

Portfolio Assignment 1, Methods 3, 2021, autumn semester

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Portfolio Assignment 1: Using mixed effects modelling to model hierarchical data

In this assignment we will be investigating the *politeness* dataset of Winter and Grawunder (2012) and apply basic methods of multilevel modelling.

```
#this is the dataset we will be exploring in this exercise
politeness <- read.csv('politeness.csv') ## read in data
```

Exercise 1 - describing the dataset and making some initial plots

1. Describe the dataset, such that someone who happened upon this dataset could understand the variables and what they contain

Exercise 1, part 1– (EH)

The politeness dataset contains the data obtained from the study of Korean formal and informal speech (<https://doi.org/10.1016/j.wocn.2012.08.006>) which investigated the fundamental frequency of male and female participants' speech in a variety of formal and informal scenarios.

The following table describes the variables in the dataset:

Variable	Description
subject	participant ID
gender	participant's gender
scenario	the experimental scenario from 1 to 7 such as "asking a favour"
attitude	either 'inf' for informal stimuli or 'pol' for formal stimuli
total_duration	duration of participant's response in seconds
f0mn	mean fundamental frequency (f0) of the participant's speech
hiss_count	number of times the participants made a noisy breath intake

Remark: The `gender`, `scenario` and `attitude` variables should be encoded as factors as they show a categorical function withing this dataset . In addition, these variables have non-unique values across participants, and are not ordered.

```
#Encoding some of the variables as factors (gender, attitude, and scenario)
```

```
politeness$attitude <- as.factor(politeness$attitude)
politeness$gender <- as.factor(politeness$gender)
politeness$scenario <- as.factor(politeness$scenario)
```

Exercise 1, part 2–(EH)

2. Create a new data frame that just contains the subject *F1* and run two linear models; one that expresses *f0mn* as dependent on *scenario* as an integer; and one that expresses *f0mn* as dependent on *scenario* encoded as a factor
 - i. Include the model matrices, X from the General Linear Model, for these two models in your report and describe the different interpretations of *scenario* that these entail
 - ii. Which coding of *scenario*, as a factor or not, is more fitting?

```
# Create a subset dataframe for subject F1 only
pf1 <- politeness[politeness$subject == "F1", ]
pf1
```

##	subject	gender	scenario	attitude	total_duration	f0mn	hiss_count
## 1	F1	F	1	pol	18.392	214.6	2
## 2	F1	F	1	inf	13.551	210.9	0
## 3	F1	F	2	pol	5.217	284.7	0
## 4	F1	F	2	inf	4.247	265.6	0
## 5	F1	F	3	pol	6.791	210.6	0
## 6	F1	F	3	inf	4.126	285.6	0
## 7	F1	F	4	pol	6.244	251.5	1
## 8	F1	F	4	inf	3.245	281.5	0
## 9	F1	F	5	pol	5.625	229.6	1
## 10	F1	F	5	inf	3.950	250.5	0
## 11	F1	F	6	pol	28.508	181.1	1
## 12	F1	F	6	inf	55.159	229.3	0
## 13	F1	F	7	inf	60.309	219.8	2
## 14	F1	F	7	pol	40.825	175.8	0

```
# make model predicting f0mn by scenario (integer)
m1<- lm(f0mn ~ as.integer(scenario), data = pf1)
```

```
# get model matrix
mm1 <- model.matrix(m1)
```

```
# make model predicting f0mn by scenario (factor)
m2 <- lm(f0mn ~ as.factor(scenario), data = pf1)
```

```
# get model matrix
mm2 <- model.matrix(m2)
```

Here is the model using "scenario" encoded as an integer

```
summary(m1)
```

```
##
## Call:
## lm(formula = f0mn ~ as.integer(scenario), data = pf1)
##
## Residuals:
##      Min       1Q   Median       3Q      Max
## -44.836 -36.807   6.686  20.918  46.421
##
## Coefficients:
##              Estimate Std. Error t value Pr(>|t|)
## (Intercept)    262.621     20.616   12.738 2.48e-08 ***
## as.integer(scenario)   -6.886      4.610   -1.494   0.161
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 34.5 on 12 degrees of freedom
## Multiple R-squared:  0.1568, Adjusted R-squared:  0.0865
## F-statistic: 2.231 on 1 and 12 DF,  p-value: 0.1611
```

```
mm1
```

```
##      (Intercept) as.integer(scenario)
## 1              1              1
## 2              1              1
## 3              1              2
## 4              1              2
## 5              1              3
## 6              1              3
## 7              1              4
## 8              1              4
## 9              1              5
## 10             1              5
## 11             1              6
## 12             1              6
## 13             1              7
## 14             1              7
## attr(,"assign")
## [1] 0 1
```

And here is the model using “scenario” encoded as a factor

```
summary(m2)
```

```
##
## Call:
## lm(formula = f0mn ~ as.factor(scenario), data = pf1)
##
## Residuals:
##      Min       1Q   Median       3Q      Max
## -37.50 -13.86   0.00  13.86  37.50
##
## Coefficients:
##              Estimate Std. Error t value Pr(>|t|)
## (Intercept)      212.75      20.35  10.453  1.6e-05 ***
## as.factor(scenario)2      62.40      28.78   2.168   0.0668 .
## as.factor(scenario)3      35.35      28.78   1.228   0.2591
## as.factor(scenario)4      53.75      28.78   1.867   0.1041
## as.factor(scenario)5      27.30      28.78   0.948   0.3745
## as.factor(scenario)6      -7.55      28.78  -0.262   0.8006
## as.factor(scenario)7     -14.95      28.78  -0.519   0.6195
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 28.78 on 7 degrees of freedom
## Multiple R-squared:  0.6576, Adjusted R-squared:  0.364
## F-statistic: 2.24 on 6 and 7 DF, p-value: 0.1576
```

mm2

```
##      (Intercept) as.factor(scenario)2 as.factor(scenario)3 as.factor(scenario)4
## 1             1             0             0             0
## 2             1             0             0             0
## 3             1             1             0             0
## 4             1             1             0             0
## 5             1             0             1             0
## 6             1             0             1             0
## 7             1             0             0             1
## 8             1             0             0             1
## 9             1             0             0             0
## 10            1             0             0             0
## 11            1             0             0             0
## 12            1             0             0             0
## 13            1             0             0             0
## 14            1             0             0             0
##      as.factor(scenario)5 as.factor(scenario)6 as.factor(scenario)7
## 1             0             0             0
## 2             0             0             0
## 3             0             0             0
## 4             0             0             0
## 5             0             0             0
## 6             0             0             0
## 7             0             0             0
## 8             0             0             0
## 9             1             0             0
## 10            1             0             0
## 11            0             1             0
## 12            0             1             0
## 13            0             0             1
## 14            0             0             1
## attr(,"assign")
## [1] 0 1 1 1 1 1 1
## attr(,"contrasts")
## attr(,"contrasts")$`as.factor(scenario)`
## [1] "contr.treatment"
```

Conclusion: The above output shows the difference in model matrices between scenario encoded as an integer and factor. The integer version treats scenario as a continuous variable, whereas the factorized version creates a regression line per scenario.

