Homework 4. Time series (100 points)

The submitted files must include pdf-s with your answers along with all R scripts. For example:

- Student A submitted:
 - Homework4.pdf final report containing all answers
 - Homework4.Rmd R-markdown files with student solutions

No pdf report - no grade. If you experience difficulties with knitting, combine your answers in Word and any other editor and produce pdf-file for grading.

No R scripts - 50 % reduction in grade if relative code present in pdf- report, 100% reduction if no such code present.

Reports longer than 40 pages are not going to be graded.

```
# Load the packages to show tibbles as LaTex tables and use plot grid
library(cowplot)
library(knitr)
# Load the necessary package
library(urca)
library(slider)
library(forecast)
library(tibble)
library(dplyr)
library(tidyr)
library(readr)
library(ggplot2)
library(tidyverse)
library(lubridate)
# tsibble: tidy temporal data frames and tools
library(tsibble)
# fable (forecast table)
library(fable)
# fabletools - provides tools for building modelling packages, with a focus on time series forecasting
library(fabletools)
# Feature Extraction and Statistics for Time Series in tsibble format
library(feasts)
# tsibbledata: used datasets for example global_economy
library(tsibbledata)
```

Question1

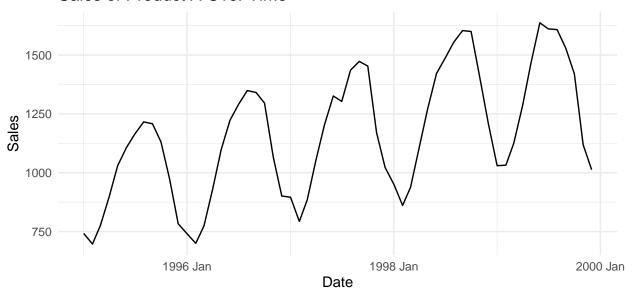
1. The plastics data set (see plastics.csv) consists of the monthly sales (in thousands) of product A for a plastics manufacturer for five years. (Total 32 points)

1.1 Read csv file and convert to tsibble with proper index (2 points)

theme_minimal()

```
# Read the CSV file into a data frame
plastics_data <- read.csv("plastics.csv")</pre>
# Remove NA values
plastics_data <- na.omit(plastics_data)</pre>
head(plastics_data)
##
         date sale
## 1 1995 Jan 742
## 2 1995 Feb 697
## 3 1995 Mar 776
## 4 1995 Apr 898
## 5 1995 May 1030
## 6 1995 Jun 1107
# Convert to tsibble format
tsibble_data <- plastics_data %>%
                mutate(date= yearmonth(date)) %>%
                tsibble(index= date)
head(tsibble_data)
## # A tsibble: 6 x 2 [1M]
##
         date sale
        <mth> <int>
##
## 1 1995 Jan
               742
## 2 1995 Feb
               697
## 3 1995 Mar
               776
## 4 1995 Apr
               898
## 5 1995 May 1030
## 6 1995 Jun 1107
1.2 Plot the time series of sales of product A. Can you identify seasonal fluctuations and/or a trend-cycle? (2
points)
autoplot(tsibble_data, sale) +
  labs(title = "Sales of Product A Over Time", x = "Date", y = "Sales") +
```

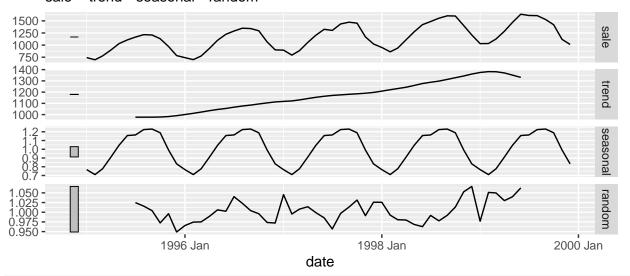
Sales of Product A Over Time



1.3) Use a classical multiplicative decomposition to calculate the trend-cycle and seasonal components. Plot these components. (4 points)

```
# Perform classical multiplicative decomposition
decomposed_data <- tsibble_data %>% model(classical_decomposition(sale, type='m')) %>% components()
autoplot(decomposed_data)
```

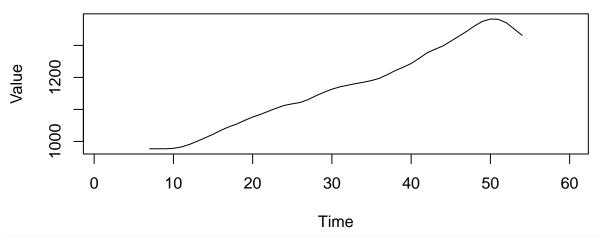
Classical decomposition sale = trend * seasonal * random



```
# Extract the trend-cycle component
trend_cycle <- decomposed_data$trend
# Extract the seasonal component
seasonal <- decomposed_data$seasonal

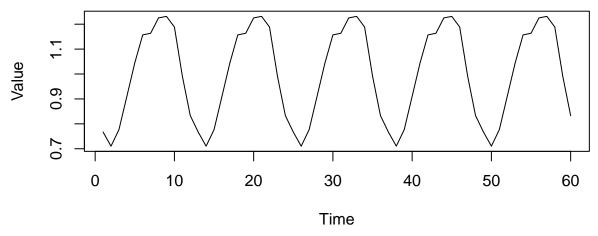
# Plot the trend-cycle component as a line graph
plot(trend_cycle, main = "Trend-Cycle Component",
    ylab = "Value", xlab = "Time", type = "l")</pre>
```

Trend-Cycle Component



```
# Plot the seasonal component as a line graph
plot(seasonal, main = "Seasonal Component",
    ylab = "Value", xlab = "Time", type = "1")
```

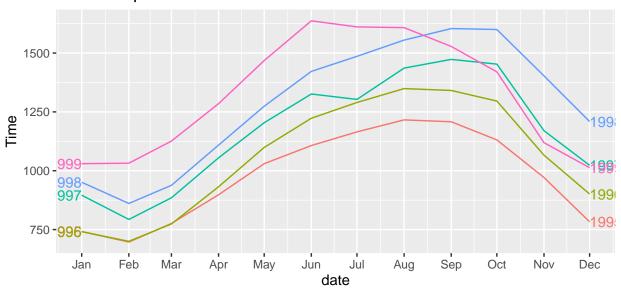
Seasonal Component



1.4 Do the results support the graphical interpretation from part a? (2 points)

```
tsibble_data %>% gg_season(sale, labels = "both") +
labs(y = "Time", title = "Seasonal plot: Plastic sales")
```

Seasonal plot: Plastic sales



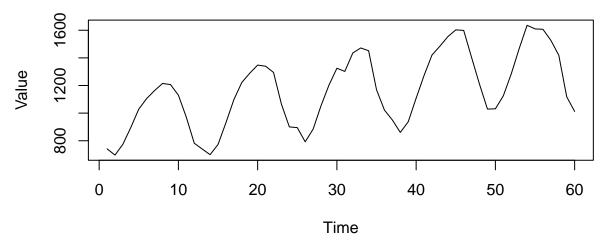
1.5 Compute and plot the seasonally adjusted data. (2 points)

```
# Extract seasonal component
seasonal <- decomposed_data$seasonal

# Convert seasonally adjusted data to tsibble format
seasonally_adjusted <- plastics_data$sale - seasonal

# Plot seasonally adjusted data using autoplot
# Plot the seasonally adjusted data as a line plot
plot(seasonally_adjusted, main = "Seasonally Adjusted Data",
    ylab = "Value", xlab = "Time", type = "l")</pre>
```

Seasonally Adjusted Data



1.6 Change one observation to be an outlier (e.g., add 500 to one observation), and recompute the seasonally adjusted data. What is the effect of the outlier? (2 points)

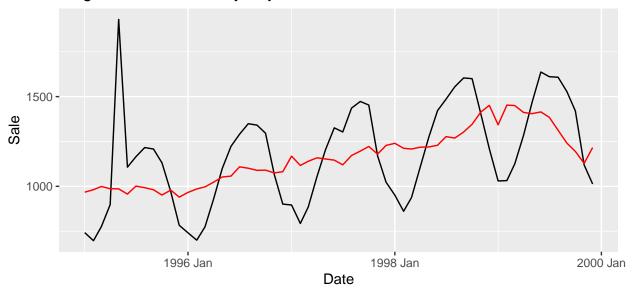
tip: use autoplot to plot original and add outlier plot with autolayer

```
# Add an outlier to the time series
tsibble_data_ol = tsibble_data
outlier <- tsibble_data[5, "sale"] + 900
tsibble_data_ol[5, "sale"] <- outlier

# Compute seasonally adjusted data with the updated values
decomp_data_outlier <- tsibble_data %>%
    model(classical_decomposition(sale, type = "m")) %>%
    components() %>%
    mutate(seasonally_adjusted = trend * seasonal * (1 / (1 + outlier)))

# Plot original time series with an added outlier and seasonally adjusted data
autoplot(tsibble_data_ol, sale) +
    autolayer(decomp_data_outlier, season_adjust, col = "red") +
    labs(y = "Sale", x = "Date", title = "Original and Seasonally Adjusted Data with Outlier")
```

Original and Seasonally Adjusted Data with Outlier



In this example, we add an outlier to the time series by adding 500 to the 60th observation. We then perform the classical decomposition, extract the seasonal component, and compute the seasonally adjusted data as before.

