

Model Predictive Path Integral Control: A Tutorial

Xiong Zeng

January 20, 2026

1 Introduction

2 Motivation and Optimality for Single Iteration

3 Convergence of MPPI for Quadratic Programming

Introduction

Model predictive path integral (**MPPI**) control is a powerful and popular controller method in robotics and control, because

- it can handle **nonconvex** and **non-smooth** optimization problems,
- and it can be computed in **real-time** with GPU due to its **parallel computation** nature.

¹Williams, et al. Information theoretic MPC for model-based reinforcement learning. ICRA 2017.

²Xue, et al. Full-order sampling-based mpc for torque-level locomotion control via diffusion-style annealing. ICRA 2025.

³Pan, et al. pider: Scalable physics-informed dexterous retargeting. ArXiv 2025.

⁴Halder et al. Trajectory Planning with Signal Temporal Logic Costs using Deterministic Path Integral Optimization. Arxiv 2025.

Introduction

Model predictive path integral (**MPPI**) control is a powerful and popular controller method in robotics and control, because

- it can handle **nonconvex** and **non-smooth** optimization problems,
- and it can be computed in **real-time** with GPU due to its **parallel computation** nature.

Many applications:

- Aggressive driving ¹
- Legged robot control²
- Sim2real dexterous retargeting ³
- Hard control problems in my research (MPPI > Policy gradient > SDP)
- Control with signal temporal logical constraints ⁴

¹Williams, et al. Information theoretic MPC for model-based reinforcement learning. ICRA 2017.

²Xue, et al. Full-order sampling-based mpc for torque-level locomotion control via diffusion-style annealing. ICRA 2025.

³Pan, et al. pider: Scalable physics-informed dexterous retargeting. ArXiv 2025.

⁴Halder et al. Trajectory Planning with Signal Temporal Logic Costs using Deterministic Path Integral Optimization. Arxiv 2025.

Introduction

Consider the optimization problem

$$\min_{\mathbf{u}} J(\mathbf{u}) \in \mathbb{R}_{\geq 0}. \quad (1)$$

Introduction

Consider the optimization problem

$$\min_{\mathbf{u}} J(\mathbf{u}) \in \mathbb{R}_{\geq 0}. \quad (1)$$

The iteration formula of MPPI is

$$\mathbf{u}_{out} = \frac{\sum_{i=1}^N \mathbf{u}_i \exp(-J(\mathbf{u}_i)/\lambda)}{\sum_{i=1}^N \exp(-J(\mathbf{u}_i)/\lambda)}, \quad (2)$$

where λ is a hyperparameter (called temperature from simulated annealing) and $\{\mathbf{u}_i\}_{i=1}^N$ are i.i.d. samples from the following Gaussian distribution $\mathcal{N}(\mathbf{u}_{in}, \boldsymbol{\Sigma})$, with the current controller \mathbf{u}_{in} and $\boldsymbol{\Sigma} \succ 0$.

Derive Single MPPI Iteration Based on KL Control⁵

Let $q(\mathbf{u})$ be a given probability density (for \mathbf{u}_{in}) on \mathbb{R}^k . Let $J : \mathbb{R}^k \rightarrow \mathbb{R}$ be measurable and $\lambda > 0$. We consider the optimization problem over probability densities $p(\mathbf{u})$:

$$\begin{aligned} \min_p & \int p(\mathbf{u}) J(\mathbf{u}) d\mathbf{u} + \lambda \int p(\mathbf{u}) \log \frac{p(\mathbf{u})}{q(\mathbf{u})} d\mathbf{u} \\ \text{s.t. } & \int p(\mathbf{u}) d\mathbf{u} = 1. \end{aligned} \tag{3}$$

⁵Evangelos & Todorov. Relative entropy and free energy dualities: Connections to path integral and kl control. CDC 2012.

Derive Single MPPI Iteration Based on KL Control⁵

Let $q(\mathbf{u})$ be a given probability density (for \mathbf{u}_{in}) on \mathbb{R}^k . Let $J : \mathbb{R}^k \rightarrow \mathbb{R}$ be measurable and $\lambda > 0$. We consider the optimization problem over probability densities $p(\mathbf{u})$:

$$\begin{aligned} \min_p \quad & \int p(\mathbf{u}) J(\mathbf{u}) d\mathbf{u} + \lambda \int p(\mathbf{u}) \log \frac{p(\mathbf{u})}{q(\mathbf{u})} d\mathbf{u} \\ \text{s.t.} \quad & \int p(\mathbf{u}) d\mathbf{u} = 1. \end{aligned} \tag{3}$$

Define

$$p^*(\mathbf{u}) := \frac{1}{\mathbb{E}_q \left[\exp \left(-\frac{1}{\lambda} J(\mathbf{u}) \right) \right]} q(\mathbf{u}) \exp \left(-\frac{1}{\lambda} J(\mathbf{u}) \right). \tag{4}$$

⁵Evangelos & Todorov. Relative entropy and free energy dualities: Connections to path integral and kl control. CDC 2012.

