

### Exercises for the Stochastic part of Mathematical Data Science WI4231

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$$\begin{aligned} 1a. E((U - U')^2) &= E(U^2 + U'^2 - 2UU') \\ &= E(U^2) + E(U'^2) - E(2UU') \quad (\text{by linearity of expectation}) \\ &= \frac{1}{12} + \left(\frac{1}{2}\right)^2 + \frac{1}{12} + \left(\frac{1}{2}\right)^2 - E(2UU') \quad (\text{using variance of standard uniform}) \\ &= \frac{1}{6} + \frac{1}{2} - 2E(U)E(U') \quad (\text{as } U \text{ and } U' \text{ are independent}) \\ &= \frac{1}{6} + \frac{1}{2} - 2 * \frac{1}{2} * \frac{1}{2} \quad (\text{as } U \text{ and } U' \text{ are independent}) \\ &= \frac{1}{6} \end{aligned}$$

1b. The moments of  $U$  and  $U'$  are given by:

$$E(U) = E(U') = \frac{1}{2}(0 + 1) = \frac{1}{2}$$

$$E(U^2) = E(U'^2) = \frac{1}{3}(0^2 + 0 * 1 + 1) = \frac{1}{3}$$

$$E(U^3) = E(U'^3) = \frac{1}{4}(0 + 1) * (0^2 + 1^2) = \frac{1}{4}$$

$$E(U^4) = E(U'^4) = \frac{1}{5}(0^4 + 0^3 * 1 + 0^2 * 1^2 + 0 * 1^3 + 1^4) = \frac{1}{5}$$

$$\begin{aligned} Var((U - U')^2) &= E((U - U')^4) - (E((U - U')^2))^2 \\ &= E((U - U')^4) - \left(\frac{1}{6}\right)^2 \\ &= E(U^4 - 4U^3U' + 6U^2U'^2 - 4UU'^3 + U'^4) - \frac{1}{36} \\ &= E(U^4) - 4E(U^3)E(U') + 6E(U^2)E(U'^2) - 4E(U)E(U'^3) + E(U'^4) - \frac{1}{36} \\ &= E(U^4) - 4E(U^3)E(U') + 6E(U^2)E(U'^2) - 4E(U)E(U'^3) + E(U'^4) - \frac{1}{36} \\ &= \frac{1}{5} - 4 * \frac{1}{4} * \frac{1}{2} + 6 * \frac{1}{3} * \frac{1}{3} - 4 * \frac{1}{2} * \frac{1}{4} + \frac{1}{5} - \frac{1}{36} \\ &= 0.038888888 \\ &\approx 0.04 \end{aligned}$$

3. Let  $Z$  be a random variable with a standard normal  $N(0,1)$ -distribution. It can be seen that  $P(|Z| \leq z)$  is twice of the normal probability from positive  $z$  and 0 otherwise as distribution is folded over.

For  $z > 0$ ,

$$P(|Z| \geq z) = 2 \int_z^{\infty} e^{-\frac{x^2}{2}} dx$$

$$\begin{aligned}
&= 2 \int_z^\infty (-x) e^{-\frac{x^2}{2}} \left(-\frac{1}{x}\right) dx \\
&= \left[ -\frac{2}{(2\pi)^{0.5} x} e^{-\frac{x^2}{2}} \right]_z^\infty - \frac{2}{(2\pi)^{0.5}} \int_z^\infty x^{-2} e^{-\frac{x^2}{2}} dx \\
&= \left[ -\sqrt{\frac{2}{\pi}} \frac{e^{-\frac{x^2}{2}}}{x} \right]_z^\infty - \sqrt{\frac{2}{\pi}} \int_z^\infty x^{-2} e^{-\frac{x^2}{2}} dx \\
&= -0 + \sqrt{\frac{2}{\pi}} \frac{e^{-\frac{z^2}{2}}}{z} - \sqrt{\frac{2}{\pi}} \int_z^\infty x^{-2} e^{-\frac{x^2}{2}} dx \\
&= \sqrt{\frac{2}{\pi}} \frac{e^{-\frac{z^2}{2}}}{z} - \sqrt{\frac{2}{\pi}} \int_z^\infty x^{-2} e^{-\frac{x^2}{2}} dx \\
&= \sqrt{\frac{2}{\pi}} \frac{e^{-\frac{z^2}{2}}}{z} + \sqrt{\frac{2}{\pi}} \int_z^\infty (-x) x^{-3} e^{-\frac{x^2}{2}} dx \\
&= \sqrt{\frac{2}{\pi}} \frac{e^{-\frac{z^2}{2}}}{z} + \left[ \frac{\sqrt{\frac{2}{\pi}} e^{-\frac{x^2}{2}}}{x^3} \right]_z^\infty + 3 \int_z^\infty x^{-4} e^{-\frac{x^2}{2}} dx \\
&= \sqrt{\frac{2}{\pi}} \frac{e^{-\frac{z^2}{2}}}{z} + 0 - \frac{\sqrt{\frac{2}{\pi}} e^{-\frac{z^2}{2}}}{z^3} + 3 \int_z^\infty x^{-4} e^{-\frac{x^2}{2}} dx \\
&= \sqrt{\frac{2}{\pi}} \frac{e^{-\frac{z^2}{2}}}{z} - \frac{\sqrt{\frac{2}{\pi}} e^{-\frac{z^2}{2}}}{z^3} + 3 \int_z^\infty x^{-4} e^{-\frac{x^2}{2}} dx \\
&= \sqrt{\frac{2}{\pi}} \frac{e^{-\frac{z^2}{2}}}{z} \left(1 - \frac{1}{z^2}\right) + 3 \int_z^\infty x^{-4} e^{-\frac{x^2}{2}} dx
\end{aligned}$$

