CSC 665: Calibration Homework

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Please complete the following set of exercises **on your own**. The homework is due **on Sep 3, in class**. You may find the following version of Taylor's Theorem in multivariate calculus helpful:

Theorem 1. Suppose f is twice differentiable in \mathbb{R}^d . Then given two points a, b in \mathbb{R}^d , there exists some t in [0,1], such that

$$f(b) = f(a) + \left\langle \nabla f(a), b - a \right\rangle + \frac{1}{2} (b - a)^{\mathsf{T}} \nabla^2 f(\xi) (b - a),$$

where $\xi = ta + (1-t)b$. Here $\nabla f(x) \triangleq (\frac{\partial f}{\partial x_1}, \dots, \frac{\partial f}{\partial x_d})$ is the gradient of f at x, and

$$\nabla^2 f(x) \triangleq \begin{bmatrix} \frac{\partial^2 f}{\partial x_1^2} & \cdots & \frac{\partial^2 f}{\partial x_1 \partial x_d} \\ & \cdots & \\ \frac{\partial^2 f}{\partial x_d \partial x_1} & \cdots & \frac{\partial^2 f}{\partial x_d^2} \end{bmatrix}$$

is the Hessian of f at x.

Problem 1

Denote by B(n, p) the binomial distribution with n being the number of trials, and p being the success probability of each trial, and denote by $N(\mu, \sigma^2)$ the normal distribution with mean μ and variance σ^2 .

- 1. Suppose Y is a random variable such that $P(Y=+1)=P(Y=-1)=\frac{1}{2}$. In addition, given Y, X has the following conditional probability distribution: given Y=-1, $X \sim \mathrm{B}(3,\frac{2}{3})$; given Y=+1, $X \sim \mathrm{B}(2,\frac{1}{3})$. Calculate:
 - (a) the joint probability table of (X,Y);
 - (b) P(Y = 1|X = 3);
 - (c) P(Y = -1|X = 1).
- 2. Suppose Y is a random variable such that $P(Y=+1)=P(Y=-1)=\frac{1}{2}$. In addition, suppose given Y, X has the following conditional probability distribution: given Y=-1, $X \sim N(\mu_-, \sigma^2)$; given Y=+1, $X \sim N(\mu_+, \sigma^2)$. Define

$$P(Y = +1|x) \triangleq \frac{P(Y = +1)p_{+1}(x)}{P(Y = +1)p_{+1}(x) + P(Y = -1)p_{-1}(x)}.$$

where p_{+1} and p_{-1} are the conditional probability density functions of X given Y = +1 and Y = -1 respectively. Show that

$$P(Y = +1|x) = \frac{1}{1 + \exp\left(-\frac{\mu_{+} - \mu_{-}}{\sigma^{2}} \cdot (x - \frac{\mu_{+} + \mu_{-}}{2})\right)}.$$

(Remark: P(Y = +1|x) has the intuitive interpretation that it is the conditional probability of Y = +1 given X = x. It can be shown rigorously that $P(Y = +1|x) = \lim_{\epsilon \to 0} P(Y = +1|X \in [x - \epsilon, x + \epsilon])$.)

Solution

1. (a) The joint probability table is

	X = 0	X = 1	X = 2	X = 3
Y = -1	$\frac{1}{54}$	$\frac{1}{9}$	$\frac{2}{9}$	$\frac{4}{27}$
Y = +1	$\frac{2}{9}$	$\frac{2}{9}$	$\frac{1}{18}$	0

(b)

$$P(Y=1|X=3) = \frac{P(Y=1,X=3)}{P(X=3)} = \frac{P(Y=1,X=3)}{P(Y=1,X=3) + P(Y=-1,X=3)} = \frac{0}{0 + \frac{4}{27}} = 0.$$

(c)

$$P(Y = -1|X = 1) = \frac{P(Y = -1, X = 1)}{P(X = 1)} = \frac{P(Y = -1, X = 1)}{P(Y = 1, X = 1) + P(Y = -1, X = 1)} = \frac{\frac{1}{9}}{\frac{1}{9} + \frac{2}{9}} = \frac{1}{3}.$$

2. Recall that by the definition of normal distribution, $p_{+1}(x) = \frac{1}{\sqrt{2\pi\sigma^2}} \exp\left\{-\frac{(x-\mu_+)^2}{2\sigma^2}\right\}$, and $p_{-1}(x) = \frac{1}{\sqrt{2\pi\sigma^2}} \exp\left\{-\frac{(x-\mu_-)^2}{2\sigma^2}\right\}$.

