

第1题答案:

解 由泊松分布的数学期望及方差知

$$E(X_i) = \mu = \lambda = 0.04,$$

$$D(X_i) = \sigma^2 = \lambda = 0.04,$$

$$E\left(\sum_{i=1}^{100} X_i\right) = 100 \times 0.04 = 4,$$

$$D\left(\sum_{i=1}^{100} X_i\right) = 100\sigma^2 = 100 \times 0.04 = 4 \text{ (这里利用了 } X_1, X_2, \dots \text{ 的相互独立性),}$$

所以,由中心极限定理知,对于任意 x ,有

$$\lim_{n \rightarrow \infty} P\left\{\frac{\sum_{i=1}^{100} X_i - n\mu}{\sqrt{n}\sigma} \leq x\right\} = \int_{-\infty}^x \frac{1}{\sqrt{2\pi}} e^{-\frac{t^2}{2}} dt.$$

从而

$$\begin{aligned} P\{Z > 3\} &= P\left\{\sum_{i=1}^{100} X_i > 3\right\} = 1 - P\left\{\sum_{i=1}^{100} X_i \leq 3\right\} \\ &= 1 - P\left\{\frac{\sum_{i=1}^{100} X_i - 4}{\sqrt{4}} \leq \frac{3-4}{2}\right\} \\ &= 1 - P\left\{\frac{\sum_{i=1}^{100} X_i - 4}{2} \leq -\frac{1}{2}\right\} \\ &\approx 1 - \int_{-\infty}^{-\frac{1}{2}} \frac{1}{\sqrt{2\pi}} e^{-\frac{t^2}{2}} dt \\ &= 1 - \Phi\left(-\frac{1}{2}\right) = \Phi\left(\frac{1}{2}\right) \\ &\approx 0.6915, \end{aligned}$$

$$\text{即 } P\{Z > 3\} \approx 0.6915.$$

第2题答案:

解 因两个戏院情况一样,故只需考虑甲戏院,即可设甲戏院需设 M 个座位,定义随机变量 ξ_i 如下:

$$\xi_i = \begin{cases} 1, & \text{若第 } i \text{ 个观众选择甲戏院,} \\ 0, & \text{否则} \end{cases} \quad (i=1, 2, \dots, 1000),$$

则 $P\{\xi_i = 1\} = P\{\xi_i = 0\} = \frac{1}{2}$, $\xi_1, \xi_2, \dots, \xi_{1000}$ 是独立同分布的随机变量.

以 ξ 表示“选择甲戏院的观众总数”,则

$$\xi = \xi_1 + \xi_2 + \dots + \xi_{1000}.$$

据题意,要决定 M 使 $P\{\xi \leq M\} \geq 99\%$.

注意到 $E(\xi_i) = \frac{1}{2}$, $D(\xi_i) = \frac{1}{4}$, 由独立同分布的中心极限定理,得

$$\begin{aligned} P\{\xi \leq M\} &= P\left\{\sum_{i=1}^{1000} \xi_i \leq M\right\} \\ &= P\left\{\frac{\sum_{i=1}^{1000} (\xi_i - 0.5)}{\frac{1}{2}\sqrt{1000}} \leq \frac{M-500}{\frac{1}{2}\sqrt{1000}}\right\} \\ &\approx \Phi\left(\frac{M-500}{5\sqrt{10}}\right) \geq 99\%. \end{aligned}$$

查标准正态分布表,得

$$\frac{M-500}{5\sqrt{10}} \geq 2.33,$$

所以

$$M \geq 2.33 \times 5\sqrt{10} + 500 \approx 537,$$

即每个戏院应设 537 个以上的座位.

第3题答案:

解 设 X 表示“投掷一枚均匀硬币 n 次,其中正面向上的次数”,则

$$X \sim b(n, 0.5), E(X) = np = 0.5n, D(X) = np(1-p) = 0.25n.$$

$$(1) P\left\{0.4 < \frac{X}{n} < 0.6\right\} = P\{0.4n < X < 0.6n\} = P\{-0.1n < X - 0.5n < 0.1n\}$$

$$= P\{|X - 0.5n| < 0.1n\} \geq 1 - \frac{D(X)}{(0.1n)^2}$$

$$= 1 - \frac{0.25n}{0.01n^2} = 1 - \frac{25}{n} \geq 0.9,$$

则 $\frac{25}{n} \leq 0.1$, 解得 $n \geq 250$.

$$(2) P\left\{0.4 < \frac{X}{n} < 0.6\right\} = P\{0.4n < X < 0.6n\}$$

$$= \Phi\left(\frac{0.6n - 0.5n}{\sqrt{0.25n}}\right) - \Phi\left(\frac{0.4n - 0.5n}{\sqrt{0.25n}}\right)$$

$$= \Phi(0.2\sqrt{n}) - \Phi(-0.2\sqrt{n})$$

$$= 2\Phi(0.2\sqrt{n}) - 1 \geq 0.9,$$

则 $2\Phi(0.2\sqrt{n}) - 1 \geq 0.95$, 查表得 $0.2\sqrt{n} \geq 1.645$, 解得 $n \geq 67.65$, 故取 $n = 68$.