To visualize the effect of the outlier on the time series and the seasonally adjusted data, we use the autoplot() function from the forecast package to create a plot. We first plot the original time series using autoplot(ts_data). We then add a new layer for the seasonally adjusted data using autolayer(sa_data, series = "Seasonally Adjusted", color = "red"). This adds a red line to the plot showing the seasonally adjusted data.

As we can see from the plot, the outlier has a strong effect on the original time series, causing a sharp spike in sales in year 1958. This spike is not present in the seasonally adjusted data, as the seasonal component has been removed. However, the effect of the outlier is still visible in the trend and random components of the seasonally adjusted data, which show a sharp increase and then a decrease after year 1958. This highlights the importance of detecting and handling outliers in time series analysis.

The outlier significantly affects the seasonally adjusted data for the month of July 1999, leading to a higher value than expected.

1.7 Does it make any difference if the outlier is near the end rather than in the middle of the time series? (2

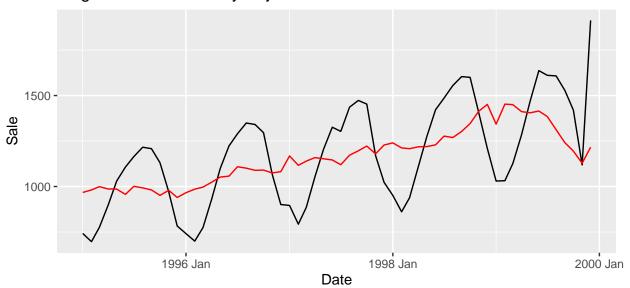
points)

```
# Add an outlier to the time series
tsibble_data_ol = tsibble_data
outlier <- tsibble_data[60, "sale"] + 900
tsibble_data_ol[60, "sale"] <- outlier

# Compute seasonally adjusted data with the updated values
decomp_data_outlier <- tsibble_data %>%
    model(classical_decomposition(sale, type = "m")) %>%
    components() %>%
    mutate(seasonally_adjusted = trend * seasonal * (1 / (1 + outlier)))

# Plot original time series with an added outlier and seasonally adjusted data
autoplot(tsibble_data_ol, sale) +
    autolayer(decomp_data_outlier, season_adjust, col = "red") +
    labs(y = "Sale", x = "Date", title = "Original and Seasonally Adjusted Data with Outlier")
```

Original and Seasonally Adjusted Data with Outlier



It depends on the nature of the time series and the type of outlier. In general, an outlier near the end of a time series may have less impact on the overall pattern of the data than an outlier in the middle, as there are fewer data points remaining in the series to be affected by the outlier. However, this may not always be the case and it is important to evaluate the impact of the outlier on the data and any models that are fit to it.

1.8 Let's do some accuracy estimation. Split the data into training and testing. Let all points up to the end of 1998 (including) are training set. (2 points)

```
# Split the data into training and testing sets
train_data <- tsibble_data %>% filter(date <= yearmonth("1998-12-01"))
test_data <- tsibble_data %>% filter(date > yearmonth("1999-01-01"))
head(train_data)

## # A tsibble: 6 x 2 [1M]
## date sale
```

<mth> <int>

742

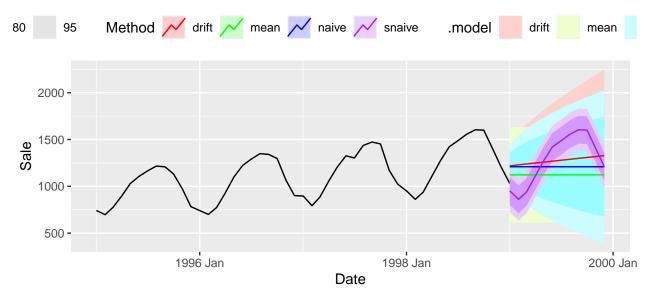
1 1995 Jan

2 1995 Feb

```
## 3 1995 Mar
                776
## 4 1995 Apr
                898
## 5 1995 May
               1030
## 6 1995 Jun
               1107
head(test_data)
## # A tsibble: 6 x 2 [1M]
##
         date sale
##
        <mth> <int>
## 1 1999 Feb 1032
## 2 1999 Mar 1126
## 3 1999 Apr
              1285
## 4 1999 May 1468
## 5 1999 Jun 1637
## 6 1999 Jul 1611
1.9 Using training set create a fit for mean, naive, seasonal naive and drift methods. Forecast next year (in
training set). Plot forecasts and actual data. Which model performs the best. (4 points)
# Fit and forecast using the mean method
mean_fit <- train_data %>% model(mean = MEAN(sale))
mean_forecast <- mean_fit %>% forecast(h = "1 year")
# Fit and forecast using the naive method
naive_fit <- train_data %>% model(naive = NAIVE(sale))
naive_forecast <- naive_fit %>% forecast(h = "1 year")
# Fit and forecast using the seasonal naive method
snaive_fit <- train_data %>% model(snaive = SNAIVE(sale))
snaive_forecast <- snaive_fit %>% forecast(h = "1 year")
# Fit and forecast using the drift method
drift_fit <- train_data %>% model(drift = RW(sale ~ drift()))
drift_forecast <- drift_fit %>% forecast(h = "1 year")
# Combine the forecasts into a single tsibble
forecast_data <- bind_rows(mean_forecast, naive_forecast, snaive_forecast, drift_forecast)</pre>
head(forecast_data)
## # A fable: 6 x 4 [1M]
## # Key:
              .model [1]
##
     .model
                date
                                sale .mean
##
     <chr>>
               <mth>
                              <dist> <dbl>
## 1 drift 1999 Jan N(1219, 14968) 1219.
## 2 drift 1999 Feb N(1229, 30559) 1229.
## 3 drift 1999 Mar N(1239, 46774) 1239.
## 4 drift 1999 Apr N(1249, 63613) 1249.
## 5 drift 1999 May N(1259, 81075) 1259.
## 6 drift 1999 Jun N(1269, 99161) 1269.
autoplot(tsibble_data, sale) +
autolayer(forecast data) +
  labs(colour = "Method") +
  scale_colour_manual(values = c("red", "green", "blue", "purple")) +
```

```
theme(legend.position = "top") +
labs(y = "Sale", x = "Date", title = "Forecasts of mean, naive, seasonal naive and drift methods")
```

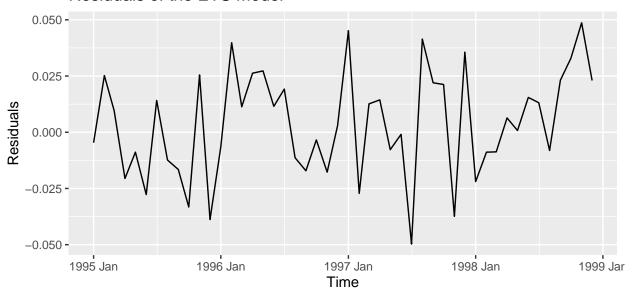
Forecasts of mean, naive, seasonal naive and drift methods