Derive Single MPPI Iteration Based on KL Control⁵

Let $q(\mathbf{u})$ be a given probability density (for \mathbf{u}_{in}) on \mathbb{R}^k . Let $J : \mathbb{R}^k \rightarrow \mathbb{R}$ be measurable and $\lambda > 0$. We consider the optimization problem over probability densities $p(\mathbf{u})$:

$$\begin{aligned} \min_p \quad & \int p(\mathbf{u}) J(\mathbf{u}) d\mathbf{u} + \lambda \int p(\mathbf{u}) \log \frac{p(\mathbf{u})}{q(\mathbf{u})} d\mathbf{u} \\ \text{s.t.} \quad & \int p(\mathbf{u}) d\mathbf{u} = 1. \end{aligned} \tag{3}$$

Define

$$p^*(\mathbf{u}) := \frac{1}{\mathbb{E}_q \left[\exp \left(-\frac{1}{\lambda} J(\mathbf{u}) \right) \right]} q(\mathbf{u}) \exp \left(-\frac{1}{\lambda} J(\mathbf{u}) \right). \tag{4}$$

Goal: prove that p^* is the unique minimizer of (3).

⁵ Evangelos & Todorov. Relative entropy and free energy dualities: Connections to path integral and kl control. CDC 2012.

Variation Calculus

Proof. It can be proved that (3) is a strictly convex infinite-dimensional optimization problem. Then, the feasible stationary point of its Lagrange function is a unique minimizer. Define Lagrangian

$$\mathcal{L}(p, \eta) = \int p(\mathbf{u}) J(\mathbf{u}) d\mathbf{u} + \lambda \int p(\mathbf{u}) \log \frac{p(\mathbf{u})}{q(\mathbf{u})} d\mathbf{u} + \eta \left(\int p(\mathbf{u}) d\mathbf{u} - 1 \right), \quad (5)$$

where $\eta \in \mathbb{R}$ is a Lagrange multiplier.

Variation Calculus

Proof. It can be proved that (3) is a strictly convex infinite-dimensional optimization problem. Then, the feasible stationary point of its Lagrange function is a unique minimizer. Define Lagrangian

$$\mathcal{L}(p, \eta) = \int p(\mathbf{u}) J(\mathbf{u}) d\mathbf{u} + \lambda \int p(\mathbf{u}) \log \frac{p(\mathbf{u})}{q(\mathbf{u})} d\mathbf{u} + \eta \left(\int p(\mathbf{u}) d\mathbf{u} - 1 \right), \quad (5)$$

where $\eta \in \mathbb{R}$ is a Lagrange multiplier. We compute stationarity with respect to $p(\mathbf{u})$. Let $p_\epsilon = p + \epsilon h$ with $\int h(\mathbf{u}) d\mathbf{u} = 0$.

Using $\frac{d}{dx}(x \log x) = 1 + \log x$, we obtain

$$\frac{d}{d\epsilon} \Big|_{\epsilon=0} \mathcal{L}(p_\epsilon, \eta) = \int h(\mathbf{u}) \left[J(\mathbf{u}) + \lambda(1 + \log p(\mathbf{u}) - \log q(\mathbf{u})) + \eta \right] d\mathbf{u}.$$

Variation Calculus

Proof. It can be proved that (3) is a strictly convex infinite-dimensional optimization problem. Then, the feasible stationary point of its Lagrange function is a unique minimizer. Define Lagrangian

$$\mathcal{L}(p, \eta) = \int p(\mathbf{u}) J(\mathbf{u}) d\mathbf{u} + \lambda \int p(\mathbf{u}) \log \frac{p(\mathbf{u})}{q(\mathbf{u})} d\mathbf{u} + \eta \left(\int p(\mathbf{u}) d\mathbf{u} - 1 \right), \quad (5)$$

where $\eta \in \mathbb{R}$ is a Lagrange multiplier. We compute stationarity with respect to $p(\mathbf{u})$. Let $p_\epsilon = p + \epsilon h$ with $\int h(\mathbf{u}) d\mathbf{u} = 0$.

