Random Convex Programs with L_1 -Regularization Sparsity and Generalization

Campi & Carè, SIAM J. Contr. Optim, 50(5), 3532-3357, 2013.

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Random Convex Program

Random Convex Program

Consider a standard random convex program (min-max)

RCP
$$\min_{x \in \mathcal{X} \subset \mathbb{R}^d} \max_{i=1,\dots,N} L(x, \delta^{(i)}),$$

where $\delta^{(i)}$, $i=1,\cdots,N$ are N scenarios sampled from Δ in an i.i.d. fashion according to probability \mathbb{P} .

RCP is equivalent to a epigraphic form

$$\min_{L \in \mathbb{R}, x \in \mathcal{X}} \ L \quad \text{ subject to } \ L(x, \delta^{(i)}) \leq L, \ i = 1, \cdots, N.$$

The optimization is to take worst-case minimization w.r.t. scenarios $\delta^{(i)}$.

Generalization Property (ϵ -level Performance Robustness)

There is a set Δ_{ϵ} with $\mathbb{P}\{\Delta_{\epsilon}\} \geq 1 - \epsilon$ such that $\max_{\delta \in \Delta_{\epsilon}} L(x_{N}^{*}, \delta) \leq L_{N}^{*}$. where $L_{N}^{*} = \max_{i=1,\cdots,N} L(x, \delta^{(i)})$, and x_{N}^{*} is the optimal solution of RCP.

RCP with L_1 -Regularization

Random Convex Program with L_1 -Regularization

$$L_1\text{-RCP} \qquad \min_{x \in \mathcal{X} \in \mathbb{R}^d} \max_{i=1,\cdots,N} \ L(x,\delta^{(i)}) \quad \textit{subject to} \quad \|Ax - b\|_1 \leq r,$$

where $A \in \mathbb{R}^{p \times d}$, $b \in \mathbb{R}^p$, $\|\cdot\|_1$ is the L_1 norm (e.g., $\|z\|_1 = \sum_{j=1}^p |z_j|$), and $r \in \mathbb{R}$ is the constraining parameter to tune the level of sparsity.

 \clubsuit L_1 -RCP is equivalent to a epigraphic form

$$\min_{L \in \mathbb{R}, x \in \bar{\mathcal{X}}} \ L \quad \text{ subject to } \quad L(x, \delta^{(i)}) \leq L, \ i = 1, \cdots, N,$$

where
$$\bar{\mathcal{X}} = \{x \in \mathcal{X} : \|Ax - b\|_1 \le r\}.$$

• Reduce the effective dimension of decision variable (i.e., dim(x) = d).

Assumption 1 (Convexity)

Function $L(x, \delta)$ is convex in x, while it has an arbitrary dependence on δ , and the optimization domain \mathcal{X} is a convex and closed set.

Examples (lasso and basalt column constraints)

Example (lasso constraint)

Letting A = I and b = 0, then the generalized constraint in L_1 -RCP reduces to the following lasso constraint

$$||x||_{1} \le r$$

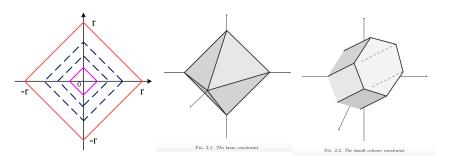
Example (basalt column constraint)

Letting A be a total variation matrix and b=0, then the generalized constraint in L_1 -RCP reduces to the following basalt column constraint

$$\begin{bmatrix} 1 & -1 & 0 & 0 & \cdots \\ 0 & 1 & -1 & 0 & \cdots \\ \vdots & \vdots & \vdots & \vdots \\ 0 & 0 & 0 & 1 & -1 \\ -1 & 0 & \cdots & 0 & 1 \end{bmatrix} \begin{bmatrix} x_1 \\ x_2 \\ \vdots \\ x_{d-1} \\ x_d \end{bmatrix} \implies \begin{bmatrix} x_2 - x_1 \\ x_3 - x_2 \\ \vdots \\ x_{d-1} \\ x_{d-1} \end{bmatrix}_{1} \le r$$

Moderate the number of jumps or switches for piecewise functions.