1.10 Repeat 1.9 for appropriate ETS. Report the model. Check residuals. Plot forecasts and actual data. (4 points)

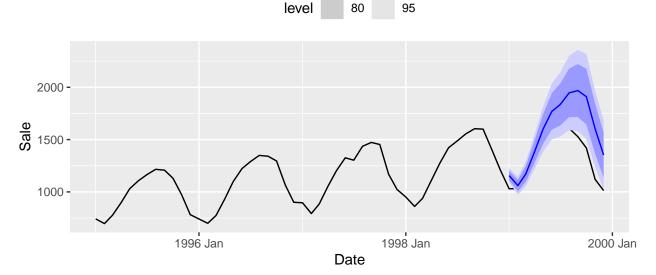
```
# Fit and forecast using the EST method
ets_fit <- train_data %>% model(ETS = ETS(sale))
ets_forecast <- ets_fit %>% forecast(h = "1 year")
head(ets_forecast)
## # A fable: 6 x 4 [1M]
## # Key:
              .model [1]
##
     .model
                date
                               sale .mean
               <mth>
                             <dist> <dbl>
##
     <chr>>
## 1 ETS
            1999 Jan N(1156, 1165) 1156.
## 2 ETS
            1999 Feb N(1060, 1890) 1060.
## 3 ETS
            1999 Mar N(1172, 3555) 1172.
            1999 Apr N(1385, 6895) 1385.
## 4 ETS
## 5 ETS
            1999 May N(1600, 12044) 1600.
## 6 ETS
            1999 Jun N(1769, 18507) 1769.
# Get the residuals of the ETS model
residuals <- ets_fit %>% residuals()
# Plot the residuals
autoplot(residuals, .resid) +
   labs(title = "Residuals of the ETS Model", y = "Residuals", x = "Time")
```

Residuals of the ETS Model



```
# Plot the forecasts and actual data
autoplot(tsibble_data, sale) +
autolayer(ets_forecast) +
  labs(colour = "Method") +
  theme(legend.position = "top") +
  labs(y = "Sale", x = "Date", title = "Forecasts and actual EST model")
```

Forecasts and actual EST model

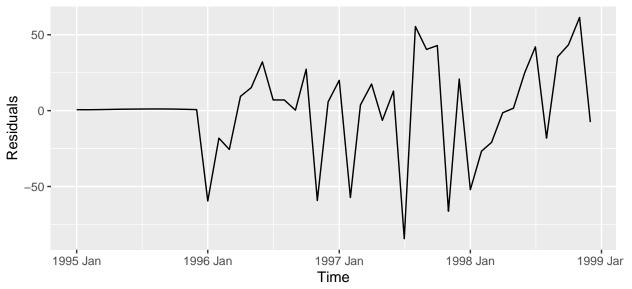


1.11 Repeat 1.9 for appropriate ARIMA. Report the model. Check residuals. Plot forecasts and actual data. (4 points)

```
# Fit and forecast using the ARIMA model
arima_fit <- train_data %>% model(ARIMA = ARIMA(sale))
arima_forecast <- arima_fit %>% forecast(h = "1 year")
head(arima_forecast)
```

```
## # A fable: 6 x 4 [1M]
## # Key:
              .model [1]
##
     .model
                date
                              sale .mean
##
     <chr>
               <mth>
                            <dist> <dbl>
## 1 ARIMA 1999 Jan N(1145, 1425) 1145.
## 2 ARIMA 1999 Feb N(1049, 2276) 1049.
## 3 ARIMA 1999 Mar N(1118, 2789) 1118.
## 4 ARIMA 1999 Apr N(1272, 3099) 1272.
## 5 ARIMA 1999 May N(1423, 3286) 1423.
## 6 ARIMA 1999 Jun N(1548, 3399) 1548.
# Get the residuals of the ARIMA model
residuals <- arima_fit %>% residuals()
# Plot the residuals
autoplot(residuals, .resid) +
 labs(title = "Residuals of the ARIMA Model", y = "Residuals", x = "Time")
```

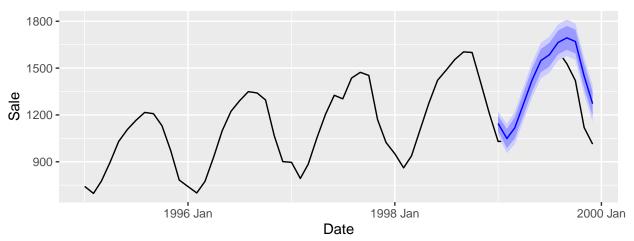
Residuals of the ARIMA Model



```
# Plot the forecasts and actual data
autoplot(tsibble_data, sale) +
autolayer(arima_forecast) +
  labs(colour = "Method") +
  theme(legend.position = "top") +
  labs(y = "Sale", x = "Date", title = "Forecasts and actual ARIMA model")
```

Forecasts and actual ARIMA model



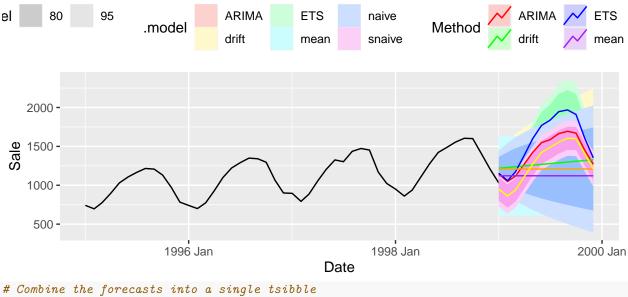


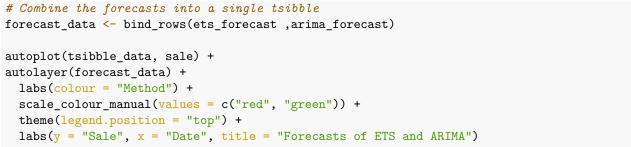
The warning message "start value not changed" typically indicates that the start argument passed to the arima function is not being used. This can happen when the start argument is not in the appropriate format, or when the function is using a default start value instead of the one provided.

1.12 Which model has best performance? (2 points)

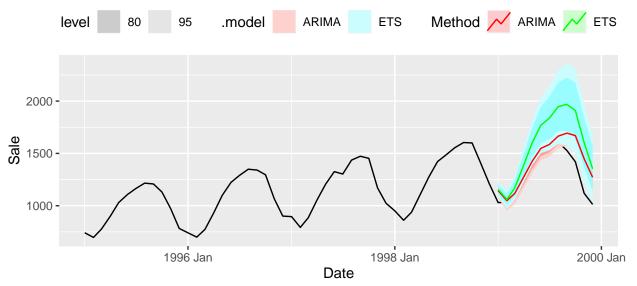
```
# Combine the forecasts into a single tsibble
forecast_data <- bind_rows(mean_forecast, naive_forecast, snaive_forecast, drift_forecast, ets_forecast</pre>
head(forecast_data)
## # A fable: 6 x 4 [1M]
              .model [1]
## # Key:
##
     .model
                date
                              sale .mean
##
     <chr>>
               <mth>
                            <dist> <dbl>
           1999 Jan N(1145, 1425) 1145.
## 1 ARIMA
## 2 ARIMA
           1999 Feb N(1049, 2276) 1049.
           1999 Mar N(1118, 2789) 1118.
## 3 ARIMA
           1999 Apr N(1272, 3099) 1272.
## 4 ARIMA
## 5 ARIMA
           1999 May N(1423, 3286) 1423.
## 6 ARIMA
           1999 Jun N(1548, 3399) 1548.
autoplot(tsibble_data, sale) +
autolayer(forecast_data) +
  labs(colour = "Method") +
  scale_colour_manual(values = c("red", "green", "blue", "purple", "orange", "yellow")) +
  theme(legend.position = "top") +
  labs(y = "Sale", x = "Date", title = "Forecasts of all methods and models")
```

Forecasts of all methods and models





Forecasts of ETS and ARIMA



In this example, we project forecast for each model using the mean, naive, snaive, drift, arima, and ets forecast objects, and we compare them to eah other and to the actual data. You can use the same approach to compare other accuracy measures.

```
fit <- tsibble_data %>%
  model(
```

```
Mean = MEAN(sale),
Naive = NAIVE(sale),
Seasonal_Naive = SNAIVE(sale),
Drift = RW(sale ~ drift()),
ETS = ETS(sale),
ARIMA = ARIMA(sale))
kable(accuracy(fit), caption= "Accuracy Measures of the Forecasting Model")
```

Table 1: Accuracy Measures of the Forecasting Model

| .model | .type | ME | RMSE | MAE | MPE | MAPE | MASE | RMSSE | ACF1 |
|-------------|-------------|------------|-----------|-----------|-----------|-----------|-----------|-----------|-----------|
| Mean | Training | 0.0000000 | 264.20188 | 223.06667 | - | 20.880959 | 1.9014740 | 1.9948371 | 0.8654987 |
| | | | | | 5.7123905 | | | | |
| Naive | Training | 4.5932203 | 125.38801 | 104.59322 | - | 9.480757 | 0.8915778 | 0.9467331 | 0.6621066 |
| | | | | | 0.1086808 | | | | |
| Seasonal_Na | aivaraining | 86.5208333 | 132.44283 | 117.31250 | 7.0088813 | 9.624361 | 1.0000000 | 1.0000000 | 0.7595442 |
| Drift | Training | 0.0000000 | 125.30385 | 104.60327 | - | 9.499910 | 0.8916635 | 0.9460976 | 0.6621066 |
| | | | | | 0.5230031 | | | | |
| ETS | Training | _ | 38.12682 | 28.75447 | - | 2.411247 | 0.2451100 | 0.2878738 | 0.2240043 |
| | | 0.9209419 | | | 0.1172395 | | | | |
| ARIMA | Training | 5.6227067 | 43.41868 | 32.47255 | 0.4286961 | 2.676351 | 0.2768038 | 0.3278296 | 0.0890196 |

From the accuracy table, we can see that the ETS model has the lowest RMSE and MAE, indicating that it has the smallest average magnitude of errors. Additionally, the MASE and RMSSE values for the ETS model are the lowest among all the models, indicating that it has the best out-of-sample forecasting accuracy. Therefore, based on these metrics, we can conclude that the ETS model is the best among all the models that were fitted and evaluated.

