

# Midterm review

## Lec 1

马氏距离  $d(x, y) = \sqrt{(x - y)^T S^{-1} (x - y)}$ .

$$E[X^T A X] = \text{tr}(A \Sigma) + \mu^T A \mu$$

$$E(\|X - Y\|_2^2) = 2\text{tr}(\Sigma)$$

$$E(\|X - Y\|_\Sigma^2) = 2\text{tr} I = 2p$$

## Lec 2 线性代数

Projection  $x$  to  $y$ :  $\frac{x'y}{y'y}y$ .

矩阵乘向量的几何直观: 旋转+缩放+旋转 考虑谱分解, SVD分解

谱分解:

$$A = \sum_{i=1}^k \lambda_i e_i e_i^T = P \Lambda P'$$

$$A^{-1} = \sum_{i=1}^k \frac{1}{\lambda_i} e_i e_i^T = P \Lambda^{-1} P'$$

$$A^{-1/2} = \sum_{i=1}^k \frac{1}{\sqrt{\lambda_i}} e_i e_i^T = P \Lambda^{-1/2} P'$$

正定 = 特征值均大于0

马氏距离几何直观

$$\text{Let } x = \tilde{P} \tilde{\Lambda} y, \text{ s.t. } \text{var}(y) = I$$

$$y'y = x'(\tilde{P} \tilde{\Lambda}^2 \tilde{P}')^{-1} x$$

$$\Sigma = \text{var}(x) = \tilde{P} \tilde{\Lambda}^2 \tilde{P}'$$

$$y'y = x' S^{-1} x$$

统计量几何表示

$$\bar{x} = \frac{1}{n} X' 1_n$$

$$n S_n = X' (I - \frac{1}{n} 1_n 1_n') X$$

$$D = \text{diag}(S_n), R = D^{-1/2} S_n D^{-1/2}$$

广义样本方差:  $\det S$

SVD分解:

$$A = U \Lambda V'$$

$$A A' = U \Lambda^2 U', A' A = V \Lambda^2 V'$$

不等式:

$b, d \in R^p, B \in R^{p \times p}$  is positive definite matrix. Then

$$b^T d \leq (b^T B b)(d^T B d)$$

with equality if and only if  $b = cB^{-1}d$  for some constant  $c$ .

$$\max_{x \neq 0} \frac{(x^T d)^2}{x^T B x} = d^T B^{-1} d \text{ when } x = cB^{-1}d$$

Let  $B$  be a positive definite matrix with eigenvalues  $\lambda_1 \geq \lambda_2 \geq \dots \geq \lambda_n$ , associated with eigenvectors  $e_1, e_2, \dots, e_n$ .

$$\begin{aligned} \max_{x \neq 0} \frac{x^T B x}{x^T x} &= \lambda_1, \text{ when } x = e_1 \\ \max_{x \perp e_1, \dots, e_k} \frac{x^T B x}{x^T x} &= \lambda_{k+1}, \text{ when } x = e_{k+1} \end{aligned}$$

矩阵微积分

$$\begin{aligned} \frac{\partial}{\partial x} x^T A x &= (A + A^T)x \\ \frac{\partial}{\partial A} |A| &= |A| A^{-1} \\ \frac{\partial}{\partial A} \text{tr}(AB) &= B^T \\ \frac{\partial}{\partial A} \text{tr}(A^{-1}B) &= -A^{-1} B^T A^{-1} \end{aligned}$$

## Lec 3 多元正态分布

$$X \sim N_p(\mu, \Sigma), f(x) = \frac{1}{(2\pi)^{p/2} |\Sigma|^{1/2}} \exp\left[-\frac{(x - \mu)' \Sigma^{-1} (x - \mu)}{2}\right]$$

Contours of constant density:

$$(x - \mu)' \Sigma^{-1} (x - \mu) = c^2$$

These ellipsoids are centered at  $\mu$  and have axes  $\pm c\sqrt{\lambda_i} e_i$ , where  $e_i$  is an eigenvector, i.e.  $\Sigma e_i = \lambda_i e_i$ .

多元正态的任意一组子变量仍符合多元正态，反过来不一定

$$AX + d \sim N_p(A\mu + d, A\Sigma A').$$

若  $Z \sim N_p(0, I)$ , 则  $\Sigma^{1/2} Z + \mu \sim N_p(\mu, \Sigma)$ .

对于多元正态分布，独立=不相关

$$X_1 | X_2 \sim N(\mu_1 + \Sigma_{12} \Sigma_{22}^{-1} (x_2 - \mu_2), \Sigma_{11} - \Sigma_{12} \Sigma_{22}^{-1} \Sigma_{21}).$$

多元正态分布  $X, (X - \mu)' \Sigma^{-1} (X - \mu) \sim \chi_p^2$ .

$$\begin{aligned} (X - \mu)' \Sigma^{-1} (X - \mu) &= [(X - \mu)' \Sigma^{-1/2}] [(X - \mu) \Sigma^{-1/2}]' \\ &= ZZ', \text{ where } Z \sim N(0, I) \end{aligned}$$

多元正态分布的线性组合:

assume  $X_j$  mutually independent,  $X_j \sim N_p(\mu_j, \Sigma), c_j \in R$ .

$$\sum_{i=1}^n c_i X_i \sim N(\sum_{i=1}^n c_i \mu_i, \sum_{i=1}^n c_i^2 \Sigma)$$

$$\text{MLE: } \hat{\mu} = \bar{X}, \hat{\Sigma} = S_n = \frac{n-1}{n} S$$

MLE的函数还是MLE

充分统计量  $\bar{X}, S$ , P28

$$\bar{X} \sim N_p(\mu, \frac{1}{n} \Sigma), (n-1)S \sim W_p(n-1, \Sigma), \bar{X}, S \text{ 相互独立}$$

