

# Homework 1

Zihan Zhao

1001103708

**1**

**(a)**

First determine  $E[Z]$

$$\begin{aligned} E[Z] &= E[(X - Y)^2] \\ &= E[(X^2 - 2XY + Y^2)] \\ &= E[X^2] - E[2XY] + E[Y^2] \\ &= E[X^2] - 2E[X]E[Y] + E[Y^2] \end{aligned}$$

Since X and Y both are variables from uniform distribution in  $[0, 1]$ ,  $E[x] = E[Y]$  and  $E[X^2] = E[Y^2]$ . Therefore,

*continue :*

$$\begin{aligned} E[Z] &= 2E[X^2] - 2E[X]^2 \\ &= 2 \int_0^1 x^2 f(x) dx - 2 \left( \int_0^1 x f(x) dx \right)^2 \end{aligned}$$

Note that  $f(x)$  is the pdf of X. So  $f(x) = \frac{1}{1-0} = 1$ . Therefore,

*continue :*

$$\begin{aligned} E[Z] &= 2 \int_0^1 x^2 dx - 2 \left( \int_0^1 x dx \right)^2 \\ &= 2 * \frac{1}{3} - 2 * \left( \frac{1}{2} \right)^2 \\ &= \frac{1}{6} = 0.16667 \end{aligned}$$

Second determine  $Var[Z]$

$$\begin{aligned} Var[Z] &= E[Z^2] - E[Z]^2 \\ &= E[(X - Y)^4] - \left( \frac{1}{6} \right)^2 \\ &= E[X^4 - 4X^3Y + 6X^2Y^2 - 4XY^3 + Y^4] - \left( \frac{1}{6} \right)^2 \\ &= E[X^4] - 4E[X^3]E[Y] + 6E[X^2]E[Y^2] - 4E[X]E[Y^3] + E[Y^4] - \left( \frac{1}{6} \right)^2 \end{aligned}$$

Since X and Y both are variables from uniform distribution in  $[0, 1]$ ,  $E[x] = E[Y]$ ,  $E[X^2] = E[Y^2]$ ,  $E[X^3] = E[Y^3]$ , and  $E[X^4] = E[Y^4]$  Therefore,

$$\begin{aligned} Var[Z] &= 2E[X^4] - 8E[X^3]E[X] + 6E[X^2]^2 - \left(\frac{1}{6}\right)^2 \\ &= 2 \int_0^1 x^4 f(x) dx - 8 \left( \int_0^1 x^3 f(x) dx \right) \left( \int_0^1 x f(x) dx \right) + 6 \left( \int_0^1 x^2 f(x) dx \right)^2 - \left(\frac{1}{6}\right)^2 \\ &= 2 * \frac{1}{5} - 8 * \frac{1}{4} * \frac{1}{2} + 6 \left(\frac{1}{3}\right)^2 - \left(\frac{1}{6}\right)^2 = \frac{7}{180} = 0.38889 \end{aligned}$$

**(b)**

First determine  $E[R]$ . Since  $X_1, \dots, Y_d$  are i.i.d,  $(X_1 - Y_1)^2, \dots, (X_d - Y_d)^2$  are independent. Therefore,