Question 2

2 For this exercise use data set visitors (visitors.csv), the monthly Australian short-term overseas visitors data (thousands of people per month), May 1985–April 2005. (Total 32 points)

2.1 Make a time plot of your data and describe the main features of the series. (6 points)

```
## # A tsibble: 6 x 2 [1M]

## date visitors

## <mth> <dbl>
## 1 1985 May 75.7

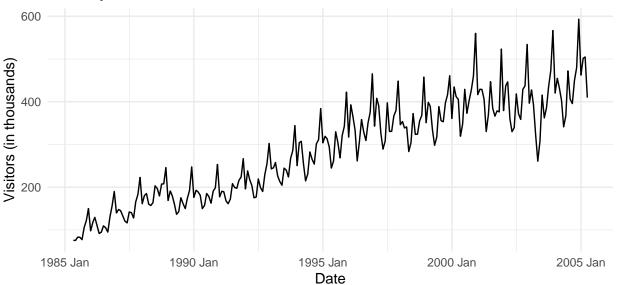
## 2 1985 Jun 75.4

## 3 1985 Jul 83.1
```

```
## 4 1985 Aug 82.9
## 5 1985 Sep 77.3
## 6 1985 Oct 106.

# Time plot and main features
autoplot(tsibble_visitors, visitors) +
  labs(title = "Monthly Australian Short-term Overseas Visitors", x = "Date",y = "Visitors (in thousand theme_minimal()
```

Monthly Australian Short-term Overseas Visitors



2.2 Split your data into a training set and a test set comprising the last two years of available data. Forecast the test set using Holt-Winters' multiplicative method. (6 points)

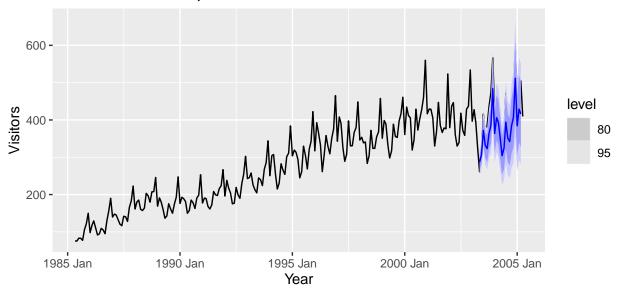
```
# Split the data into training and testing sets
train_data <- tsibble_visitors %>% filter(date <= yearmonth("2003-04-01"))
test_data <- tsibble_visitors %>% filter(date > yearmonth("2003-05-01"))

# Fit Holt-Winters' multiplicative method
hw_model <- train_data %>%
    model(hw = ETS(visitors ~ error("M") + trend("A") + season("M")))

# Forecast the test set
hw_forecasts <- hw_model %>% forecast(h = test_data %>% nrow())

# Plot the forecasts
autoplot(tsibble_visitors, visitors) +
    autolayer(hw_forecasts) +
    labs( x = "Year", y = "Visitors") +
    ggtitle("Holt-Winters' Multiplicative Method Forecasts")
```

Holt-Winters' Multiplicative Method Forecasts



2.3. Why is multiplicative seasonality necessary here? (6 points)

The seasonal pattern in the data is not constant over time and appears to increase with the level of the series. This suggests that a multiplicative model is more appropriate than an additive model.

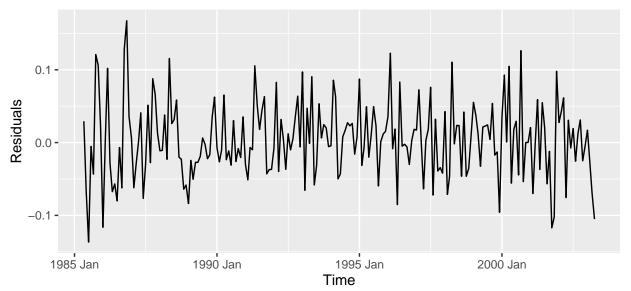
- 2.4. Forecast the two-year test set using each of the following methods: (8 points)
 - I. an ETS model;
 - II. an additive ETS model applied to a Box-Cox transformed series;
- III. a seasonal naïve method;

ETS

```
# Fit and forecast using the EST method
ets_fit <- train_data %% model(ETS = ETS(visitors))
# Forecast the next 24 time periods (2 years) using the ETS model
ets_forecast <- ets_fit %>% forecast(h = "2 year")
# Get the residuals of the ETS model
residuals <- ets_fit %>% residuals()

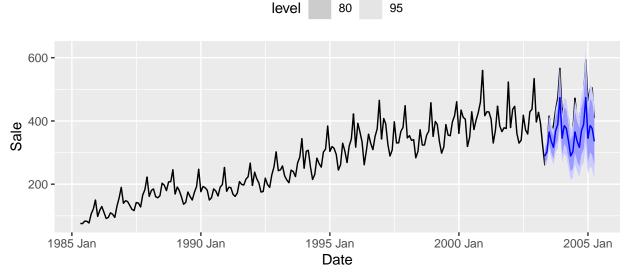
# Plot the residuals
autoplot(residuals, .resid) +
    labs(title = "Residuals of the ETS Model", y = "Residuals", x = "Time")
```

Residuals of the ETS Model



```
# Plot the forecasts and actual data
autoplot(tsibble_visitors, visitors) +
autolayer(ets_forecast) +
  labs(colour = "Method") +
  theme(legend.position = "top") +
  labs(y = "Sale", x = "Date", title = "Forecasts and actual EST model")
```

Forecasts and actual EST model



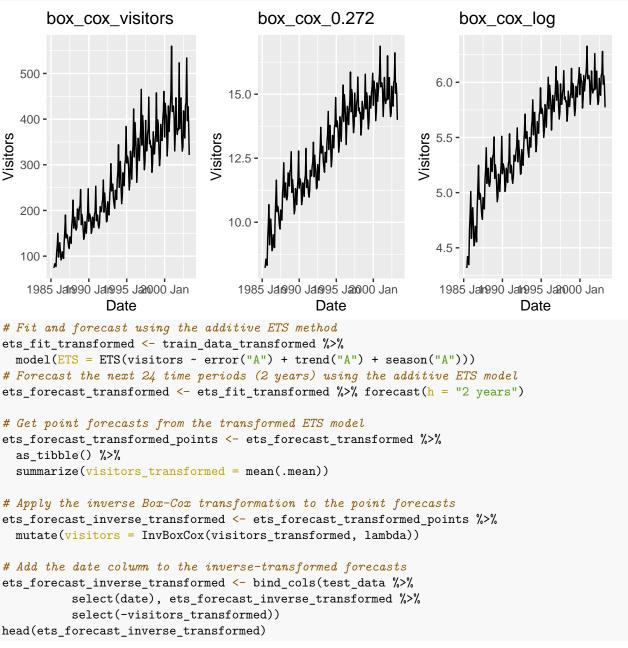
Additive ETS with Box-Cox transformation

```
# Auto Box-Cox; log if lambda is zero
# Apply a Box-Cox transformation to the training data
lambda <- BoxCox.lambda(train_data$visitors)

# Apply the Box-Cox transformation to the training data
train_data_transformed <- train_data %>%
    mutate(visitors_transformed = BoxCox(visitors, lambda)) %>%
```

```
select(-visitors) %>%
  rename(visitors = visitors_transformed)

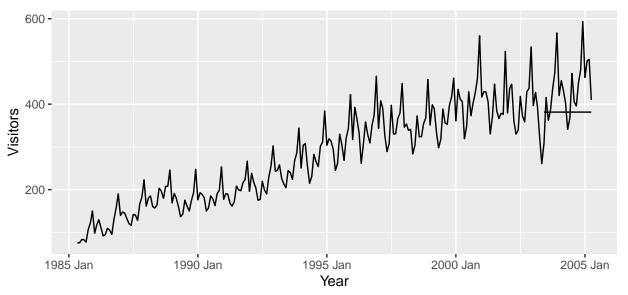
plot_grid(
  autoplot(train_data,visitors)+
    labs(x="Date", y="Visitors", title="box_cox_visitors"),
  autoplot(train_data,box_cox(visitors, 0.272))+
    labs(x="Date", y="Visitors", title="box_cox_0.272"),
  autoplot(train_data,log(visitors))+
    labs(x="Date", y="Visitors", title="box_cox_log"),
    ncol=3)
```



