CSCE 633: Machine Learning

Lecture 8: Model Selection and Regularization

Texas A&M University

9-11-18

Last Time

- Logistic Regression
- Gradient Descent

Goals of this lecture

- Ridge Regression
- Lasso Regularization
- Measures of Model Performance

Shrinkage

BZB effect B

 Train a model fitting all p predictors that constrains or regularizes the coefficients

• Two best methods for this, Ridge Regression and the Lasso

• Recall the least squares fit for $\beta_0, \beta_1, \cdots, \beta_p$ minimizes

•
$$RSS = \sum_{i=1}^{n} (y_i - \beta_0 - \sum_{j=1}^{p} \beta_j x_{ij})^2$$

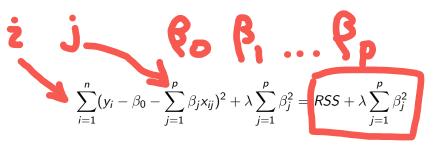
- Recall the least squares fit for $\beta_0, \beta_1, \dots, \beta_p$ minimizes
- $RSS = \sum_{i=1}^{n} (y_i \beta_0 \sum_{i=1}^{p} \beta_i x_{ij})^2$
- With Ridge Regression, we modify the equation that we want to minimize by adding a penalty (paying a price) for using predictors

- Recall the least squares fit for $\beta_0, \beta_1, \cdots, \beta_p$ minimizes
- $RSS = \sum_{i=1}^{n} (y_i \beta_0 \sum_{i=1}^{p} \beta_i x_{ij})^2$
- With Ridge Regression, we modify the equation that we want to minimize by adding a penalty (paying a price) for using predictors

$$\sum_{i=1}^{n} (y_i - \beta_0 - \sum_{j=1}^{p} \beta_j x_{ij})^2 + \lambda \sum_{j=1}^{p} \beta_j^2 = RSS + \lambda \sum_{j=1}^{p} \beta_j^2$$

ullet Where $\lambda \geq 0$ is the tuning parameter

因為RSS是正的,我們想找到 最小的RSS,如果Penalty是 負的,沒辦法找 Penalty的原因是,我們不 需要那麼多參數,5000個 參數和5個參數差不多,如 果只用5個參數,就需要 Penalty



- Where $\lambda > 0$ is the tuning parameter
- Ridge Regression creates a tradeoff, we still want coefficients that reduce *RSS* but now we have a shrinkage penalty.
- This penalty is small if β_0, \dots, β_p are close to 0
- Where least squares creates a single set of coefficients, Ridge Regression creates a set $\hat{\beta}^R_{\lambda}$ for each λ

mean value

How we know the penalty should keep positive?

$$\sum_{i=1}^{n} (y_i - \beta_0 - \sum_{j=1}^{p} \beta_j x_{ij})^2 + \lambda \sum_{j=1}^{p} \beta_j^2 = RSS + \lambda \sum_{j=1}^{p} \beta_j^2$$

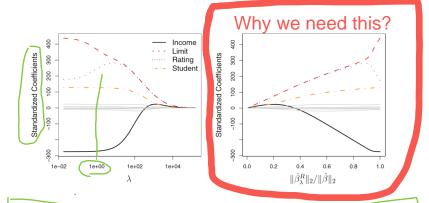
- Where least squares creates a single set of coefficients, Ridge Regression creates a set $\hat{\beta}_{\lambda}^{R}$ for each λ
- Selecting the right λ is key (done via cross-validation)
- Note, penalty is not assigned to the intercept β_0 since β_0 is the measure of the mean value of the response when

$$x_{i1} = x_{i2} = \cdots = x_{ip} = 0$$

• If we assume columns of X have been centered (meaning each column has a mean of 0) then intercept is $\hat{\beta}_0 = \bar{y} = \sum_{i=1}^n \frac{y_i}{n}$

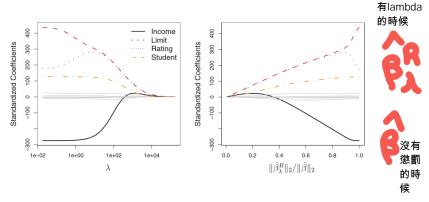
h T

Ridge Regression and Credit Data

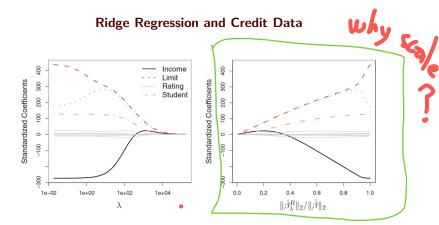


- Each line is one of the ten variables as a function of λ . Solid black line is the income variable
- We can see when $\lambda = 0$ we get standard least squares.
- When $\lambda = \infty$ we have the null model