Using $\frac{d}{dx}(x \log x) = 1 + \log x$, we obtain

$$\frac{d}{d\epsilon} \Big|_{\epsilon=0} \mathcal{L}(p_\epsilon, \eta) = \int h(\mathbf{u}) \left[J(\mathbf{u}) + \lambda(1 + \log p(\mathbf{u}) - \log q(\mathbf{u})) + \eta \right] d\mathbf{u}.$$

For this to vanish for all admissible h , the bracket must be constant:

$$J(\mathbf{u}) + \lambda(1 + \log p(\mathbf{u}) - \log q(\mathbf{u})) = c. \quad (6)$$

Solve the Stationarity Condition

Rearranging,

$$\log p(\mathbf{u}) = \log q(\mathbf{u}) - \frac{1}{\lambda} J(\mathbf{u}) + \frac{c - \lambda}{\lambda}.$$

Solve the Stationarity Condition

Rearranging,

$$\log p(\mathbf{u}) = \log q(\mathbf{u}) - \frac{1}{\lambda} J(\mathbf{u}) + \frac{c - \lambda}{\lambda}.$$

Exponentiating,

$$p(\mathbf{u}) = e^{\frac{c-\lambda}{\lambda}} q(\mathbf{u}) \exp\left(-\frac{1}{\lambda} J(\mathbf{u})\right).$$

Solve the Stationarity Condition

Rearranging,

$$\log p(\mathbf{u}) = \log q(\mathbf{u}) - \frac{1}{\lambda} J(\mathbf{u}) + \frac{c - \lambda}{\lambda}.$$

Exponentiating,

$$p(\mathbf{u}) = e^{\frac{c-\lambda}{\lambda}} q(\mathbf{u}) \exp\left(-\frac{1}{\lambda} J(\mathbf{u})\right).$$

Imposing normalization $\int p = 1$ gives

$$e^{\frac{c-\lambda}{\lambda}} = \frac{1}{\mathbb{E}_q\left[\exp\left(-\frac{1}{\lambda} J(\mathbf{u})\right)\right]},$$

Solve the Stationarity Condition

Rearranging,

$$\log p(\mathbf{u}) = \log q(\mathbf{u}) - \frac{1}{\lambda} J(\mathbf{u}) + \frac{c - \lambda}{\lambda}.$$

Exponentiating,

$$p(\mathbf{u}) = e^{\frac{c-\lambda}{\lambda}} q(\mathbf{u}) \exp\left(-\frac{1}{\lambda} J(\mathbf{u})\right).$$

Imposing normalization $\int p = 1$ gives

$$e^{\frac{c-\lambda}{\lambda}} = \frac{1}{\mathbb{E}_q\left[\exp\left(-\frac{1}{\lambda} J(\mathbf{u})\right)\right]},$$

so

$$p(\mathbf{u}) = \frac{1}{\mathbb{E}_q\left[\exp\left(-\frac{1}{\lambda} J(\mathbf{u})\right)\right]} q(\mathbf{u}) \exp\left(-\frac{1}{\lambda} J(\mathbf{u})\right) = p^*(\mathbf{u}).$$

From Optimal Distribution to Optimal Controller

KL-control yields an *optimal distribution* $p^*(\mathbf{u})$, but controller implementation must output a deterministic value $\bar{\mathbf{u}}$. Adopt the posterior squared-loss criterion

$$\bar{\mathbf{u}}^* \in \arg \min_{\bar{\mathbf{u}} \in \mathbb{R}^{mH}} \mathbb{E}_{p^*} [\|\mathbf{u} - \bar{\mathbf{u}}\|_2^2]. \quad (7)$$

From Optimal Distribution to Optimal Controller

KL-control yields an *optimal distribution* $p^*(\mathbf{u})$, but controller implementation must output a deterministic value $\bar{\mathbf{u}}$. Adopt the posterior squared-loss criterion

$$\bar{\mathbf{u}}^* \in \arg \min_{\bar{\mathbf{u}} \in \mathbb{R}^{mH}} \mathbb{E}_{p^*} [\|\mathbf{u} - \bar{\mathbf{u}}\|_2^2]. \quad (7)$$

Expand the risk:

$$\mathbb{E}_{p^*} [\|\mathbf{u} - \bar{\mathbf{u}}\|_2^2] = \mathbb{E}_{p^*} [\|\mathbf{u}\|_2^2] - 2\bar{\mathbf{u}}^\top \mathbb{E}_{p^*} [\mathbf{u}] + \|\bar{\mathbf{u}}\|_2^2. \quad (8)$$

