

## 7 Moment Generating Functions

### 7.1 Moment Generating Functions

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#### Moment Generating Functions

We introduce a useful tool for later arguments.

**Moment Generating Function (mgf):** The mgf for a r.v.  $X$  is

$$M_X(t) = E[e^{tX}].$$

The mgf encaptures all moments of  $X$ :

**Theorem 4.1:** Let  $X$  be a r.v. with mgf  $M_X(t)$ . If exchanging expectation and differentiation is valid, then

$$E(X^n) = M_X^{(n)}(0)$$

where  $M_X^{(n)}(0)$  is the  $n$ th derivative of  $M_X(t)$  evaluated at  $t = 0$ .

**Proof:** under the assumption, we get

$$M_X^{(n)}(t) = E[X^n e^{tX}] \implies M_X^{(n)}(0) = E(X^n).$$

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#### Moment Generating Function (ctd.)

$$\Pr(X = n) = (1-p)^{n-1}p$$

**Example:**  $X \sim \text{Geom}(p)$ . Then, letting  $q = 1 - p$ , for  $qe^t < 1$ , or  $t < -\log(q)$ <sup>1</sup>

$$M_X(t) = Ee^{tX} = \sum_{k=1}^{\infty} q^{k-1} p e^{tk} = \frac{p}{q} \sum_{k=1}^{\infty} (qe^t)^k =$$

Geometric series:

$$a/(1-r) = a + ar + ar^2 + ar^3 + \dots$$

$$\text{In this case, } a = 1, r = qe^t \implies \frac{p}{q} \left( \frac{1}{1 - qe^t} - 1 \right).$$

**Theorem 4.2:** If  $M_X(t) = M_Y(t)$  for  $-\delta < t < \delta$  for some  $\delta > 0$ , then  $X$  and  $Y$  have the same distribution. (without proof)

**Theorem 4.3:** If  $X \perp Y$  then  $M_{X+Y}(t) = M_X(t)M_Y(t)$

**Solution:** exercise.

### 7.2 Chernoff Bounds

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#### Chernoff Bounds

Use Markov's inequality for  $e^{tX}$ , substituting  $E(e^{tX}) = M_X(t)$

$$\Pr(X \geq a) = \Pr(e^{tX} \geq e^{ta}) \leq \frac{E(e^{tX})}{e^{ta}}.$$

**Right tail:** This is true for all  $t > 0$  – use the one with the sharpest bound:

$$\Pr(X \geq a) \leq \min_{t>0} \frac{M_X(t)}{e^{ta}}.$$

<sup>1</sup> the constraint on  $t$  was missing in the lecture

**Left tail:** For  $\Pr(X \leq a)$  use  $t < 0$  and then again Markov's inequ:

$$\Pr(X \leq a) = \Pr(e^{tX} \geq e^{ta}) \leq \min_{t<0} \frac{M_X(t)}{e^{ta}}$$

Clever use of these allows many useful bounds that are better than Chebyshev bounds.

### 7.3 Examples

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#### Examples

1. Find  $M_X(t)$  for  $X \sim \text{Bern}(p)$ , i.e.,  $p_X(1) = p$ .

$$\text{Solution: } M_X(t) = (1-p) \cdot 1 + p \cdot e^t = 1 + p(e^t - 1).$$

Using  $1 + x \leq e^x$  we get

$$M_X(t) \leq e^{p(e^t - 1)},$$

which is sometimes a useful bound.

Writing  $q = 1 - p$  we get alternatively  $M_X(t) = q + pe^t$ .

2. Find  $M_Y(t)$  for  $Y \sim \text{Bin}(n, p)$ .

**Solution:** Use  $Y = \sum_{i=1}^n X_i$ , with  $X_i \sim \text{Bern}(p)$ , and therefore, letting  $q = 1 - p$ , we get  $M_Y(t) = (q + pe^t)^n$ .

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3. If  $X \sim \text{Poi}(\lambda)$ , show that for  $i < \lambda$ ,

$$\Pr(X \leq i) \leq \frac{e^{-\lambda}(e\lambda)^i}{i^i}.$$

**Hint:** If  $X \sim \text{Poi}(\lambda)$ , then  $M_X(t) = e^{\lambda(e^t - 1)}$ .

**Solution:** First,

$$\Pr(X \leq a) \leq M(t)/e^{ta} = e^{\lambda(e^t - 1) - ta},$$

for  $t < 0$ .

To find the lowest bound we minimize  $\lambda(e^t - 1) - ta$ , to find  $e^t = a/\lambda \implies t = \log(a/\lambda) < 0$  (for  $a < \lambda$ ).

Substituting in the Chernoff bound we get

$$\Pr(X \leq a) \leq e^{\lambda(a/\lambda - 1)} \left( \frac{\lambda}{a} \right)^a = \frac{e^{-\lambda}(e\lambda)^a}{a^a}$$

as claimed.

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For those following along in the book: We will

- skip the rest of §4, including §4.2.1 and beyond,

- skip §5-7, and
- continue with §8 in the next unit.

Some material (§4.3, §6.1 and §7.1) will be briefly introduced in homeworks.