Ridge Regression and Credit Data



- Income, limit, rating, and student have the largest coefficients.
- Note, in some steps individual estimates might actually grow because of relative important
- Right hand side of figure, we scale the x-axis as $\frac{\|\hat{\beta}\|_2}{\|\hat{\beta}\|_2}$ where $\|\hat{\beta}\|_2$ is the ℓ_2 norm of the least squares coefficient estimates

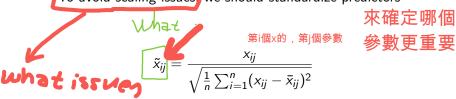


- Right hand side of figure, we scale the x-axis as $\frac{\|\hat{\beta}_{\lambda}^{R}\|_{2}}{\|\hat{\beta}\|_{2}}$ where $\|\hat{\beta}\|_{2}$ is the ℓ_{2} norm of the least squares coefficient estimates
- That value ranges from 1 when $\lambda=0$ to 0 when $\lambda=\infty$ what
- Thus, the x-axis represents amounts coefficient estimates have been shrunk to 0

B Mortazavi CSE

Scaling

- Scaling is now an important part of data we need to consider.
- Whereas, with least squares, if X_j was scaled by some constant c then least squares β would have been scaled by $\frac{1}{c}$, this is not true for Ridge Regression
- $x_j \hat{\beta}_{j,\lambda}^R$ will depend on λ and scaling of x_j and perhaps even the scaling of other predictors.
- To avoid scaling issues we should standardize predictors



Where the denominator is the estimated standard deviation of

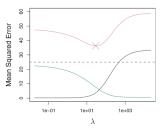
B Mortazavi CSE

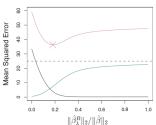
Normalization

- One important step in numeric data is normalization of the data to help these techniques
- ullet one common technique is to center and scale each predictor X_i
- This causes all the predictors to have mean 0 and standard deviation of 1

$$\tilde{x}_j = \frac{x_j - \bar{x}_j}{\sigma_i^2}$$

- 参數越多,約 flexible
- Rooted in the bias-variance trade-off
- As λ increases, flexibility of ridge regression fit decreases, decreasing variance but increasing bias.





- Simulated data of p = 45, n = 50, black is bias, green is variance, purple is test error
- $\lambda=30$ is the optimal solution and MSE of least squares is almost as high as null-model

LASSO

 Ridge Regression has one obvious disadvantage. Unlike subset methods, ridge regression still fits all p predictors.



- The penalty $\lambda \sum_j \beta_j^2$ will shrink all coefficients but none will hit 0 exactly
- This may not be a problem for accuracy but is for interpretability
- For example with our credit data set, Ridge Regression will still use all 10 predictors, even if it finds that income, limit, rating, and student are the most important.

LASSO Regularization

$$+\lambda \sum_{j=1}^{p} \beta_{j}^{2}$$

$$\sum_{i=1}^{n} (y_i - \beta_0 - \sum_{j=1}^{p} \beta_j x_{ij})^2 + \lambda \sum_{j=1}^{p} |\beta_j| = RSS + \lambda \sum_{j=1}^{p} |\beta_j|$$

- Creates a set $\hat{\beta}^L_{\lambda}$ for each λ
- \bullet We use the ℓ_1 norm instead of ℓ_2

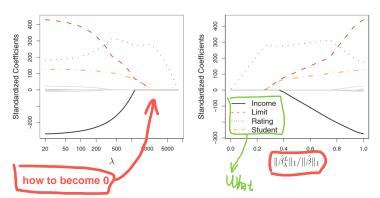




- Lasso shrinks coefficients but the ℓ_1 penalty drives coefficients to 0 when λ is sufficiently large
- This means Lasso performs variable selection!