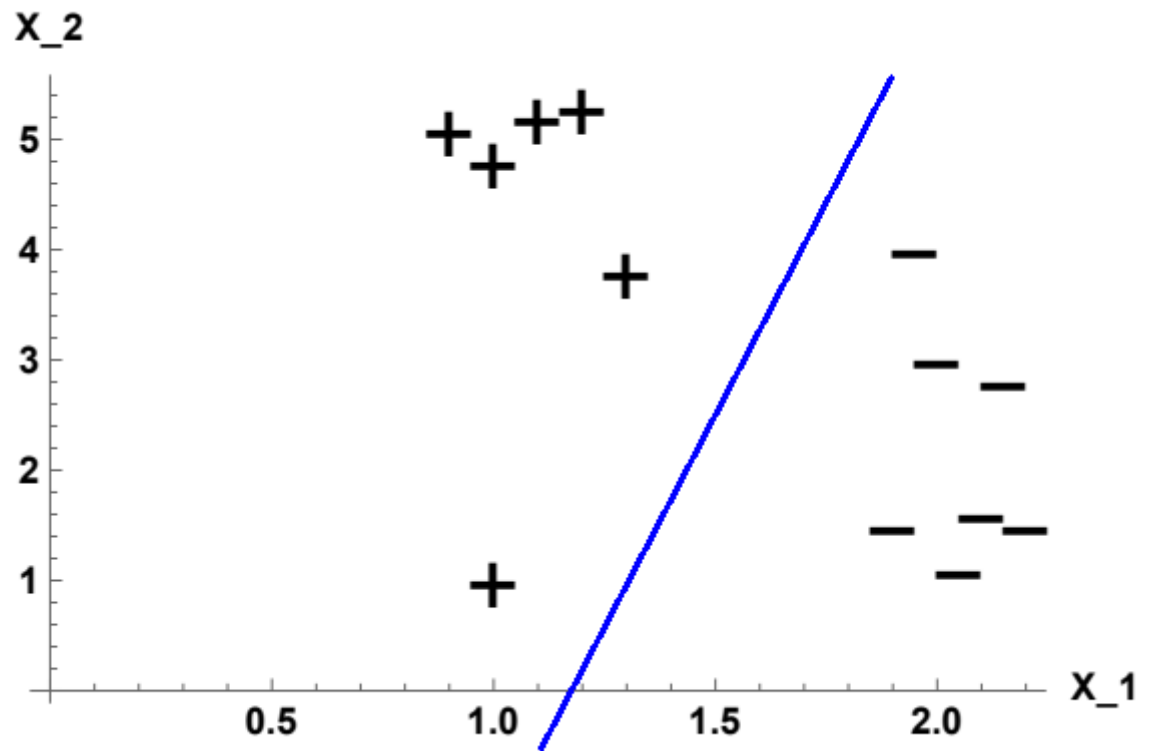
$$\begin{aligned} E[R] &= E\left[\sum_{i=1}^d Z_i\right] \\ &= \sum_{i=1}^d E[Z_i] \\ &= dE[Z] = \frac{d}{6} \end{aligned}$$

Second determine  $Var[R]$ . Since  $Z_1, \dots, Z_d$  are independent,  $Cov(Z_i, Z_j) = 0$  where i and j are arbitrary picks from 1 to d. Therefore,

$$\begin{aligned} Var[R] &= Var\left[\sum_{i=1}^d Z_i\right] \\ &= \sum_{i=1}^d Var[Z_i] + 2 \sum_{i < j} Cov[Z_i, Z_j] \\ &= dVar[Z] = \frac{7d}{180} \end{aligned}$$

2

(a)



(b)

