

# CSC 665: Probability review

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August 29, 2019

## 1 Probability review

1. **Probability.** Denote by  $\mathbb{P}(A)$  the probability of event  $A$ ; (e.g. throwing a die,  $A = \{ \text{number 6 is up} \}$ ,  $\mathbb{P}(A) = 1/6$ .)

Probability satisfies additivity: if  $A \cap B = \emptyset$ , i.e. they are mutually exclusive, then  $\mathbb{P}(A \cup B) = \mathbb{P}(A) + \mathbb{P}(B)$ . It also satisfies subadditivity: for general  $A, B$ ,  $\mathbb{P}(A \cup B) \leq \mathbb{P}(A) + \mathbb{P}(B)$ .

Events  $A$  and  $B$  are called independent if  $\mathbb{P}(A \cap B) = \mathbb{P}(A) \cdot \mathbb{P}(B)$ .

Union bound:  $\mathbb{P}(A_1 \cup A_2 \cup \dots \cup A_n) \leq \mathbb{P}(A_1) + \dots + \mathbb{P}(A_n)$ .

2. **Expectation.** For a random variable  $X$ , denote by its expectation  $\mathbb{E}[X]$ . Specifically, if  $X$  takes value in a discrete set  $S$ , with probability mass function  $p$ , then

$$\mathbb{E}[X] \triangleq \sum_{x \in S} x \cdot p(x);$$

If  $X$  is continuous and has probability density function  $p$ , then

$$\mathbb{E}[X] \triangleq \int_{\mathbb{R}} x \cdot p(x) dx.$$

3. **Indicator function.** Denote by indicator function

$$\mathbf{1}(A) \triangleq \begin{cases} 1 & A \text{ is true,} \\ 0 & A \text{ is false.} \end{cases}$$

As  $\mathbf{1}(A)$  only takes values 0 and 1, we immediately get from the definition of expectation that

$$\mathbb{E}\mathbf{1}(A) = 1 \times \mathbb{P}(A) + 0 \times \mathbb{P}(\bar{A}) = \mathbb{P}(A).$$

4. **Linearity of expectation.** Suppose  $X, Y$  are two (possibly dependent) random variables. Then,  $\mathbb{E}[X + Y] = \mathbb{E}[X] + \mathbb{E}[Y]$ .

Is  $\mathbb{E}[XY] = \mathbb{E}[X]\mathbb{E}[Y]$ ? This is not true in general (consider  $(X, Y)$  as having the joint distribution of taking  $(-1, -1)$  and  $(+1, +1)$  with probability 0.5.) However, if  $X$  and  $Y$  are independent, then

$$\mathbb{E}[XY] = \mathbb{E}[X]\mathbb{E}[Y]$$

is true. Furthermore, if  $X$  and  $Y$  are independent, then for any functions  $f$  and  $g$ ,

$$\mathbb{E}f(X)g(Y) = \mathbb{E}[f(X)]\mathbb{E}[g(Y)].$$

5. **Variance.** Recall that the variance of a random variable  $X$  (with mean  $\mu$ ) is defined as:

$$\text{Var}(X) \triangleq \mathbb{E}(X - \mu)^2.$$

By linearity of expectation,

$$\mathbb{E}(X - \mu)^2 = \mathbb{E}X^2 - \mathbb{E}2X \cdot \mu + \mu^2 = \mathbb{E}X^2 - \mu^2.$$

How does  $\text{Var}(X+Y)$  relate to  $\text{Var}(X)$  and  $\text{Var}(Y)$ ? Again, there is no equation relationship in general. However a notable fact is that if  $X$  and  $Y$  are independent, then  $\text{Var}(X+Y) = \text{Var}(X) + \text{Var}(Y)$ . This is because,

$$\text{Var}(X+Y) = \mathbb{E}(X+Y - \mathbb{E}X - \mathbb{E}Y)^2 = \mathbb{E}(X - \mathbb{E}X)^2 + \mathbb{E}(Y - \mathbb{E}Y)^2 + 2\mathbb{E}(X - \mathbb{E}X)(Y - \mathbb{E}Y) = \text{Var}(X) + \text{Var}(Y).$$

6. **Jensen's Inequality.** Recall that a convex function  $f$  is one that for all  $x_1, x_2$  in  $\mathbb{R}$ , and  $t \in [0, 1]$ ,

$$f(tx_1 + (1-t)x_2) \leq tf(x_1) + (1-t)f(x_2).$$

Useful facts:

- (a) A twice differentiable function is convex if and only if its second derivative is always nonnegative. (This provides a practical way to check convexity.)
- (b) If  $f$  is differentiable, then for any  $x, y$ ,  $f(y) \geq f(x) + f'(x)(y - x)$ . That is,  $f$  is always above its first-order approximation. (For twice differentiable  $f$ , this is a direct consequence of Taylor's Theorem:  $f(y) = f(x) + f'(x)(y - x) + \frac{f''(\xi)}{2}(y - x)^2$  for some  $\xi$  between  $x$  and  $y$ .)

**Theorem 1.** Suppose  $f$  is a convex function, and  $X$  is a random variable. Then

$$f(\mathbb{E}X) \leq \mathbb{E}f(X).$$

*Proof.* We only show the inequality when  $f$  is differentiable. Denote by  $\mu \triangleq \mathbb{E}X$ . Observe that for all  $x$ ,  $f(x) \geq f(\mu) + f'(\mu)(x - \mu)$ . Taking expectation on both sides, we get that  $\mathbb{E}f(X) \geq f(\mu) + \mathbb{E}f'(\mu)(X - \mu) = f(\mu) = f(\mathbb{E}X)$ .  $\square$

7. **Markov's inequality:** a positive random variable with bounded mean should not take large values too often.

**Theorem 2** (Markov's Inequality). Suppose  $X$  is a nonnegative random variable. Then for any  $a > 0$ ,  $\mathbb{P}[X \geq a] \leq \frac{\mathbb{E}X}{a}$ .

*Proof.* Observe that for any positive  $x$ ,  $x \geq a\mathbf{1}(x \geq a)$ . Therefore,

$$\mathbb{E}X \geq \mathbb{E}a\mathbf{1}(X \geq a) = a\mathbb{P}(X \geq a).$$

The proof is concluded by dividing both sides by  $a$ .  $\square$

8. **Chebyshev's Inequality:** a random variable with a bounded variance should not deviate from its mean too often.

**Theorem 3** (Chebyshev's Inequality). Suppose  $X$  is a random variable with mean  $\mu$  and variance  $v > 0$ . Then for any  $b > 0$ ,  $\mathbb{P}[|X - \mu| \geq b] \leq \frac{v}{b^2}$ .

*Proof.* Applying Markov's Inequality to the random variable  $Y = (X - \mu)^2$  and  $a = b^2$ , we get

$$\mathbb{P}((X - \mu)^2 \geq b^2) \leq \frac{\mathbb{E}Y}{b^2}.$$

The proof is concluded by noting that the event  $\{|X - \mu| \geq b\}$  is the same as the event  $\{(X - \mu)^2 \geq b^2\}$ , and the fact that  $\mathbb{E}Y = v$ .  $\square$