Lasso and Credit Data



- Lasso picks rating, then student and limit together, then income. Eventually all others would enter as you approach least squares fit
- Where ridge selects coefficients/shrinkage, lasso produces models with any number of variables

Another Formulation

$$\min_{\beta} \sum_{i=1}^{n} (y_i - \beta_0 - \sum_{i=1}^{p} \beta_j x_{ij})^2$$

Subject to $\sum_{j=1}^{p} |\beta_j| \le s$ for Lasso and Subject to $\sum_{j=1}^{p} \beta_j^2 \le s$ for Ridge Regression

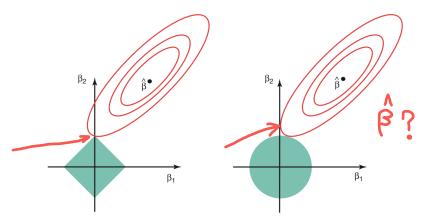
- If p=2 Lasso solution falls within the diamond $|eta_1|+|eta_2|\leq s$
- Circle for Ridge $\beta_1^2 + \beta_2^2 \le s$





B Mortazavi CSE

Another Formulation



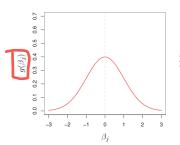
- 橢圓
- Ellipses are increasing RSS from the least squares solution
- if the λ allows enough to include RSS that is the fit found
- Because Lasso will intersect at a corner, while Ridge somewhere in between on circle, is why Lasso sets some coefficients to 0 while

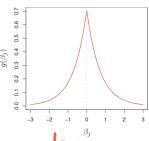
Ridge just shrinks them

Final Notes

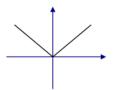
- Lasso is better if small set of predictors dominates response
- Ridge is better if all predictors contribute somewhat equally
- Cannot tell in advance, need cross-validation to give us an idea
- Lasso shrinks very differently than Ridge, known as soft thresholding

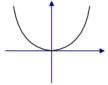
Ridge assumes the density function of the posterior probabilities of β are Gaussian (most coefficients are somewhere near 0), while Lasso assumes Laplacian (most coefficients centered at 0)





More about sparsity





$$0.5 \times (x-v)^2 + \lambda |x|$$

$$0.5 \times (x-v)^2 + \lambda x^2$$

If
$$v \ge \lambda$$
, $x=v-\lambda$
If $v \le -\lambda$, $x=v+\lambda$

Else, x=0

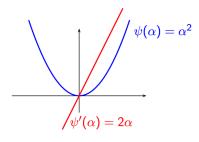
$$x=v/(1+2\lambda)$$

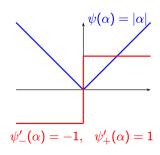
Nondifferentiable at 0

Differentiable at 0



More about sparsity





The gradient of the ℓ_2 -norm vanishes when α get close to 0. On its differentiable part, the norm of the gradient of the ℓ_1 -norm is constant.

How to Solve LASSO

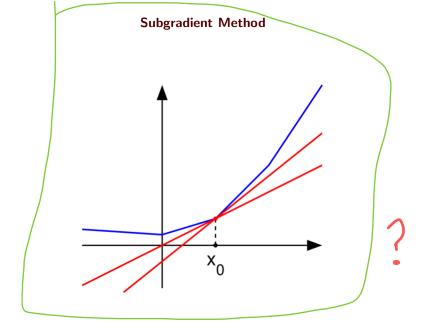
Rewrite the optimization problem: what why

$$\frac{\min_{\beta} \frac{1}{2} \|y - X\beta\|_2^2}{\|y - X\beta\|_2^2} + \lambda \|\beta\|_1$$

Challenges:



- The optimization if non-smooth.
- Subgradient Method
 - Subgradients are easy to derive and implement
 - Convergence needs carefully chosen step sizes
 - Convergence is weak theoretically



How to Solve LASSO



Fast I_1 minimization algorithms:

- Iterative Shrinkage Thresholding Algorithm (ISTA)
- Proximal Gradient Method (PGM)
- Alternating Direction Methods of Multpliers

Iterative Shrinkage Thresholding Algorithm (ISTA)

ISTA considers the LASSO model as a special case of the composite objective function:

$$\min_{\beta} F(\beta) = f(\beta) + g(\beta),$$

where f is a smooth and convex function, and g is the regularization term that is not necessarily smooth nor convex. Here $f(\beta) = \frac{1}{2} ||y - X\beta||_2^2$.

