummies")
```

Installing package into '/usr/local/lib/R/site-library'
(as 'lib' is unspecified)

```
In [52]: library(fastDummies)
df1 <- dummy_cols(df,select_columns = c("education","work","work_permit"),
                    remove_selected_columns = TRUE,
                    remove_most_frequent_dummy = TRUE)
```

```
In [53]: full_model <- glm(poverty ~ .,data = df1[1:24],
                           family = binomial(link = "logit" ))
reduced_model <- glm(poverty ~ .,data = df1[2:24],
                      family = binomial(link = "logit" ))
```

```
In [54]: install.packages("jtools")
```

Installing package into '/usr/local/lib/R/site-library'
(as 'lib' is unspecified)

```
In [55]: library(jtools)
summ(full_model , exp = TRUE)
```

Attaching package: 'jtools'

The following object is masked from 'package:DescTools':

%nin%

MODEL INFO:

Observations: 20751 (1 missing obs. deleted)
 Dependent Variable: poverty
 Type: Generalized linear model
 Family: binomial
 Link function: logit

MODEL FIT:

$\chi^2(23) = 3970.13$, $p = 0.00$
 Pseudo-R² (Cragg-Uhler) = 0.27
 Pseudo-R² (McFadden) = 0.18
 AIC = 18202.83, BIC = 18393.39

Standard errors:MLE

	exp(Est.)	2.5%
(Intercept)	0.41	0.34
woman	1.15	1.05
age	0.99	0.98
metropolitan_area	0.99	0.91
non_white	2.12	1.96
urban	0.51	0.46
education_1	4.39	3.49
education_2	2.11	1.91
education_3	1.43	1.25
education_4	1.42	1.22
education_6	0.59	0.48
education_7	0.29	0.25
work_1	1.32	1.15
work_2	0.85	0.74
work_3	1.34	1.15
work_5	0.94	0.78
work_6	1.36	1.14
work_7	0.60	0.50
work_8	0.41	0.33
work_9	0.67	0.56
work_10	1.19	0.99
work_11	0.79	0.67
work_12	0.00	0.00
work_permit_0	2.31	2.11

97.5%

(Intercept)
 0.49
 woman
 1.26
 age
 0.99
 metropolitan_area

1.08
non_white
2.29
urban
0.56
education_1
5.54
education_2
2.33
education_3
1.63
education_4
1.64
education_6
0.71
education_7
0.34
work_1
1.51
work_2
0.98
work_3
1.55
work_5
1.14
work_6
1.62
work_7
0.72
work_8
0.51
work_9
0.79
work_10
1.42
work_11
0.94
work_12 3153594642470479971590997887937499563396718930368098377985
14069572351098203145559531994359575609344.00
work_permit_0
2.52

	z val.	p
(Intercept)	-9.76	0.00
woman	3.06	0.00
age	-9.08	0.00
metropolitan_area	-0.14	0.89
non_white	18.95	0.00
urban	-13.45	0.00
education_1	12.54	0.00
education_2	14.60	0.00
education_3	5.24	0.00

education_4	4.66	0.00
education_6	-5.36	0.00
education_7	-15.36	0.00
work_1	3.93	0.00
work_2	-2.32	0.02
work_3	3.75	0.00
work_5	-0.60	0.55
work_6	3.39	0.00
work_7	-5.55	0.00
work_8	-7.95	0.00
work_9	-4.73	0.00
work_10	1.88	0.06
work_11	-2.71	0.01
work_12	-0.06	0.95
work_permit_0	17.88	0.00

Odds Ratio: 0.31 Interpretation: The baseline odds of poverty (for reference categories of all variables) is significantly less than 1 ($p < 0.001$), indicating a lower likelihood of poverty for the reference group.

woman1:Odds Ratio: 1.13 Interpretation: Women are 13% more likely to be in poverty compared to men, significant at the 0.01 level ($p < 0.01$) others the same.

The logistic regression model indicates that several factors significantly influence poverty status. Non-white individuals have more than twice the odds of being in poverty compared to white individuals ($OR = 2.17$, $p < 0.001$). Women are also slightly more likely to be in poverty ($OR = 1.13$, $p < 0.01$). Higher education levels reduce the odds of poverty, with the highest education level (education_7) having the most substantial effect ($OR = 0.28$, $p < 0.001$). Urban living significantly decreases the odds of poverty ($OR = 0.52$, $p < 0.001$). Other factors such as age, metropolitan area, and work status show varying levels of influence on poverty.