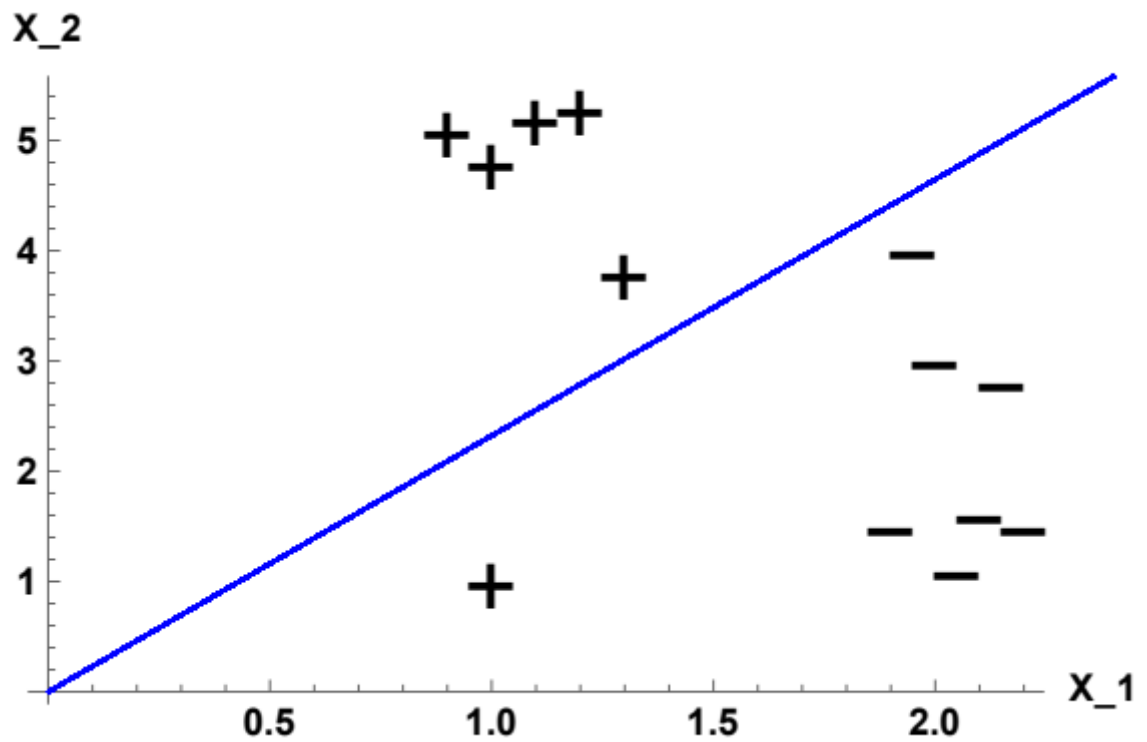
No, it is not unique.

(c)

0. The left side of the boundary marks "+" and the right side marks "-". All "+" points are in left side and all "-" points are in right side.

(d)

As  $\lambda$  approaches  $\infty$ ,  $J(w)$  heavily depends on  $\lambda\omega_0^2$ . In order to minimize  $\lambda\omega_0^2$ ,  $\omega_0$  should be 0. So the boundary should pass through  $(0,0)$ . The sketch is shown below.