From Optimal Distribution to Optimal Controller

KL-control yields an *optimal distribution* $p^*(\mathbf{u})$, but controller implementation must output a deterministic value $\bar{\mathbf{u}}$. Adopt the posterior squared-loss criterion

$$\bar{\mathbf{u}}^* \in \arg \min_{\bar{\mathbf{u}} \in \mathbb{R}^{mH}} \mathbb{E}_{p^*} [\|\mathbf{u} - \bar{\mathbf{u}}\|_2^2]. \quad (7)$$

Expand the risk:

$$\mathbb{E}_{p^*} [\|\mathbf{u} - \bar{\mathbf{u}}\|_2^2] = \mathbb{E}_{p^*} [\|\mathbf{u}\|_2^2] - 2\bar{\mathbf{u}}^\top \mathbb{E}_{p^*} [\mathbf{u}] + \|\bar{\mathbf{u}}\|_2^2. \quad (8)$$

Differentiate w.r.t. $\bar{\mathbf{u}}$:

$$\nabla_{\bar{\mathbf{u}}} = -2\mathbb{E}_{p^*} [\mathbf{u}] + 2\bar{\mathbf{u}}. \quad (9)$$

Setting $\nabla_{\bar{\mathbf{u}}} = 0$ gives

$$\bar{\mathbf{u}}^* = \mathbb{E}_{p^*} [\mathbf{u}].$$

Since the Hessian is $2I \succ 0$, this minimizer is unique.

Convergence of MPPI for Quadratic Programming

Consider a QP

$$\min_{\mathbf{u}} J(\mathbf{u}) = \mathbf{u}^\top \mathbf{D} \mathbf{u} + \mathbf{u}^\top \mathbf{d}, \quad (10)$$

with $\mathbf{D} \succ 0$. Draw N i.i.d. samples $\mathbf{u}_i \sim q = \mathcal{N}(\mathbf{u}_{in}, \Sigma)$ with $\Sigma \succ 0$.

⁶Yi et al. CoVO-MPC: Theoretical Analysis of Sampling-based MPC and Optimal Covariance Design. L4DC 2024.

Convergence of MPPI for Quadratic Programming

Consider a QP

$$\min_{\mathbf{u}} J(\mathbf{u}) = \mathbf{u}^\top \mathbf{D} \mathbf{u} + \mathbf{u}^\top \mathbf{d}, \quad (10)$$

with $\mathbf{D} \succ 0$. Draw N i.i.d. samples $\mathbf{u}_i \sim q = \mathcal{N}(\mathbf{u}_{in}, \Sigma)$ with $\Sigma \succ 0$. Consider the MPPI update

$$\hat{\mathbf{u}}_{out} = \frac{\sum_{i=1}^N \mathbf{u}_i \exp(-J(\mathbf{u}_i)/\lambda)}{\sum_{i=1}^N \exp(-J(\mathbf{u}_i)/\lambda)} \text{ and } \mathbf{u}_{out} = \frac{\mathbb{E}_q \left[\mathbf{u} \exp\left(-\frac{1}{\lambda} J(\mathbf{u})\right) \right]}{\mathbb{E}_q \left[\exp\left(-\frac{1}{\lambda} J(\mathbf{u})\right) \right]}. \quad (11)$$

⁶Yi et al. CoVO-MPC: Theoretical Analysis of Sampling-based MPC and Optimal Covariance Design. L4DC 2024.

Convergence of MPPI for Quadratic Programming

Consider a QP

$$\min_{\mathbf{u}} J(\mathbf{u}) = \mathbf{u}^\top \mathbf{D} \mathbf{u} + \mathbf{u}^\top \mathbf{d}, \quad (10)$$

with $\mathbf{D} \succ 0$. Draw N i.i.d. samples $\mathbf{u}_i \sim q = \mathcal{N}(\mathbf{u}_{in}, \Sigma)$ with $\Sigma \succ 0$. Consider the MPPI update

$$\hat{\mathbf{u}}_{out} = \frac{\sum_{i=1}^N \mathbf{u}_i \exp(-J(\mathbf{u}_i)/\lambda)}{\sum_{i=1}^N \exp(-J(\mathbf{u}_i)/\lambda)} \text{ and } \mathbf{u}_{out} = \frac{\mathbb{E}_q \left[\mathbf{u} \exp\left(-\frac{1}{\lambda} J(\mathbf{u})\right) \right]}{\mathbb{E}_q \left[\exp\left(-\frac{1}{\lambda} J(\mathbf{u})\right) \right]}. \quad (11)$$

Theorem 1.⁶ Consider the minimizer $\mathbf{u}^* = \frac{-1}{2} \mathbf{d} \mathbf{D}^{-1}$ of (10). We have

$$\hat{\mathbf{u}}_{out} \xrightarrow{P} \mathbf{u}_{out}, \quad (12)$$

and

$$\mathbf{u}_{out} - \mathbf{u}^* \xrightarrow{P} (\mathbf{I} + 2\lambda \Sigma \mathbf{D})^{-1} (\mathbf{u}_{in} - \mathbf{u}^*). \quad (13)$$

⁶Yi et al. CoVO-MPC: Theoretical Analysis of Sampling-based MPC and Optimal Covariance Design. L4DC 2024.

