

# Smoothed Nonparametric Derivative Estimation Using Weighted Difference Quotients

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Given a differentiable function  $m : \mathbb{R} \rightarrow \mathbb{R}$ , its derivative is defined as:

$$m'(x) = \lim_{h \rightarrow 0} \frac{m(x+h) - m(x)}{h}.$$

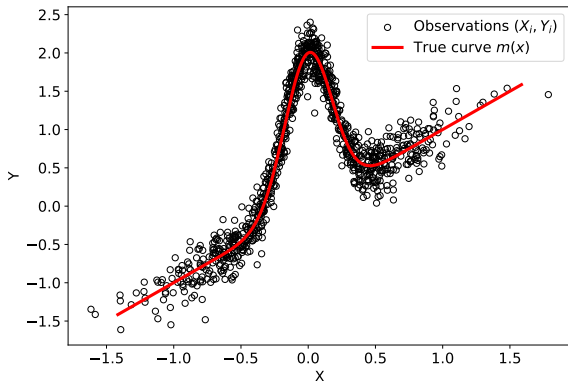
Studying  $m'(x)$  has significant impacts in Statistics:

- Explore the structures in curves ([Chaudhuri and Marron, 1999](#)).
- Signify the changing trend in time series ([Rondonotti et al., 2007](#)).
- Correct the bias for density/regression estimators to conduct valid inference ([Eubank and Speckman, 1993](#); [Calonico et al., 2018](#); [Cheng and Chen, 2019](#)).
- ...

**Data setting:**

$$Y_i = m(X_i) + e_i, \quad \text{with } X_i \in [a, b] \text{ for } i = 1, \dots, n,$$

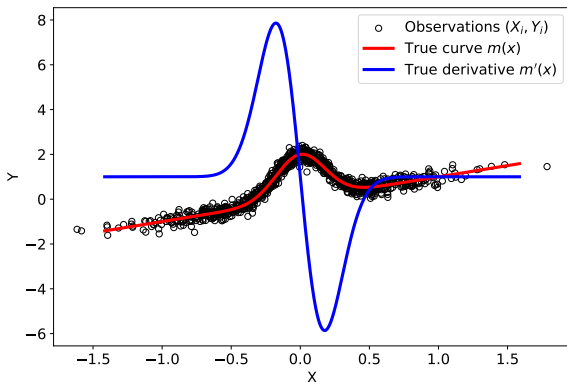
with  $e_i$  being independent of  $X_i$  and  $\mathbb{E}(e_i) = 0$ ,  $\text{Var}(e_i) = \sigma_e^2 < \infty$ .



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**Question:** How do we estimate  $m'(x)$  from the data  $\mathcal{D} = \{(X_i, Y_i)\}_{i=1}^n$ ?

**Parametric Methods:** Assume  $m(x)$  lying in some parametric family  $\{f(x; \theta) : \theta \in \Theta\}$  and fit

$$\hat{\theta} \in \arg \min_{\theta \in \Theta} \sum_{i=1}^n [Y_i - f(X_i; \theta)]^2 \quad \implies \quad \hat{m}'(x) = f'(x; \hat{\theta}).$$

- *Drawback:* It is difficult to propose a correct family  $\{f(x; \theta) : \theta \in \Theta\}$ .

**Nonparametric Methods:** Make no model assumptions on  $m(x)$  and estimate  $m'(x)$  from the data  $\mathcal{D}$ .

- **Smoothing Splines (Zhou and Wolfe, 2000):** Estimate  $m'(x)$  through the derivative of smoothing splines (*i.e.*, piecewise polynomial curves).
- **Local Polynomial Regression (Fan and Gijbels, 1996):** It solves the weighted least-square problem at each query point  $x$ :<sup>1</sup>

$$\hat{\beta} = \arg \min_{\beta \in \mathbb{R}^{p+1}} \sum_{i=1}^n \left[ Y_i - \sum_{j=0}^p \beta_j (X_i - x)^j \right]^2 K \left( \frac{X_i - x}{h} \right)$$

and estimate  $m'(x)$  as  $\hat{m}'(x) = \hat{\beta}_1$ .

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<sup>1</sup> $K$  is a symmetric kernel and  $h > 0$  is the bandwidth parameter.

We order the data  $\mathcal{D} = \{(X_i, Y_i)\}_{i=1}^n$  according to the increasing order of  $X_i, i = 1, \dots, n$ :

$$Y_i = m(X_{(i)}) + e_i, \quad i = 1, \dots, n.$$

The first-order difference quotients are defined as (Müller et al., 1987; Härdle, 1990):

$$\hat{q}^{(1)}(X_{(i)}) = \frac{Y_i - Y_{i-1}}{X_{(i)} - X_{(i-1)}}, \quad i = 1, \dots, n.$$

Under some conditions, it estimates  $m'(X_{(i)})$  in the asymptotic rate:

$$\hat{q}^{(1)}(X_{(i)}) - m'(X_{(i)}) = O_P(n^2).$$

**Drawback:** The variance of  $\hat{q}^{(1)}(X_{(i)})$  scales *quadratically* with respect to the sample size  $n$ .

To reduce the variance, [Iserles \(2009\)](#); [Charnigo et al. \(2011\)](#) considered

$$\hat{Y}_i^{(1)} \equiv \hat{Y}_i^{(1)}(X_{(i)}) = \sum_{j=1}^k w_j \left( \frac{Y_{i+j} - Y_{i-j}}{X_{(i+j)} - X_{(i-j)}} \right)$$

for  $k+1 \leq i \leq n-k$  and  $k \leq \frac{(n-1)}{2}$ .

- The weights  $w_j, j = 1, \dots, k$  with  $\sum_{j=1}^k w_j = 1$  are chosen to minimize the variance  $\text{Var} \left( \hat{Y}_i^{(1)} | X_{(1)}, \dots, X_{(n)} \right)$ .
- The asymptotic rate of convergence becomes

$$\hat{Y}_i^{(1)} - m'(X_{(i)}) = O\left(\frac{k}{n}\right) + O_P\left(\frac{n^2}{k^3}\right).$$

**Drawback:** It only estimates  $m'$  at the design points  $X_{(1)} \leq \dots \leq X_{(n)}$ .



De Brabanter et al. (2013) proposed using local polynomial regression to smooth out the noisy derivative estimates  $\hat{Y}_i^{(1)}, i = k + 1, \dots, n - k$ .

- Noises  $\tilde{e}_i, i = k + 1, \dots, n - k$  are *correlated* for the noisy derivative estimates

$$\hat{Y}_i^{(1)} = m'(X_{(i)}) + \tilde{e}_i, \quad i = 1, \dots, n.$$

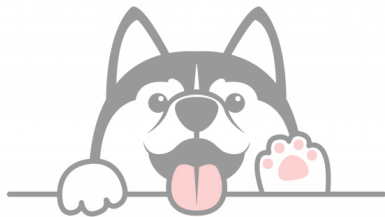
- The kernel  $K$  needs to be bimodal and satisfies  $K(0) = 0$ .

**Drawback:** Their method only works for equispaced design, *i.e.*,

$$X_{(i)} = a + \frac{i \cdot (b - a)}{n - 1}, \quad i = 1, \dots, n.$$

In this paper (Liu and De Brabanter, 2020), the author will extend the above framework to random designs.

# Thank you!



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