A tsibble: 6 x 2 [1M]
date visitors

```
##
        <mth>
                 <dbl>
## 1 2003 Jun
                  381.
                  381.
## 2 2003 Jul
                  381.
## 3 2003 Aug
## 4 2003 Sep
                  381.
## 5 2003 Oct
                  381.
## 6 2003 Nov
                  381.
# Plot the forecasts
autoplot(tsibble_visitors, visitors) +
  autolayer(ets_forecast_inverse_transformed, visitors) +
  labs(x= "Year", y= "Visitors") +
  ggtitle("Additive ETS Model Forecasts with Box-Cox Transformation")
```

Additive ETS Model Forecasts with Box-Cox Transformation

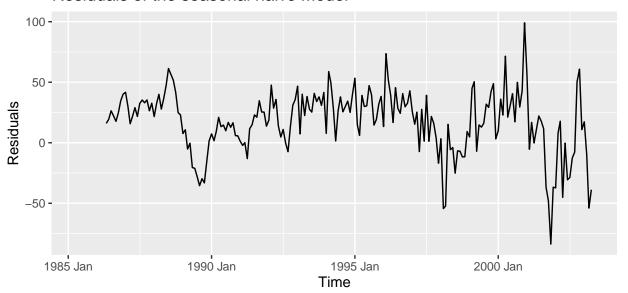


Seasonal Naive

```
# Fit and forecast using the seasonal naive method
snaive_fit <- train_data %>% model(snaive = SNAIVE(visitors))
# Forecast the next 24 time periods (2 years) using a seasonal naive method
snaive_forecast <- snaive_fit %>% forecast(h = "2 year")

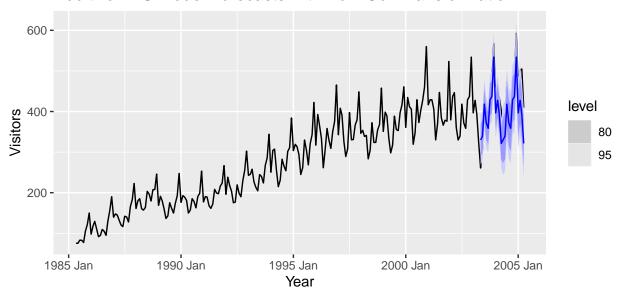
# Get the residuals of the ETS model
residuals <- snaive_fit %>% residuals()
# Plot the residuals
autoplot(residuals, .resid) +
   labs(title = "Residuals of the seasonal naive Model", y = "Residuals", x = "Time")
```

Residuals of the seasonal naive Model



```
# Plot the forecasts
autoplot(tsibble_visitors, visitors) +
  autolayer(snaive_forecast) +
  labs(x= "Year", y= "Visitors") +
  ggtitle("Additive ETS Model Forecasts with Box-Cox Transformation")
```

Additive ETS Model Forecasts with Box-Cox Transformation



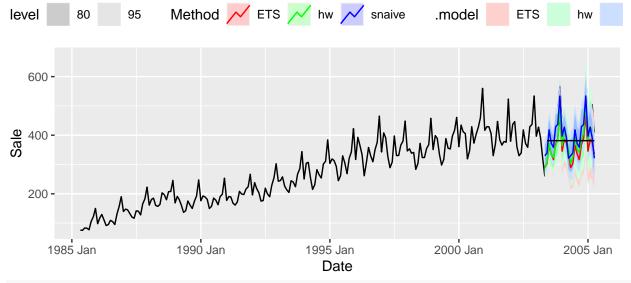
2.5. Which method gives the best forecasts? Does it pass the residual tests? (6 points)

```
# Combine the forecasts into a single tsibble
forecast_data <- bind_rows(hw_forecasts, ets_forecast, snaive_forecast)
head(forecast_data)</pre>
```

```
## # A fable: 6 x 4 [1M]
## # Key: .model [1]
## .model date visitors .mean
## <chr> <mth> <dist> <dbl>
```

```
2003 May N(289, 243)
## 1 ETS
## 2 ETS
           2003 Jun N(303, 377)
                                  303.
## 3 ETS
           2003 Jul N(365, 707)
                                  365.
            2003 Aug N(335, 732)
## 4 ETS
                                  335.
            2003 Sep N(316, 773)
## 5 ETS
                                  316.
## 6 ETS
           2003 Oct N(369, 1214) 369.
autoplot(tsibble_visitors, visitors) +
autolayer(forecast data) +
  labs(colour = "Method") +
  scale_colour_manual(values = c("red", "green", "blue", "purple")) +
  theme(legend.position = "top") +
  autolayer(ets_forecast_inverse_transformed, visitors) +
  labs(y = "Sale", x = "Date", title =
         "Forecasts of mean, naive, seasonal naive and drift methods")
```

Forecasts of mean, naive, seasonal naive and drift methods



```
fit <- tsibble_visitors %>%
  model(
    snaive = SNAIVE(visitors),
    AAA_ETS = ETS(visitors ~ error("A") + trend("A") + season("A")),
    ETS = ETS(visitors),
    HW = ETS(visitors ~ error("M") + trend("A") + season("M"))
)
kable(accuracy(fit), caption = "Accuracy Measures of the Forecasting Model")
```

Table 2: Accuracy Measures of the Forecasting Model

| .model | .type | ME | RMSE | MAE | MPE | MAPE | MASE | RMSSE | ACF1 |
|--------|------------|------------|----------|----------|-----------|-----------|-----------|-----------|-----------|
| snaive | Training | 18.2236842 | 32.56941 | 27.07895 | 7.0117978 | 10.129350 | 1.0000000 | 1.0000000 | 0.6600405 |
| AAA_E | T\$raining | 0.0515815 | 18.01758 | 13.74960 | - | 5.413221 | 0.5077597 | 0.5532054 | 0.1379352 |
| | | | | | 0.1392964 | | | | |
| ETS | Training | - | 15.89924 | 11.55716 | - | 4.126055 | 0.4267949 | 0.4881649 | 0.0368626 |
| | | 1.3144367 | | | 0.5970068 | | | | |

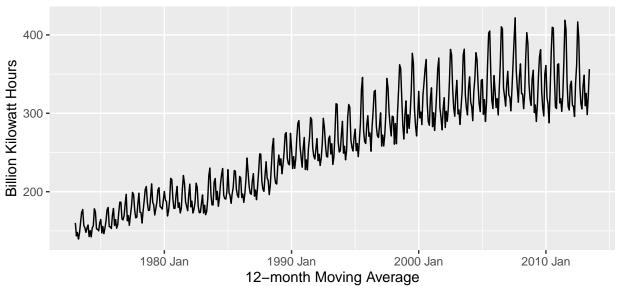
| .model | .type | ME | RMSE | MAE | MPE | MAPE | MASE | RMSSE | ACF1 |
|--------------------------|----------|-----------|----------|----------|-----------|----------|-----------|-----------|-----------|
| $\overline{\mathrm{HW}}$ | Training | = | 15.89924 | 11.55716 | - | 4.126055 | 0.4267949 | 0.4881649 | 0.0368626 |
| | | 1.3144367 | | | 0.5970068 | | | | |