Pictorial interpretation



- The contour of lasso is a diamond in \mathbb{R}^2 .
- As r increases, search domain enlarges, optimal value $(\min_{x \in \mathbb{R}^d} L_N^*)$ improves, optimal solution x_N^* loses generalization property.
- Recall some constrained optimization problems
 - $\min_{x} \|b Ax\|_2$ s.t. $\|x\|_0 \le r$ (Best subset selection) $\Leftrightarrow \min_{x} \|x\|_0$ s.t. $\|b - Ax\|_2 \le t$
 - $\min_{x} \|b Ax\|_2$ s.t. $\|x\|_1 \le r$ (Lasso)

q-dimensional subspace

- ullet q is a user-chosen "complexity barrier" and satisfies q < d.
- Optimal q-dimensional subspace: \mathbb{Z}^{opt} : Set some rows of Ax - b as zero (i.e., $a_h^{\top} - b_h = 0$), that is,

$$\begin{bmatrix} a_{11}x_1 + a_{12}x_2 + \dots + a_{1d}x_d - b_1 \\ a_{21}x_1 + a_{22}x_2 + \dots + a_{2d}x_d - b_2 \\ \vdots \\ a_{p1}x_1 + a_{p2}x_2 + \dots + a_{pd}x_d - b_p \end{bmatrix} = \begin{bmatrix} a_1^\top - b_1 \\ a_2^\top - b_2 \\ \vdots \\ a_p^\top - b_p \end{bmatrix}$$

so that

$$\min_{x \in \mathcal{Z}^{opt} \cap \mathcal{X}} \max_{i=1,\cdots,N} L(x,\delta^{(i)}) \leq \min_{x \in \mathcal{Z} \cap \mathcal{X}} \max_{i=1,\cdots,N} L(x,\delta^{(i)})$$

- Two requirements for a suitable selection of q:
 - Guarantee adequate generalization properties;
 - Allow for a satisfactory optimal cost.

Algorithm (L_1 -RCA)

Random convex algorithm with L_1 -regularization (L_1 -RCA).

- (a) Let s be the dimension of the affine subspace of \mathbb{R}^d identified by relation Ax b = 0. Select an integer q with s < q < d. Initialize r = 0
- (b) Let $x_N^*(r)$ be the optimal solution path of L_1 -RCP as r is increased. For all values of $r \geq 0$, evaluate which components of $Ax_N^*(r) b$ are zero, and let H(r) be the index set of the zero components of $Ax_N^*(r) b$; thus, if, for example, the first two components of $Ax_N^*(r) b$ are zero, we have $H(r) = \{1, 2\}$. Further, define $\mathcal{Z}(r) := \{x : a_h^T x b_h = 0, h \in H(r)\}$, where $a_h^T x b_h$ is the hth component of Ax b; that is, $\mathcal{Z}(r)$ is the affine subspace of \mathbb{R}^d preserving the null components of $Ax_N^*(r) b$.

Set \bar{r} to be the largest r such that $\dim(\mathcal{Z}(r)) = q$.

(c) Solve

$$\min_{x \in \mathcal{Z}(\bar{r}) \cap \mathcal{X}} \max_{i=1,...,N} L(x, \delta^{(i)}),$$

and let x_N^* and L_N^* be the optimal solution and the optimal value of this problem.

Assumptions

Assumption 2 (Existence and Uniqueness)

W.p. 1 w.r.t. the multisample δ , any RCP considered here admits a unique solution.

Assumption 3

W.p. 1 w.r.t. the multisample δ , when function $m(r) = \dim(\mathcal{Z}(r))$ increases, it does so one unit at a time, that is, it does not have jumps up of 2 or more units, and $m(\infty)$: $\lim_{r\to\infty} m(r) = d$.

