

参数估计

$$X \Rightarrow (X_1, \dots, X_n) \Rightarrow (x_1, \dots, x_n)$$

参数估计的思想

总体 X 分布已知, 但含未知参数 θ , 对 θ 进行估计

方式一: 点估计 $\hat{\theta} = \varphi(X_1, \dots, X_n)$ — 估计量

$\hat{\theta} = \varphi(x_1, \dots, x_n)$ — 估计值

方式二: 区间估计, $\forall \alpha, \underline{\theta} = \varphi_1(X_1, \dots, X_n), \bar{\theta} = \varphi_2(X_1, \dots, X_n)$

$$P\{\underline{\theta} < \theta < \bar{\theta}\} = 1 - \alpha$$

$(\underline{\theta}, \bar{\theta})$ 为 θ 的置信度为 $1 - \alpha$ 的置信区间

参数的点估计

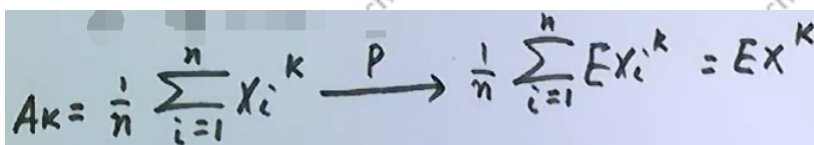
矩估计法

(一) 矩估计法

1. 理论基础 — 大数定律

$$X \Rightarrow (X_1, \dots, X_n)$$

$$A_k = \frac{1}{n} \sum_{i=1}^n X_i^k \text{ 依概率收敛于 } \frac{1}{n} \sum_{i=1}^n EX_i^k = EX^k$$


$$A_k = \frac{1}{n} \sum_{i=1}^n X_i^k \xrightarrow{P} \frac{1}{n} \sum_{i=1}^n EX_i^k = EX^k$$

2. 矩估计法步骤: (X 分布已知, $\begin{cases} X \text{ 离散} \\ X \text{ 连续} \end{cases}$)

case1. 含 θ :

若 $EX = ?$ 含 θ , 令 $\bar{X} = EX \Rightarrow \hat{\theta} = ?$

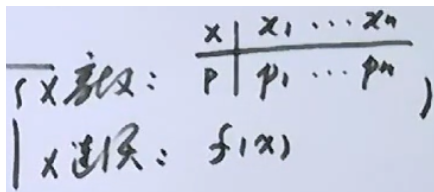
若 $EX = ?$ 不含 θ , $EX^2 = ?$ 令 $EX^2 = A_2 = \frac{1}{n} \sum_{i=1}^n X_i^2 \Rightarrow \hat{\theta} = ?$

$(\frac{1}{n} \sum_{i=1}^n X_i^2 \text{ 依概率收敛于 } EX^2)$

case2. 含 θ_1, θ_2 :

1. 求 $EX = ? EX^2 = ?$

2. 令 $\begin{cases} EX = \bar{X} \\ EX^2 = A_2 = \frac{1}{n} \sum_{i=1}^n X_i^2 \end{cases} \Rightarrow \hat{\theta}_1 = ?, \hat{\theta}_2 = ?$


$$\begin{array}{l} X \text{ 离散: } \begin{array}{c|c} x & x_1 \dots x_n \\ \hline p & p_1 \dots p_n \end{array} \\ X \text{ 连续: } f(x) \end{array}$$

设总体 $X \sim \begin{pmatrix} 0 & 1 & 2 \\ \theta & 2\theta & 1-3\theta \end{pmatrix}$, (X_1, X_2, \dots, X_n) 为来自总体 X 的简单随机样本, 求参数 θ 的矩估计量.