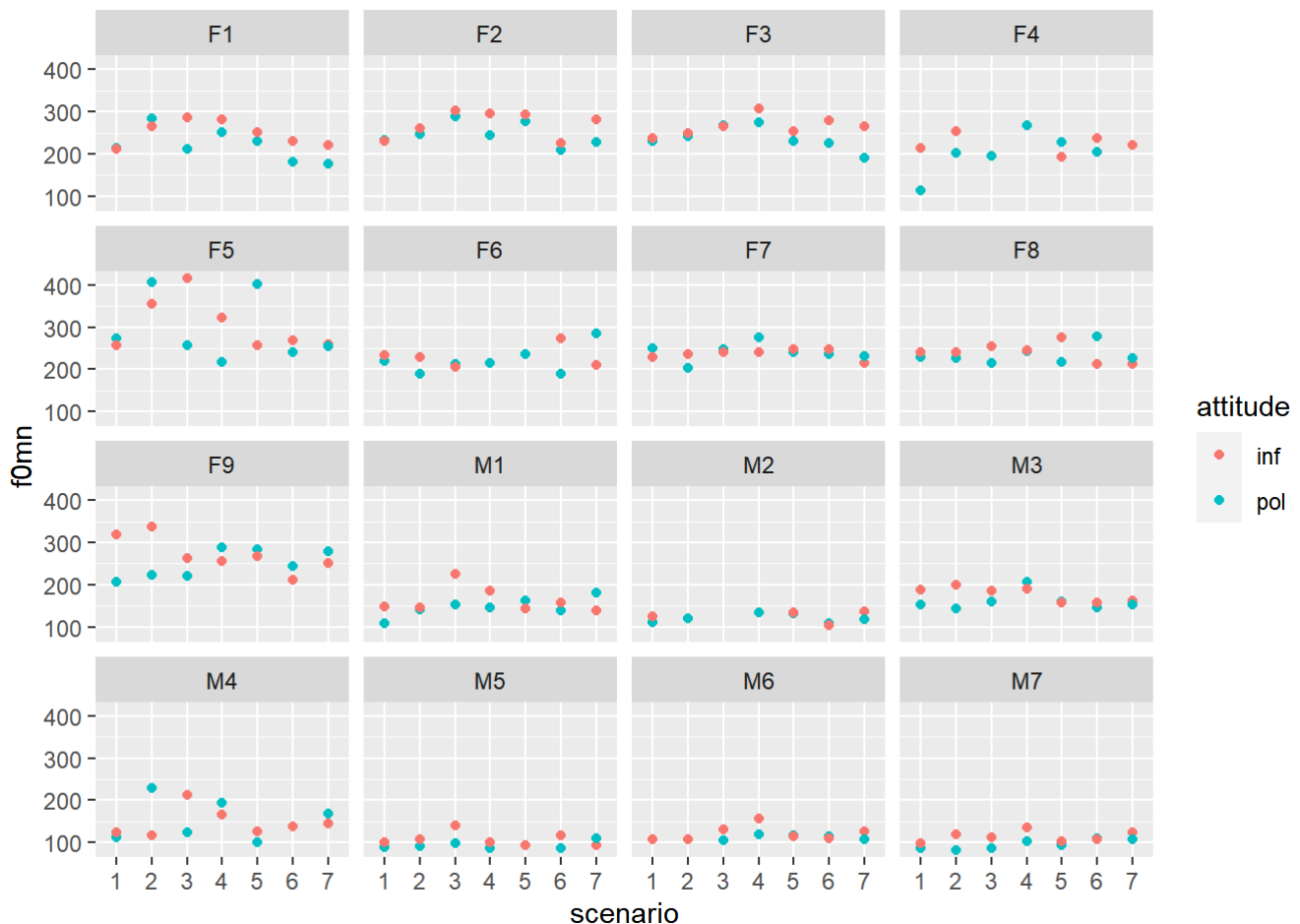
For this dataset, scenario should be a factor, since the scenarios are not a continuous variable and depending on the prescribed scenario, the participants may have a different *f0* (mean fundamental frequency of speech), and we are interested in following the trajectory of *f0* across scenarios not as a variable that consistently decreases or increases, but a separate regression line showing the changes in *f0* between the 7 different scenarios. And in order, to be able to see that crucial difference we need to consider the 'scenario' variable as a factor when we run a model predicting *f0* across scenarios.

Exercise 1, part 3 – (EH)

3. Make a plot that includes a subplot for each subject that has *scenario* on the x-axis and *f0mn* on the y-axis and where points are colour coded according to *attitude*
 - i. Describe the differences between subjects

```
politeness %>% ggplot(aes(scenario, f0mn, color = attitude)) +
  geom_point() +
  facet_wrap(vars(subject))
```

```
## Warning: Removed 12 rows containing missing values (geom_point).
```



...