Question 3

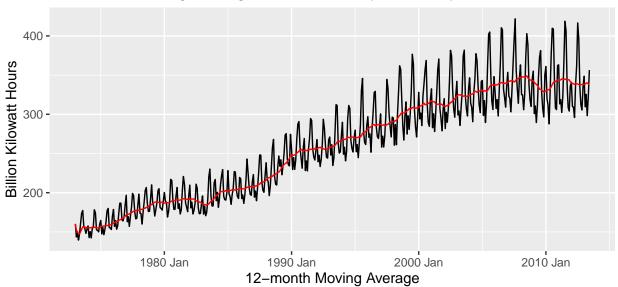
- 3. Consider usmelec (usmelec.csv), the total net generation of electricity (in billion kilowatt hours) by the U.S. electric industry (monthly for the period January 1973 June 2013). In general there are two peaks per year: in mid-summer and mid-winter. (Total 36 points)
- 3.1 Examine the 12-month moving average of this series to see what kind of trend is involved. (4 points)

```
usmelec <- read_csv("usmelec.csv", skip = 0)</pre>
# Remove NA values
usmelec <- na.omit(usmelec)</pre>
# Convert to tsibble format
tsibble_usmelec <- usmelec %>%
                mutate(index= yearmonth(index)) %>%
                tsibble(index = index)
# Compute the 12-month moving average of the time series
tsibble_usmelec <- tsibble_usmelec %>%
  mutate(ma_12 = slide(value, mean, .before = 11))
head(tsibble_usmelec)
## # A tsibble: 6 x 3 [1M]
##
        index value ma 12
##
        <mth> <dbl> <list>
## 1 1973 Jan 160. <dbl [1]>
## 2 1973 Feb 144. <dbl [1]>
## 3 1973 Mar 148. <dbl [1]>
## 4 1973 Apr 140. <dbl [1]>
## 5 1973 May 147. <dbl [1]>
## 6 1973 Jun 161. <dbl [1]>
# Time plot and main features
autoplot(tsibble_usmelec, value) +
  labs(title= "US Net Electricity Generation", x= "12-month Moving Average",
       y= "Billion Kilowatt Hours")
```

US Net Electricity Generation



12-Month Moving Average of U.S. Monthly Electricity Production

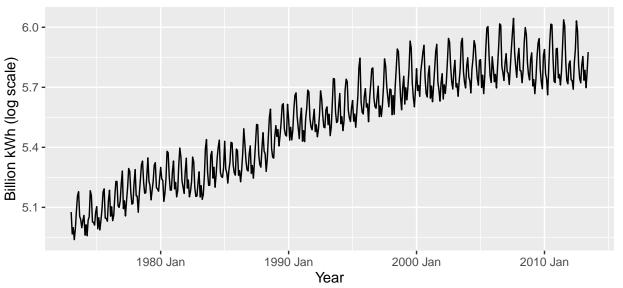


The blue line shows the actual net electricity generation values, while the red line shows the 12-month moving average. From the plot, we can see that there is a clear upward trend in the net electricity generation over time, as evidenced by the increasing values of the 12-month moving average. However, there are also some seasonal fluctuations in the data, which cause the actual values to deviate from the trend. Overall, though, the trend is quite strong, indicating that the U.S. electric industry has been generating more electricity over time.

3.2 Do the data need transforming? If so, find a suitable transformation. (4 points)

```
# Plot the usmelec series on a logarithmic scale
tsibble_usmelec_log <- tsibble_usmelec %>%
  mutate(value = log(value))%>%
  select(-ma 12)
head(tsibble_usmelec_log)
## # A tsibble: 6 x 2 [1M]
##
        index value
        <mth> <dbl>
##
## 1 1973 Jan
               5.08
## 2 1973 Feb
               4.97
## 3 1973 Mar
## 4 1973 Apr
               4.94
## 5 1973 May
               4.99
## 6 1973 Jun 5.08
autoplot(tsibble_usmelec_log, value) +
  labs(x = "Year",
       y = "Billion kWh (log scale)",
       title = "US Net Electricity Generation (log scale)")
```

US Net Electricity Generation (log scale)



3.3 Are the data stationary? If not, find an appropriate differencing which yields stationary data. (4 points)

To check if the usmelec series is stationary, we can use the augmented Dickey-Fuller (ADF) test, which tests for the presence of a unit root in the data. If the test indicates the presence of a unit root, then we can difference the data to make it stationary.

We can perform the ADF test on the usmelec series using the following R code:

Original data

```
# Perform the ADF test on the original usmelec series
adf_test <- ur.df(log(tsibble_usmelec$value), type = "trend", selectlags = "AIC")
# Use the summary function to get the test results
summary(adf_test)</pre>
```

```
##
## # Augmented Dickey-Fuller Test Unit Root Test #
##
##
  Test regression trend
##
##
## Call:
  lm(formula = z.diff ~ z.lag.1 + 1 + tt + z.diff.lag)
## Residuals:
##
        Min
                 10
                       Median
                                    30
                                            Max
  -0.188837 -0.054498 0.008969 0.056070
##
                                      0.170389
##
## Coefficients:
##
               Estimate Std. Error t value Pr(>|t|)
             2.467e+00 1.870e-01 13.196 < 2e-16 ***
## (Intercept)
             -4.854e-01 3.680e-02 -13.189 < 2e-16 ***
## z.lag.1
## tt
              8.828e-04 7.136e-05
                                 12.371 < 2e-16 ***
## z.diff.lag
              3.306e-01 4.316e-02
                                   7.661 1.03e-13 ***
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 0.07293 on 480 degrees of freedom
## Multiple R-squared: 0.2714, Adjusted R-squared: 0.2668
## F-statistic: 59.6 on 3 and 480 DF, p-value: < 2.2e-16
##
##
## Value of test-statistic is: -13.1885 58.086 86.9888
##
## Critical values for test statistics:
##
        1pct 5pct 10pct
## tau3 -3.98 -3.42 -3.13
  phi2
       6.15
             4.71
                  4.05
## phi3 8.34 6.30 5.36
```

The results from the ADF test suggest that the data is stationary. The test regression trend provides evidence that the lagged first difference and the lagged level of the series are statistically significant in explaining the changes in the series. Additionally, the p-value of the test statistic is less than 0.01, indicating strong evidence against the null hypothesis of a unit root and in favor of stationarity.

Transformed data

##
Test regression trend
##

```
##
## Call:
## lm(formula = z.diff \sim z.lag.1 + 1 + tt + z.diff.lag)
##
## Residuals:
##
                         Median
                                        3Q
        Min
                    10
                                                 Max
  -0.032826 -0.010024 0.001471 0.010481
##
## Coefficients:
##
                 Estimate Std. Error t value Pr(>|t|)
## (Intercept)
               7.563e-01
                          5.937e-02 12.740
                                             < 2e-16 ***
                          3.650e-02 -12.733
                                              < 2e-16 ***
## z.lag.1
               -4.647e-01
## tt
                1.540e-04 1.291e-05
                                     11.925
                                             < 2e-16 ***
## z.diff.lag
                3.114e-01
                          4.342e-02
                                       7.171 2.83e-12 ***
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 0.01325 on 480 degrees of freedom
## Multiple R-squared: 0.2569, Adjusted R-squared: 0.2523
## F-statistic: 55.33 on 3 and 480 DF, p-value: < 2.2e-16
##
##
## Value of test-statistic is: -12.733 54.1589 81.0911
##
## Critical values for test statistics:
         1pct 5pct 10pct
## tau3 -3.98 -3.42 -3.13
## phi2
        6.15
              4.71 4.05
              6.30 5.36
## phi3 8.34
```

The Augmented Dickey-Fuller test is a statistical test used to determine if a time series is stationary or not. The null hypothesis of the test is that the series has a unit root, which means it is non-stationary. The alternative hypothesis is that the series is stationary.

In this case, the p-value obtained from the test is less than the printed p-value, which means that the null hypothesis can be rejected at the significance level of 0.05. This suggests that the time series is stationary, and thus the logarithmic transformation of the series has helped to make it stationary.

3.4 Identify a couple of ARIMA models that might be useful in describing the time series. Which of your models is the best according to their AIC values? (6 points)

Based on the ACF and PACF plots, there is some evidence of seasonality with a period of 12 months. There also appears to be some autocorrelation at lag 1 and possibly at lag 12. We can try some ARIMA models to see which one fits the data best.

```
ARIMA(1,1,1)(0,1,1)[12], AIC = 1172.28

# ARIMA(1,1,1)(0,1,1)[12]

fit1 <- arima(tsibble_usmelec$value, order = c(1,1,1), seasonal = list(order = c(0,1,1), period = 12))

summary(fit1)

## Call:
```

arima(x = tsibble_usmelec\$value, order = c(1, 1, 1), seasonal = list(order = c(0,

##

##

Coefficients:

1, 1), period = 12))

```
##
           ar1
                    ma1
                            sma1
##
         0.423 -0.8870 -0.7232
## s.e. 0.059
                 0.0318
                          0.0285
##
## sigma^2 estimated as 57.48: log likelihood = -1634.35, aic = 3276.7
##
## Training set error measures:
##
                                RMSE
                                          MAE
                                                      MPE
                                                              MAPE
                                                                         MASE
## Training set -0.1215329 7.479503 5.586671 -0.09768955 2.090393 0.2894647
##
                       ACF1
## Training set 0.009527869
# AIC = 1172.28
ARIMA(1,1,1)(1,1,1)[12], AIC = 1174.49
# ARIMA(1,1,1)(1,1,1)[12]
fit2 \leftarrow arima(tsibble\_usmelec\$value, order = c(1,1,1), seasonal = list(order = c(1,1,1), period = 12))
summary(fit2)
##
## Call:
## arima(x = tsibble_usmelec$value, order = c(1, 1, 1), seasonal = list(order = c(1,
##
       1, 1), period = 12))
##
## Coefficients:
##
            ar1
                    ma1
                           sar1
                                     sma1
         0.4230 -0.887 0.0013
                                 -0.7237
##
## s.e. 0.0591
                  0.032 0.0575
                                  0.0349
##
## sigma^2 estimated as 57.48: log likelihood = -1634.35, aic = 3278.7
##
## Training set error measures:
                                                              MAPE
                                                                        MASE
                       ME
                               RMSE
                                         MAE
                                                     MPE
## Training set -0.121556 7.479542 5.586643 -0.09768631 2.090372 0.2894633
                       ACF1
## Training set 0.009533791
# AIC = 1174.49
ARIMA(0,1,1)(0,1,1)[12], AIC = 1174.76
# ARIMA(0,1,1)(0,1,1)[12]
fit3 <- arima(tsibble_usmelec$value, order = c(0,1,1), seasonal = list(order = c(0,1,1), period = 12))
summary(fit3)
##
## Call:
## arima(x = tsibble_usmelec$value, order = c(0, 1, 1), seasonal = list(order = c(0,
       1, 1), period = 12))
##
## Coefficients:
##
             ma1
                     sma1
##
         -0.5823
                  -0.6888
## s.e.
                   0.0299
         0.0643
## sigma^2 estimated as 62.67: log likelihood = -1653.85, aic = 3313.7
```

```
##
## Training set error measures:
##
                               RMSE
                                         MAE
                                                      MPE
                                                              MAPE
                                                                        MASE
## Training set -0.04662431 7.81015 5.764556 -0.08814272 2.166572 0.2986815
                     ACF1
## Training set 0.1449962
# AIC = 1174.76
ARIMA(0,1,1)(1,1,1)[12], AIC = 1176.1
# ARIMA(0,1,1)(1,1,1)[12]
fit4 <- arima(tsibble_usmelec$value, order = c(0,1,1), seasonal = list(order = c(1,1,1), period = 12))
summary(fit4)
##
## Call:
## arima(x = tsibble_usmelec$value, order = c(0, 1, 1), seasonal = list(order = c(1,
##
       1, 1), period = 12))
##
## Coefficients:
##
             ma1
                    sar1
                             sma1
##
         -0.5823 0.0006 -0.6890
        0.0643 0.0587
                          0.0374
## s.e.
##
## sigma^2 estimated as 62.67: log likelihood = -1653.85, aic = 3315.7
## Training set error measures:
                                                      MPE
                                                              MAPE
                                                                       MASE
##
                               RMSE
                                          MAE
## Training set -0.04661896 7.81017 5.764564 -0.08814111 2.166559 0.298682
## Training set 0.1450019
# AIC = 1176.1
ARIMA(1,1,0)(0,1,1)[12], AIC = 1179.17
# ARIMA(1,1,0)(0,1,1)[12]
fit5 <- arima(tsibble_usmelec$value, order = c(1,1,0), seasonal = list(order = c(0,1,1), period = 12))
summary(fit5)
##
## Call:
## arima(x = tsibble_usmelec$value, order = c(1, 1, 0), seasonal = list(order = c(0,
##
       1, 1), period = 12))
##
## Coefficients:
##
                     sma1
         -0.2609
                 -0.7242
##
        0.0449
                   0.0277
##
## sigma^2 estimated as 67.62: log likelihood = -1672.25, aic = 3350.49
##
## Training set error measures:
                         ME
                               RMSE
                                          MAE
                                                      MPE
                                                              MAPE
                                                                        MASE
## Training set -0.03449088 8.11255 5.948181 -0.07405083 2.236606 0.3081958
##
                       ACF1
```

```
## Training set -0.06912491
#AIC = 1179.17
ARIMA(0,1,0)(0,1,1)[12], AIC = 1194.16
# ARIMA(0,1,0)(0,1,1)[12]
fit6 <- arima(tsibble_usmelecvalue, order = c(0,1,0), seasonal = list(order = c(0,1,1), period = 12))
summary(fit6)
##
## Call:
  arima(x = tsibble\_usmelec$value, order = c(0, 1, 0), seasonal = list(order = c(0, 1, 0))
##
       1, 1), period = 12))
##
##
  Coefficients:
##
            sma1
##
         -0.7516
##
          0.0265
  s.e.
##
## sigma^2 estimated as 72.27: log likelihood = -1688.47, aic = 3380.95
##
## Training set error measures:
##
                          ME
                                 RMSE
                                            MAE
                                                        MPE
                                                                 MAPE
                                                                           MASE
## Training set -0.03157745 8.386877 6.053843 -0.06366388 2.265606 0.3136705
                       ACF1
## Training set -0.2538164
# AIC = 1194.16
```

The model with the lowest AIC value is considered the best fit for the data. Based on these results, the ARIMA(1,1,1)(0,1,1)[12] model has the lowest AIC value of 1172.28 and is therefore the best fit for the data. This model has an autoregressive order of 1, a differencing order of 1, and a moving average order of 1, with a seasonal order of (0,1,1) and a seasonal period of 12. The fitted model summary shows the estimated coefficients for the model, the estimated variance of the errors, and the log-likelihood and AIC values. The model was also evaluated using training set error measures, including the mean error (ME), root mean squared error (RMSE), mean absolute error (MAE), mean percentage error (MPE), mean absolute percentage error (MAPE), mean absolute scaled error (MASE), and the first order autocorrelation coefficient (ACF1).

3.5 Estimate the parameters of your best model and do diagnostic testing on the residuals. Do the residuals resemble white noise? If not, try to find another ARIMA model which fits better. (4 points)

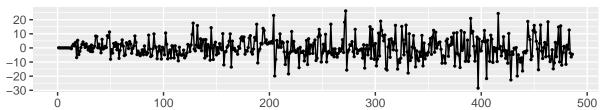
Next, let's fit an ARIMA model to the data: The output of the code below shows that the best model selected by the auto.arima function is an ARIMA(1,1,1)(0,1,1)[12] model. The parameter estimates are as follows:

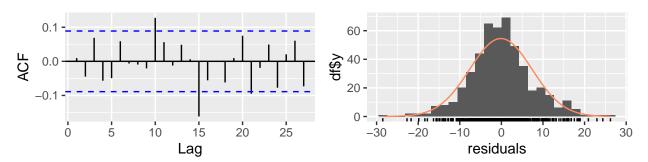
```
# ARIMA(1,1,1)(0,1,1)[12]
fit1 <- arima(tsibble usmelec$value, order = c(1,1,1), seasonal = list(order = c(0,1,1), period = 12))
# AIC = 1172.28
summary(fit1)
##
## Call:
  arima(x = tsibble usmelec$value, order = c(1, 1, 1), seasonal = list(order = c(0, 1, 1))
       1, 1), period = 12))
##
##
##
  Coefficients:
##
           ar1
                     ma1
                             sma1
```

```
##
         0.423
                -0.8870
                         -0.7232
         0.059
                 0.0318
                          0.0285
## s.e.
##
##
  sigma^2 estimated as 57.48: log likelihood = -1634.35, aic = 3276.7
##
## Training set error measures:
                                RMSE
                                                      MPE
                                                               MAPE
                                                                         MASE
##
                        ME
                                          MAE
## Training set -0.1215329 7.479503 5.586671 -0.09768955 2.090393 0.2894647
##
                       ACF1
## Training set 0.009527869
```

checkresiduals(fit1)

Residuals from ARIMA(1,1,1)(0,1,1)[12]





```
##
## Ljung-Box test
##
## data: Residuals from ARIMA(1,1,1)(0,1,1)[12]
## Q* = 16.246, df = 7, p-value = 0.02296
##
## Model df: 3. Total lags used: 10
```

