

例1. 设 $X_1, X_2, \dots, X_n$  独立同分布于 $N(\mu, \sigma^2)$ , 记

$$\bar{X} = \frac{1}{n} \sum_{k=1}^n X_k, \quad S^2 = \frac{1}{n-1} \sum_{k=1}^n (X_k - \bar{X})^2, \quad \bar{Y} = \frac{1}{n} \sum_{k=1}^n (X_k - \mu)^2,$$

求  $E(S^2)$ ;  $D(\bar{Y})$ .

例2. 设随机变量 $X_1, \dots, X_n$ 独立同分布于正态分布 $N(\mu, \sigma^2)$ , 若

$$\bar{X} = \frac{1}{n} \sum_{k=1}^n X_k, \quad Z_1 = X_1 - \bar{X}, \quad Z_2 = X_2 + \bar{X},$$

求 $Z_1, Z_2$ 的相关系数.

例3. 设 $X, Y$ 独立同分布于标准正态分布 $N(0, 1)$ , 求  $E(\max(X, Y))$ .

例4. 设 $X_1, X_2, \dots, X_n$  独立同分布于  $f(x) = \begin{cases} \frac{1}{2\theta}, & -\theta \leq x \leq \theta, \\ 0, & \text{else} \end{cases}, \quad \theta > 0$

求  $E(\max\{|X_1|, |X_2|, \dots, |X_n|\})$ ;  $D(\max\{|X_1|, |X_2|, \dots, |X_n|\})$ .

例5. 设 $X_1, X_2, \dots, X_n$  独立同分布于  $f(x) = \begin{cases} \frac{1}{\theta} e^{-\frac{x-\mu}{\theta}}, & x \geq \mu, \\ 0, & x < \mu \end{cases}, \quad \theta > 0$

求  $E(\min\{X_1, X_2, \dots, X_n\})$ .

例1. 设 $X_1, X_2, \dots, X_n$  独立同分布于 $N(\mu, \sigma^2)$ , 记

$$\bar{X} = \frac{1}{n} \sum_{k=1}^n X_k, \quad S^2 = \frac{1}{n-1} \sum_{k=1}^n (X_k - \bar{X})^2, \quad \bar{Y} = \frac{1}{n} \sum_{k=1}^n (X_k - \mu)^2,$$

求  $E(S^2)$ ;  $D(\bar{Y})$ .

$$\begin{aligned} \text{解: } E(S^2) &= \frac{1}{n-1} \sum_{k=1}^n E[(X_k - \bar{X})^2] \\ &= \frac{1}{n-1} \sum_{k=1}^n \{D(X_k - \bar{X}) + [E(X_k - \bar{X})]^2\} \end{aligned}$$

$X_1, X_2, \dots, X_n$  独立同分布于  $N(\mu, \sigma^2)$

$$D(X_k - \bar{X}) = D(X_k) + D(\bar{X}) - 2cov(X_k, \bar{X})$$

$$D(X_k) = \sigma^2 \quad D(\bar{X}) = \frac{\sigma^2}{n}$$

$$Cov(X_k, \bar{X}) = Cov(X_k, \frac{1}{n} \sum_{i=1}^n X_i) = \frac{1}{n} \sum_{i=1}^n Cov(X_k, X_i)$$

由于  $X_1, X_2, \dots, X_n$  相互独立, 故  $Cov(X_k, X_i) = 0$  ( $k \neq i$ )

$$Cov(X_k, \bar{X}) = \frac{1}{n} Cov(X_k, X_k) = \frac{1}{n} D(X_k) = \frac{\sigma^2}{n}$$

$$D(X_k) = \sigma^2 \quad D(\bar{X}) = \frac{\sigma^2}{n} \quad \text{Cov}(X_k, \bar{X}) = \frac{\sigma^2}{n}$$

于是得

$$D(X_k - \bar{X}) = D(X_k) + D(\bar{X}) - 2\text{cov}(X_k, \bar{X}) = \frac{(n-1)\sigma^2}{n}$$

$$E(S^2) = \frac{1}{n-1} \sum_{k=1}^n \{D(X_k - \bar{X}) + [E(X_k - \bar{X})]^2\} = \sigma^2$$

$$X_1, \dots, X_n \text{ i.i.d. } \sim N(\mu, \sigma^2), \bar{Y} = \frac{1}{n} \sum_{k=1}^n (X_k - \mu)^2$$

•  $X_1, \dots, X_n \text{ i.i.d. } \sim N(\mu, \sigma^2)$   
 记  $Y_k = \frac{X_k - \mu}{\sigma}, k = 1, \dots, n$