- If $g(\beta) = \lambda \|\beta\|_2^2$: Ridge regression.
- If $g(\beta) = \lambda \|\beta\|_1$: LASSO.

ISTA using Hessian Matrix Approximation

Estimate $f(\beta)$ using its Taylor expansion to the second order around β^k :

around
$$\beta^k$$
:
$$\beta^{k+1} = \underset{\beta}{\operatorname{argmin}_{\beta}} \{ f(\beta^k) + \nabla f(\beta^k)^T (\beta - \beta^k) + \frac{1}{2} (\beta - \beta^k)^T \nabla^2 f(\beta^k) (\beta - \beta^k) + g(\beta) \}$$

$$\approx \underset{\beta}{\operatorname{argmin}_{\beta}} \{ \nabla f(\beta^k)^T (\beta - \beta^k) + \frac{\alpha^k}{2} \|\beta - \beta^k\|_2^2 + g(\beta) \}$$

$$= \underset{\beta}{\operatorname{argmin}_{\beta}} \{ \frac{\alpha^k}{2} \|\beta - \gamma^k\|_2^2 + g(\beta) \}$$

where

$$\gamma^k = \beta^k - \frac{1}{\alpha^k} \nabla f(\beta^k).$$

ISTA using Hessian Matrix Approximation

Specifically for LASSO, where $g(\beta) = \|\beta\|_1$, the last optimization step is seperable:

$$\beta^{k+1} = \operatorname{argmin}_{\beta} \{ \sum_{i} \frac{\alpha^{k}}{2} (\beta_{i} - \gamma_{i}^{k})_{2}^{2} + \lambda |\beta_{i}| \} \}$$

The problem consists of multiple independent 1-D problems that have explicit solution:

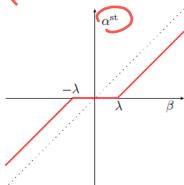
$$eta_i^{k+1} = soft(\gamma_i^k, \frac{\lambda}{\alpha^k})$$
 數學而已

$$\begin{array}{lll} \operatorname{soft}(u,a) & \doteq & \operatorname{sgn}(u) \max\{|u|-a,0\} \\ & = & \begin{cases} \operatorname{sgn}(u)(|u|-a) & \text{if lul>a} \\ 0 & \text{otherwise} \end{cases} \end{array}$$

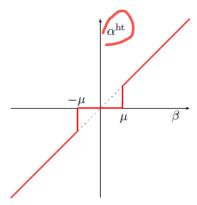
Soft Thresholding Function



When will it soft?



(a) Soft-thresholding operator, $\alpha^{\text{st}} = \text{sign}(\beta) \max(|\beta| - \lambda, 0).$



(b) Hard-thresholding operator $\alpha^{\rm ht} = \mathbf{1}_{|\beta| \geq \mu} \beta$.

ISTA using Proximal Gradient Method

At each step, perform a gradient descent step on $f(\beta)$ without considering the non-smooth regularization:

$$\gamma^{k} = \beta^{k} - \alpha^{k} \nabla f(\beta^{k}) = \beta^{k} + \alpha^{k} X^{T} (y - X \beta^{k})$$

Then combine the regularization by solving the following proximity problem:

$$\beta^{k+1} = \operatorname{argmin}_{\beta} \{ \frac{1}{2\alpha^k} \|\beta - \gamma^k\|_2^2 + \lambda \|\beta\|_1 \}$$

which will induce exactly the same solution:

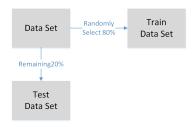
$$\beta^{k+1} = soft(\beta^k + \alpha^k X^T (y - X\beta^k), \alpha^k \lambda).$$

Need select a small enough step size α^k

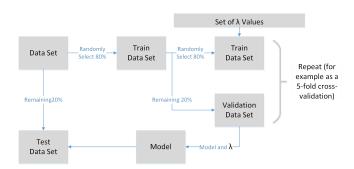
Lasso: Continued

- Selecting λ
- Further Lasso Solvers and Elastic Net

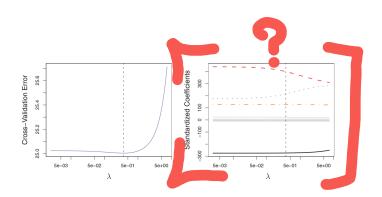
- Need to pick best λ (or s in the alternative formulation) for best estimation
- ullet We can run a cross-validation over a grid of λ values
- We pick the lambda with the smallest error