Termination of L_1 -RCA

For r=0, $\|Ax_N^*(0)-b\|_1=0$ so that $Ax_N^*(0)-b=0$ which entails that m(0)=s. Thus m(r) goes from s to d, when it increases, it does so one unit at a time. Hence, an r exists where m(r)=q. Moreover, the sup \bar{r} takes $m(\bar{r})=q$. After \bar{r} is determined in (b), then (c) generates x_N^* and L_N^* and terminates the ALGO.

Theory

THEOREM 3.2

For L_1 -RCA algorithm, if it takes sample complexity

$$N \geq rac{2}{\epsilon} igg[\ln rac{1}{eta} + q + (p-d+q) \ln igg(rac{p \cdot e}{p-d+q} igg) igg].$$

Under Assumptions 1, 2, 3, for all multisample δ with the exception of a set whose probability \mathbb{P}^N is at most β :

There is a set Δ_ϵ with $\mathbb{P}\{\Delta_\epsilon\} \geq 1 - \epsilon$ such that

$$\max_{\delta \in \Delta_{\epsilon}} L(x_N^*(\boldsymbol{\delta}), \delta) \le L_N^*(\boldsymbol{\delta})$$

• This theorem is equivalent to a more general result, that is,

$$\binom{p}{d-q}\sum_{i=0}^{q} \binom{N}{i} \epsilon^{i} (1-\epsilon)^{N-i} \leq \beta.$$

• Special case: take p = d, then the term w.r.t N is as follows:

$$N \geq rac{2}{\epsilon} igg[\ln rac{1}{eta} + q igg(1 + \ln rac{d \cdot e}{q} igg) igg].$$

Compared with previous results

Table 3.1

Values for N obtained using formula (3.15) (1st line in italic) and formula (3.8) (2nd through 9th lines); $\beta = 10^{-10}$, p = d = 2000.

	$\epsilon=1\%$	$\epsilon=2\%$	$\epsilon = 3\%$	$\epsilon = 4\%$	$\epsilon = 5\%$	$\epsilon=6\%$	$\epsilon=7\%$	$\epsilon=8\%$	$\epsilon = 9\%$	$\epsilon = 10\%$
	229735	114793	76478	57321	45826	38163	32689	28584	25390	22836
q = 1	3403	1693	1123	838	668	554	472	411	364	325
q = 2	4427	2203	1462	1091	869	720	614	535	473	424
q = 3	5403	2689	1784	1332	1060	879	750	653	578	517
q = 4	6346	3158	2096	1564	1245	1033	881	767	679	608
q = 5	7264	3615	2399	1791	1426	1182	1009	878	777	696
q = 10	11594	5771	3829	2859	2276	1888	1610	1402	1240	1111
q = 15	15644	7786	5167	3858	3072	2548	2173	1893	1674	1500
q = 20	19506	9709	6443	4810	3831	3177	2711	2361	2088	1870

• Previous result: $\sum\limits_{i=0}^d \binom{N}{i} \epsilon^i (1-\epsilon)^{N-i} \leq \beta$.

• For example: $\epsilon = 5\%$, $N_{old} = 45826$, q = 10, $N_{new} = 2276$.

[Campi & Garatti, SIAM J. Optim., 2008]

Role of Probability \mathbb{P} , confidence β and risk ϵ

Role of Probability \mathbb{P} :

- \bullet Training samples $\{\delta^{(i)}\}_{i=1}^N$ are generated according to probability $\mathbb{P}.$
- Generalization property refers to sampling a new scenario δ again according to probability \mathbb{P} ;
- Verify whether $L(x_N^*(\delta), \delta) \leq L_N^*(\delta)$.
- ♠ QUES:
- What happens if the testing probability and the verification probability do not coincide?
 - Ambiguity set (ACCP: ambiguous chance-constrained program)
 - Prohorov metric [Erdoğan & Iyengar, Math. Program., 2006]
 - Wasserstein metric (DRO: distributionally robust optimization)

[Esfahani & Kuhn, Math. Program., 2017]

Role of confidence β and risk ϵ :

- ullet For practical appeal of method, confidence eta should take small.
- As scenarios N tends to infinity, the risk ϵ tends to zero.