Step 1: Write p^* exactly

Proof. Define

$$Z := \mathbb{E}_q \left[\exp \left(-\frac{1}{\lambda} J(\mathbf{u}) \right) \right] = \int_{\mathbb{R}^k} q(\mathbf{u}) \exp \left(-\frac{1}{\lambda} J(\mathbf{u}) \right) d\mathbf{u}. \quad (14)$$

Then the optimal distribution satisfies

$$p^*(\mathbf{u}) = \frac{1}{Z} q(\mathbf{u}) \exp \left(-\frac{1}{\lambda} J(\mathbf{u}) \right), \quad \mathbf{u}_{out} = \mathbb{E}_{p^*}[\mathbf{u}]. \quad (15)$$

Step 1: Write p^* exactly

Proof. Define

$$Z := \mathbb{E}_q \left[\exp \left(-\frac{1}{\lambda} J(\mathbf{u}) \right) \right] = \int_{\mathbb{R}^k} q(\mathbf{u}) \exp \left(-\frac{1}{\lambda} J(\mathbf{u}) \right) d\mathbf{u}. \quad (14)$$

Then the optimal distribution satisfies

$$p^*(\mathbf{u}) = \frac{1}{Z} q(\mathbf{u}) \exp \left(-\frac{1}{\lambda} J(\mathbf{u}) \right), \quad \mathbf{u}_{out} = \mathbb{E}_{p^*}[\mathbf{u}]. \quad (15)$$

Because $q(\mathbf{u}) = \mathcal{N}(\mathbf{u}_{in}, \boldsymbol{\Sigma})$, then

$$q(\mathbf{u}) = \frac{1}{(2\pi)^{k/2} |\boldsymbol{\Sigma}|^{1/2}} \exp \left(-\frac{1}{2} (\mathbf{u} - \mathbf{u}_{in})^\top \boldsymbol{\Sigma}^{-1} (\mathbf{u} - \mathbf{u}_{in}) \right). \quad (16)$$

Step 1: Write p^* exactly

Proof. Define

$$Z := \mathbb{E}_q \left[\exp \left(-\frac{1}{\lambda} J(\mathbf{u}) \right) \right] = \int_{\mathbb{R}^k} q(\mathbf{u}) \exp \left(-\frac{1}{\lambda} J(\mathbf{u}) \right) d\mathbf{u}. \quad (14)$$

Then the optimal distribution satisfies

$$p^*(\mathbf{u}) = \frac{1}{Z} q(\mathbf{u}) \exp \left(-\frac{1}{\lambda} J(\mathbf{u}) \right), \quad \mathbf{u}_{out} = \mathbb{E}_{p^*}[\mathbf{u}]. \quad (15)$$

Because $q(\mathbf{u}) = \mathcal{N}(\mathbf{u}_{in}, \boldsymbol{\Sigma})$, then

$$q(\mathbf{u}) = \frac{1}{(2\pi)^{k/2} |\boldsymbol{\Sigma}|^{1/2}} \exp \left(-\frac{1}{2} (\mathbf{u} - \mathbf{u}_{in})^\top \boldsymbol{\Sigma}^{-1} (\mathbf{u} - \mathbf{u}_{in}) \right). \quad (16)$$

Using $J(\mathbf{u}) = \mathbf{u}^\top \mathbf{D} \mathbf{u} + \mathbf{u}^\top \mathbf{d}$, we obtain the exact product

$$q(\mathbf{u}) \exp \left(-\frac{1}{\lambda} J(\mathbf{u}) \right) = \frac{1}{(2\pi)^{k/2} |\boldsymbol{\Sigma}|^{1/2}} \exp \left(-\frac{1}{2} \mathbf{u}^\top \mathbf{A} \mathbf{u} + \mathbf{u}^\top \mathbf{b} - c_0 \right), \quad (17)$$

where

$$\mathbf{A} := \boldsymbol{\Sigma}^{-1} + \frac{2}{\lambda} \mathbf{D}, \quad \mathbf{b} := \boldsymbol{\Sigma}^{-1} \mathbf{u}_{in} - \frac{1}{\lambda} \mathbf{d}, \quad c_0 := \frac{1}{2} \mathbf{u}_{in}^\top \boldsymbol{\Sigma}^{-1} \mathbf{u}_{in}. \quad (18)$$