We can visually observe that there are baseline differences between the male and female subjects' mean fundamental frequency of speech, where the males' f0 is consistently lower, across scenario and attitude. Between the different scenarios there is variability in the f0 values for both male and female subjects depending on both the scenario type and the attitude (informal or formal). There is a consistent tendency across scenario type and gender for the mean fundamental frequency of speech to be slightly higher when the attitude is informal as opposed to formal. This visual information interpreted from this plot is consistent with the results of Winter and Grawunder (2012)

Exercise 2 - comparison of models

```
mixed.model <- lmer(formula=..., data=...)
example.formula <- formula(dep.variable ~ first.level.variable + (1 | second.level.variable))
```

Exercise 2, Part 1 – (VK)

1. Build four different models and do some comparisons

```
# the single level model
m3 <- lm(formula = f0mn ~ gender, data = politeness)

# a two-level model where each scenario has a unique intercept
m4 <- lmer(formula = f0mn ~ gender + (1 | scenario), data = politeness)
# a two-level model that has models subject as intercept
m5 <- lmer(formula = f0mn ~ gender + (1 | subject), data = politeness)

# a two-level model that incorporate intercepts for both subject and scenario
m6 <- lmer(formula = f0mn ~ gender +
            (1 | subject) + (1 | scenario), data = politeness)

#comparing AIC and Deviance values for all the models
AIC(m3)
```

```
## [1] 2163.971
```

```
AIC(m4)
```

```
## [1] 2152.314
```

```
AIC(m5)
```

```
## [1] 2099.626
```

```
AIC(m6)
```

```
## [1] 2092.482
```

```
deviance(m3)
```

```
## [1] 327033.6
```

```
deviance(m4)
```

```
## Warning in deviance.merMod(m4): deviance() is deprecated for REML fits;
## use REMLcrit for the REML criterion or deviance(.,REML=FALSE) for deviance
## calculated at the REML fit
```

```
## [1] 2144.314
```

```
deviance(m5)
```

```
## Warning in deviance.merMod(m5): deviance() is deprecated for REML fits;  
## use REMLcrit for the REML criterion or deviance(.,REML=FALSE) for deviance  
## calculated at the REML fit
```

```
## [1] 2091.626
```

```
deviance(m6)
```

```
## Warning in deviance.merMod(m6): deviance() is deprecated for REML fits;  
## use REMLcrit for the REML criterion or deviance(.,REML=FALSE) for deviance  
## calculated at the REML fit
```

```
## [1] 2082.482
```

```
anova(m4, m5, m6)
```

```
## refitting model(s) with ML (instead of REML)
```

```
## Data: politeness  
## Models:  
## m4: f0mn ~ gender + (1 | scenario)  
## m5: f0mn ~ gender + (1 | subject)  
## m6: f0mn ~ gender + (1 | subject) + (1 | scenario)  
##      npar    AIC    BIC logLik deviance  Chisq Df Pr(>Chisq)  
## m4      4 2162.3 2175.7 -1077.1   2154.3  
## m5      4 2112.1 2125.5 -1052.0   2104.1 50.2095  0  
## m6      5 2105.2 2122.0 -1047.6   2095.2  8.8725  1  0.002895 **  
## ---  
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

```
piecewiseSEM::rsquared(c(m4, m5, m6))
```

```
##   Response  family    link method  Marginal Conditional  
## 1    f0mn gaussian identity   none 0.6779555  0.6967788  
## 2    f0mn gaussian identity   none 0.6681651  0.7899229  
## 3    f0mn gaussian identity   none 0.6677206  0.8077964
```

The single level model performs the worst and this makes sense as we do not expect all participants to have the same f0 as their voices have naturally occurring differences (not just ones predicted by gender). There are differences that might be explained by the scenario or individual subject (as we observed in the plot above), however, neither of those are taken into account when using a single level model.

So then using two-level models that explain variance by either taking scenario or subject as random intercepts, is definitely an improvement to help explain more of the scenario/attitude based differences, not just the gender differences.

Consequently, Of the three multi-level models, it is model m6, which includes random intercepts for both subject and scenario, that has the most explained variance with for the entire model $R^2 \approx 0.81$ or 81%.

Additionally, we see that model m6 has the lowest AIC and deviance.

Exercise 2, part 2 and 3 – (LR)

2. Why is our single-level model bad?

- create a new data frame that has three variables, *subject*, *gender* and *f0mn*, where *f0mn* is the average of all responses of each subject, i.e. averaging across *attitude* and *_scenario_*
- build a single-level model that models *f0mn* as dependent on *gender* using this new dataset
- make Quantile-Quantile plots, comparing theoretical quantiles to the sample quantiles) using `qqnorm` and `qqline` for the new single-level model and compare it to the old single-level model (from 1).i). Which model's residuals (ϵ) fulfil the assumptions of the General Linear Model better?)
- Also make a quantile-quantile plot for the residuals of the multilevel model with two intercepts. Does it look alright?

3. Plotting the two-intercepts model

- Create a plot for each subject, (similar to part 3 in Exercise 1), this time also indicating the fitted value for each of the subjects for each for the scenarios (hint use `fixef` to get the “grand effects” for each gender and `ranef` to get the subject- and scenario-specific effects)

