

as in (3.1) but with  $[n(1 - b)]$  in place of  $[n/2]$ . In this case, the maximum finite sample breakdown point would be attained in  $b = 0.5 - k_n/n$  which is very close to our choice of  $b = 0.5$  when  $k_n/n$  is small.

## 4 Influence function

Consider an estimate  $\hat{\boldsymbol{\theta}}_n$  depending on a sample  $\mathbf{Z} = \{\mathbf{z}_1, \dots, \mathbf{z}_n\}$  of i.i.d. variables in  $\mathbb{R}^k$  with distribution  $H_{\boldsymbol{\theta}}$ , where  $\boldsymbol{\theta} \in \Theta \subset \mathbb{R}^m$ . Let  $T$  be an estimating functional of  $\boldsymbol{\theta}$  such that  $\mathbf{T}(H_n) = \hat{\boldsymbol{\theta}}_n$ , where  $H_n$  is the corresponding empirical distribution. Suppose that  $\mathbf{T}$  is Fisher consistent, i.e.  $\mathbf{T}(H_{\boldsymbol{\theta}}) = \boldsymbol{\theta}$ . The *influence function* of  $\mathbf{T}$ , introduced by Hampel [9], measures the effect on the functional of a small fraction of point mass contamination. If  $\delta_{\mathbf{z}}$  denotes the probability distribution that assigns mass 1 to  $\mathbf{z}$ , then the influence function is defined by

$$IF(\mathbf{z}, \mathbf{T}, \boldsymbol{\theta}) = \lim_{\varepsilon \rightarrow 0} \frac{\mathbf{T}((1 - \varepsilon)H_{\boldsymbol{\theta}} + \varepsilon\delta_{\mathbf{z}}) - \mathbf{T}(H_{\boldsymbol{\theta}})}{\varepsilon} = \left. \frac{\partial \mathbf{T}((1 - \varepsilon)H_{\boldsymbol{\theta}} + \varepsilon\delta_{\mathbf{z}})}{\partial \varepsilon} \right|_{\varepsilon=0},$$

In our case,  $\mathbf{z} = (\mathbf{y}', \mathbf{x}')'$  satisfy the linear model (1.1),  $\boldsymbol{\theta} = (\mathbf{B}_0, \boldsymbol{\Sigma}_0)$  and  $H_{\boldsymbol{\theta}} = H_0$ . Let  $\mathbf{T}_{0,1}$ ,  $\mathbf{T}_{0,2}$  be the functional estimates associated to the initial estimates  $\tilde{\mathbf{B}}_n$  and  $\tilde{\boldsymbol{\Sigma}}_n$ , and  $\mathbf{T}_1$ ,  $\mathbf{T}_2$  the functional estimates corresponding to the MM-estimates  $\hat{\mathbf{B}}_n$  and  $\hat{\boldsymbol{\Sigma}}_n$ . Then, according to (2.10) and (2.11), given a distribution function  $H$  of  $(\mathbf{y}', \mathbf{x}')'$ , the pair  $(\mathbf{T}_1(H), \mathbf{T}_2(H))$  is the value of  $(\mathbf{B}, \boldsymbol{\Sigma})$  satisfying

$$\begin{aligned} E_H W(d(\mathbf{B}, \boldsymbol{\Sigma})) \hat{\mathbf{u}}(\mathbf{B}) \mathbf{x}' &= \mathbf{0}, \\ \boldsymbol{\Sigma} &= q \frac{E_H W(d(\mathbf{B}, \boldsymbol{\Sigma})) \hat{\mathbf{u}}(\mathbf{B}) \hat{\mathbf{u}}(\mathbf{B})'}{E_{H_0} \psi_1(d(\mathbf{B}, \boldsymbol{\Sigma})) d(\mathbf{B}, \boldsymbol{\Sigma})}, \end{aligned}$$

and

$$\boldsymbol{\Sigma} = S(H)^2 \boldsymbol{\Gamma}, \text{ with } |\boldsymbol{\Gamma}| = 1,$$