Step 2: Complete the square and normalize

Complete the square with $\mu := \mathbf{A}^{-1}\mathbf{b}$:

$$-\frac{1}{2}\mathbf{u}^\top \mathbf{A}\mathbf{u} + \mathbf{u}^\top \mathbf{b} = -\frac{1}{2}(\mathbf{u} - \mu)^\top \mathbf{A}(\mathbf{u} - \mu) + \frac{1}{2}\mathbf{b}^\top \mathbf{A}^{-1}\mathbf{b}. \quad (19)$$

Step 2: Complete the square and normalize

Complete the square with $\mu := \mathbf{A}^{-1}\mathbf{b}$:

$$-\frac{1}{2}\mathbf{u}^\top \mathbf{A}\mathbf{u} + \mathbf{u}^\top \mathbf{b} = -\frac{1}{2}(\mathbf{u} - \mu)^\top \mathbf{A}(\mathbf{u} - \mu) + \frac{1}{2}\mathbf{b}^\top \mathbf{A}^{-1}\mathbf{b}. \quad (19)$$

Hence

$$p^*(\mathbf{u}) = \frac{1}{Z} \cdot \frac{1}{(2\pi)^{k/2} |\Sigma|^{1/2}} \exp\left(\frac{1}{2}\mathbf{b}^\top \mathbf{A}^{-1}\mathbf{b} - c_0\right) \exp\left(-\frac{1}{2}(\mathbf{u} - \mu)^\top \mathbf{A}(\mathbf{u} - \mu)\right). \quad (20)$$

Step 2: Complete the square and normalize

Complete the square with $\mu := \mathbf{A}^{-1}\mathbf{b}$:

$$-\frac{1}{2}\mathbf{u}^\top \mathbf{A}\mathbf{u} + \mathbf{u}^\top \mathbf{b} = -\frac{1}{2}(\mathbf{u} - \mu)^\top \mathbf{A}(\mathbf{u} - \mu) + \frac{1}{2}\mathbf{b}^\top \mathbf{A}^{-1}\mathbf{b}. \quad (19)$$

Hence

$$p^*(\mathbf{u}) = \frac{1}{Z} \cdot \frac{1}{(2\pi)^{k/2} |\boldsymbol{\Sigma}|^{1/2}} \exp\left(\frac{1}{2}\mathbf{b}^\top \mathbf{A}^{-1}\mathbf{b} - c_0\right) \exp\left(-\frac{1}{2}(\mathbf{u} - \mu)^\top \mathbf{A}(\mathbf{u} - \mu)\right). \quad (20)$$

After a long derivation,

$$Z = \int_{\mathbb{R}^k} q(\mathbf{u}) \exp\left(-\frac{1}{\lambda} J(\mathbf{u})\right) d\mathbf{u} = \frac{|\mathbf{A}^{-1}|^{1/2}}{|\boldsymbol{\Sigma}|^{1/2}} \exp\left(\frac{1}{2}\mathbf{b}^\top \mathbf{A}^{-1}\mathbf{b} - \frac{1}{2}\mathbf{u}_{in}^\top \boldsymbol{\Sigma}^{-1}\mathbf{u}_{in}\right). \quad (21)$$

Step 2: Complete the square and normalize

Complete the square with $\mu := \mathbf{A}^{-1}\mathbf{b}$:

$$-\frac{1}{2}\mathbf{u}^\top \mathbf{A}\mathbf{u} + \mathbf{u}^\top \mathbf{b} = -\frac{1}{2}(\mathbf{u} - \mu)^\top \mathbf{A}(\mathbf{u} - \mu) + \frac{1}{2}\mathbf{b}^\top \mathbf{A}^{-1}\mathbf{b}. \quad (19)$$

Hence

$$p^*(\mathbf{u}) = \frac{1}{Z} \cdot \frac{1}{(2\pi)^{k/2} |\Sigma|^{1/2}} \exp\left(\frac{1}{2}\mathbf{b}^\top \mathbf{A}^{-1}\mathbf{b} - c_0\right) \exp\left(-\frac{1}{2}(\mathbf{u} - \mu)^\top \mathbf{A}(\mathbf{u} - \mu)\right). \quad (20)$$

After a long derivation,

$$Z = \int_{\mathbb{R}^k} q(\mathbf{u}) \exp\left(-\frac{1}{\lambda} J(\mathbf{u})\right) d\mathbf{u} = \frac{|\mathbf{A}^{-1}|^{1/2}}{|\Sigma|^{1/2}} \exp\left(\frac{1}{2}\mathbf{b}^\top \mathbf{A}^{-1}\mathbf{b} - \frac{1}{2}\mathbf{u}_{in}^\top \Sigma^{-1}\mathbf{u}_{in}\right). \quad (21)$$