$$EX = 2 - 4\theta$$

$$\text{令 } EX = \bar{X} \Rightarrow \hat{\theta} = \frac{2 - \bar{X}}{4}$$

设总体 $X \sim N(\mu, \sigma^2)$, (X_1, X_2, \dots, X_n) 为来自总体 X 的简单随机样本,

(1) 设 $\mu = 2$, 求参数 σ^2 的矩估计量; (2) 设 μ 未知, 求参数 σ^2 的矩估计量.

$$\textcircled{1} EX = 2, EX^2 = DX + (EX)^2 = \sigma^2 + 4$$

$$\text{令 } EX^2 = A_2 \Rightarrow \hat{\sigma}^2 = A_2 - 4$$

$$\textcircled{2} EX = \mu, EX^2 = \sigma^2 + \mu^2$$

$$\text{令 } \begin{cases} EX = \bar{X} \\ EX^2 = A_2 \end{cases} \Rightarrow \hat{\mu} = \bar{X}, \hat{\sigma}^2 = A_2 - \mu^2$$

$$\therefore \hat{\sigma}^2 = A_2 - \bar{X}^2$$

最大似然估计法

X 总体分布已知

$$(X_1, \dots, X_n) \Rightarrow (x_1, \dots, x_n)$$

1. 总体 X 离散型:

1° 分布律

$$2^\circ L = P\{X_1 = x_1\} \cdots P\{X_n = x_n\} = P\{X = X_1\} \cdots P\{X = X_n\}$$

3° $\ln L = \dots$

4° case1. 含 θ :

$$\text{令 } \frac{d}{d\theta} \ln L = \dots = 0 \Rightarrow \hat{\theta} = ?$$

case2. 含 θ_1, θ_2 :

$$\text{令 } \begin{cases} \frac{\partial}{\partial \theta_1} \ln L = \dots = 0 \\ \frac{\partial}{\partial \theta_2} \ln L = \dots = 0 \end{cases} \Rightarrow \hat{\theta}_1 = ? \hat{\theta}_2 = ?$$

2. $X \sim f(x)$:

$$1^\circ L = f(x_1) \cdots f(x_n)$$

2° $\ln L = \dots$

3° case1. 含 θ :

$$\text{令 } \frac{d}{d\theta} \ln L = \dots = 0 \Rightarrow \hat{\theta} = ?$$

case2. 含 θ_1, θ_2 :

$$\text{令 } \begin{cases} \frac{\partial}{\partial \theta_1} \ln L = \dots = 0 \\ \frac{\partial}{\partial \theta_2} \ln L = \dots = 0 \end{cases} \Rightarrow \hat{\theta}_1 = ? \hat{\theta}_2 = ?$$

X	P
...	...
...	...
...	...

设总体 $X \sim \begin{pmatrix} -1 & 0 & 2 \\ \theta & \theta & 1-2\theta \end{pmatrix}$, (X_1, X_2, X_3, X_4) 为来自总体的简单随机样本, 其观察值为 $(0, 2, -1, 0)$, 求参数 θ 的最大似然估计值.

X	-1	0	2
P	θ	θ	$1-2\theta$

$$1. L = P\{X=0\}P\{X=2\}P\{X=-1\}P\{X=0\} = \theta^3(1-2\theta)$$

$$2. \ln L = 3 \ln \theta + \ln(1-2\theta)$$

$$3. \text{令 } \frac{d}{d\theta} \ln L = \frac{3}{\theta} - \frac{2}{1-2\theta} = 0 \Rightarrow 3-8\theta = 0$$

$$\Rightarrow \hat{\theta} = \frac{3}{8}$$

设总体 $X \sim N(\mu, \sigma^2)$, (X_1, X_2, \dots, X_n) 为来自总体 X 的简单随机样本.