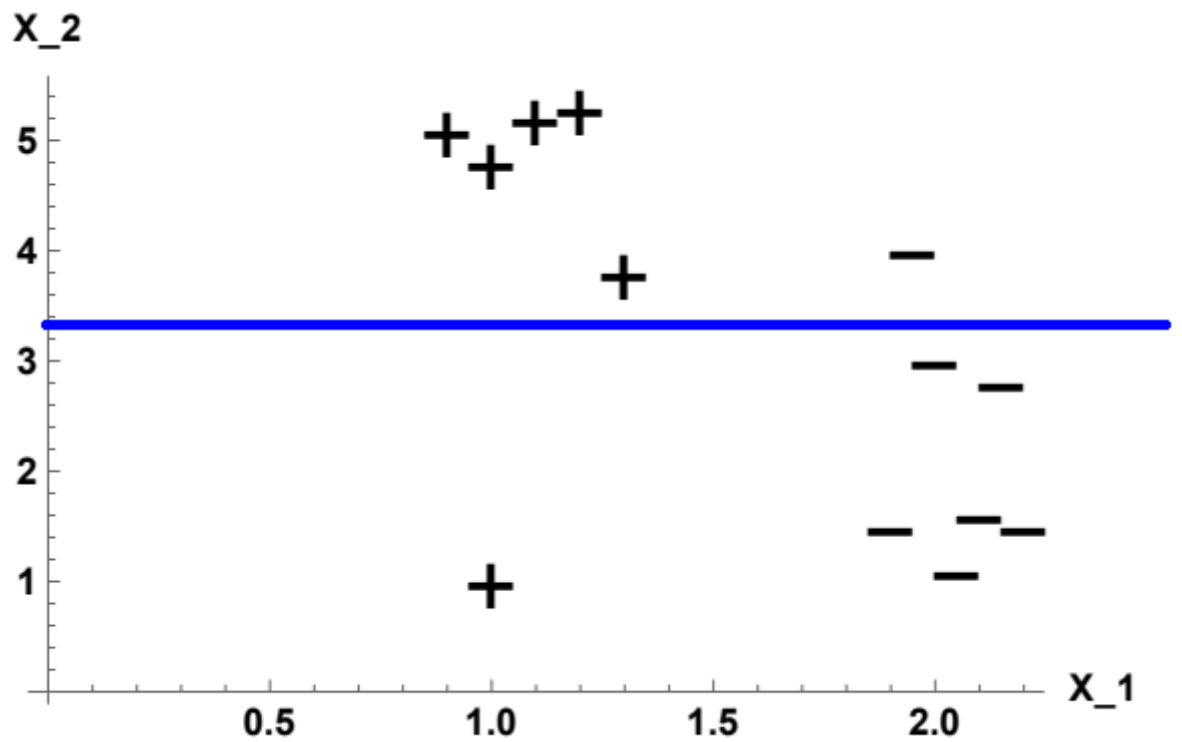


(e)

1. The upper side of the boundary marks "+" and the down side marks "-". All "+" points are in left side except one point is at down side. All "-" points are in right side. So the error is 1.

(f)

Similar with (d), in order to minimize  $\lambda\omega_1^2$ ,  $\omega_1$  should be 0. So the boundary line is  $\omega_0 + \omega_2x_2 = 0$ . It should be a horizontal line. The sketch is shown below.

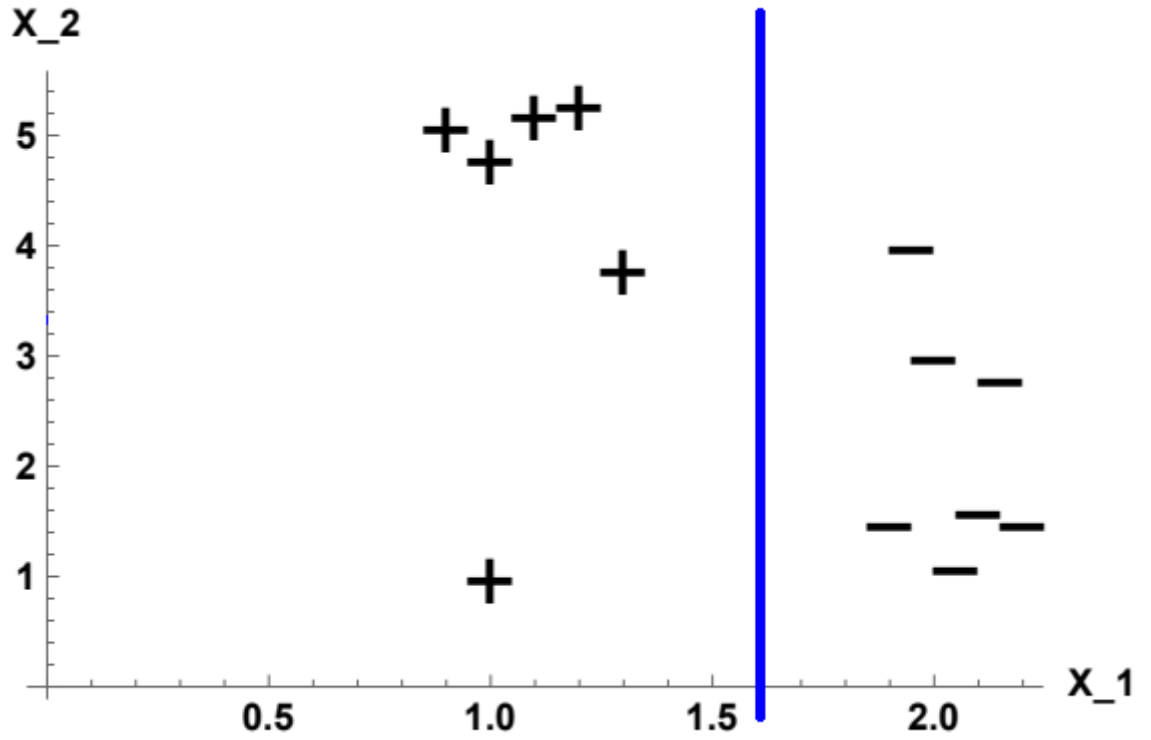


(g)

2. The upper side of the boundary marks "+" and the down side marks "-". All "+" points are in left side except one point is at down side. All "-" points are in right side except one point is at upper side. So the error is 2.