Therefore, $p_{-1}(x)/p_{+1}(x) = \exp\left\{-\frac{(x-\mu_{-})^2 - (x-\mu_{+})^2}{2\sigma^2}\right\} = \exp\left\{-\frac{(\mu_{+} - \mu_{-})(x - \frac{\mu_{+} + \mu_{-}}{2})}{\sigma^2}\right\}.$

In addition, $P(Y = +1) = P(Y = -1) = \frac{1}{2}$. Therefore,

$$P(Y = +1|x) = \frac{P(Y = +1)p_{+1}(x)}{P(Y = +1)p_{+1}(x) + P(Y = -1)p_{-1}(x)}$$

$$= \frac{p_{+1}(x)}{p_{+1}(x) + p_{-1}(x)}$$

$$= \frac{1}{1 + \frac{p_{-1}(x)}{p_{+1}(x)}}$$

$$= \frac{1}{1 + \exp\left\{-\frac{(\mu_{+} - \mu_{-})(x - \frac{\mu_{+} + \mu_{-}}{2})}{\sigma^{2}}\right\}}.$$

Problem 2

1. Suppose D = U([0,1]), i.e. the uniform distribution over the [0,1] interval. Consider a set of samples $S = (X_1, \ldots, X_n)$ drawn identically and independently from distribution D.

Write a program that plots the empirical cumulative distribution function (CDF) of the sample S, that is,

$$F_n(t) = \frac{1}{n} \sum_{i=1}^n \mathbf{1}(x_i \le t), t \in \mathbb{R},$$

where

$$\mathbf{1}(A) = \begin{cases} 1 & A \text{ is true} \\ 0 & A \text{ is false} \end{cases}$$

is the indicator function. You may use any programming languages you like (e.g. Python and Jupyter notebook).

Draw two sets of samples S_1 and S_2 of size n = 5. Plot F_n^1 , the CDF of S_1 , and plot F_n^2 , the CDF of S_2 . Are they different? Why?

- 2. Repeat the same experiment in 1 for n = 100 and n = 1000. Do the F_n^1 and F_n^2 functions become closer as n increases?
- 3. In the above experiment, as n goes to infinity, what function does F_n converge to? Can you derive a formula for that function (denoted as F)?
- 4. Suppose D is the the standard normal distribution N(0,1), what function does F_n converge to?

Solution

1. See Figure 1 below. (The exact results depend on the random points drawn - they can vary case by case.) They are different. As S_1 and S_2 are two different samples, their induced CDFs would also be different.

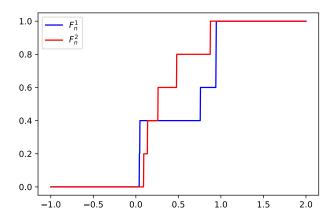


Figure 1: A random draw of F_n^1 and F_n^2 with sample size n=5.

- 2. See Figures 2 and 3 below. Yes, F_n^1 and F_n^2 do become closer as n increases. (Of course, in some extremely unlikely draws of random samples, they can become farther away as n increases.)
- 3. By the (weak) law of large numbers, for every t in \mathbb{R} , $F_n(t) = \frac{1}{n} \sum_{i=1}^n \mathbf{1}(x_i \leq t) \to F(t) = \mathbb{P}(X \leq t)$ (in probability).

$$F(t) = \mathbb{P}(X \le t) = \begin{cases} 0 & t \le 0 \\ t & 0 < t < 1 \\ 1 & t \ge 1 \end{cases}$$

4. When X is drawn from the standard normal distribution,

$$F(t) = \mathbb{P}(X \le t) = \int_{-\infty}^{t} \frac{1}{\sqrt{2\pi}} e^{-\frac{x^2}{2}} dx.$$

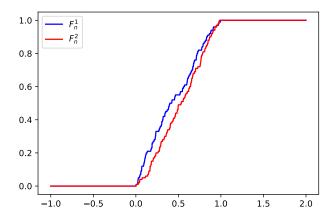


Figure 2: A random draw of ${\cal F}_n^1$ and ${\cal F}_n^2$ with sample size n=100.