Example: Minimax Regression

Minimax Regression

A signal s(t) is obtained as the composition of 200 sinusoids,

$$s(t) = \sum_{j=1}^{200} \alpha_j \sin(jt), \ \hat{s}(t) = \sum_{j=1}^{200} x_j \sin(jt) \ \rightarrow \ \hat{s}(t) = \sum_{k=1}^{7} x_{j_k} \sin(j_k t)$$

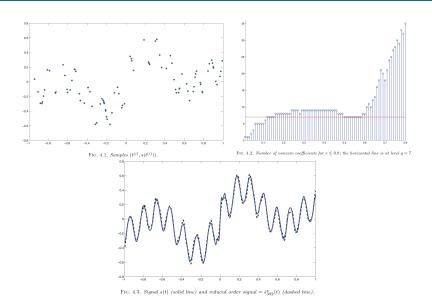
- Take $\alpha_1 = \alpha_5 = \alpha_8 = \alpha_{45} = 0.2$, $\sum_{j \neq 1,5,8,45} \alpha_j = 1$ and $\sum_{j=1}^{200} \alpha_j = 1$.
- Gather N=332 samples of $(t^{(i)},s(t^{(i)}))$, where $t^{(i)}\sim U(-\pi,\pi)$.
- Select q = 7 nonzero coefficients x_{j_k} with frequencies $j_1 = 1$, $j_2 = 5$, $j_3 = 8$, $j_4 = 41$, $j_5 = 45$, $j_6 = 109$, $j_7 = 127$.

Consider L_1 -RCP ALGO as follows

$$\min_{x \in \mathbb{R}^{200}} \max_{i=1,\cdots,332} |s(t^{(i)}) - \hat{s}(t^{(i)})|, \quad \text{s.t. } \|x\|_1 \leq r, \\ \text{(reduced order)} \ \Rightarrow \ \min_{x_{j_1},\cdots,x_{j_7}} \max_{i=1,\cdots,332} |s(t^{(i)}) - \sum_{k=1}^{7} x_{j_k} \sin(x_{j_k} t^{(i)})|,$$

The obtained optimal solutions are: $x_{j_1}^* = 0.1909$, $x_{j_2}^* = 0.1964$, $x_{j_3}^* = 0.2033$, $x_{j_4}^* = 0.0187$, $x_{j_5}^* = 0.2059$, $x_{j_6}^* = 0.0271$, $x_{j_7}^* = 0.0184$, and cost $L_{332}^* = 0.0649$.

Numerical Experiments



A posterior evaluation

PROPOSITION 4.1

Let x_N^* be the solution obtained with L_1 -RCA. Take

$$M \geq rac{1}{\epsilon'} \ln rac{1}{eta'}$$

i.i.d. samples $\delta^{(N+1)}, \cdots, \delta^{(N+M)}$ distributed according to $\mathbb P$ and independent of $\delta^{(1)}, \cdots, \delta^{(N)}$ and let

$$L^* = \max_{i=N+1,\cdots,N+M} L(x_N^*, \delta^{(i)}).$$

Then, with confidence $1 - \beta'$ w.r.t. the multisample $\delta^{(N+1)}, \dots, \delta^{(N+M)}$, relation

$$L(x_N^*,\delta) \leq L^*$$

holds with probability at least $1-\epsilon'$ w.r.t. random choices of δ .