```
# scenario x f0mn y, attitude = color
ff <- fixef(m6)
ff
```

```
## (Intercept)      genderM
##      246.7650    -115.1746
```

```
rf <- ranef(m6)
rf <- as.data.frame(rf)
rf
```

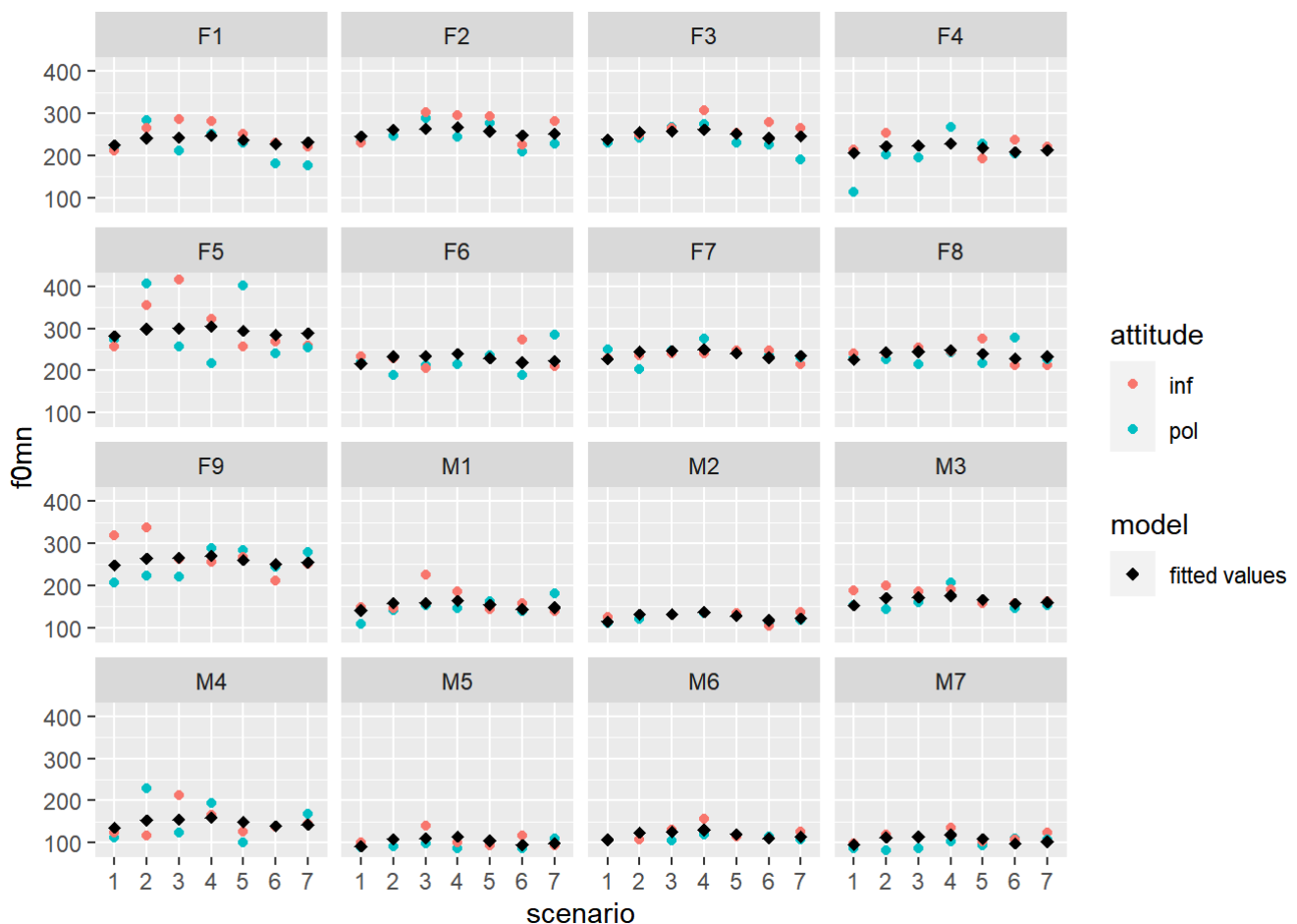
```
##      grpvar      term grp  condval  condsd
## 1 subject (Intercept) F1 -10.490356 8.280794
## 2 subject (Intercept) F2  10.251809 8.280794
## 3 subject (Intercept) F3   3.795129 8.280794
## 4 subject (Intercept) F4 -29.495270 9.095403
## 5 subject (Intercept) F5  47.093999 8.280794
## 6 subject (Intercept) F6 -18.396273 8.794359
## 7 subject (Intercept) F7  -6.976691 8.280794
## 8 subject (Intercept) F8  -8.521934 8.280794
## 9 subject (Intercept) F9  12.739587 8.280794
## 10 subject (Intercept) M1  21.052117 8.280794
## 11 subject (Intercept) M2  -5.462358 9.453009
## 12 subject (Intercept) M3  33.561535 8.280794
## 13 subject (Intercept) M4  16.093337 8.524543
## 14 subject (Intercept) M5 -28.267430 8.280794
## 15 subject (Intercept) M6 -12.640202 8.794541
## 16 subject (Intercept) M7 -24.336998 8.280794
## 17 scenario (Intercept) 1 -11.595496 5.488728
## 18 scenario (Intercept) 2   5.321218 5.532205
## 19 scenario (Intercept) 3   6.795658 5.586194
## 20 scenario (Intercept) 4  11.348815 5.578013
## 21 scenario (Intercept) 5   1.411037 5.488705
## 22 scenario (Intercept) 6  -8.622136 5.489258
## 23 scenario (Intercept) 7  -4.659096 5.488058
```

```

politeness$effect_gender <- 0.0
politeness[politeness$gender == "F", ]$effect_gender <- ff[1]
politeness[politeness$gender == "M", ]$effect_gender <- ff[1] + ff[2]
politeness$intercept_subject <- left_join(politeness, rf, by = c("subject" = "grp"),
  copy = TRUE, keep = FALSE)$condval
politeness$intercept_scenario <- left_join(politeness, rf, by = c("scenario" = "grp"
), copy = TRUE, keep = FALSE)$condval
politeness$predicted <- politeness$effect_gender + politeness$intercept_subject + pol
iteness$intercept_scenario
politeness %>% ggplot(aes(scenario, f0mn, color = attitude)) +
  geom_point() +
  geom_point(aes(y = predicted, shape = "fitted values"), color = "black", size = 2
) +
  scale_shape_manual(name = "model", values = c(18)) +
  facet_wrap(vars(subject))

```

```
## Warning: Removed 12 rows containing missing values (geom_point).
```



```
deviance(m3)
```

```
## [1] 327033.6
```

```
deviance(m4, REML = FALSE)
```

```
## [1] 2154.33
```

```
deviance(m5, REML = FALSE)
```

```
## [1] 2104.175
```

```
deviance(m6, REML = FALSE)
```

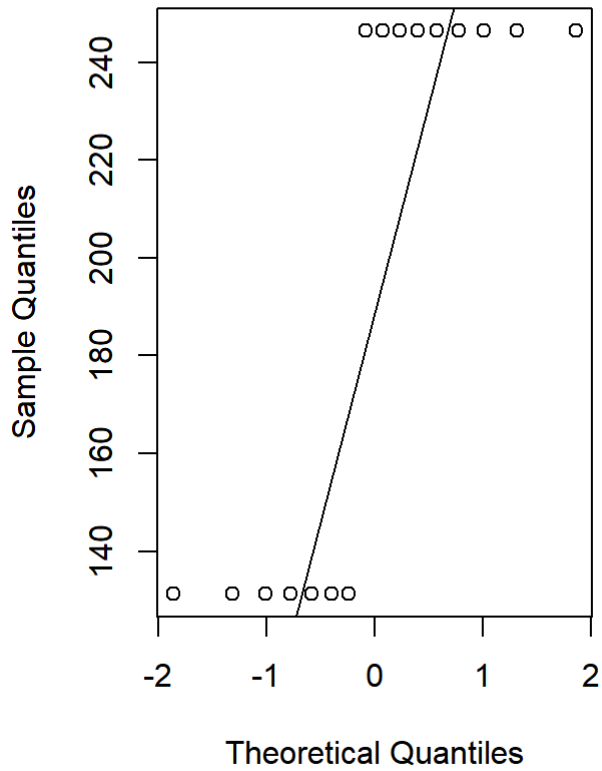
```
## [1] 2095.279
```

