Homework 1

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(a)

First determine E[Z]

$$\begin{split} E[Z] &= E[(X-Y)^2] \\ &= E[(X^2 - 2XY + Y^2] \\ &= E[X^2] - E[2XY] + E[Y^2] \\ &= E[X^2] - 2E[X]E[Y] + E[Y^2] \end{split}$$

Since X and Y both are variables from uniform distribution in [0,1], E[x] = E[Y] and $E[X^2] = E[Y^2]$. Therefore,

continue:

$$E[Z] = 2E[X^{2}] - 2E[X]^{2}$$
$$= 2 \int_{0}^{1} x^{2} f(x) dx - 2(\int_{0}^{1} x f(x) dx)^{2}$$

Note that f(x) is the pdf of X. So $f(x) = \frac{1}{1-0} = 1$. Therefore,

continue:

$$E[Z] = 2 \int_0^1 x^2 dx - 2(\int_0^1 x dx)^2$$
$$= 2 * \frac{1}{3} - 2 * (\frac{1}{2})^2$$
$$= \frac{1}{6} = 0.16667$$

Second determine Var[Z]

$$\begin{split} Var[Z] &= E[Z^2] - E[Z]^2 \\ &= E[(X - Y)^4] - (\frac{1}{6})^2 \\ &= E[X^4 - 4X^3Y + 6X^2Y^2 - 4XY^3 + Y^4] - (\frac{1}{6})^2 \\ &= E[X^4] - 4E[X^3]E[Y] + 6E[X^2]E[Y^2] - 4E[X]E[Y^3] + E[Y^4] - (\frac{1}{6})^2 \end{split}$$

Since X and Y both are variables from uniform distribution in [0,1], E[x] = E[Y], $E[X^2] = E[Y^2]$, $E[X^3] = E[Y^3]$, and $E[X^4] = E[Y^4]$ Therefore,

$$\begin{split} Var[Z] &= 2E[X^4] - 8E[X^3]E[X] + 6E[X^2]^2 - (\frac{1}{6})^2 \\ &= 2\int_0^1 x^4 f(x) dx - 8(\int_0^1 x^3 f(x) dx)(\int_0^1 x f(x) dx) + 6(\int_0^1 x^2 f(x) dx)^2 - (\frac{1}{6})^2 \\ &= 2*\frac{1}{5} - 8*\frac{1}{4}*\frac{1}{2} + 6(\frac{1}{3})^2 - (\frac{1}{6})^2 = \frac{7}{180} = 0.038889 \end{split}$$

(b)

First determine E[R]. Since $X_1, ..., Y_d$ are i.i.d, $(X_1 - Y_1)^2, ..., (X_d - Y_d)^2$ are independent. Therefore,

$$E[R] = E\left[\sum_{i=1}^{d} Z_i\right]$$
$$= \sum_{i=1}^{d} E[Z_i]$$
$$= dE[Z] = \frac{d}{6}$$

Second determine Var[R]. Since $Z_1, ..., Z_d$ are independent, $Cov(Z_i, Z_j) = 0$ where i and j are arbitrary picks from 1 to d. Therefore,

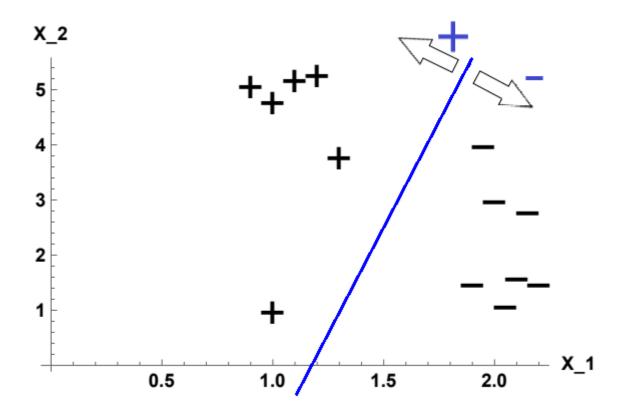
$$Var[R] = Var\left[\sum_{i=1}^{d} Z_i\right]$$

$$= \sum_{i=1}^{d} Var[Z_i] + 2\sum_{i< j}^{d} Cov[Z_i, Z_j]$$

$$= dVar[Z] = \frac{7d}{180}$$

 $\mathbf{2}$

(a)