$\longrightarrow Y_1, \dots, Y_n \text{ i.i.d. } \sim N(0, 1)$

$$\longrightarrow D(Y_k^2) = E(Y_k^4) - [E(Y_k^2)]^2 = 3 - 1 = 2$$

$$D(\bar{Y}) = \frac{\sigma^4}{n^2} \sum_{k=1}^n D\left[\left(\frac{X_k - \mu}{\sigma}\right)^2\right] = \frac{\sigma^4}{n^2} \sum_{k=1}^n D(Y_k^2) = \frac{2\sigma^4}{n}$$

$$X_1, \dots, X_n \text{ 独立 } \longrightarrow \text{cov}(X_i, X_j) = 0 \text{ for } i \neq j$$

例2. 设随机变量  $X_1, \dots, X_n$  独立同分布于正态分布  $N(\mu, \sigma^2)$ , 若

$$\bar{X} = \frac{1}{n} \sum_{k=1}^n X_k, \quad Z_1 = X_1 - \bar{X}, \quad Z_2 = X_2 + \bar{X},$$

求  $Z_1, Z_2$  的相关系数.

提示:  $\bullet \quad X_1, \dots, X_n$  同分布  $\longrightarrow \text{cov}(X_1, \bar{X}) = \text{cov}(X_2, \bar{X}) = \frac{\sigma^2}{n}$

$$\text{cov}(Z_1, Z_2) = \text{cov}(X_1 - \bar{X}, X_2 + \bar{X}) = -\text{cov}(\bar{X}, \bar{X}) = -D(\bar{X}) = -\frac{\sigma^2}{n}$$

$$\text{cov}(Z_1, Z_2) = -\frac{\sigma^2}{n}$$

$$D(Z_1) = D(X_1 - \bar{X}) = D(X_1) + D(\bar{X}) - 2\text{cov}(X_1, \bar{X}) = \frac{(n-1)\sigma^2}{n}$$

$$D(Z_2) = D(X_2 + \bar{X}) = D(X_2) + D(\bar{X}) + 2\text{cov}(X_2, \bar{X}) = \frac{(n+3)\sigma^2}{n}$$

$$\rho_{Z_1 Z_2} = \frac{\text{cov}(Z_1, Z_2)}{\sqrt{D(Z_1)}\sqrt{D(Z_2)}} = -\frac{1}{\sqrt{(n-1)(n+3)}}$$



$$f_X(x) = \frac{1}{\sqrt{2\pi}} e^{-\frac{x^2}{2}} \quad f_Y(y) = \frac{1}{\sqrt{2\pi}} e^{-\frac{y^2}{2}}$$


例3. 设 $X, Y$ 独立同分布于标准正态分布 $N(0, 1)$ , 求  $E(\max(X, Y))$ .

$$\begin{aligned} \text{解法一: } E(\max(X, Y)) &= \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} \max(x, y) f_X(x) f_Y(y) dx dy \\ &= \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} \max(x, y) \frac{1}{2\pi} e^{-\frac{x^2+y^2}{2}} dx dy \end{aligned}$$

$$E(\max(X, Y)) = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} \max(x, y) \frac{1}{2\pi} e^{-\frac{x^2+y^2}{2}} dx dy$$

$$E(\max(X, Y)) = \iint_{\{x \geq y\}} x \frac{1}{2\pi} e^{-\frac{x^2+y^2}{2}} dx dy + \iint_{\{x < y\}} y \frac{1}{2\pi} e^{-\frac{x^2+y^2}{2}} dx dy$$

$I_1$  

$I_2$  

$$I_1 = \iint_{\{x \geq y\}} x \frac{1}{2\pi} e^{-\frac{x^2+y^2}{2}} dx dy$$

则  $I_1 = \int_{-\infty}^{\infty} \int_y^{\infty} x \frac{1}{2\pi} e^{-\frac{x^2+y^2}{2}} dx dy = \int_{-\infty}^{\infty} \frac{e^{-\frac{y^2}{2}}}{2\pi} \left( \int_y^{\infty} x e^{-\frac{x^2}{2}} dx \right) dy$

$$= \int_{-\infty}^{\infty} \frac{e^{-\frac{y^2}{2}}}{2\pi} \left( \int_y^{\infty} e^{-\frac{x^2}{2}} d\frac{x^2}{2} \right) dy = \int_{-\infty}^{\infty} \frac{e^{-\frac{y^2}{2}}}{2\pi} e^{-\frac{y^2}{2}} dy$$