(h)

Similar with (d), in order to minimize  $\lambda\omega_2^2$ ,  $\omega_2$  should be 0. So the boundary line is  $\omega_0 + \omega_1x_1 = 0$ . It should be a vertical line. The sketch is shown below.



(i)

0. The left side of the boundary marks "+" and the right side marks "-". All "+" points are in left side and all "-" points are in right side.

**3**

(a)

Since  $p(x)$  is possibility mass,  $0 \leq p(x) \leq 1$ . Therefore,

$$\begin{aligned} p(x) &\leq 1 \\ \frac{1}{p(x)} &\leq 1 \\ \log\left(\frac{1}{p(x)}\right) &\geq 1 \geq 0 \end{aligned}$$

since  $p(x) \geq 0$ ,

$$\begin{aligned} p(x) \log\left(\frac{1}{p(x)}\right) &\geq 0 \\ H(X) = \sum_x p(x) \log\left(\frac{1}{p(x)}\right) &\geq 0 \end{aligned}$$

(b)

In order to prove  $KL(p||q) \geq 0$ , it is equivalent to prove  $\sum_x^X p(x) \log(\frac{p(x)}{q(x)}) \geq 0$ . Let's define new variable  $Y$  which is  $y = f(x) = \frac{q(x)}{p(x)}$ . Then let's define a new function  $\phi(y) = -\log y$ , which is convex function. According to Jensen's Inequality,

$$\begin{aligned} E[\phi(Y)] &\geq \phi(E[Y]) \\ \sum_x^X p(x) \phi(f(x)) &\geq \phi(E[Y]) \\ -\sum_x^X p(x) \log(\frac{q(x)}{p(x)}) &\geq \phi(E[Y]) \\ \sum_x^X p(x) \log(\frac{p(x)}{q(x)}) &\geq \phi(E[Y]) \\ KL(p||q) &\geq \phi(E[Y]) \\ KL(p||q) &\geq \phi(E_x[f(x)]) \\ KL(p||q) &\geq \phi(\sum_x^X p(x) \frac{q(x)}{p(x)}) \\ KL(p||q) &\geq \phi(\sum_x^X q(x)) \\ KL(p||q) &\geq \phi(1) \\ KL(p||q) &\geq -\log(1) \\ KL(p||q) &\geq 0 \end{aligned}$$

(c)

In order to show that

$$I[Y; X] = KL(p(x, y)||p(x)p(y))$$

On left side,

$$\begin{aligned} I[Y; X] &= H(Y) - H(Y|X) \\ &= \sum_y p(y) \log(\frac{1}{p(y)}) - (-\sum_y \sum_x p(x, y) \log(p(y|x))) \\ &= \sum_y p(y) \log(\frac{1}{p(y)}) + \sum_y \sum_x p(x, y) \log(p(y|x)) \end{aligned}$$

On right side,

$$KL(p(x, y)||p(x)p(y)) = \sum_y \sum_x p(x, y) \log(\frac{p(x, y)}{p(x)p(y)})$$

note that  $p(y|x) = \frac{p(x,y)}{p(x)}$

$$\begin{aligned}
KL(p(x,y)||p(x)p(y)) &= \sum_y \sum_x p(x,y) \log\left(\frac{p(y|x)}{p(y)}\right) \\
&= \sum_y \sum_x (p(x,y)(\log(p(y|x)) + \log\left(\frac{1}{p(y)}\right))) \\
&= \sum_y \sum_x (p(x,y)(\log(p(y|x))) + \sum_y \sum_x (p(x,y) \log\left(\frac{1}{p(y)}\right))) \\
&= \sum_y \sum_x (p(x,y)(\log(p(y|x))) + \sum_y (p(y) \log\left(\frac{1}{p(y)}\right))) \\
&= I[Y; X]
\end{aligned}$$

Therefore right side is equal to left side.

Done.