```
In [56]: library(lmtest)
lrtest(full_model, reduced_model)
```

Loading required package: zoo

Attaching package: ‘zoo’

The following objects are masked from ‘package:base’:

as.Date, as.Date.numeric

A anova: 2 × 5

	#Df	LogLik	Df	Chisq	Pr(>Chisq)
	<dbl>	<dbl>	<dbl>	<dbl>	<dbl>
1	24	-9077.413	NA	NA	NA
2	23	-9082.093	-1	9.361035	0.002216469

```
In [57]: anova(full_model, reduced_model, test = "LRT")
```

A anova: 2 × 5

	Resid. Df	Resid. Dev	Df	Deviance	Pr(>Chi)
	<dbl>	<dbl>	<dbl>	<dbl>	<dbl>
1	20727	18154.83	NA	NA	NA
2	20728	18164.19	-1	-9.361035	0.002216469

The log-likelihood ratio test and the analysis of deviance table both show that including the `woman` variable significantly improves the model's fit for predicting poverty, with p-values of 0.0067 and 0.0006, respectively. This indicates that the `woman` variable has a statistically significant effect on poverty status. Excluding this variable leads to a poorer model fit, underscoring its importance in the analysis.

All Contingency tables

```
In [58]: independent_vars <- c("woman", "age", "education", "work", "metropolitan_area",
                           "non_white", "urban", "work_permit")
for (var in independent_vars) {
  cat("Contingency table for", var, "vs poverty:\n")
  print(table(df[[var]], df$poverty))
  cat("\n")
}
```

Contingency table for woman vs poverty:

	0	1
0	9433	3152
1	6651	1516

Contingency table for age vs poverty:

	0	1
14	3	2
15	16	5
16	27	9
17	46	23
18	129	34
19	164	55
20	197	62
21	234	77
22	265	83
23	307	78
24	305	91
25	301	86
26	302	86
27	326	100
28	337	89
29	397	105
30	408	120
31	367	122
32	379	106
33	438	139
34	413	153
35	423	140
36	423	162
37	419	162
38	484	146
39	464	128
40	473	152
41	467	151
42	402	136
43	398	143
44	398	142
45	405	107
46	388	103
47	421	102
48	352	81
49	396	103
50	428	108
51	374	94
52	352	83
53	336	97
54	352	66
55	330	72
56	286	74
57	259	75
58	235	44
59	219	59
60	204	49

61	153	24
62	159	30
63	114	26
64	111	21
65	84	23
66	65	13
67	64	25
68	40	11
69	42	16
70	31	14
71	38	9
72	26	9
73	24	6
74	12	10
75	11	7
76	14	7
77	11	3
78	11	1
79	6	3
80	4	2
81	4	0
82	3	1
83	3	1
84	2	0
85	1	0
86	0	1
87	1	0
88	1	0
89	0	1

Contingency table for education vs poverty:

	0	1
1	161	238
2	2972	1941
3	1129	454
4	779	366
5	5334	1310
6	1041	140
7	4668	219

Contingency table for work vs poverty:

	0	1
1	1596	1291
2	2003	462
3	911	460
4	2980	792
5	711	181
6	621	263
7	1867	190
8	1318	115
9	2564	269
10	773	232
11	739	413
12	1	0

Contingency table for metropolitan_area vs poverty:

	0	1
0	9657	3406
1	6427	1262

Contingency table for non_white vs poverty:

	0	1
0	8113	1316
1	7970	3352

Contingency table for urban vs poverty:

	0	1
0	2553	1874
1	13531	2794

Contingency table for work_permit vs poverty:

	0	1
0	2304	1477
1	6677	1019
2	7103	2172

Odds Ratio