$$= \frac{1}{2\pi} \int_{-\infty}^{\infty} e^{-y^2} dy = \frac{1}{2\sqrt{\pi}}$$

$$I_2 = \iint_{\{x < y\}} y \frac{1}{2\pi} e^{-\frac{x^2+y^2}{2}} dx dy$$

同理

$$I_2 = \int_{-\infty}^{\infty} \int_x^{\infty} y \frac{1}{2\pi} e^{-\frac{x^2+y^2}{2}} dy dx = \frac{1}{2\pi} \int_{-\infty}^{\infty} e^{-\frac{x^2}{2}} \left( \int_x^{\infty} y e^{-\frac{y^2}{2}} dy \right) dx$$

$$= \frac{1}{2\pi} \int_{-\infty}^{\infty} e^{-\frac{x^2}{2}} e^{-\frac{x^2}{2}} dx = \frac{1}{2\sqrt{\pi}}$$

故

$$I_1 = I_2 = \frac{1}{2\sqrt{\pi}} \quad \text{则} \quad E(\max(X, Y)) = I_1 + I_2 = \frac{1}{\sqrt{\pi}}$$

$$E(X \pm Y) = E(X) \pm E(Y) \quad D(X \pm Y) = D(X) + D(Y) \text{ (独立)}$$

例3. 设 $X, Y$ 独立同分布于标准正态分布 $N(0, 1)$ , 求  $E(\max(X, Y))$ .

解法二:  $\max(X, Y) = \frac{1}{2}(X + Y + |X - Y|)$

$$E(\max(X, Y)) = \frac{1}{2}E(X + Y) + \frac{1}{2}E(|X - Y|)$$

$$\overset{i.i.d.}{X, Y \sim N(0, 1)} \implies X + Y \sim N(0, 2) \quad X - Y \sim N(0, 2)$$

$$\max(X, Y) = \frac{1}{2}(X + Y + |X - Y|)$$

$$Z = X - Y \sim N(0, 2)$$

$$E(|Z|) = \int_{-\infty}^{\infty} |x| \frac{1}{2\sqrt{\pi}} e^{-\frac{x^2}{4}} dx = 2 \int_0^{\infty} x \frac{1}{2\sqrt{\pi}} e^{-\frac{x^2}{4}} dx = \frac{1}{\sqrt{\pi}}$$

$$E(\max(X, Y)) = \frac{1}{\sqrt{\pi}}$$

# 期望与方差的计算

例4 设  $X_1, X_2, \dots, X_n$  独立同分布于  $f(x) = \begin{cases} \frac{1}{2\theta}, & -\theta \leq x \leq \theta, \\ 0, & \text{else} \end{cases}, \quad \theta > 0$

求  $E(\max\{|X_1|, |X_2|, \dots, |X_n|\})$ ;  $D(\max\{|X_1|, |X_2|, \dots, |X_n|\})$ .

分析: 令  $M = \max\{|X_1|, |X_2|, \dots, |X_n|\}$

$$\begin{array}{ccccccc} \text{Step 4} & E(M) = \int_{-\infty}^{\infty} x \underbrace{f_M(x)}_{\text{Step 3}} dx & \longleftarrow & f_M(x) = \underbrace{F'_M(x)}_{\text{Step 2}} & \longleftarrow & F_M(x) = \underbrace{[F_{|X_i|}(x)]^n}_{\text{Step 1}} & \end{array}$$

# 期望与方差的计算

例4 设  $X_1, X_2, \dots, X_n$  独立同分布于  $f(x) = \begin{cases} \frac{1}{2\theta}, & -\theta \leq x \leq \theta, \\ 0, & \text{else} \end{cases}, \quad \theta > 0$

求  $E(\max\{|X_1|, |X_2|, \dots, |X_n|\})$ ;  $D(\max\{|X_1|, |X_2|, \dots, |X_n|\})$ .

解: Step 1. 求  $F_{|X_i|}(x)$   $X_i \sim f(x) = \begin{cases} \frac{1}{2\theta}, & -\theta \leq x \leq \theta, \\ 0, & \text{else} \end{cases}, \quad i = 1, \dots, n$

$$\longrightarrow F_{|X_i|}(x) = P(|X_i| \leq x) = \begin{cases} 0, & x < 0 \\ \frac{x}{\theta}, & 0 \leq x \leq \theta, \\ 1, & x \geq \theta \end{cases}, \quad i = 1, \dots, n$$



$$F_{|X_i|}(x) = \begin{cases} 0, & x < 0 \\ \frac{x}{\theta}, & 0 \leq x \leq \theta, \\ 1, & x \geq \theta \end{cases} \quad i = 1, \dots, n$$