(1) 若 $\mu = 1$, 求参数 σ^2 的最大似然估计量;

(2) 若 μ 未知, 求参数 μ, σ^2 的最大似然估计量.

$$\textcircled{1} X \sim f(x) = \frac{1}{\sqrt{2\pi}\sigma} e^{-\frac{(x-1)^2}{2\sigma^2}}$$

$$1. L = f(x_1) \cdots f(x_n) = (2\pi)^{-\frac{n}{2}} e^{-\frac{1}{2\sigma^2} \sum_{i=1}^n (x_i-1)^2}$$

$$2. \ln L = -\frac{n}{2} \ln 2\pi - \frac{n}{2} \ln \sigma^2 - \frac{1}{2\sigma^2} \sum_{i=1}^n (x_i-1)^2$$

$$3. \frac{d}{d\sigma^2} \ln L = -\frac{n}{2\sigma^2} + \frac{1}{2\sigma^4} \sum_{i=1}^n (x_i-1)^2 = 0$$

$$\Rightarrow \hat{\sigma}^2 = \frac{1}{n} \sum_{i=1}^n (x_i-1)^2 \text{ (最大似然估计法)}$$

$\therefore \sigma^2$ 的最大似然估计量为

$$\hat{\sigma}^2 = \frac{1}{n} \sum_{i=1}^n (X_i-1)^2$$

$$\textcircled{2} L = f(x_1) \cdots f(x_n) = (2\pi)^{-\frac{n}{2}} (\sigma^2)^{-\frac{n}{2}} e^{-\frac{1}{2\sigma^2} \sum_{i=1}^n (x_i-\mu)^2}$$

$$\ln L = -\frac{n}{2} \ln 2\pi - \frac{n}{2} \ln \sigma^2 - \frac{1}{2\sigma^2} \sum_{i=1}^n (x_i-\mu)^2$$

$$\text{由 } \begin{cases} \frac{\partial}{\partial \mu} \ln L = -\frac{1}{\sigma^2} \sum_{i=1}^n (x_i-\mu) = 0 \\ \frac{\partial}{\partial \sigma^2} \ln L = -\frac{n}{2\sigma^2} + \frac{1}{2\sigma^4} \sum_{i=1}^n (x_i-\mu)^2 = 0 \end{cases}$$

$$\Rightarrow \hat{\mu} = \bar{x}, \hat{\sigma}^2 = \frac{1}{n} \sum_{i=1}^n (x_i-\bar{x})^2 \text{ (估计法)}$$

$\therefore \mu, \sigma^2$ 的最大似然估计量为

$$\hat{\mu} = \bar{X}, \hat{\sigma}^2 = \frac{1}{n} \sum_{i=1}^n (X_i-\bar{X})^2$$

设总体 X 的概率密度为

$$f(x) = \begin{cases} \theta x^{\theta-1}, & 0 < x < 1, \\ 0, & \text{其他} \end{cases}$$

(X_1, X_2, \dots, X_n) 为来自总体 X 的简单随机样本.

(1) 求参数 θ 的矩估计量; (2) 求参数 θ 的最大似然估计量.

$$\textcircled{1} EX = \int_0^1 \theta x^\theta dx = \frac{\theta}{\theta+1}$$

$$\text{令 } EX = \bar{X} \Rightarrow 1 + \frac{1}{\theta} = \frac{1}{\bar{X}} \Rightarrow \hat{\theta} = \frac{\bar{X}}{1 - \bar{X}}$$

$$\textcircled{2} L = f(x_1) \cdots f(x_n) = \theta^n (x_1 \cdots x_n)^{\theta-1} (0 < x_i < 1, 1 \leq i \leq n)$$

$$\ln L = n \ln \theta + (\theta - 1) \sum_{i=1}^n \ln x_i$$

$$\text{令 } \frac{d}{d\theta} \ln L = \frac{n}{\theta} + \sum_{i=1}^n \ln x_i = 0 \Rightarrow \hat{\theta} = -\frac{n}{\sum_{i=1}^n \ln x_i}$$