```
politeness_aggregated <- politeness[!is.na(politeness$f0mn), ] %>% group_by(subject)
%>% summarize(subject = subject[1], gender = gender[1], f0mn = mean(f0mn))
politeness_aggregated
```

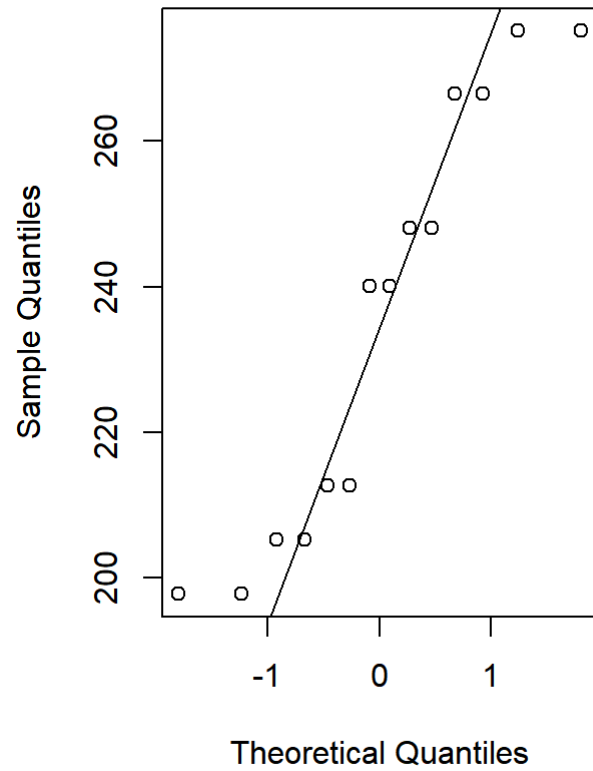
```
## # A tibble: 16 x 3
##   subject gender  f0mn
##   <chr>    <fct> <dbl>
## 1 F1      F      235.
## 2 F2      F      258.
## 3 F3      F      251.
## 4 F4      F      212.
## 5 F5      F      299.
## 6 F6      F      225.
## 7 F7      F      239.
## 8 F8      F      237.
## 9 F9      F      261.
## 10 M1     M      155.
## 11 M2     M      122.
## 12 M3     M      169.
## 13 M4     M      150.
## 14 M5     M      100.
## 15 M6     M      118.
## 16 M7     M      104.
```

```
m7 <- lm(f0mn ~ gender, data = politeness_aggregated)
par(mfrow=c(1,2))
qqnorm(fitted.values(m7))
qqline(fitted.values(m7))
qqnorm(fitted.values(m2))
qqline(fitted.values(m2))
```

Normal Q-Q Plot

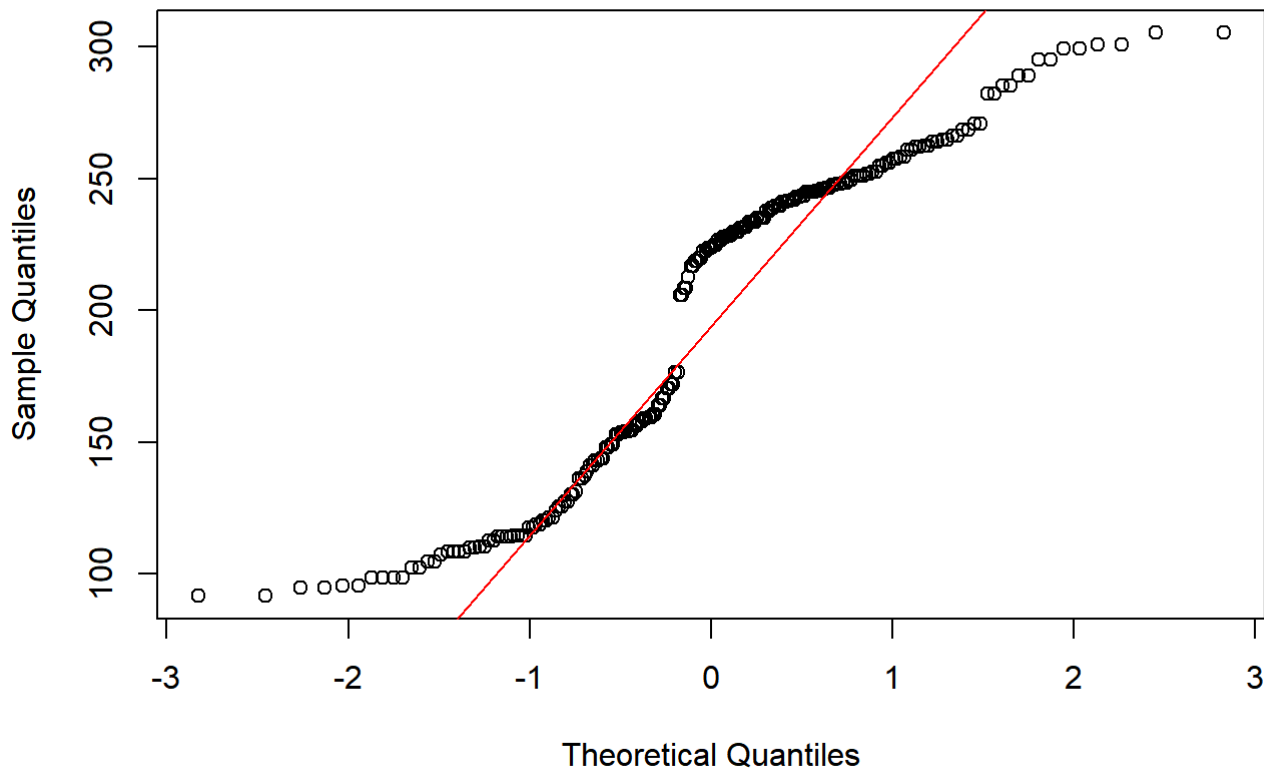


Normal Q-Q Plot