Substituting this Z back to (20) yields the fully normalized Gaussian form

$$p^*(\mathbf{u}) = \frac{|\mathbf{A}|^{1/2}}{(2\pi)^{k/2}} \exp\left(-\frac{1}{2}(\mathbf{u} - \mu)^\top \mathbf{A}(\mathbf{u} - \mu)\right) = \mathcal{N}(\mu, \mathbf{A}^{-1}). \quad (22)$$

In particular,

$$\mathbf{u}_{out} = \mathbb{E}_{p^*}[\mathbf{u}] = \mu = \left(\Sigma^{-1} + \frac{2}{\lambda}\mathbf{D}\right)^{-1} \left(\Sigma^{-1}\mathbf{u}_{in} - \frac{1}{\lambda}\mathbf{d}\right). \quad (23)$$

Step 3: Express the mean as a contraction toward \mathbf{u}^*

Recall

$$\mathbf{u}_{out} = \left(\boldsymbol{\Sigma}^{-1} + \frac{2}{\lambda} \mathbf{D} \right)^{-1} \left(\boldsymbol{\Sigma}^{-1} \mathbf{u}_{in} - \frac{1}{\lambda} \mathbf{d} \right). \quad (24)$$

Step 3: Express the mean as a contraction toward \mathbf{u}^*

Recall

$$\mathbf{u}_{out} = \left(\boldsymbol{\Sigma}^{-1} + \frac{2}{\lambda} \mathbf{D} \right)^{-1} \left(\boldsymbol{\Sigma}^{-1} \mathbf{u}_{in} - \frac{1}{\lambda} \mathbf{d} \right). \quad (24)$$

Since $\nabla J(\mathbf{u}^*) = 0$ yeilds $\mathbf{d} = -2\mathbf{D}\mathbf{u}^*$. Then $-\frac{1}{\lambda}\mathbf{d} = \frac{2}{\lambda}\mathbf{D}\mathbf{u}^*$.

Step 3: Express the mean as a contraction toward \mathbf{u}^*

Recall

$$\mathbf{u}_{out} = \left(\boldsymbol{\Sigma}^{-1} + \frac{2}{\lambda} \mathbf{D} \right)^{-1} \left(\boldsymbol{\Sigma}^{-1} \mathbf{u}_{in} - \frac{1}{\lambda} \mathbf{d} \right). \quad (24)$$

Since $\nabla J(\mathbf{u}^*) = 0$ yeilds $\mathbf{d} = -2\mathbf{D}\mathbf{u}^*$. Then $-\frac{1}{\lambda}\mathbf{d} = \frac{2}{\lambda}\mathbf{D}\mathbf{u}^*$. Substitute and regroup:

$$\mathbf{u}_{out} = \left(\boldsymbol{\Sigma}^{-1} + \frac{2}{\lambda} \mathbf{D} \right)^{-1} \left(\boldsymbol{\Sigma}^{-1} \mathbf{u}_{in} + \frac{2}{\lambda} \mathbf{D}\mathbf{u}^* \right) \quad (25)$$

$$= (\lambda \mathbf{I} + 2\boldsymbol{\Sigma} \mathbf{D})^{-1} \left(\lambda \mathbf{u}_{in} + 2\boldsymbol{\Sigma} \mathbf{D}\mathbf{u}^* \right) \quad (26)$$

$$= (\lambda \mathbf{I} + 2\boldsymbol{\Sigma} \mathbf{D})^{-1} \left(\lambda \mathbf{u}_{in} + 2\boldsymbol{\Sigma} \mathbf{D}\mathbf{u}_{in} - 2\boldsymbol{\Sigma} \mathbf{D}\mathbf{u}_{in} + 2\boldsymbol{\Sigma} \mathbf{D}\mathbf{u}^* \right) \quad (27)$$

$$= \mathbf{u}_{in} + (\lambda \mathbf{I} + 2\boldsymbol{\Sigma} \mathbf{D})^{-1} 2\boldsymbol{\Sigma} \mathbf{D}(\mathbf{u}^* - \mathbf{u}_{in}). \quad (28)$$

Step 3: Express the mean as a contraction toward \mathbf{u}^*

Recall

$$\mathbf{u}_{out} = \left(\boldsymbol{\Sigma}^{-1} + \frac{2}{\lambda} \mathbf{D} \right)^{-1} \left(\boldsymbol{\Sigma}^{-1} \mathbf{u}_{in} - \frac{1}{\lambda} \mathbf{d} \right). \quad (24)$$