$\therefore \theta$ 的最大似然估计量为

$$\hat{\theta} = -\frac{n}{\sum_{i=1}^n \ln x_i}$$

$$\text{总体 } X \sim f(x) = \begin{cases} 2e^{-2(x-\theta)}, & x > \theta \\ 0, & x \leq \theta \end{cases}$$

$$(X_1, X_2, X_3) \Rightarrow (x_1, x_2, x_3)$$

$\textcircled{1} \theta$ 的矩估计量 $\hat{\theta}$; $\textcircled{2} \theta$ 的最大似然估计量 $\hat{\theta}$;

$\textcircled{3} \textcircled{2}$ 中 $\hat{\theta}$, 求 $E\hat{\theta}$.

$$\begin{aligned} \textcircled{1} EX &= \int_{\theta}^{+\infty} 2xe^{-2(x-\theta)} dx = \frac{1}{2} \int_{\theta}^{+\infty} [2(x-\theta) + 2\theta] e^{-2(x-\theta)} d[2(x-\theta)] \\ &= \frac{1}{2} \int_0^{+\infty} (x+2\theta) e^{-x} dx = \frac{1}{2} (1+2\theta) = \theta + \frac{1}{2} \end{aligned}$$

$$\text{令 } EX = \bar{X} \Rightarrow \hat{\theta} = \bar{X} - \frac{1}{2}$$

$$\textcircled{2} L = f(x_1)f(x_2)f(x_3) = 2^3 \cdot e^{-2(x_1+x_2+x_3)+6\theta}$$

$$(\theta < x_i, i = 1, 2, 3)$$

$$\ln L = 3 \ln 2 - 2(x_1 + x_2 + x_3) + 6\theta$$

$$\text{令 } \frac{d}{d\theta} \ln L = 6 > 0 \Rightarrow L(\theta) \uparrow$$

$$\Rightarrow \hat{\theta} = \min\{x_1, x_2, x_3\}$$

θ 的最大似然估计量为

$$\hat{\theta} = \min\{x_1, x_2, x_3\}$$

$$\textcircled{3} 1. \text{ 总体 } X \text{ 的分布函数 } F(x) = \int_{-\infty}^x f(t) dt$$

$$x < \theta : F(x) = 0$$

$$x \geq \theta : F(x) = \int_{\theta}^x 2e^{-2(t-\theta)} dt = 1 - e^{-2(x-\theta)}$$

$$F(x) = \begin{cases} 1 - e^{-2(x-\theta)}, & x \geq \theta \\ 0, & x < \theta \end{cases}$$

$$2. F_{\hat{\theta}}(x) = P\{\hat{\theta} \leq x\} = 1 - P\{\hat{\theta} < x\}$$

$$= 1 - P\{x_1 > x\}P\{x_2 > x\}P\{x_3 > x\} = 1 - P\{X > x\}$$

$$= 1 - [1 - F(x)]^3 = \begin{cases} 1 - e^{-6(x-\theta)}, & x \geq \theta \\ 0, & x < \theta \end{cases}$$

$$3. f_{\hat{\theta}}(x) = \begin{cases} 6e^{-6(x-\theta)}, & x > \theta \\ 0, & x \leq \theta \end{cases}$$

$$4. E\hat{\theta} = \int_{\theta}^{+\infty} 6xe^{-6(x-\theta)} dx = \frac{1}{6} \int_{\theta}^{+\infty} [6(x-\theta) + 6\theta] e^{-6(x-\theta)} d[6(x-\theta)]$$

$$= \frac{1}{6} \int_0^{+\infty} (x+6\theta) e^{-x} dx = \frac{1}{6} (1+6\theta) = \theta + \frac{1}{6}$$

估计量的评价标准

1.无偏性

$X \Rightarrow (X_1, \cdots, X_n)$ (含 θ)

$\hat{\theta} = \varphi(X_1, \cdots, X_n)$ 为 θ 的估计量

若 $E\hat{\theta} = \theta$, 称 $\hat{\theta} = \varphi(X_1, \cdots, X_n)$ 为 θ 的无偏估计量.