```
par(mfrow=c(1,1))  
qqnorm(fitted.values(m6))  
qqline(fitted.values(m6), col = "red")
```

Normal Q-Q Plot



Assessing the QQ-plots of the single-level models it seems that the aggregated model m7's residuals are worse off than those of model m2. The residuals of model m2 are better dispersed along the line - however it still doesn't look fantastic. The QQ-plot of the multilevel model m6 looks better than any of the single level ones, with data points closer to the line and more evenly dispersed on both sides of the line.

Looking at the plot for the observed and the fitted values it looks as if the model m6 performs reasonably as well.

Exercise 3 - now with attitude

Exercise 3, part 1 –(VK)

1. Carry on with the model with the two unique intercepts fitted (*scenario* and *subject*) but now build a new model that has *attitude* as a main effect besides *gender*. After create a separate model that besides the main effects of *attitude* and *gender* also include their interaction

```
#making a model with two unique intercepts and having attitude and gender as main effects
m8 <- lmer(formula = f0mn ~ gender + attitude + (1 | subject) + (1 | scenario), data = politeness)
summary(m8)
```

```
## Linear mixed model fit by REML ['lmerMod']
## Formula: f0mn ~ gender + attitude + (1 | subject) + (1 | scenario)
## Data: politeness
##
## REML criterion at convergence: 2065.1
##
## Scaled residuals:
##      Min       1Q   Median       3Q      Max
## -2.8511 -0.6081 -0.0602  0.4329  3.8745
##
## Random effects:
## Groups   Name      Variance Std.Dev.
## subject  (Intercept) 585.6    24.20
## scenario (Intercept) 106.7     10.33
## Residual                882.7    29.71
## Number of obs: 212, groups:  subject, 16; scenario, 7
##
## Fixed effects:
##              Estimate Std. Error t value
## (Intercept)  254.398      9.597  26.507
## genderM      -115.437     12.881  -8.962
## attitudepol  -14.819      4.096  -3.618
##
## Correlation of Fixed Effects:
##              (Intr) gendrM
## genderM      -0.587
## attitudepol  -0.220  0.006
```

```
fitted.values(m8)
```

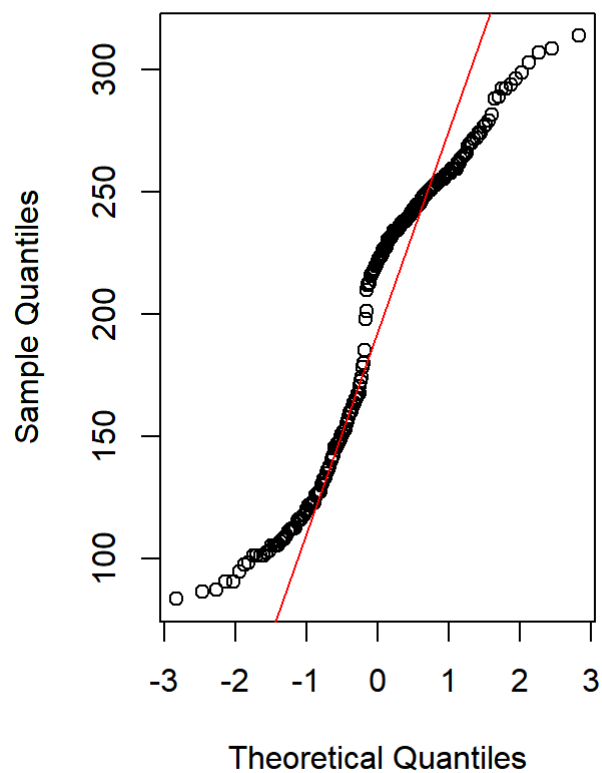
##	1	2	3	4	5	6	7	8
##	216.47929	231.29853	234.27048	249.08973	236.09087	250.91011	241.21606	256.03530
##	9	10	11	12	13	14	15	16
##	230.41934	245.23858	219.58594	234.40518	238.54395	223.72471	237.34031	252.15956
##	17	18	19	20	21	22	23	24
##	255.13151	269.95075	256.95189	271.77114	262.07708	276.89632	251.28036	266.09960
##	25	26	27	28	29	30	31	32
##	240.44696	255.26621	259.40498	244.58573	230.84663	245.66588	248.63783	263.45707
##	33	34	35	36	37	38	39	40
##	250.45822	265.27746	255.58340	270.40264	244.78668	259.60593	233.95328	248.77253
##	41	42	43	44	45	46	47	49
##	252.91130	238.09206	198.00782	212.82706	215.79901	230.61826	217.61940	222.74458
##	51	52	53	54	55	57	58	59
##	211.94787	226.76711	201.11447	215.93371	220.07248	274.39362	289.21286	292.18481
##	60	61	62	63	64	65	66	67
##	307.00406	294.00520	308.82444	299.13038	313.94963	288.33367	303.15291	277.50027
##	68	69	70	71	72	73	74	75
##	292.31951	296.45828	281.63904	209.70329	224.52253	227.49448	242.31372	229.31487
##	76	77	79	81	82	83	84	85
##	244.13411	234.44005	223.64333	212.80994	227.62918	231.76795	216.94871	220.01309
##	86	87	88	89	90	91	92	93
##	234.83233	237.80428	252.62353	239.62467	254.44391	244.74985	259.56910	233.95314
##	94	95	96	97	98	99	100	101
##	248.77238	223.11974	237.93898	242.07775	227.25851	218.45899	233.27823	236.25019
##	102	103	104	105	106	107	108	109
##	251.06943	238.07057	252.88982	243.19576	258.01500	232.39904	247.21828	221.56564
##	110	111	112	113	114	115	116	117
##	236.38488	240.52366	225.70441	239.84235	254.66159	257.63354	272.45278	259.45393
##	118	119	120	121	122	123	124	125
##	274.27317	264.57911	279.39835	253.78240	268.60164	242.94900	257.76824	261.90701
##	126	127	128	129	130	131	132	133
##	247.08777	133.00242	147.82166	150.79361	165.61285	152.61400	167.43324	157.73918
##	134	135	136	137	138	139	140	141
##	172.55843	146.94247	161.76171	136.10907	150.92831	155.06708	140.24784	107.79414
##	142	143	147	149	150	151	152	153
##	122.61338	125.58533	132.53090	121.73419	136.55343	110.90079	125.72003	129.85880
##	154	155	156	157	158	159	160	161
##	115.03956	145.58352	160.40276	163.37471	178.19395	165.19510	180.01434	170.32028
##	162	163	164	165	166	167	168	169
##	185.13952	159.52357	174.34281	148.69017	163.50941	167.64818	152.82894	127.46793
##	170	171	172	173	174	175	176	177
##	142.28717	145.25913	160.07837	147.07951	161.89876	152.20470	167.02394	141.40798
##	178	180	181	182	183	184	185	186
##	156.22722	145.39382	149.53260	134.71335	83.40025	98.21950	101.19145	116.01069
##	187	188	189	190	191	192	193	194
##	103.01184	117.83108	108.13702	122.95626	97.34030	112.15955	86.50690	101.32615
##	195	196	198	200	201	202	203	204
##	105.46492	90.64568	112.78128	130.57248	117.57362	132.39287	122.69881	137.51805
##	205	206	207	208	209	210	211	212
##	111.90209	126.72133	101.06869	115.88793	120.02671	105.20746	87.35321	102.17245
##	213	214	215	216	217	218	219	220
##	105.14440	119.96365	106.96479	121.78403	112.08998	126.90922	101.29326	116.11250
##	221	222	223	224				
##	90.45986	105.27910	109.41787	94.59863				

