

## 9.1 Testing Hypotheses: Uniform Example

$$X \sim Unif(0, \theta)$$

We want to test

$$H_0 : \theta \geq 2$$

$$H_A : \theta < 2$$

with  $\alpha_\theta = 0.10$  based on a random sample of size 8. We'll base our test on  $\hat{\theta}_{MLE} = \max(X_1, \dots, X_8) = X_{(8)}$ . We want to find  $\pi(\theta|\delta)$ .

Reject  $H_0$  if  $X_{(8)} \leq C$  (since we are looking for evidence that  $\theta < 2$ ).

We want  $\Pr(\text{Reject } H_0 | H_0 \text{ true}) = .10$  or  $\Pr(X_{(8)} \leq C | \theta = 2) = .10$

## Finding the pdf for an order statistic

First we need the pdf for  $X_{(8)}$ . Begin with cdf of  $X_{(8)}$ .

$$\begin{aligned}F_{X_{(8)}}(x) &= [Pr(X \leq x)]^8 \\f_{X_{(8)}}(x) &= 8F(x)^{n-1}f(x) \\&= 8 \left(\frac{x}{\theta}\right)^7 \frac{1}{\theta} \\&= \frac{8\theta^7}{\theta^8}\end{aligned}$$

If  $\theta = 2$  then  $f_{X_{(8)}}(x) = \frac{1}{32}x^7$  for  $0 \leq x \leq 2$ .

## Finding the Critical Region

So

$$\begin{aligned}Pr(X_{(8)} \leq C | H_0(\text{true}) = .10 &\rightarrow \int_0^c \frac{1}{32} x^7 dx = .10 \\&\rightarrow \frac{1}{256} x^8 \Big|_0^c \\&\rightarrow c = (256)^{1/8} = 1.50\end{aligned}$$

So the Critical Region (Rejection Region) is  $\{x : x_{(8)} \leq 1.50\}$ .

What is the power of the test when  $\theta = 1.7$ ?

$$Pr(\text{Reject } H_0 | \theta = 1.7) = \int_0^{1.5} \frac{8}{1.7^8} x^7 dx =$$

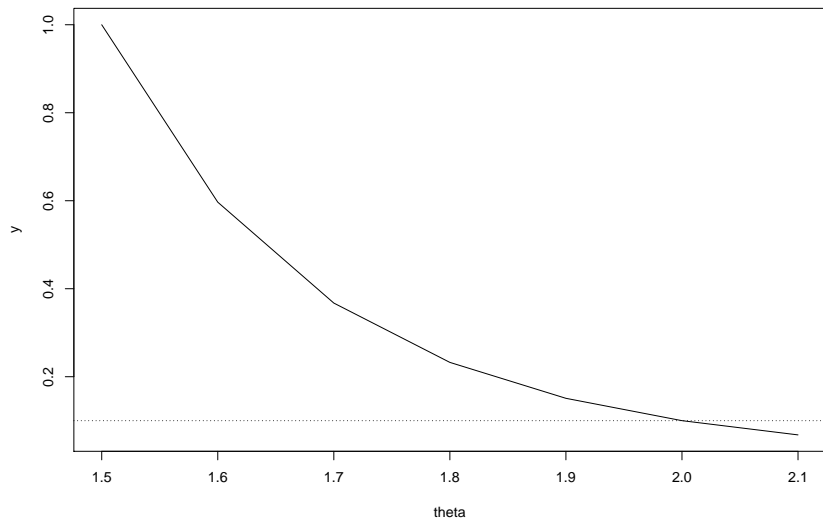
```
## [1] 0.3673996
```

```
## [1] 1.00000000 0.59671947 0.36739962 0.23256804 0.150904
```

```
## [7] 0.06776036
```

$\theta$	Power
1.5	1.00
1.6	0.597
1.7	0.367
1.8	0.233
1.9	0.151
2.0	0.100
2.1	0.068

## Plot of Power vs $\theta$



## Normal Example

$X \sim N(\mu, \sigma^2)$  with  $\sigma^2$  known.

$$H_0 : \mu = \mu_0; H_A : \mu \neq \mu_0$$

$$\pi(\mu|\delta) = Pr(\text{Reject } H_0|\mu)$$

Let  $\delta$  be the test that rejects when  $|Z| > 1.96$

$$\begin{aligned} &= Pr\left(\frac{\bar{X}_n - \mu_0}{\sigma/\sqrt{n}} > 1.96|\mu\right) + Pr\left(\frac{\bar{X}_n - \mu_0}{\sigma/\sqrt{n}} < -1.96|\mu\right) \\ &= Pr(\bar{X} > \mu_0 + 1.96 \times \sigma/\sqrt{n}|\mu) + Pr(\bar{X} < \mu_0 - 1.96 \times \sigma/\sqrt{n}|\mu) \\ &= Pr\left(Z > \frac{\mu_0 + 1.96 \times \sigma/\sqrt{n} - \mu}{\sigma/\sqrt{n}}\right) + Pr\left(Z < \frac{\mu_0 - 1.96 \times \sigma/\sqrt{n} - \mu}{\sigma/\sqrt{n}}\right) \\ &= Pr\left(Z > \frac{\mu_0 - \mu}{\sigma/\sqrt{n}} + 1.96\right) + Pr\left(Z < \frac{\mu_0 - \mu}{\sigma/\sqrt{n}} - 1.96\right) \end{aligned}$$

## Bernoulli Example (Example 9.1.9)

$$X_1, \dots, X_{10} \sim \text{Bernoulli}(p)$$

$$H_0 : p \leq .3$$

$$H_A : p > .3$$

Test statistic:  $Y = \sum X_i \rightarrow \text{Reject } H_0 \text{ if } Y > C \text{ where}$

$$Y \sim \text{Binomial}(n, p); \quad n = 10$$

## Size of Binomial Test

$$\alpha(\delta) = \sup_{p \in \Omega_0} \pi(p|\delta) = \sup_{p \in \Omega_0} Pr(Y \geq c|p) = Pr(Y \geq c|p = 0.3)$$

In general  $Pr(Y = y|p = .3) = \binom{10}{y}.3^y.7^{10-y}$

$$Pr(Y = 10|p = .3) = \binom{10}{10}.3^{10}.7^0 = 0.0000059$$

$$Pr(Y = 9|p = .3) = \binom{10}{9}.3^9.7^1 = 0.00014$$



## Using R

```
y=seq(5,10)
d=dbinom(y,10,.3)
p=1-pbinom(y-1,10,.3)
cbind(y,d,p)
```

##		y	d	p
##	[1,]	5	0.1029193452	0.1502683326
##	[2,]	6	0.0367569090	0.0473489874
##	[3,]	7	0.0090016920	0.0105920784
##	[4,]	8	0.0014467005	0.0015903864
##	[5,]	9	0.0001377810	0.0001436859
##	[6,]	10	0.0000059049	0.0000059049

In order to keep the size of the test to at most .1 (or 0.05 for that matter) we must choose  $c > 5$ .

## P-value

P-value: The smallest level  $\alpha_0$  such that we would reject the null hypothesis at a level  $\alpha_0$  with the observed data.

If  $\delta = \text{Reject } H_0 \text{ when } T \geq c$  ( $\delta_t$  : test that rejects  $H_0$  if  $T \geq t$ ).

$$\text{pvalue} = \sup_{\theta \in \Omega_0} \pi(\theta | \delta_t) = \sup_{\theta \in \Omega_0} \Pr(T \geq t | \theta)$$

### Bernoulli Example

If we observe  $Y = 8$  then

$$\begin{aligned} \text{pvalue} &= \sup_{p \leq .3} \Pr(Y \geq 8 | p) \\ &= \Pr(Y \geq 8 | p = .3) \\ &= 0.0016 \end{aligned}$$

# Equivalence of Hypthesis test and Confidence Intervals

A coefficient  $\gamma$  confidence set can be thought of as a set of  $H_0$ s that would be accepted at significance level  $1 - \gamma$ .

## Example 9.1.15

$X_1, \dots, X_n \sim N(\mu, \sigma^2)$  with  $\mu, \sigma^2$  unknown.

$$H_0 : \mu \leq \mu_0; \{\Omega_0 : \mu \leq \mu_0 \text{ and } \sigma^2 > 0\}$$

$$H_A : \mu > \mu_0; \{\Omega_1 : \mu > \mu_0 \text{ and } \sigma^2 > 0\}$$

Coefficient  $\gamma$  Confidence Interval:

$$\left( \bar{X}_n \pm c\sigma'/\sqrt{n} \right)$$

where  $c$  is the  $\frac{1+\gamma}{2}$  quantile of  $t_{n-1}$  distribution.

## Normal CIs

We can use the CI to find a level  $\alpha_0 = 1 - \gamma$  test of  $H_0 : \mu = \mu_0; H_A : \mu \neq \mu_0$

Reject  $H_0$  if  $\mu_0 \notin (\bar{X}_n \pm c\sigma'/\sqrt{n})$

Which is equivalent to  $(|\bar{X}_n - \mu_0| \geq \pm c\sigma'/\sqrt{n})$

Or

$$\left| \frac{\bar{X}_n - \mu_0}{\sigma'/\sqrt{n}} \right| \geq c$$

Which looks like a t-test!