We see we can continuously integrate the remaining the integral by parts to get smaller order of  $z$ . Hence, we can conclude the following as needed.

$$P(|Z| \geq z) = \sqrt{\frac{2}{\pi}} \frac{e^{-\frac{z^2}{2}}}{z} \left(1 + O\left(\frac{1}{z^2}\right)\right)$$

b. For  $Z_1, \dots, Z_p$  i. i. d with  $N(0,1)$  standard Gaussian distribution and  $\alpha > 0$ , probability one of the random variable,  $Z_i$ , is greater than  $\sqrt{\alpha \log(p)}$  is

$$P(|Z_i| \geq \sqrt{\alpha \log(p)})$$

probability that one of the random variable,  $Z_i$ , is not greater than  $\sqrt{\alpha \log(p)}$  is

$$1 - P(|Z_i| \geq \sqrt{\alpha \log(p)})$$

probability that none of the independent random variables is not greater than  $\sqrt{\alpha \log(p)}$  is

$$\left(1 - P(|Z_i| \geq \sqrt{\alpha \log(p)})\right)^p$$

Hence, when  $p \rightarrow \infty$

$$\begin{aligned}
P\left(\max_{j=1, \dots, p} |Z_j| \geq \sqrt{\alpha \log(p)}\right) &= 1 - \left(1 - P(|Z_1| \geq \sqrt{\alpha \log(p)})\right)^p \\
&= 1 - \left(1 - \sqrt{\frac{2}{\pi}} \frac{e^{-\frac{\sqrt{\alpha \log(p)}^2}{2}}}{\sqrt{\alpha \log(p)}} \left(1 + O\left(\frac{1}{\sqrt{\alpha \log(p)}^2}\right)\right)\right)^p
\end{aligned}$$

$$\begin{aligned}
&= 1 - \left( 1 - \sqrt{\frac{2}{\pi}} \frac{e^{-\frac{\alpha \log(p)}{2}}}{\sqrt{\alpha \log(p)}} (1 + O\left(\frac{1}{\alpha \log(p)}\right)) \right)^p \\
&= 1 - \left( 1 - \sqrt{\frac{2}{\alpha \pi}} \frac{e^{\log(p)^{-\frac{\alpha}{2}}}}{\sqrt{\log(p)}} (1 + O\left(\frac{1}{\log(p)}\right)) \right)^p \\
&= 1 - \left( 1 - \sqrt{\frac{2}{\alpha \pi}} \frac{p^{-\frac{\alpha}{2}}}{\sqrt{\log(p)}} + O\left(\frac{p^{-\frac{\alpha}{2}}}{(\log(p))^{\frac{3}{2}}}\right) \right)^p \\
&= 1 - \left( 1 - \left( \sqrt{\frac{2}{\alpha \pi}} \frac{p^{1-\frac{\alpha}{2}}}{\sqrt{\log(p)}} + O\left(\frac{p^{1-\frac{\alpha}{2}}}{(\log(p))^{\frac{3}{2}}}\right) \right) * \frac{1}{p} \right)^p \\
&= 1 - \exp\left(-\sqrt{\frac{2}{\alpha \pi}} \frac{p^{1-\frac{\alpha}{2}}}{\sqrt{\log(p)}} + O\left(\frac{p^{1-\frac{\alpha}{2}}}{(\log(p))^{\frac{3}{2}}}\right)\right)
\end{aligned}$$