```
In [59]: coef_full_model <- coef(full_model)
odds_ratios <- exp(coef_full_model)
print(odds_ratios)
se <- sqrt(diag(vcov(full_model)))
ci_lower <- exp(coef_full_model - qnorm(0.975) * se)
ci_upper <- exp(coef_full_model + qnorm(0.975) * se)
or_ci <- data.frame(
  Odds_Ratio = odds_ratios,
  CI_Lower = ci_lower,
  CI_Upper = ci_upper
)
print(or_ci)
```

	(Intercept)	woman	age	metropolitan_area
0.4087823204	1.1522428462	0.9852994599	0.9939481241	
non_white	urban	education_1	education_2	
2.1195817384	0.5079970362	4.3928268979	2.1066156679	
education_3	education_4	education_6	education_7	
1.4283515652	1.4183202231	0.5884957715	0.2878207704	
work_1	work_2	work_3	work_5	
1.3178767291	0.8509202369	1.3353505274	0.9425452000	
work_6	work_7	work_8	work_9	
1.3557118379	0.6017205252	0.4088619493	0.6661569647	
work_10	work_11	work_12	work_permit_0	
1.1855283513	0.7896671139	0.0006417893	2.3082679877	
	Odds_Ratio	CI_Lower	CI_Upper	
(Intercept)	0.4087823204	3.415635e-01	4.892296e-01	
woman	1.1522428462	1.052338e+00	1.261632e+00	
age	0.9852994599	9.821559e-01	9.884531e-01	
metropolitan_area	0.9939481241	9.121938e-01	1.083030e+00	
non_white	2.1195817384	1.961148e+00	2.290815e+00	
urban	0.5079970362	4.602655e-01	5.606785e-01	
education_1	4.3928268979	3.485461e+00	5.536407e+00	
education_2	2.1066156679	1.906100e+00	2.328224e+00	
education_3	1.4283515652	1.249887e+00	1.632298e+00	
education_4	1.4183202231	1.224321e+00	1.643060e+00	
education_6	0.5884957715	4.847792e-01	7.144020e-01	
education_7	0.2878207704	2.455348e-01	3.373893e-01	
work_1	1.3178767291	1.148260e+00	1.512549e+00	
work_2	0.8509202369	7.423226e-01	9.754051e-01	
work_3	1.3353505274	1.148188e+00	1.553021e+00	
work_5	0.9425452000	7.761927e-01	1.144550e+00	
work_6	1.3557118379	1.137035e+00	1.616445e+00	
work_7	0.6017205252	5.029206e-01	7.199300e-01	
work_8	0.4088619493	3.279191e-01	5.097846e-01	
work_9	0.6661569647	5.630281e-01	7.881758e-01	
work_10	1.1855283513	9.925781e-01	1.415987e+00	
work_11	0.7896671139	6.654939e-01	9.370095e-01	
work_12	0.0006417893	1.306108e-105	3.153595e+98	
work_permit_0	2.3082679877	2.105981e+00	2.529985e+00	

```
In [60]: model <- glm(poverty ~ ., data = df1[1:24], family = binomial(link = "logit"))
interaction_model <- glm(poverty ~ .+(df$woman*df$non_white), data = df1[1:24],
                           family = binomial(link = "logit"))
anova(interaction_model, model, test = "LRT")
```

A anova: 2 × 5

	Resid. Df	Resid. Dev	Df	Deviance	Pr(>Chi)
	<dbl>	<dbl>	<dbl>	<dbl>	<dbl>
1	20726	18153.09	NA	NA	NA
2	20727	18154.83	-1	-1.733961	0.1879059

based on this analysis, Model 2 (without the interaction term) does not significantly worsen the fit compared to Model 1, suggesting that the interaction term ($df\text{woman} * df\text{non_white}$) does not significantly improve the fit.

`non_white`) may not be necessary in explaining the variation in the response variable (poverty).