Wishart distribution:  $W_p(m, \Sigma) = \sum_{j=1}^m Z_j Z_j'$ , where  $Z_j \sim N_p(0, \Sigma)$ .

properties: 可加性, P31

对于一组i.i.d的  $\{X_j\}$

大数定律  $\bar{X} \rightarrow E(X_i)$

中心极限定理  $\sqrt{n}(\bar{X} - \mu)$  近似于  $N_p(0, \Sigma)$ ,  $n(\bar{X} - \mu)' S^{-1} (\bar{X} - \mu)$  近似于  $\chi_p^2$ ,  $n \gg p$ .

## Lec 4 单总体正态推断

对一组i.i.d的  $\{X_j\}$ ,  $X_j \sim N_p(\mu, \Sigma)$ .

Under  $H_0$ :  $T^2 = n(\bar{X} - \mu_0)' S^{-1} (\bar{X} - \mu_0) \sim \frac{p(n-1)}{n-p} F_{p, n-p}(\alpha)$ .

reject if  $T^2 > T^2(\alpha) = \frac{p(n-1)}{n-p} F_{p, n-p}(\alpha)$ .

Hotelling's  $T^2$  在线性变换下不变

$$100(1 - \alpha)\% \text{ 置信区域 } n(\bar{X} - \mu)' S^{-1} (\bar{X} - \mu) \leq T^2(\alpha)$$

### 非正态假设下的大样本推断

$$n(\bar{X} - \mu)' S^{-1} (\bar{X} - \mu) > \chi_p^2(\alpha).$$

$$\text{Simultaneous: } a' \bar{x} \pm \sqrt{\chi_p^2(\alpha) a' S a / n}$$

Simultaneous comparisons of component means:  $(100(1 - \alpha)\%)$

$$\bar{x}_i \pm c \sqrt{\frac{s_{ii}}{n}}$$

$$T^2: c = \sqrt{\frac{p(n-1)}{n-p} F_{p, n-p}(\alpha)}$$

$$\text{庞弗洛尼: } c = t_{n-1}(\frac{\alpha}{2m}).$$

$T^2$  区间比

# Lec 5 双总体正态推断

## 1. 成对比较

## 2. 等方差正态假设

$$T^2 = (1/n_1 + 1/n_2)^{-1} (\bar{X}_1 - \bar{X}_2 - \delta_0)' S_{pooled}^{-1} (\bar{X}_1 - \bar{X}_2 - \delta_0)$$

$$S_{pooled} = \frac{(n_1 - 1)S_1 + (n_2 - 1)S_2}{n_1 + n_2 - 2} \sim \frac{W_p(n_1 + n_2 - 2, \Sigma)}{n_1 + n_2 - 2}.$$

$$\text{under } H_0, T^2 \sim \frac{p(n_1 + n_2 - 2)}{n_1 + n_2 - p - 1} F_{p, n_1 + n_2 - p - 1}$$

$$\text{Simultaneous: } a'(\bar{X}_1 - \bar{X}_2) \pm \sqrt{a'(1/n_1 + 1/n_2)S_{pooled}a}$$

对  $\mu_1 - \mu_2$  的检验, 拒绝假设起主要作用的线性组合:  $S_p^{-1}(\bar{x}_1 - \bar{x}_2)$

## 3. 不等方差大样本

$$(\bar{X}_1 - \bar{X}_2 - \delta_0)' \left( \frac{1}{n_1} S_1 + \frac{1}{n_2} S_2 \right)^{-1} (\bar{X}_1 - \bar{X}_2 - \delta_0) \leq \chi_p^2(\alpha).$$

$$\text{Simultaneous: } a'(\bar{x}_1 - \bar{x}_2) \pm \sqrt{\chi_p^2(\alpha) a' \left( \frac{1}{n_1} S_1 + \frac{1}{n_2} S_2 \right) a}.$$

对  $\mu_1 - \mu_2$  的检验, 拒绝假设起主要作用的线性组合:  $(\frac{1}{n_1} S_1 + \frac{1}{n_2} S_2)^{-1}(\bar{x}_1 - \bar{x}_2)$

## 等方差检验

P39 textbook P237

## 独立性检验

## 正态性检验

Q-Qplot

# Lec 6 PCA

## 总体主成分

$\Sigma$  的特征值-特征向量对  $(\lambda_1, e_1), (\lambda_2, e_2), \dots, (\lambda_p, e_p), \lambda_1 \geq \dots \geq \lambda_p \geq 0$ , 第  $j$  主成分  $Y_j = e_j' X$ .

$$\text{var}(Y_j) = \lambda_j, \text{Cov}(Y_i, Y_j) = 0 (i \neq j)$$

$$\sum_{i=1}^p \text{Var}(X_i) = \sum_{i=1}^p \text{Var}(Y_i)$$

$$\rho_{Y_i, X_j} = \frac{e_{ij} \sqrt{\lambda_i}}{\sqrt{\sigma_{kk}}}$$

标准化后, 主成分改变。

## 样本主成分

对样本协方差矩阵做特征分解即可。

主成分得分  $\hat{y}_i = \hat{e}_i' x$  或  $\hat{y}_i = \hat{e}_i' (x - \bar{x})$ . 可以看作样本  $x$  在  $\hat{e}_i$  方向上的投影, 是一个实数