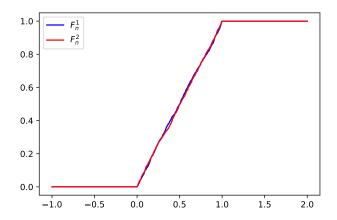


Figure 3: A random draw of F_n^1 and F_n^2 with sample size n=1000.

Problem 3

1. Define function

$$h(x) \triangleq x \ln x + (1-x) \ln(1-x), x \in (0,1).$$

Show that for any p and q in (0,1),

$$h(p) - h(q) - h'(q)(p - q) = p \ln \frac{p}{q} + (1 - p) \ln \frac{1 - p}{1 - q}.$$

(Remark: the expression on the right hand side is often called the binary relative entropy, denoted as kl(p,q); the h function is often called the negative binary entropy.)

2. Suppose 0 . Use Taylor's Theorem to show that

$$\mathrm{kl}(p,q) \geq 2(p-q)^2.$$

Furthermore, show that

$$\mathrm{kl}(p,q) \geq \frac{(p-q)^2}{2q}.$$

3. Define the *m*-dimensional probability simplex Δ^{m-1} as $\{p \in \mathbb{R}^m : \text{ for all } i, p_i \geq 0, \sum_{i=1}^m p_i = 1\}$. For two vectors p, q in Δ^{m-1} , define the negative entropy of p as:

$$H(p) \triangleq \sum_{i=1}^{m} p_i \ln p_i,$$

and the relative entropy between p and q as:

$$\mathrm{KL}(p,q) \triangleq \sum_{i=1}^{m} p_i \ln \frac{p_i}{q_i}.$$

Verify that

$$H(p) - H(q) - \langle \nabla H(q), p - q \rangle = KL(p, q).$$

4. Using Taylor's Theorem, show that for any p, q in Δ^{m-1} , $\mathrm{KL}(p,q) \geq 0$. Furthermore, show that $\mathrm{KL}(p,q) \geq \frac{1}{2} (\sum_{i=1}^m |p_i - q_i|)^2$.

Hint: at some point, you may want to use the following variant of Cauchy-Schwarz inequality:

$$(\sum_{i=1}^{m} y_i) \cdot (\sum_{i=1}^{m} \frac{x_i^2}{y_i}) \ge (\sum_{i=1}^{m} |x_i|)^2.$$

Solution

1. Observe that $h'(x) = (1 + \ln x) - (1 + \ln(1 - x)) = \ln x - \ln(1 - x)$. Therefore, for any p and q in (0, 1),

$$\begin{array}{lcl} h(p)-h(q)-h'(q)(p-q) & = & h(p)-(h(q)+h'(q)(p-q)) \\ & = & p\ln p + (1-p)\ln(1-p) - (q\ln q + (1-q)\ln(1-q) + (p-q)(\ln q - \ln(1-q)) \\ & = & p\ln p + (1-p)\ln(1-p) - (p\ln q + (1-p)\ln(1-q)) \\ & = & p\ln\frac{p}{q} + (1-p)\ln\frac{1-p}{1-q}. \end{array}$$

2. Observe that h(x) is twice differentiable on (0,1). Then, by Talyor's Theorem, for any p, q,

$$kl(p,q) = h(p) - h(q) - h'(q)(p-q) = \frac{1}{2}h''(\xi)(p-q)^2,$$

for some $\xi = (1-t)p + tq$, where $t \in [0,1]$. As both p and q are in $(0,1), \xi \in (p,q) \subset (0,1)$.