Performance robustness cost ϵ_ℓ

Assessment of Robustness-loss curve

Let $\ell=q+1,\cdots,q+h$, $\alpha=p-d+q$, and

$$\epsilon_{\ell} = \frac{\ell}{\mathit{N}} + \frac{\mathit{g} - 1 + \sqrt{\mathit{g}^2 + 2(\ell - 1)\mathit{g}}}{\mathit{N}}, \quad \mathit{g} = \ln\left[\frac{1}{\beta} \cdot \left(\frac{\mathit{p} \cdot \mathit{e}}{\alpha}\right)^{\alpha}\right],$$

where h is an arbitrary integer chosen by the user such that $q + h \leq N$. To easy notation, denote x^* as $x_N^*(\delta)$. Define

$$L_{\epsilon_\ell}^* = \max\{L \text{ such that } L(x^*, \delta^{(i)}) \ge L \text{ for } \ell \text{ scenarios } \delta^{(i)}\}.$$

Thus, $L^*_{\epsilon_\ell}$ are the values $L(x^*, \delta^{(i)})$ listed in decreasing order of magnitude.

- The first term of ϵ_ℓ is the *empirical probability* of the scenarios that greater than or equal to $L_{\epsilon_\ell}^*$.
- The second term $\frac{g+\sqrt{g^2+2(\ell-1)g}}{N}$ of ϵ_ℓ is the adjustment term accounting for the mismatch between empirical and real probability.

Results

THEOREM 5.1

The statement $L(x^*, \delta) \leq L_{\epsilon_{\ell}}^*$ holds with probability at least $1 - \epsilon_{\ell}$ is true simultaneously for all $\ell = q + 1, \dots, q + h$ with confidence $1 - h\beta$.

$$\Leftrightarrow \quad \mathbb{P}^{N}\{\boldsymbol{\delta}: \mathbb{P}\{L(\boldsymbol{x}^{*},\boldsymbol{\delta}) > L_{\epsilon_{\ell}}^{*}\} > \epsilon_{\ell}\} \leq \left(\begin{smallmatrix} \rho \\ d-q \end{smallmatrix}\right) \Sigma_{i=0}^{\ell-1} \left(\begin{smallmatrix} N \\ i \end{smallmatrix}\right) \epsilon_{\ell}^{i} (1-\epsilon_{\ell})^{N-i}.$$

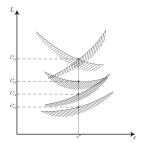


Fig. 5.1. Visualization of $L_{\epsilon_{\ell}}^{*}$ for q=1. Each constraint represents the region where $L(x, \delta^{(i)})$ for some $\delta^{(i)}$.

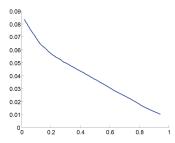


Fig. 5.3. Robustness-loss curve: $L_{\epsilon_{\ell}}^{*}$ (vertical axis) vs. ϵ_{ℓ} (horizontal axis). ℓ is in the range $8, \ldots, 6007$.

Conclusions

Take home messages:

- L_1 -regularization shrinks the number of optimization variables.
- Induce a L_1 -RCP algorithm.
- Enhance the generalization properties of the RCP.
- Perform a novel finite-sample guarantee:

$$\mathbb{P}^{N}\{\boldsymbol{\delta}: \mathbb{P}\{L(\boldsymbol{x}^{*},\boldsymbol{\delta}) > L^{*}\} > \epsilon\} \leq \binom{p}{d-q} \sum_{i=0}^{q} \binom{N}{i} \epsilon^{i} (1-\epsilon)^{N-i}.$$

ullet Does not require any knowledge of probability measure ${\mathbb P}$ (unknown).

Improvement

- Even use L_1 -RCA, N scales as $\frac{1}{\epsilon} \cdot d$.
- Fast algorithm gives the form of sample complexity N as $\frac{1}{\epsilon} + d$.

[Carè, Garatti and Campi, Operations Research, 2014]