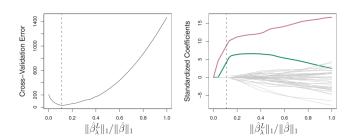
LASSO: Our Credit Example



- Sometimes Lasso does not do better than Least Squares Solution
- Small λ selected here



LASSO: Synthetic Example



• Sometimes Lasso does a lot better than Least Squares Solution



LASSO: Elastic Net

- Last week, we discussed the differences between Ridge Regression and Lasso
- We also discussed how it is not immediately obvious which would be better, sometimes need cross-validation to test
- if n > p but variables are correlated, ridge empirically does better than lasso
- if p > n lasso cannot select more than n variables before it saturates.
- A mix between Lasso and Ridge exists, called Elastic Net





B Mortazavi CSE

Vanilla Elastic Net

New Objective Function is

用它來決定用哪個 Penalty

$$J(\beta, \lambda_1, \lambda_2) = \|y - X\beta\|^2 + \lambda_2 \|\beta\|_2^2 + \lambda_1 \|\beta\|_1$$

- The objective now has a penalty that is from ridge regression and a penalty that is from lasso
- It turns out this doesn't predict really well, unless the optimal solution is found by ridge or by lasso
- This is because some solution in the middle has coefficients penalized by both λ_1 and λ_2

Vanilla Elastic Net

New Objective Function is

$$J(\beta, \lambda_1, \lambda_2) = \|y - X\beta\|^2 + \lambda_2 \|\beta\|_2^2 + \lambda_1 \|\beta\|_1$$

- The objective now has a penalty that is from ridge regression and a penalty that is from lasso
- It turns out this doesn't predict really well, unless the optimal solution is found by ridge or by lasso
- This is because some solution in the middle has coefficients penalized by both λ_1 and λ_2
- To fix it, we adjust the optimal solution. So, first, we solve the vanilla version



LARS-Elastic Net

First we re-write X as

方便differentiation!



$$\tilde{X} = \frac{1}{\sqrt{1+\lambda_2}} \begin{pmatrix} X \\ \sqrt{\lambda_2} I_p \end{pmatrix}$$



Where I_p is the identity matrix and

$$\tilde{y} = \begin{pmatrix} y \\ 0_{p \times 1} \end{pmatrix}$$



Then we solve for β like a normal lasso problem

$$\tilde{\beta} = \operatorname{argmin}_{\tilde{\beta}} \|\tilde{y} - \tilde{X}\tilde{\beta}\|^2 + \frac{\lambda_1}{\sqrt{1 + \lambda_2}} \|\tilde{\beta}\|_1$$

So
$$\beta = \frac{\tilde{\beta}}{\sqrt{1+\lambda_2}}$$

Improved Elastic Net

Then we solve for β like a normal lasso problem

$$\tilde{\beta} = \operatorname{argmin}_{\tilde{\beta}} \lVert \tilde{\mathbf{y}} - \tilde{X} \tilde{\beta} \rVert^2 + \frac{\lambda_1}{\sqrt{1 + \lambda_2}} \lVert \tilde{\beta} \rVert_1$$

So
$$\beta = \frac{\tilde{\beta}}{\sqrt{1+\lambda_2}}$$

- So now we want to undo one of the penalties so coefficients aren't double penalized
- for simplicity we undo the λ_2 penalty (ℓ_2)

$$\hat{\beta} = \sqrt{1 + \lambda_2} \tilde{\beta}$$

Measures of Classification Accuracy

- Confusion Matrices
- Class labels versus Probabilities
- Measurements

Confusion Matrix

		True default status					
		No	Yes	Total			
Predicted	No	9,644	252	9,896			
$default\ status$	Yes	23	81	104			
	Total	9,667	333	10,000			

- If we denote true positives as TP, false positives as FP, true negatives as TN, and false negatives as FN then we can say Accuracy is
- $ACC = \frac{TP+TN}{TP+FN+TN+FP}$
- What is the accuracy here?