(b)

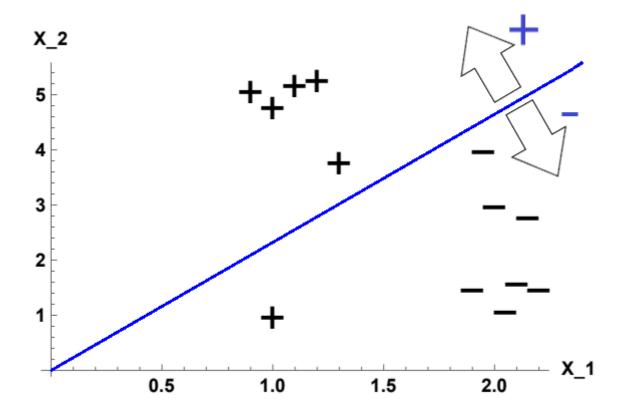
No, it is not unique.

(c)

0. The left side of the boundary marks "+" and the right side marks "-". All "+" points are in left side and all "-" points are in right side.

(d)

As λ approaches ∞ , J(w) heavily depends on $\lambda\omega_0^2$. In order to minimize $\lambda\omega_0^2$, ω_0 should be 0. So the boundary should pass through (0,0). The sketch is shown below.

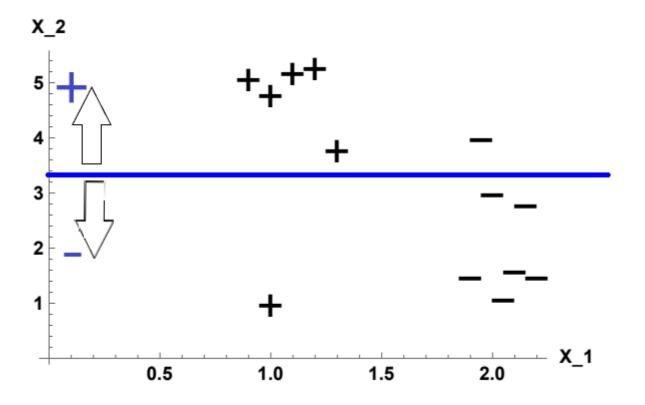


(e)

1. The upper side of the boundary marks "+" and the down side marks "-". All "+" points are in left side except one point is at down side. All "-" points are in right side. So the error is 1.

(f)

Similar with ??, in order to minimize $\lambda \omega_1^2$, ω_1 should be 0. So the boundary line is $\omega_0 + \omega_2 x_2 = 0$. It should be a horizontal line. The sketch is shown below.

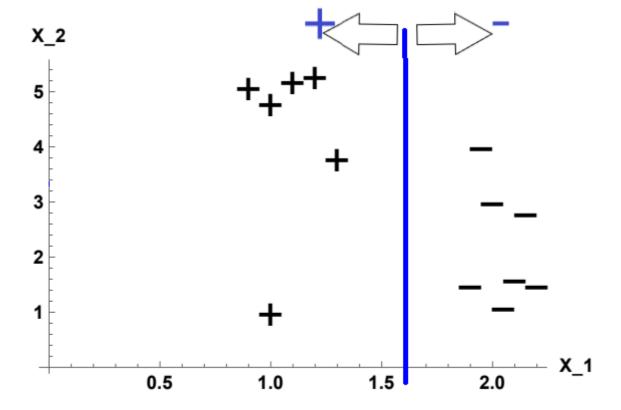


(g)

2. The upper side of the boundary marks "+" and the down side marks "-". All "+" points are in left side except one point is at down side. All "-" points are in right side except one point is at upper side. So the error is 2.

