Problem 1

For OLS,

$$argmin \sum_{i=1}^{n} (y_i - f(x_i))^2 = argmin(y - X\beta)^T (y - X\beta)$$
$$\nabla_{\beta} (y - X\beta)^T (y - X\beta) = \nabla_{\beta} (y^T y + \beta^T X^T X \beta - \beta^T X^T y - y^T X \beta)$$
$$= 2(X^T X \beta - X^T y)$$

Let $\nabla_{\beta}(y - X\beta)^T(y - X\beta) = 0$, we get $\hat{\beta} = (X^T X)^{-1} X^T y$.

Problem 2

Since $y = X\beta + \epsilon$,

$$Cov(\hat{\beta}) = Cov[(X^T X)^{-1} X^T y]$$

$$= Cov[\beta + (X^T X)^{-1} X^T \epsilon]$$

$$= (X^T X)^{-1} X^T Cov(\epsilon) X(X^T X)^{-1}$$

$$= (X^T X)^{-1} \sigma^2$$

Problem 3

$$SS_{tot} = \sum_{i=1}^{n} (y_i - \bar{y})^2 = y^T y - n\bar{y}^2$$
 (1)

Since $E(y - \hat{y}) = E(y - X\beta) = E(\epsilon) = 0, \sum_{i=1}^{n} y_i = \sum_{i=1}^{n} \hat{y_i}$.

$$SS_{reg} = \sum_{i=1}^{n} (\hat{y}_i - \bar{y})^2$$

$$= \hat{y}^T \hat{y} + n\bar{y}^2 - 2\bar{y} \sum_{i=1}^{n} \hat{y}_i$$

$$= \hat{y}^T \hat{y} - n\bar{y}^2$$
(2)

Notice that $\hat{y}^T(y - \hat{y}) = \beta^T X^T [y - X(X^T X)^{-1} X^T y] = 0.$

$$SS_{res} = \sum_{i=1}^{n} (y_i - \hat{y})^2$$

$$= y^T y - \hat{y}^T \hat{y} + \hat{y}^T (y - \hat{y})$$

$$= y^T y - \hat{y}^T \hat{y}$$
(3)

Form (1), (2) and (3), we get $SS_{tot} = SS_{reg} + SSres$.

Problem 4

(1). Assume that ϵ and $f(x^*) - \hat{f}(x^*)$ are independent.

$$\begin{split} E[Y - \hat{f}(x^*)]^2 &= E[f(x^*) - \hat{f}(x^*) + \epsilon]^2 \\ &= E[f(x^*) - \hat{f}(x^*)]^2 + E(\epsilon^2) + 2E(\epsilon)E[f(x^*) - \hat{f}(x^*)] \\ &= E[f(x^*) - \hat{f}(x^*)]^2 + E(\epsilon^2) \end{split}$$

Since $E(\epsilon^2)$ is constant, minimizing $E[Y-\hat{f}(x^*)]^2$ is equivalent to minimizing $E[f(x^*)-\hat{f}(x^*)]^2$. We take the expectation with regard to r.m. x^* .

$$\begin{split} E[f(x^*) - \hat{f}(x^*)]^2 &= E[f(x^*) - E(\hat{f}(x^*)) + E(\hat{f}(x^*)) - \hat{f}(x^*)]^2 \\ &= (f(x^*) - E\hat{f}(x^*))^2 + E[E(\hat{f}(x^*)) - \hat{f}(x^*)]^2 \\ &\quad + 2[f(x^*) - E(\hat{f}(x^*))]E[E(\hat{f}(x^*)) - \hat{f}(x^*)] \\ &= (f(x^*) - E\hat{f}(x^*))^2 + Var[\hat{f}(x^*)]^2 \end{split}$$