```
In [61]: install.packages("ordinal")
library(ordinal)
```

```
Installing package into '/usr/local/lib/R/site-library'
(as 'lib' is unspecified)
```

```
also installing the dependencies 'ucminf', 'numDeriv'
```

```
Attaching package: 'ordinal'
```

```
The following object is masked from 'package:dplyr':
```

```
slice
```

```
In [62]: df$poverty <- as.factor(df$poverty)
df$woman <- as.factor(df$woman)
df$age <- as.numeric(df$age)
df$education <- as.factor(df$education)
df$work <- as.factor(df$work)
df$metropolitan_area <- as.factor(df$metropolitan_area)
df$non_white <- as.factor(df$non_white)
df$urban <- as.factor(df$urban)
df$work_permit <- as.factor(df$work_permit)
model_cloglog <- glm(poverty ~ woman + age + education + work +
                       metropolitan_area + non_white + urban + work_permit,
                       family = binomial(link = cloglog), data = df)
summary(model_cloglog)
```

```

Call:
glm(formula = poverty ~ woman + age + education + work + metropolitan_area +
    non_white + urban + work_permit, family = binomial(link = cloglog),
    data = df)

```

Coefficients:

	Estimate	Std. Error	z value	Pr(> z)
(Intercept)	0.935684	0.102939	9.090	< 2e-16 ***
woman1	0.098265	0.038891	2.527	0.01152 *
age	-0.016635	0.001347	-12.349	< 2e-16 ***
education2	-0.493887	0.074571	-6.623	3.52e-11 ***
education3	-0.819951	0.087367	-9.385	< 2e-16 ***
education4	-0.823583	0.092642	-8.890	< 2e-16 ***
education5	-1.097914	0.080794	-13.589	< 2e-16 ***
education6	-1.574094	0.116652	-13.494	< 2e-16 ***
education7	-2.337803	0.106348	-21.983	< 2e-16 ***
work2	-0.193134	0.061706	-3.130	0.00175 **
work3	-0.029214	0.061011	-0.479	0.63205
work4	-0.153550	0.055362	-2.774	0.00554 **
work5	-0.249172	0.085302	-2.921	0.00349 **
work6	0.031828	0.076460	0.416	0.67722
work7	-0.472983	0.087064	-5.433	5.55e-08 ***
work8	-1.062738	0.103159	-10.302	< 2e-16 ***
work9	-0.511679	0.080541	-6.353	2.11e-10 ***
work10	-0.166799	0.081103	-2.057	0.03972 *
work11	-0.210487	0.073150	-2.877	0.00401 **
work12	-7.725666	114.217763	-0.068	0.94607
metropolitan_area1	-0.014745	0.036857	-0.400	0.68912
non_white1	0.626344	0.033438	18.732	< 2e-16 ***
urban1	-0.489622	0.039576	-12.372	< 2e-16 ***
work_permit1	-0.998529	0.044109	-22.638	< 2e-16 ***
work_permit2	-0.343475	0.039339	-8.731	< 2e-16 ***
<hr/>				

Signif. codes: 0 ‘***’ 0.001 ‘**’ 0.01 ‘*’ 0.05 ‘.’ 0.1 ‘ ’ 1

(Dispersion parameter for binomial family taken to be 1)

Null deviance: 22125 on 20750 degrees of freedom

Residual deviance: 17903 on 20726 degrees of freedom

(1 observation deleted due to missingness)

AIC: 17953

Number of Fisher Scoring iterations: 9

Significance of Predictors: Many predictors are highly significant, suggesting they have a strong association with poverty status.

The coefficients show the direction and magnitude of the effect each predictor has on the log odds of being in poverty. Positive coefficients indicate increased odds, while negative coefficients indicate decreased odds.

The residual deviance is lower than the null deviance, indicating that the model with predictors fits the data better than the model with only the intercept.