Confusion Matrix

		True default status					
		No	Yes	Total			
Predicted	No	9,644	252	9,896			
$default\ status$	Yes	23	81	104			
	Total	9,667	333	10,000			

- If we denote true positives as TP, false positives as FP, true negatives as TN, and false negatives as FN then we can say Accuracy is
- $ACC = \frac{TP+TN}{TP+FN+TN+FP}$
- What is the accuracy here? 97.25%

Confusion Matrix: Problems with Accuracy?

		True	
		No	Yes
Predicted	No	50	50
	Yes	0	0

- If we denote true positives as TP, false positives as FP, true negatives as TN, and false negatives as FN then we can say Accuracy is
- $ACC = \frac{TP + TN}{TP + FN + TN + FP}$

Confusion Matrix with Class Imbalance: Problems with Accuracy?

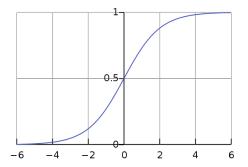
		True	
		No	Yes
Predicted	No	900	100
	Yes	0	0

- If we denote true positives as TP, false positives as FP, true negatives as TN, and false negatives as FN then we can say Accuracy is
- $ACC = \frac{TP + TN}{TP + FN + TN + FP}$

Confusion Matrix: Measurements in Addition to Accuracy

		True default status						
		No	Yes	Total				
Predicted	No	9,644	252	9,896				
$default\ status$	Yes	23	81	104				
	Total	9,667	333	10,000				

- If we denote true positives as TP, false positives as FP (Type I Error), true negatives as TN, and false negatives as FN (Type II Error) then we can say Accuracy is
- $ACC = \frac{TP + TN}{TP + FN + TN + FP}$
- Recall = Sensitivity = True Positive Rate = $\frac{TP}{TP+FN}$
- Specificity = True Negative Rate = $\frac{TN}{TN+FP} = 1 FPR$
- False Positive Rate (FPR) = $\frac{FP}{FP+TN}$
- Precision = Positive Predictive Value = $\frac{TP}{TP+FP}$
- F1 Score = $2 \times \frac{Precision \times Recall}{Precision + Recall}$
- Can calculate these per-class and average together or total across all classes



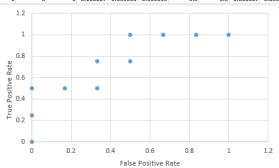
If we pick a decision threshold of p(x) > 0.5 what happens?

Predicted	Ground	Truth
2%	0	
3%	0	
5%	1	
1%	0	
15%	0	
25%	1	
24%	1	
13%	0	
8%	0	
12%	1	

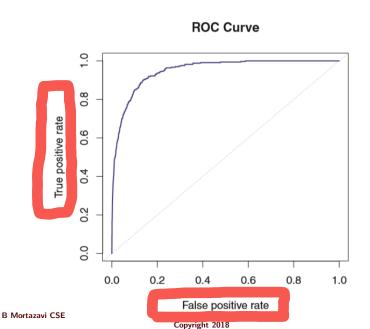
Predicted	TRUE	50	25	24	15	13	12	8	5	3	2	1
1	0	0	0	0	0	0	0	0	0	0	0	1
2	0	0	0	0	0	0	0	0	0	0	1	1
3	0	0	0	0	0	0	0	0	0	1	1	1
5	1	0	0	0	0	0	0	0	1	1	1	1
8	0	0	0	0	0	0	0	1	1	1	1	1
12	1	0	0	0	0	0	1	1	1	1	1	1
13	0	0	0	0	0	1	1	1	1	1	1	1
15	0	0	0	0	1	1	1	1	1	1	1	1
24	1	0	0	1	1	1	1	1	1	1	1	1
25	1	0	1	1	1	1	1	1	1	1	1	1

	50	25	24	15	13	12	8	5	3	2	1
TP	0	1	2	2	2	3	3	4	4	4	4
FP	0	0	0	1	2	2	3	3	4	5	6
TN	6	6	6	5	4	4	3	3	2	1	0
FN	4	3	2	2	2	1	1	0	0	0	0
TPR	0	0.25	0.5	0.5	0.5	0.75	0.75	1	1	1	1
FPR	0	0	0	0.166667	0.333333	0.333333	0.5	0.5	0.666667	0.833333	1