Now, observe that for all x in (0,1), $h''(x) = \frac{1}{x} + \frac{1}{1-x} = \frac{1}{x(1-x)} \ge \frac{1}{((1-x+x)/2)^2} = 4$, This implies that

$$kl(p,q) = \frac{1}{2}h''(\xi)(p-q)^2 \ge \frac{1}{2} \cdot 4(p-q)^2 = 2(p-q)^2.$$

The second inequality follows from the fact that $h''(\xi) = \frac{1}{\xi(1-\xi)} \ge \frac{1}{\xi} \ge \frac{1}{q}$. This implies that

$$kl(p,q) = \frac{1}{2}h''(\xi)(p-q)^2 \ge \frac{(p-q)^2}{2q}.$$

3. Observe that $\nabla H(q) = (\frac{\partial H}{\partial x_1}, \dots, \frac{\partial H}{\partial x_m}) = (1 + \ln q_1, \dots, 1 + \ln q_m).$

$$H(p) - H(q) - \langle \nabla H(q), p - q \rangle = \sum_{i=1}^{m} p_i \ln p_i - \sum_{i=1}^{m} q_i \ln q_i - \sum_{i=1}^{m} (p_i - q_i)(1 + \ln q_i)$$

$$= \sum_{i=1}^{m} (p_i \ln p_i - q_i \ln q_i - (p_i - q_i) \ln q_i) - \sum_{i=1}^{m} (p_i - q_i)$$

$$= \sum_{i=1}^{m} p_i \ln \frac{p_i}{q_i} - \sum_{i=1}^{m} (p_i - q_i),$$

As both p_i and q_i are in Δ^{m-1} , $\sum_{i=1}^m p_i = \sum_{i=1}^m q_i = 1$. Therefore,

$$H(p) - H(q) - \langle \nabla H(q), p - q \rangle = \sum_{i=1}^{m} p_i \ln \frac{p_i}{q_i} = KL(p, q).$$

4. Observe that H(p) is twice differentiable on $(0,1)^m$. Then, by Talyor's Theorem, for any p, q in $(0,1)^m$,

$$H(p) - H(q) - \langle \nabla H(q), p - q \rangle = \frac{1}{2} (p - q)^{\top} \nabla^2 H(\xi) (p - q).$$

for some $\xi = (1 - t)p + tq$ and $t \in [0, 1]$.

Here, the entries of Hessian of H can be written as the following:

$$(\nabla^2 \mathbf{H}(x))_{i,j} = \frac{\partial^2 f}{\partial x_i \partial x_j} = \begin{cases} \frac{1}{x_i}, & i = j \\ 0, & i \neq j \end{cases}.$$

Now, for p and q in Δ^{m-1} , we have that

$$KL(p,q) = H(p) - H(q) - \langle \nabla H(q), p - q \rangle$$

$$= \frac{1}{2} (p-q)^{\top} \nabla^{2} H(\xi) (p-q)$$

$$= \frac{1}{2} \sum_{i=1}^{m} \frac{(p_{i} - q_{i})^{2}}{\xi_{i}}, \qquad (1)$$

for some $\xi = (1-t)p + tq$ and $t \in [0,1]$. Specifically, all ξ_i 's are nonnegative, and $\sum_{i=1}^m \xi_i = (1-t)\sum_{i=1}^m p_i + t\sum_{i=1}^m q_i = 1$. As all the terms in the sum are nonnegative, $\mathrm{KL}(p,q) \geq 0$.

Now, applying the (variant of) Cauchy-Schwarz Inequality, we get

$$\sum_{i=1}^{m} \frac{(p_i - q_i)^2}{\xi_i} = \left(\sum_{i=1}^{m} \xi_i\right) \cdot \left(\sum_{i=1}^{m} \frac{(p_i - q_i)^2}{\xi_i}\right)$$

$$\geq \left(\sum_{i=1}^{m} |p_i - q_i|\right)^2.$$

Plugging this inequality into Equation (1), we conclude that $KL(p,q) \ge \frac{1}{2} (\sum_{i=1}^{m} |p_i - q_i|)^2$.