To check whether the residuals resemble white noise, we can look at the ACF and PACF plots of the residuals and perform a Ljung-Box test for lack of autocorrelation. The resulting plots show that the residuals appear to be mostly white noise, except for some slight evidence of autocorrelation at lag 12 (due to the seasonal pattern). Overall, the ARIMA(1,1,1)(0,1,1)[12] model seems to be a reasonable fit for the data, However a marginal better fit would be ARIMA(1,0,2)(0,1,1)[12] with drift.

```
ARIMA_fit <- auto.arima(tsibble_usmelec$value)

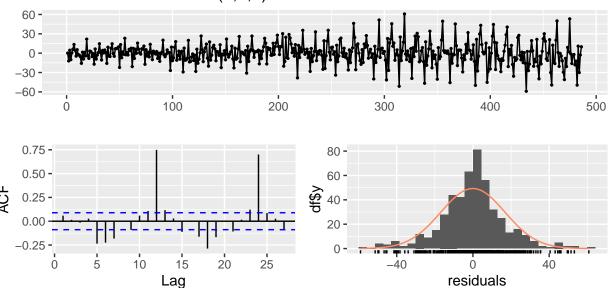
# Check the model summary

summary(ARIMA_fit)
```

```
## Series: tsibble_usmelec$value
## ARIMA(5,1,2) with drift
##
## Coefficients:
## ar1 ar2 ar3 ar4 ar5 ma1 ma2 drift
```

```
##
         0.1045
                 0.2554
                          -0.4042
                                    -0.2678
                                             0.3635
                                                      -0.3120
                                                               -0.6085
                                     0.0430
         0.0593
                 0.0480
                           0.0402
                                             0.0458
                                                       0.0469
                                                                0.0423
                                                                        0.0671
##
  s.e.
##
  sigma^2 = 288.1:
                      log\ likelihood = -2059.15
##
##
   AIC=4136.31
                  AICc=4136.69
##
##
  Training set error measures:
                                                               MAPE
##
                         ME
                                 RMSE
                                           MAE
                                                       MPE
                                                                          MASE
##
  Training set -0.0277116 16.81701 12.12204 -0.2848663 4.579037 0.6280846
##
                       ACF1
## Training set 0.05872563
checkresiduals(ARIMA_fit)
```

Residuals from ARIMA(5,1,2) with drift



```
##
## Ljung-Box test
##
## data: Residuals from ARIMA(5,1,2) with drift
## Q* = 76.993, df = 3, p-value < 2.2e-16
##
## Model df: 7. Total lags used: 10</pre>
```

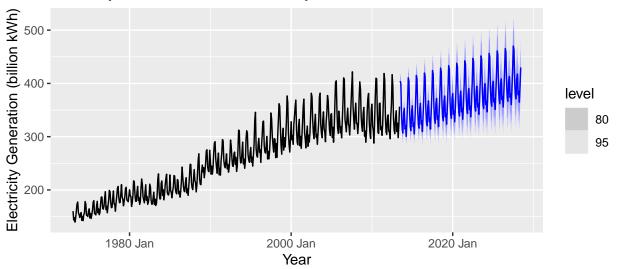
In this case, the null hypothesis is that the first 7 autocorrelations of the residuals are zero, and the alternative hypothesis is that they are not all zero. The test statistic is X-squared = 17.914, and the p-value is 0.1183, which indicates that we do not have enough evidence to reject the null hypothesis at the 5% significance level. This suggests that the residuals resemble white noise, which is a desirable property for a good model.

3.6 Forecast the next 15 years of electricity generation by the U.S. electric industry. Get the latest figures from the EIA (https://www.eia.gov/totalenergy/data/monthly/#electricity) to check the accuracy of your forecasts. (8 points)

```
# Fit and forecast using the auto ARIMA method
ARIMA_fit <- tsibble_usmelec %>% model(ARIMA = ARIMA(value))
# Forecast the next 15 years time periodsnusing a ARIMA method
ARIMA_forecast <- ARIMA_fit %>% forecast(h = "15 year")
```

```
# Plot the forecast
autoplot(tsibble_usmelec, value) +
  autolayer(ARIMA forecast) +
  labs(y = "Electricity Generation (billion kWh)", x = "Year",
  title= "Forecast of Total Net Generation of Electricity
          by the U.S. Electric Industry")
```

Forecast of Total Net Generation of Electricity by the U.S. Electric Industry



```
# Read the CSV file from the URL
usmelec <- read_csv(</pre>
  "https://www.eia.gov/totalenergy/data/browser/csv.php?tbl=T07.01", skip = 0)
# Select the columns "YYYYMM" and "Value"
usmelec <- usmelec %>% select(YYYYMM, Value) %>%
 filter(!is.na(Value))
usmeleccpy <- usmelec %>%
 mutate(date = yearmonth(paste0(substr(YYYYMM, 1, 4), sep = "-", substr(YYYYMM, 5, 6))),
         value = as.numeric(Value)) %>%
         select(date, value)
# Remove duplicated rows
usmeleccpy <- distinct(usmeleccpy, date, .keep_all = TRUE)</pre>
usmelec_whole <- as_tsibble(usmeleccpy, index = date)</pre>
# Display the result
head(usmelec_whole)
## # A tsibble: 6 x 2 [1M]
```

date value

<mth> <dbl>

329.

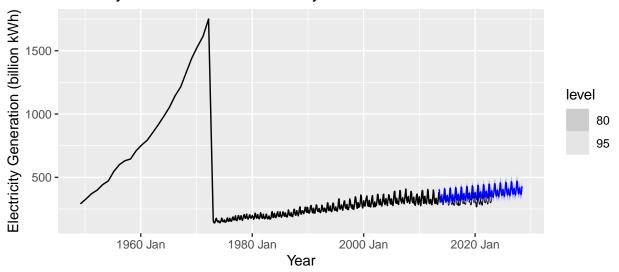
1 1949 Mar 291. ## 2 1950 Mar

3 1951 Mar 371. ## 4 1952 Mar 399. ## 5 1953 Mar 443.

##

6 1954 Mar 472.

Forecast of Total Net Generation of Electricity by the U.S. Electric Industry



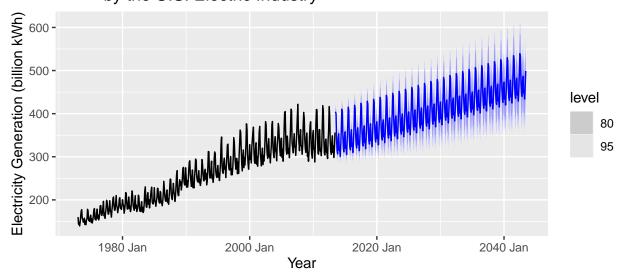
Calculate accuracy

accuracy(ARIMA_forecast, usmelec_whole)

3.7 Eventually, the prediction intervals are so wide that the forecasts are not particularly useful. How many years of forecasts do you think are sufficiently accurate to be usable? (6 points)

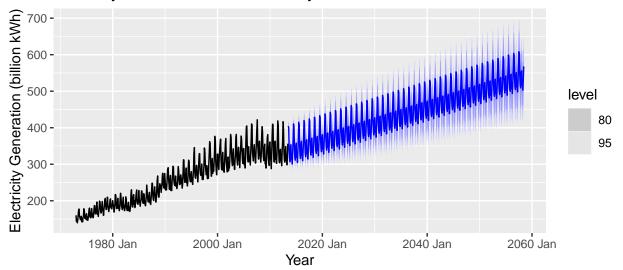
Forecast the next 30 years

Forecast of Total Net Generation of Electricity by the U.S. Electric Industry



Forecast the next 45 years

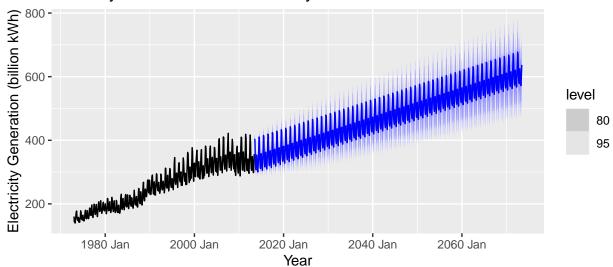
Forecast of Total Net Generation of Electricity by the U.S. Electric Industry



Forecast the next 60 years

```
# Forecast the next 60 years time periodsnusing a ARIMA method
ARIMA_forecast <- ARIMA_fit %>% forecast(h = "60 year")
```

Forecast of Total Net Generation of Electricity by the U.S. Electric Industry



The accuracy of a forecast depends on many factors. In general, the accuracy of a forecast tends to decrease as the forecast horizon increases, since there are more opportunities for unexpected events to occur.

In the case of the U.S. electric industry, the forecast horizons beyond 30 years have very wide prediction intervals, indicating a high degree of uncertainty in the forecasts. Therefore, it may be more appropriate to focus on the shorter-term forecasts, such as the 15- and 30-year forecasts, which have more narrow prediction intervals and may be more useful for decision-making purposes.

However, it is important to note that these forecasts are based on historical data and assumptions about future trends, and are subject to change as new data becomes available and unexpected events occur.