```

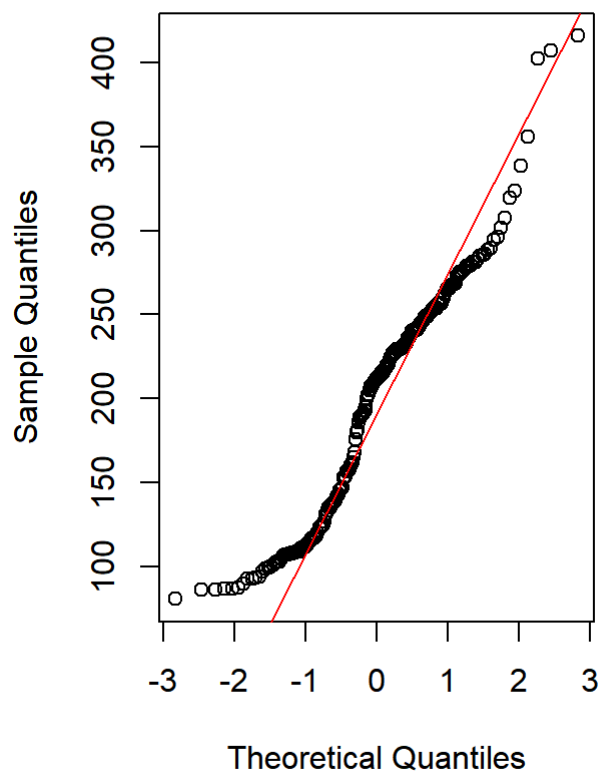
par(mfrow=c(1,2))
qqnorm(fitted.values(m8))
qqline(fitted.values(m8), col = "red")
qqnorm(politeness$f0mn)
qqline(politeness$f0mn, col = "red")

```

Normal Q-Q Plot



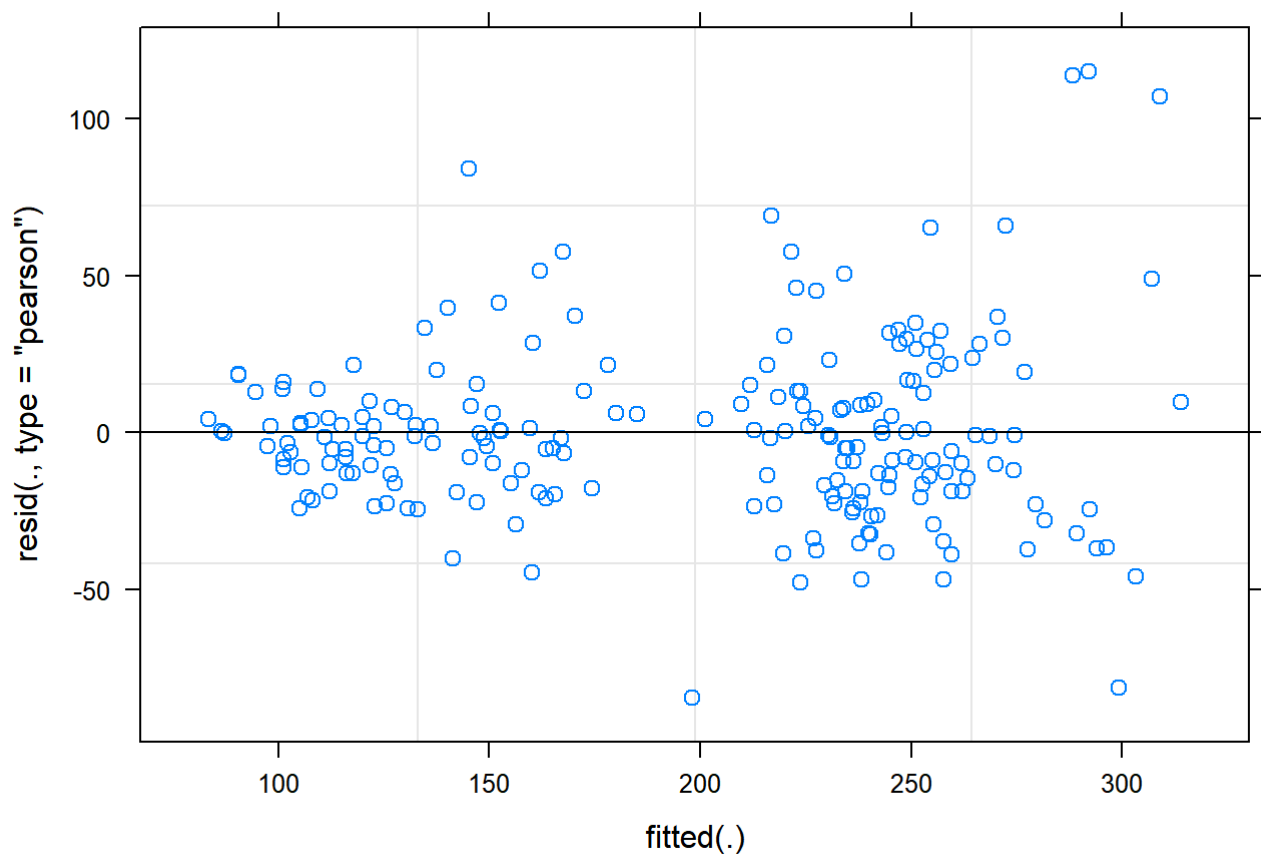
Normal Q-Q Plot