注 一般情况下,估计概率时,用中心极限定理要比用契比雪夫不等式精确得多.

第4题答案:

证 令 $X = \frac{1}{n} \sum_{i=1}^n X_i$, 则

$$E(X) = E\left(\frac{1}{n} \sum_{i=1}^n X_i\right) = \frac{1}{n} \sum_{i=1}^n E(X_i),$$

$$D(X) = D\left(\frac{1}{n} \sum_{i=1}^n X_i\right) = \frac{1}{n^2} D\left(\sum_{i=1}^n X_i\right).$$

由契比雪夫不等式知,对任意给定的 $\epsilon > 0$, 有

$$P\{|X - E(X)| < \epsilon\} \geq 1 - \frac{D(X)}{\epsilon^2}. \quad (*)$$

根据假设条件有

$$\lim_{n \rightarrow \infty} D(X) = \lim_{n \rightarrow \infty} \frac{1}{n^2} D\left(\sum_{i=1}^n X_i\right) = 0.$$

在(*)式两边取极限,得

$$\lim_{n \rightarrow \infty} P\{|X - E(X)| < \epsilon\} \geq 1.$$

由于概率不可能大于1,故有

$$\lim_{n \rightarrow \infty} P\{|X - E(X)| < \epsilon\} = 1,$$

即

$$\lim_{n \rightarrow \infty} P\left\{\left|\frac{1}{n} \sum_{i=1}^n X_i - \frac{1}{n} \sum_{i=1}^n E(X_i)\right| < \epsilon\right\} = 1.$$

第5题答案:

解 (1) 因为

$$X_i \sim N(0, 1), i = 1, 2, \dots, n,$$

所以

$$X_1 \sim N(0, 1), \sum_{i=2}^n X_i^2 \sim \chi^2(n-1),$$

从而

$$\frac{\sqrt{n-1}X_1}{\sqrt{\sum_{i=2}^n X_i^2}} = \frac{X_1}{\sqrt{\sum_{i=2}^n \frac{X_i^2}{n-1}}} \sim t(n-1).$$

(2) 因为 $X_1 \sim N(0, 1), X_2 \sim N(0, 1)$, 所以

$$X_1 - X_2 \sim N(0, 2), \frac{X_1 - X_2}{\sqrt{2}} \sim N(0, 1), X_3^2 + X_4^2 \sim \chi^2(2),$$

从而

$$\frac{X_1 - X_2}{(X_3^2 + X_4^2)^{\frac{1}{2}}} = \frac{\frac{X_1 - X_2}{\sqrt{2}}}{\sqrt{\frac{X_3^2 + X_4^2}{2}}} \sim t(2).$$

(3) 由题意知

$$X_1 - X_2 \sim N(0, 2), \frac{X_1 - X_2}{\sqrt{2}} \sim N(0, 1), \left(\frac{X_1 - X_2}{\sqrt{2}}\right)^2 \sim \chi^2(1),$$

$$X_3 + X_4 \sim N(0, 2), \frac{X_3 + X_4}{\sqrt{2}} \sim N(0, 1), \left(\frac{X_3 + X_4}{\sqrt{2}}\right)^2 \sim \chi^2(1),$$

所以

$$\left(\frac{X_1 - X_2}{X_3 + X_4}\right)^2 = \frac{(X_1 - X_2)^2}{(X_3 + X_4)^2} = \frac{\frac{(X_1 - X_2)^2}{\sqrt{2}}}{\frac{(X_3 + X_4)^2}{\sqrt{2}}} \sim F(1, 1).$$

(4) 由题意知

$$\sum_{i=1}^3 X_i^2 \sim \chi^2(3), \sum_{i=4}^n X_i^2 \sim \chi^2(n-3),$$

所以

$$\frac{\left(\frac{n}{3} - 1\right) \sum_{i=1}^3 X_i^2}{\sum_{i=4}^n X_i^2} = \frac{\frac{1}{3} \sum_{i=1}^3 X_i^2}{\frac{1}{n-3} \sum_{i=4}^n X_i^2} \sim F(3, n-3).$$

第6题答案:

证 不妨设总体的方差为 σ^2 , 则

$$\text{Corr}(x_i - \bar{x}, x_j - \bar{x}) = \frac{\text{Cov}(x_i - \bar{x}, x_j - \bar{x})}{\sqrt{\text{Var}(x_i - \bar{x})} \sqrt{\text{Var}(x_j - \bar{x})}}.$$

由 $\text{Cov}(x_i - \bar{x}, x_j - \bar{x}) = \text{Cov}(x_i, x_j) - \text{Cov}(x_i, \bar{x}) - \text{Cov}(x_j, \bar{x}) + \text{Cov}(\bar{x}, \bar{x})$, 由于

$$\text{Cov}(x_i, x_j) = 0, \text{Cov}(\bar{x}, \bar{x}) = \frac{\sigma^2}{n},$$

$$\text{Cov}(x_i, \bar{x}) = \text{Cov}(x_j, \bar{x}) = \text{Cov}\left(x_i, \frac{1}{n} \sum_{i=1}^n x_i\right) = \frac{\sigma^2}{n},$$

因而 $\text{Cov}(x_i - \bar{x}, x_j - \bar{x}) = -\frac{\sigma^2}{n},$

$$\text{Var}(x_i - \bar{x}) = \text{Var}(x_j - \bar{x}) = \text{Var}(x_1 - \bar{x}) = \text{Var}\left(\frac{(n-1)x_1 - x_2 - \dots - x_n}{n}\right)$$

$$= \frac{(n-1)^2 \sigma^2 + (n-1) \sigma^2}{n^2}$$

$$= \frac{(n-1) \sigma^2}{n},$$

所以 $\text{Corr}(x_i - \bar{x}, x_j - \bar{x}) = -(n-1)^{-1}.$

第7题答案:

分析 (X_1, \dots, X_n) 是取自 $X \sim b(1, p)$ 的样本, 故 $X_i \sim b(1, p)$ 且相互独立,

由两点分布的可加性, $\sum_{i=1}^n X_i \sim b(n, p)$.

由于 $EX = p, DX = p(1-p)$. 由一般总体抽样分布的极限分布结论, $U =$

$\frac{\bar{X} - p}{\sqrt{p(1-p)}} \sqrt{n}$ 渐近服从 $N(0, 1)$, 从而 \bar{X} 渐近服从 $N\left(p, \frac{p(1-p)}{n}\right)$

答 $b(n, p), N\left(p, \frac{p(1-p)}{n}\right)$

第8题答案:

分析 $X_i \sim N(0, \sigma^2)$ 且相互独立, 故 $X_1 + X_2 + X_3 \sim N(0, 3\sigma^2), X_4 + X_5 + X_6 \sim N(0, 3\sigma^2)$ 且相互独立, 从而

$$\left(\frac{X_1 + X_2 + X_3}{\sqrt{3}\sigma}\right)^2 \sim \chi^2(1), \left(\frac{X_4 + X_5 + X_6}{\sqrt{3}\sigma}\right)^2 \sim \chi^2(1)$$

且相互独立, 故有

$$\frac{1}{3\sigma^2} Y = \left(\frac{X_1 + X_2 + X_3}{\sqrt{3}\sigma}\right)^2 + \left(\frac{X_4 + X_5 + X_6}{\sqrt{3}\sigma}\right)^2 \sim \chi^2(2).$$

即应有 $C = \frac{1}{3\sigma^2}$.

答 $\frac{1}{3\sigma^2}, 2$.

第9题答案:

分析 由于 $D(2X_2 - X_1) = 4DX_2 + DX_1 = 5\sigma^2$, 故 (A) 不正确.

由 $\frac{(n-1)S^2}{\sigma^2} \sim \chi^2(n-1)$, 故 (C) 也不正确.

由 $\frac{\bar{X} - \mu}{S/\sqrt{n}} \sim t(n-1)$, 故 (D) 也不正确.

(B) 是正确的, 因为 $T = \frac{\bar{X} - \mu}{S/\sqrt{n}} \sim t(n-1)$, 故 $T^2 = \frac{n(\bar{X} - \mu)^2}{S^2} \sim F(1, n-1)$.

ABCD 即对应 (1)(2)(3)(4)

第十题:

因为 $X \sim \chi^2(m)$, $E(X) = m$, $D(X) = 2m$

则

$$E(\bar{X}) = E(X) = m$$

由于样本的独立性

$$D(\bar{X}) = \frac{1}{n^2} \sum_{i=1}^m D(X_i) = \frac{1}{n^2} \times n \times D(X_i) = \frac{1}{n} \cdot 2m = \frac{2m}{n}$$