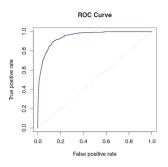
	50	25	24	15	13	12	8	5	3	2	1
TP	0	1	2	2	2	3	3	4	4	4	4
FP	0	0	0	1	2	2	3	3	4	5	6
TN	6	6	6	5	4	4	3	3	2	1	0
FN	4	3	2	2	2	1	1	0	0	0	0
TPR	0	0.25	0.5	0.5	0.5	0.75	0.75	1	1	1	1
FPR	0	0	0	0.166667	0.333333	0.333333	0.5	0.5	0.666667	0.833333	1



Receiver Operating Characteristic Curve



ROC and Measurements



- By predicting probability instead of label can generate ROC Curve
- Can choose threshold on ROC curve that optimizes some threshold-specific Measurement
- Can also plot Precision-Recall Curve
- Area Under ROC Curve (AUROC) is a measurement of how well your model separates classes (without taking decision threshold into account)
 - Can also calculate AUPRC

Problem Solving: Gradient Descent

$$MSE(\beta_0, \beta_1) = J(\beta_0, \beta_1) = J(\beta) = \frac{1}{n} \sum_{i=1}^{n} (y_i - (\beta_0 + \beta_1 x_i))^2$$

- Optimize β_0 and β_1 by Gradient Descent
- Remember, Update Rules $\beta_0^{t+1} = \beta_0^t \alpha \frac{\partial}{\partial \beta_0} J(\beta_0, \beta_1)$
- Remember, Update Rules $\beta_1^{t+1} = \beta_1^t \alpha \frac{\partial}{\partial \beta_1} J(\beta_0, \beta_1)$
- if $\boldsymbol{\beta} = (\beta_0, \beta_1)$ then
- $\boldsymbol{\beta}^{t+1} = \boldsymbol{\beta}^k \alpha \nabla J(\boldsymbol{\beta})$

Problem Solving: Jacobian

$$\nabla J(\beta)$$

- If we have a series of functions f that map input x to an output y (such as linear regression)
- The Jacobian is a matrix such that $J_{ij}=rac{\partial f_i}{\partial x_j}$
- The Hessian is a matrix such that $H_{ij}=rac{\partial^2 f}{\partial x_i\partial x_j}$

Problem Solving: Jacobian

$$f(x,y) = \begin{bmatrix} x^2y \\ 5x + \sin(y) \end{bmatrix}$$

$$\nabla f = \begin{bmatrix} \frac{\partial f_1}{\partial x} & \frac{\partial f_1}{\partial y} \\ \frac{\partial f_2}{\partial x} & \frac{\partial f_2}{\partial y} \end{bmatrix} = \begin{bmatrix} 2xy & x^2 \\ 5 & \cos(y) \end{bmatrix}$$

Problem Solving: Gradient Descent

$$MSE(\beta_0, \beta_1) = J(\beta_0, \beta_1) = J(\beta) = \frac{1}{2n} \sum_{i=1}^{n} ((\beta_0 + \beta_1 x_i) - y_i)^2$$

Our Objective

$$\min_{\beta_0,\beta_1} J(\beta_0,\beta_1) = \min_{\beta} J(\beta)$$

Then our update rule at time t+1, with stopping criteria ϵ is

$$\beta^{t+1} = \beta^t - \alpha \nabla J(\beta)$$

$$\nabla J(\beta) = \begin{bmatrix} \frac{\partial J(\beta)}{\partial \beta_0} & \frac{\partial J(\beta)}{\partial \beta_1} \end{bmatrix}$$
$$\|\nabla J(\beta)\| < \epsilon$$

Problem Solving: Gradient Descent

$$\nabla J(\beta) = \begin{bmatrix} \frac{\partial J(\beta)}{\partial \beta_0} & \frac{\partial J(\beta)}{\partial \beta_1} \end{bmatrix}$$

$$= \begin{bmatrix} \frac{1}{n} \sum_{i=1}^{n} (\beta_0 + \beta_1 x_i - y_i) & \frac{1}{n} \sum_{i=1}^{n} (\beta_0 + \beta_1 x_i - y_i) x_i \end{bmatrix}$$

Takeaways and Next Time

- Model selection finds optimal solutions by using a subset of predictors
- Regularization minimizes RSS subject to a price for the predictors used
- Model Performance and Problem Solving
- Next Time: Measures of Performance with Logistic Regression