$$\text{令 } M = \max\{|X_1|, |X_2|, \dots, |X_n|\}$$

Step 2. 求  $F_M(x)$

$$F_M(x) = [F_{|X_i|}(x)]^n = \frac{x^n}{\theta^n}, \quad 0 \leq x \leq \theta$$

Step 3. 求  $f_M(x)$

$$f_M(x) = F'_M(x) = \frac{nx^{n-1}}{\theta^n}, \quad 0 \leq x \leq \theta$$

$$f_M(x) = \frac{nx^{n-1}}{\theta^n}, \quad 0 \leq x \leq \theta$$

**Step 4. 求 $E(M)$ 和 $D(M)$**

$$E(M) = E(\max\{|X_1|, |X_2|, \dots, |X_n|\}) = \int_{-\infty}^{\infty} xf_M(x) dx = \frac{n}{n+1} \theta$$

$$E(M^2) = \int_{-\infty}^{\infty} x^2 f_M(x) dx = \frac{n}{n+2} \theta^2$$

$$D(M) = E(M^2) - [E(M)]^2 = \frac{n}{(n+2)(n+1)^2} \theta^2$$

# 期望与方差的计算

例5. 设  $X_1, X_2, \dots, X_n$  独立同分布于  $f(x) = \begin{cases} \frac{1}{\theta} e^{-\frac{x-\mu}{\theta}}, & x \geq \mu, \\ 0, & x < \mu \end{cases}, \quad \theta > 0$

求  $E(\min\{X_1, X_2, \dots, X_n\})$ .

分析: 令  $N = \min\{X_1, X_2, \dots, X_n\}$

$$\begin{array}{ccccccc} \text{Step 4} & E(N) = \int_{-\infty}^{\infty} x \underbrace{f_N(x)}_{\text{Step 3}} dx & \longleftarrow & f_N(x) = \underbrace{F'_N(x)}_{\text{Step 2}} & \longleftarrow & F_N(x) = 1 - [1 - \underbrace{F_{X_i}(x)}_{\text{Step 1}}]^n \end{array}$$

$$X_1, X_2, \dots, X_n \text{ 独立同分布于 } f(x) = \begin{cases} \frac{1}{\theta} e^{-\frac{x-\mu}{\theta}}, & x \geq \mu, \\ 0, & x < \mu \end{cases}, \quad \theta > 0$$

解: Step 1.  $F_{X_i}(x)$

$$X_i \sim f_{X_i}(x) = \begin{cases} \frac{1}{\theta} e^{-\frac{x-\mu}{\theta}}, & x \geq \mu, \\ 0, & x < \mu \end{cases}, \quad i = 1, \dots, n$$

$$\longrightarrow X_i \sim F_{X_i}(x) = \begin{cases} 0, & x < \mu \\ \int_{\mu}^x \frac{1}{\theta} e^{-\frac{t-\mu}{\theta}} dt, & x \geq \mu \end{cases}, \quad i = 1, \dots, n$$

$$= \begin{cases} 0, & x < \mu \\ 1 - e^{-\frac{x-\mu}{\theta}}, & x \geq \mu \end{cases}, \quad i = 1, \dots, n$$

$$X_i \sim F_{X_i}(x) = \begin{cases} 0, & x < \mu \\ 1 - e^{-\frac{x-\mu}{\theta}}, & x \geq \mu \end{cases}, \quad i = 1, \dots, n$$

解: Step 2. 求  $F_N(x)$

$$F_N(x) = 1 - [1 - F_{X_i}(x)]^n = 1 - e^{-\frac{n(x-\mu)}{\theta}}, \quad x \geq \mu$$

Step 3. 求  $f_N(x)$

$$f_N(x) = F'_N(x) = \frac{n}{\theta} e^{-\frac{n(x-\mu)}{\theta}}, \quad x \geq \mu$$

$$N = \min\{X_1, X_2, \dots, X_n\} \sim f_N(x) = \frac{n}{\theta} e^{-\frac{n(x-\mu)}{\theta}}, \quad x \geq \mu$$

Step 4. 求  $E(N)$

$$E(N) = E(\min\{X_1, X_2, \dots, X_n\}) = \int_{-\infty}^{\infty} x f_N(x) dx = \int_{\mu}^{\infty} x \frac{n}{\theta} e^{-\frac{n(x-\mu)}{\theta}} dx$$

$$\underline{\underline{t = \frac{n(x-\mu)}{\theta}}} \quad \int_0^{\infty} \left(t + \frac{n}{\theta}\mu\right) e^{-t} \frac{\theta}{n} dt = \int_0^{\infty} t e^{-t} \frac{\theta}{n} dt + \int_0^{\infty} \mu e^{-t} dt = \mu + \frac{\theta}{n}$$