$$F_{X_1, X_2}(x_1, x_2) = P\{X_1 \le x_1, X_2 \le x_2\}$$

$$= \sum_{X_1 \le x_1} \sum_{X_2 \le x_2} P(x_1, x_2)$$

where  $p(x_1, x_2) = P(X_1 = x_1, X_2 = x_2)$ .

The joint(bivariate) PMF uniquely defines cdf. It also is charaterized by the two properties

(i) 
$$0 \le p(x_1, x_2) \le 1$$
  
(ii)  $\sum_{all x_1} \sum_{all x_2} p(x_1, x_2) = 1$ 

### Theorem(Prove):

For all reals, a < b, c < d,

$$P\{a \le x_1 \le b, c \le x_2 \le d\} = F_{x_1, x_2}(b, d) - F_{x_1, x_2}(a, d) - F_{x_1, x_2}(b, c) + F_{x_1, x_2}(a, c)$$

A random vector  $(x_1, x_2)$  with space D is of continuous type if its CDF  $F_{x_1, x_2}(x_1, x_2)$  is continuous. The function can be expressed as

$$F_{x_1,x_2}(x_1,x_2) = \int_{-\infty}^{x_1} \int_{-\infty}^{x_2} f_{x_1,x_2}(t_1,t_2) dt_1 dt_2$$

 $\forall t_1, t_2 \in \mathbb{R}$ 

At points of continuity of  $f_{x_1,x_2}(x_1,x_2)$  we have

$$\frac{\partial^2 F_{x_1, x_2}(x_1, x_2)}{\partial x_1 x_2} = f(x_1, x_2)$$

### **PDF Properties**

A PDF if characterized by two properties:

(i) 
$$f_{x_1,x_2}(x_1, x_2) \ge 0$$
  
(ii)  $\int \int_X f(x_1, x_2) dx_1 dx_2 = 1$ 

 $Note: P\{(X_1, X_2) \in X\}$  is the volume under the surface

 $z = f_{x_1, x_2}(x_1, x_2)$  over the set X.

#### **Theorem**

For a  $Rvec(x_1, x_2)$  with  $F_{x_1, x_2}(x_1, x_2)$ 

$$(i)F(-\infty,\infty) = F(x_1,-\infty) = F(-\infty,\infty) = 0$$

$$(ii)F(\infty,\infty)=1$$

$$(iii) for x_1 \le a, x_2 \le b, F(a, b) - F(a, x_2) - F(x_1, b) + F(x_1, x_2) \ge 0$$

## **Conditional Expectation**

$$E[g(X_2|X_1 = x_1)] = \sum_{x_2} g(x_2)p(x_2|x_1)$$

if discrete and

$$E[g(X_2|X_1 = x_1)] = \int_{-\infty}^{\infty} g(x_2)f(x_2|x_1)dx_2$$

if continuous

### Measure of skewness

$$\alpha_3 = \frac{\mu_3}{\mu_2^{3/2}}$$

where skewness meaures the lack of symmetry in the pdf. The third moment is related to this.

#### **Kurtosis**

$$k = \alpha_4 = \frac{\mu_4}{\mu_2^2}$$

where kurtosis meaures the peakness of flatness of distribution. The 4th moment is related to this.

5.4 - 5.6

### **Independent Random Variables and Functions**

CDF:

$$F(x) = P(X \le x)$$

where  $b \in \mathbb{R}$ 

$$\implies (-\infty, b] = (-\infty, a] + (a, b]$$

$$\iff F(b) = F(a) + P\{(a, b]\}$$

$$\implies P(a < X \le b) = F(b) - F(a)$$

$$I_{n} = \{x : a - \frac{1}{n} < x \le a + \frac{1}{n}\},$$

$$P(X = a) = P(\cap I_{n})$$

$$\lim_{n \to \infty} P(I_{n}) = \lim_{n \to \infty} \{F(a + \frac{1}{n}) - F(a - \frac{1}{n})\}$$

$$= F(a^{+}) - F(a^{-})$$

### **Definition 5.8**

Let  $Y_1$  have distribution  $F_1(y_1)$  and  $Y_2$  have distribution  $F_2(y_2)$ , and  $Y_1$  and  $Y_2$  have joint distribution function  $F(y_1, y_2)$ . Then  $Y_1$  and  $Y_2$  are said to be *independent* if and only if

$$F(y_1, y_2) = F_1(y_1)F_2(y_2)$$

for every pair of real numbers  $(y_1, y_2)$ . If  $Y_1$  and  $Y_2$  are not independent, they are dependent.

Discrete RVec, X(orX)

$$P(\mathbb{X}\in B)=\sum p_X(x)\ for\ B\in \mathcal{B}$$
 
$$p_X(x)=p(x_1,x_2,\ldots,x_k)=(X_1=x_1,X_2=x_2,\ldots,X_k=x_k)$$
 n particular,  $k=4$ : Bivariate

In particular, k = 4; Bivariate

1. 
$$P\{(X_1, X_2) \in A\} = \sum_{(x_1, x_2) \in A} p(x_1, x_2)$$

2. 
$$\sum_{(x_1, x_2) \in \mathbb{R}^2} p(x_1, x_2) = p\{(X_1, X_2) \in \mathbb{R}^2\} = 1$$

3. 
$$E(g(X_1, X_2)) = \sum g(x_1, x_2) p(x_1, x_2)$$

## Marginals(Discrete)

$$p_{X_1}(x_1) = P(X_1 = x_1) = P(X_1 = x_2, -\infty < x_2 < \infty) = \sum_{x_2 \in \mathbb{R}} p(x_1, x_2)$$
$$p_{X_2}(x_2) = \sum_{x_2 \in \mathbb{R}} p(x_1, x_2)$$