(h)

Similar with ??, in order to minimize $\lambda \omega_2^2$, ω_2 should be 0. So the boundary line is $\omega_0 + \omega_1 x_1 = 0$. It should be a vertical line. The sketch is shown below.



(i)

0. The left side of the boundary marks "+" and the right side marks "-". All "+" points are in left side and all "-" points are in right side.

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(a)

Since p(x) is possibility mass, $0 \le p(x) \le 1$. Therefore,

$$p(x) \le 1$$

$$\frac{1}{p(x)} \le 1$$

$$\log(\frac{1}{p(x)}) \ge 1 \ge 0$$

since $p(x) \ge 0$,

$$p(x)\log(\frac{1}{p(x)}) \ge 0$$

$$H(X) = \sum_{x} p(x)\log(\frac{1}{p(x)}) \ge 0$$

(b)

In order to prove $KL(p||q) \ge 0$, it is equivalent to prove $\sum_{x}^{X} p(x) \log(\frac{p(x)}{q(x)}) \ge 0$. Let's define new variable Y which is $y = f(x) = \frac{q(x)}{p(x)}$. Then let's define a new function $\phi(y) = -\log y$, which is convex function. According to Jensen's Inequality,

$$E[\phi(Y)] \ge \phi(E[Y])$$

$$\sum_{x}^{X} p(x)\phi(f(x)) \ge \phi(E[Y])$$

$$-\sum_{x}^{X} p(x)\log(\frac{q(x)}{p(x)}) \ge \phi(E[Y])$$

$$\sum_{x}^{X} p(x)\log(\frac{p(x)}{q(x)}) \ge \phi(E[Y])$$

$$KL(p||q) \ge \phi(E[Y])$$

$$KL(p||q) \ge \phi(E_x[f(x)])$$

$$KL(p||q) \ge \phi(\sum_{x}^{X} p(x)\frac{q(x)}{p(x)})$$

$$KL(p||q) \ge \phi(\sum_{x}^{X} q(x))$$

$$KL(p||q) \ge \phi(1)$$

$$KL(p||q) \ge -\log(1)$$

$$KL(p||q) \ge 0$$

(c)

In order to show that

$$I[Y;X] = KL(p(x,y)||p(x)p(y))$$

On left side,

$$\begin{split} I[Y;X] &= H(Y) - H(Y|X) \\ &= \sum_{y} p(y) \log(\frac{1}{p(y)}) - (-\sum_{y} \sum_{x} p(x,y) \log(p(y|x))) \\ &= \sum_{y} p(y) \log(\frac{1}{p(y)}) + \sum_{y} \sum_{x} p(x,y) \log(p(y|x)) \end{split}$$

On right side,

$$KL(p(x,y)||p(x)p(y)) = \sum_{y} \sum_{x} p(x,y) \log(\frac{p(x,y)}{p(x)p(y)})$$

note that $p(y|x) = \frac{p(x,y)}{p(x)}$

$$\begin{split} KL(p(x,y)||p(x)p(y)) &= \sum_{y} \sum_{x} p(x,y) \log(\frac{p(y|x)}{p(y)}) \\ &= \sum_{y} \sum_{x} (p(x,y) (\log(p(y|x)) + \log(\frac{1}{p(y)}))) \\ &= \sum_{y} \sum_{x} (p(x,y) (\log(p(y|x))) + \sum_{y} \sum_{x} (p(x,y) \log(\frac{1}{p(y)}))) \\ &= \sum_{y} \sum_{x} (p(x,y) (\log(p(y|x))) + \sum_{y} (p(y) \log(\frac{1}{p(y)}))) \\ &= I[Y;X] \end{split}$$

Therefore right side is equal to left side. Done.