Since $\nabla J(\mathbf{u}^*) = 0$ yeilds $\mathbf{d} = -2\mathbf{D}\mathbf{u}^*$. Then $-\frac{1}{\lambda}\mathbf{d} = \frac{2}{\lambda}\mathbf{D}\mathbf{u}^*$. Substitute and regroup:

$$\mathbf{u}_{out} = \left(\boldsymbol{\Sigma}^{-1} + \frac{2}{\lambda} \mathbf{D} \right)^{-1} \left(\boldsymbol{\Sigma}^{-1} \mathbf{u}_{in} + \frac{2}{\lambda} \mathbf{D}\mathbf{u}^* \right) \quad (25)$$

$$= (\lambda \mathbf{I} + 2\boldsymbol{\Sigma}\mathbf{D})^{-1} \left(\lambda \mathbf{u}_{in} + 2\boldsymbol{\Sigma}\mathbf{D}\mathbf{u}^* \right) \quad (26)$$

$$= (\lambda \mathbf{I} + 2\boldsymbol{\Sigma}\mathbf{D})^{-1} \left(\lambda \mathbf{u}_{in} + 2\boldsymbol{\Sigma}\mathbf{D}\mathbf{u}_{in} - 2\boldsymbol{\Sigma}\mathbf{D}\mathbf{u}_{in} + 2\boldsymbol{\Sigma}\mathbf{D}\mathbf{u}^* \right) \quad (27)$$

$$= \mathbf{u}_{in} + (\lambda \mathbf{I} + 2\boldsymbol{\Sigma}\mathbf{D})^{-1} 2\boldsymbol{\Sigma}\mathbf{D}(\mathbf{u}^* - \mathbf{u}_{in}). \quad (28)$$

Therefore,

$$\mathbf{u}_{out} - \mathbf{u}^* = (\mathbf{I} + 2\lambda\boldsymbol{\Sigma}\mathbf{D})^{-1} (\mathbf{u}_{in} - \mathbf{u}^*).$$

Step 4: Convergence in probability and Slutsky

Let the MPPI estimator be

$$\mathbf{u}_{out}^{(N)} = \frac{\frac{1}{N} \sum_{i=1}^N \mathbf{u}_i w_i}{\frac{1}{N} \sum_{i=1}^N w_i}, \quad w_i := \exp\left(-\frac{1}{\lambda} J(\mathbf{u}_i)\right), \quad \mathbf{u}_i \stackrel{\text{i.i.d.}}{\sim} q. \quad (29)$$

Step 4: Convergence in probability and Slutsky

Let the MPPI estimator be

$$\mathbf{u}_{out}^{(N)} = \frac{\frac{1}{N} \sum_{i=1}^N \mathbf{u}_i w_i}{\frac{1}{N} \sum_{i=1}^N w_i}, \quad w_i := \exp\left(-\frac{1}{\lambda} J(\mathbf{u}_i)\right), \quad \mathbf{u}_i \stackrel{\text{i.i.d.}}{\sim} q. \quad (29)$$

Then by the weak LLN,

$$\frac{1}{N} \sum_{i=1}^N w_i \xrightarrow{p} \mathbb{E}_q[w], \quad \frac{1}{N} \sum_{i=1}^N \mathbf{u}_i w_i \xrightarrow{p} \mathbb{E}_q[\mathbf{u}w]. \quad (30)$$

Step 4: Convergence in probability and Slutsky

Let the MPPI estimator be

$$\mathbf{u}_{out}^{(N)} = \frac{\frac{1}{N} \sum_{i=1}^N \mathbf{u}_i w_i}{\frac{1}{N} \sum_{i=1}^N w_i}, \quad w_i := \exp\left(-\frac{1}{\lambda} J(\mathbf{u}_i)\right), \quad \mathbf{u}_i \stackrel{\text{i.i.d.}}{\sim} q. \quad (29)$$

Then by the weak LLN,

$$\frac{1}{N} \sum_{i=1}^N w_i \xrightarrow{P} \mathbb{E}_q[w], \quad \frac{1}{N} \sum_{i=1}^N \mathbf{u}_i w_i \xrightarrow{P} \mathbb{E}_q[\mathbf{u}w]. \quad (30)$$

Then, Slutsky's theorem yields

$$\hat{\mathbf{u}}_{out} \xrightarrow{P} \frac{\mathbb{E}_q[\mathbf{u}w]}{\mathbb{E}_q[w]} = \mathbb{E}_{p^*}[\mathbf{u}] = \mathbf{u}_{out}. \quad (31)$$