## Conditional(Discrete)

$$P_{X_1|X_2=x_2} = P(X_1|X_2=x_2)$$

### **Continuous**

CDF:

$$F_X(x) = P(X_1 \le x_1, X_2 \le x_2, \dots, X_k \le x_k)$$

PDF:

$$f_X(x) = \frac{\partial^2 F(x_1, x_2, \dots, x_k)}{\partial x_1 \partial x_2 \dots \partial x_k} = fX_1, X_2, \dots, X_k(x_1, x_2, \dots, x_k)$$

and

$$P\{X \in B\} = \int_{X_1} \int_{X_2} \dots \int_{X_k} f(x_1, x_2, \dots, x_k) dx_1, dx_2, \dots, dx_k$$

Again, k = 2(bivariate)

(i) 
$$P\{(X_1, X_2) \in A\} = \int \int_A f(x_1, x_2) dx_1 dx_2$$

$$(ii) \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} f(x_1, x_2) dx_1 dx_2 = 1$$

(iii) 
$$E(X_1 X_2) = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} g(x_1, x_2) f(x_1, x_2) dx_1 dx_2$$

## Marginals(Continuous)

$$f_{X_1}(x_1) = \int_{-\infty}^{\infty} f(x_1, x_2) dx_2$$
$$f_{X_2}(x_2) = \int_{-\infty}^{\infty} f(x_1, x_2) dx_1$$

#### Theorem 5.4

If  $Y_1$  and  $Y_2$  are discrete random variables with joint probability function  $p(y_1, y_2)$  and marginal probability functions  $p_1(y_1)$  and  $p_2(y_2)$ , respectively, then  $Y_1$  and  $Y_2$  are independent if and only if

$$p(y_1, y_2) = p_1(y_1)p_2(y_2)$$

for all pairs of real numbers  $(y_1,y_2)$ . The corollary for this is that

$$E(Y_1Y_2) = E(Y_1)E(Y_2)$$

# **Multinomial Probability Distribution**

Properties:

- 1. The experiment consists of *n* identical trials.
- 2. The outcome of each trial falls into one of k classes of cells.

- 3. The probability that the outcome of a single trial falls into cell i, is  $p_i$ ,  $i=1,2,\ldots,k$  and remains the same from trial to trial. Notice that  $p_1+p_2+p_3+\ldots+p_k=1$ .
- 4. The trials are independent.
- 5. The random variables of interest are  $Y_1, Y_2, \dots, Y_k$ , where  $Y_i$  equals the number of trials for which the outcome falls into cell i. Notice that  $Y_1 + Y_2 + Y_3 + \dots + Y_k = n$ .

The joint probability distribution for  $Y_1, Y_2, \ldots, Y_k$  is given

$$p(y_1, y_2, \dots, y_k) = \frac{n!}{y_1! y_2! \dots y_k!} p_1^{y_1} p_2^{y_2} \dots p_k^{y_k}$$

where

$$\sum_{i=1}^{k} p_i = 1 \text{ and } \sum_{i=1}^{k} y_i = n$$

Other properties of the Multinomial Distribution are

1. 
$$E(Y_i) = np_i$$

$$2. V(Y_i) = n p_i q_i$$

3. 
$$Cov(Y_s, Y_t) = -np_s p_t$$
, if  $s \neq t$ 

# **Conditional Expectations**

For *Jointly* continuous,

$$E\bigg(g(Y_1|Y_2=y_2)\bigg) = \int_{-\infty}^{\infty} g(y_1)f(y_1|y_2)dy_1$$

For *Jointly* discrete,

$$E\bigg(g(Y_1|Y_2=y_2)\bigg) = \sum_{\forall y_1} g(y_1)p(y_1|y_2)$$

Also, the following are true

(i) 
$$E(Y_1) = E\left(E(Y_1|Y_2)\right)$$
  
(ii)  $V(Y_1) = E\left(V(Y_1|Y_2)\right) + V\left(E(Y_1|Y_2)\right)$ 

#### The Covariance of Two Random Variables

If two random variables are not independent, they are dependent or somewhat related.

#### **Definition 5.10**

If  $Y_1$  and  $Y_2$  are two random variables with means  $\mu_1$  and  $\mu_2$ , respectively, the *covariance* of  $Y_1$  and  $Y_2$  is

$$Cov(Y_1, Y_2) = E[(Y_1 - \mu_1)(Y_2 - \mu_2)]$$

#### **Theorem**

If two random variables  $X_1$  and  $X_2$  are independent, then

$$Cov(X_1, X_2) = 0$$

The converse may not be true.

#### **The Correlation Coefficient**

$$\rho = \frac{\text{Cov}(X_1, X_2)}{\sigma_1 \sigma_2}$$

 $\rho=\frac{\mathrm{Cov}(X_1,X_2)}{\sigma_1\sigma_2}$  The correlation coefficient measures the strength of linearity between  $X_1$  and  $X_2$ . Also,

(*i*) 
$$0 \le \rho \le 1$$

(*ii*) 
$$Corr(X_1, X_2) = Corr(X_2, X_1)$$

#### **Theorem**

If  $X_1$  and  $X_2$  are two random variables and a and b are two constants, then

$$Var(aX_1 + bX_2) = a^2 Var(X_1) + b^2 Var(X_2) \pm 2ab Cov(X_1, X_2)$$

Corollary(Prove!!!):

(i) 
$$Cov(aX_1, bX_2) = ab Cov(X_1, X_2)$$

(ii) 
$$Cov(X_1 + Y, Y) = Cov(X_1, Y) + Cov(Y, Y)$$

$$(iii)$$
 Cov $(X, aX + b)$  = Cov $(X, aX)$  = Var $(Y)$  =  $a$  Var $(X)$ 

In [ ]: