To calculate the posterior, the following still holds:

$$P(\beta|Y,\alpha,\sigma_{\eta}^2,\sigma_{\varepsilon}^2) = \frac{P(Y,\alpha,\sigma_{\eta}^2,\sigma_{\varepsilon}^2|\beta)P(\beta)}{P(Y,\alpha,\sigma_{\eta}^2,\sigma_{\varepsilon}^2)}$$

and

$$\begin{split} P(Y,\alpha,\sigma_{\eta}^2,\sigma_{\varepsilon}^2|\beta) = & P(y_1,...,y_J,\alpha_1,...,\alpha_J,\sigma_{\eta}^2,\sigma_{\varepsilon}^2|\beta) \\ = & P(y_J|y_1,...,y_{T-1},\alpha_1,...,\alpha_J,\sigma_{\eta}^2,\sigma_{\varepsilon}^2\beta) \\ & \times P(y_1,...,y_{T-1},\alpha_1,...,\alpha_J,\sigma_{\eta}^2,\sigma_{\varepsilon}^2|\beta) \\ = & P(y_J|\alpha_J,\sigma_{\varepsilon}^2\beta)P(\alpha_{T-1}|y_1,...,y_{T-1},\alpha_1,...,\alpha_{T-1},\sigma_{\eta}^2,\sigma_{\varepsilon}^2,\beta) \\ & \times P(y_1,...,y_{T-1},\alpha_1,...,\alpha_{T-1},\sigma_{\eta}^2,\sigma_{\varepsilon}^2|\beta) \\ = & P(y_J|\alpha_J,\sigma_{\varepsilon}^2\beta)P(\alpha_{T-1}|\alpha_{T-1},\sigma_{\eta}^2) \\ & \times P(y_1,...,y_{T-1},\alpha_1,...,\alpha_{T-1},\sigma_{\eta}^2,\sigma_{\varepsilon}^2|\beta) \\ = & P(\sigma_{\varepsilon}^2)P(\sigma_{\eta}^2)\bigg(\prod_{k=0}^T P(\alpha_k|\alpha_{k-1},\sigma_{\eta}^2)\bigg)\prod_{j=1}^T P(y_j|\alpha_j,\sigma_{\varepsilon}^2,\beta) \\ & \propto \prod_{j=1}^T P(y_j|\alpha_j,\sigma_{\varepsilon}^2,\beta) \end{split}$$

For simplicity we can further write,

$$\begin{split} -2logP(Y,\alpha,\sigma_{\eta}^2,\sigma_{\varepsilon}^2|\beta) \propto & \sum_{j=1}^T (y_j - \alpha_j - X_j\beta)' \Sigma_{\varepsilon}(y_j - \alpha_j - X_j\beta) \\ \text{where,} \end{split}$$

$$y_j = \begin{bmatrix} y_{1j} \\ y_{2j} \end{bmatrix}, \quad \alpha_j = \begin{bmatrix} \alpha_{1j} \\ \alpha_{2j} \end{bmatrix}, \quad X_j \beta = \begin{bmatrix} X_j \beta_1 \\ X_j \beta_2 \end{bmatrix}, \quad \Sigma_\varepsilon = \begin{bmatrix} \sigma_{\varepsilon 1}^2 & & & \\ & \ddots & & \\ & & \sigma_{\varepsilon 2}^2 & \\ & & \ddots \end{bmatrix}$$

Because Σ_{ε} is diagnol we can rewrite this portion as,

$$\begin{split} -2logP(Y,\alpha,\sigma_{\eta}^{2},\sigma_{\varepsilon}^{2}|\beta) \propto & \frac{\sum_{j=1}^{T}(y_{j1}-\alpha_{j1}-X_{j}\beta_{1})^{2}}{\sigma_{\varepsilon1}^{2}} + \frac{\sum_{j=1}^{T}(y_{j2}-\alpha_{j2}-X_{j}\beta_{2})^{2}}{\sigma_{\varepsilon2}^{2}} \\ \propto & \frac{\beta_{1}'(\sum_{j=1}^{T}X_{j}'X_{j})\beta_{1} - 2\sum_{j=1}^{T}(y_{j1}-\alpha_{j1})'X_{j}\beta_{1}}{\sigma_{\varepsilon1}^{2}} + \frac{\beta_{2}'(\sum_{j=1}^{T}X_{j}'X_{j})\beta_{2} - 2\sum_{j=1}^{T}(y_{j2}-\alpha_{j2})'X_{j}\beta_{2}}{\sigma_{\varepsilon2}^{2}} \end{split}$$

Similarly for the prior we have,

$$-2logP(\beta) \propto \frac{(\beta_1-\theta_1)^2}{\sigma_\beta^2} + \frac{(\beta_2-\theta_2)^2}{\sigma_\beta^2}$$

So for the posterior of β ,

$$\begin{split} &\frac{\beta_1'(\sum_{j=1}^T X_j' X_j)\beta_1 - 2\sum_{j=1}^T (y_{j1} - \alpha_{j1})' X_j \beta_1}{\sigma_{\varepsilon 1}^2} + \frac{(\beta_1 - \theta_1)^2}{\sigma_{\beta}^2} \\ &+ \frac{\beta_2'(\sum_{j=1}^T X_j' X_j)\beta_2 - 2\sum_{j=1}^T (y_{j2} - \alpha_{j2})' X_j \beta_2}{\sigma_{\varepsilon 2}^2} + \frac{(\beta_2 - \theta_2)^2}{\sigma_{\beta}^2} \end{split}$$

which we have previously seen has the form

$$\frac{(\beta_{1}-{\Sigma_{1}}^{-1}B_{1})'\Sigma_{1}(\beta_{1}-{\Sigma_{1}}^{-1}B_{1})}{\sigma_{\varepsilon_{1}}^{2}\sigma_{\beta}^{2}}+\frac{(\beta_{2}-{\Sigma_{2}}^{-1}B_{2})'\Sigma_{2}(\beta_{2}-{\Sigma_{2}}^{-1}B_{2})}{\sigma_{\varepsilon_{2}}^{2}\sigma_{\beta}^{2}}$$

Where, $B_r = \sigma_{\beta}^2 \left(\sum_{j=1}^T y_{jr} - \alpha_{jr}\right)' X_j - \sigma_{\varepsilon r}^2 \theta_r$ and $\Sigma_r = \left(\sigma_{\beta}^2 \sum_{j=1}^T X_j' X_j\right) + \sigma_{\varepsilon r}^2 I_p$. This means we have the following posterior distribution,

$$\begin{bmatrix} \beta_1 \\ \beta_2 \end{bmatrix} \sim N(\begin{bmatrix} \Sigma_1^{-1}B_1 \\ \Sigma_2^{-1}B_2 \end{bmatrix}, \begin{bmatrix} \sigma_{\varepsilon 1}^2 \sigma_{\beta}^2 \Sigma_1^{-1} & 0 \\ 0 & \sigma_{\varepsilon 1}^2 \sigma_{\beta}^2 \Sigma_2^{-1} \end{bmatrix})$$