```

plot(m8)

```

```
#a model that additionally includes the interaction between the main effects (attitud  
e and gender)  
m9 <- lmer(formula = f0mn ~ gender + attitude + gender:attitude + (1 | subject) + (1  
| scenario), data = politeness)  
summary(m9)
```

```
## Linear mixed model fit by REML ['lmerMod']
## Formula: f0mn ~ gender + attitude + gender:attitude + (1 | subject) +
##      (1 | scenario)
##      Data: politeness
##
## REML criterion at convergence: 2058.6
##
## Scaled residuals:
##      Min       1Q   Median       3Q      Max
## -2.8120 -0.5884 -0.0645  0.4014  3.9100
##
## Random effects:
##   Groups      Name      Variance Std.Dev.
##  subject (Intercept) 584.4      24.17
##  scenario (Intercept) 106.4      10.32
##  Residual              885.5      29.76
## Number of obs: 212, groups:  subject, 16; scenario, 7
##
## Fixed effects:
##              Estimate Std. Error t value
## (Intercept)      255.618      9.761  26.186
## genderM           -118.232     13.531  -8.738
## attitudepol       -17.192      5.423  -3.170
## genderM:attitudepol  5.544      8.284   0.669
##
## Correlation of Fixed Effects:
##              (Intr) gendrM atttdp
## genderM      -0.606
## attitudepol  -0.286  0.206
## gendrM:tttdp  0.187 -0.309 -0.654
```

The model m9 can be read as following: The intercept for women/inf are $\approx 256Hz$, when we look at men with the same attitude their pitch drops by $\approx 118Hz$. Overall a polite attitude will result in a drop in pitch by $\approx 17Hz$, however for men it will only be $-17.2 + 5.5 \approx 11.6Hz$. Korean women's relative drop in pitch is therefore larger than male's in a polite situation according to this sample.

Exercise 3, part 2 –(KV)

2. Compare the three models (1. gender as a main effect; 2. gender and attitude as main effects; 3. gender and attitude as main effects and the interaction between them.

```
#comparing the models
anova(m6, m8, m9)
```

```
## refitting model(s) with ML (instead of REML)
```

```
## Data: politeness
## Models:
## m6: f0mn ~ gender + (1 | subject) + (1 | scenario)
## m8: f0mn ~ gender + attitude + (1 | subject) + (1 | scenario)
## m9: f0mn ~ gender + attitude + gender:attitude + (1 | subject) + (1 | scenario)
##      npar      AIC      BIC logLik deviance  Chisq Df Pr(>Chisq)
## m6      5 2105.2 2122.0 -1047.6   2095.2
## m8      6 2094.5 2114.6 -1041.2   2082.5 12.6868   1 0.0003683 ***
## m9      7 2096.0 2119.5 -1041.0   2082.0  0.4551   1 0.4998998
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

```
piecewiseSEM::rsquared(c(m6, m8, m9))
```

```
##      Response  family      link method  Marginal Conditional
## 1      f0mn gaussian identity   none 0.6677206   0.8077964
## 2      f0mn gaussian identity   none 0.6782542   0.8196777
## 3      f0mn gaussian identity   none 0.6782490   0.8192531
```

Exercise 3, part 3 –(KV)

3. Choose the model that you think describe the data the best - and write a short report on the main findings based on this model.

```
summary(m8)
```

```
## Linear mixed model fit by REML ['lmerMod']
## Formula: f0mn ~ gender + attitude + (1 | subject) + (1 | scenario)
## Data: politeness
##
## REML criterion at convergence: 2065.1
##
## Scaled residuals:
##      Min       1Q   Median       3Q      Max
## -2.8511 -0.6081 -0.0602  0.4329  3.8745
##
## Random effects:
## Groups   Name                Variance Std.Dev.
## subject  (Intercept)    585.6      24.20
## scenario (Intercept)   106.7      10.33
## Residual                    882.7      29.71
## Number of obs: 212, groups:  subject, 16; scenario, 7
##
## Fixed effects:
##              Estimate Std. Error t value
## (Intercept)   254.398      9.597  26.507
## genderM       -115.437     12.881  -8.962
## attitudepol   -14.819      4.096  -3.618
##
## Correlation of Fixed Effects:
##              (Intr) gendrM
## genderM      -0.587
## attitudepol  -0.220  0.006
```

This dataset consists of the basic demographic information of 16 Korean participants, and their observed pitch in different situations that require either an informal or polite(formal) attitude.

The model we chose as the best is m8. Consistent with the conclusions of the original oauthor's study, this model showed that women on average have a higher pitch than men BUT it also suggested a negative relationship between *attitude* and pitch with p-values<0.001. It would seem that both Korean men's and women's frequency of voice drops when having a polite attitude.

Subjects and scenarios should have different intercepts because it would be assumed they would all have different baselines, (and therefore need different intercepts to account for this). Different subjects will naturally already speak at a different pitch level, so separate intercepts allows us to account for these differences. The different scenarios may also need separate intercepts as certain scenarios may result in participants lowering or raising their pitch to meet the appropriate ambiance of the scenario. Once again, by including separate intercepts for scenarios then we should be accounting for these differences in our models.

Furthermore the output of the summary function shows that more variance is explained by the random effect of the subject than that of the scenario, further strengthening the choice of multilevel modelling.

And here's a QQ-plot of our chosen model

```
qqnorm(fitted.values(m8))
qqline(fitted.values(m8), col = "red")
```

Normal Q-Q Plot

