



## Objective Bayesian v3\_20230905

### Motivation

(1) Let the data speak for themselves. (2) Cromwell's rule. 🌙🧀🧑

“... the search of priors with a minimal impact on the corresponding posterior analysis, ...” (p. 627)

Consonni, G., Fouskakakis, D., Liseo, B., & Ntzoufras, I. (2018). Prior distributions for objective Bayesian analysis. *Bayesian Analysis*, 13, 627–679.

**Types of Priors**, no unanimity, often a misnomer (p. 628)

- Noninformative (**Objective**)

e.g., **Jeffreys**,  $\pi(\theta) \propto \sqrt{|\mathcal{I}(\theta)|}$ ;

cannot be directly used on the discrete parameter space;  
(How to interpret?) invariant to transformations of the parameters;  
observations can be independent or dependent.

$B(n, p)$ .  $\pi(p) = \text{Beta}(0.5, 0.5)$ , U-shape, **conjugate**

$\mathcal{N}(\mu, \sigma^2)$ .  $\pi(\mu) \propto 1$

**Stan**: implicit

**JAGS**: `mu ~ dunif(-100, 100)`

$Y = (X - a)/(b - a) \sim \text{Beta}(1, 1)$ , maybe slightly faster

`mu ~ dnorm(0, precision  $\tau = 0.0001$ )`

$$\pi(\sigma^2) \propto \frac{1}{\sigma^2}, \quad \pi(\tau) \propto \frac{1}{\tau}, \quad \pi(\sigma) \propto \frac{1}{\sigma}, \quad \pi(\ln \sigma) \propto 1$$

**Stan**: `target += -log(sigma);`

**JAGS**: `sigma ~ dgamma(0.0001, rate = 0.0001)`

`log.sigma ~ dunif(-10, 10)`

e.g., (Bernardo's) **reference**, maximizing the expected Kullback–Leibler divergence;

observations are independent and identically distributed.

Is there always a reference prior? Is it unique?

- Weakly informative (**Vague**), diffuse

- Informative (**Subjective**)

*Examples:*

Flat: You know nothing, often **improper** (may not integrate or sum to 1)

- Super-vague but proper (not usually recommended):  $\mathcal{N}(0, 1000^2)$

- Weakly informative:  $\mathcal{N}(0, 10^2)$

- Generic weakly informative:  $\mathcal{N}(0, 1)$

- Specific informative (context-based):  $\mathcal{N}(0.4, 0.5^2)$  or a scaling  $0.4 + 0.5 \cdot \mathcal{N}(0, 1)$

An improper prior may lead to an improper posterior.

e.g., Haldane,  $\pi(p) = p^{-1}(1 - p)^{-1}$  for  $B(n, p)$ .

**Default**, e.g., g-priors (and **hyperpriors**) in the *anovaBF* function of the “*BayesFactor*” R package.

$$g \sim \text{Scale-inv-}\chi^2(\nu, h^2), \quad g \sim \text{Inv-Gamma}(\nu/2, \text{scale} = \nu h^2/2), \quad g^{-1} \sim \text{Gamma}(\nu/2, \text{rate} = \nu h^2/2)$$

Prefer to choose priors for biologically meaningful parameters, e.g.,  $p$  instead of  $\text{logit}(p)$  (Meredith, 2021).

**Virtues** for data analysis (p. 629) + **Sensitivity analysis**, Bayesian model checking and model evaluation

Instead of objectivity, think about:

Transparency

Consensus

Impartiality

Correspondence to observable reality

Instead of subjectivity, think about:

Multiple perspectives

Context dependence

Example: mean of population distribution of  $\ln(\text{BVA}_i^{\text{latent}}/50)$ , centered at 0 because the mean of the boundary value analysis (BVA) values in the population should indeed be near 50. We set the prior standard deviation to 0.2 which is close to  $\ln(60/50) \approx 0.18$  to indicate that we are pretty sure the mean is between 40 and 60 (Gelman & Hennig, 2017, p. 979).

**Criteria for Objective Bayesian Model Selection** (p. 640-641)

• Basic

• Predictive matching

Consistency

• Model selection consistency

• Information consistency

• Intrinsic consistency

Invariance

• Measurement invariance

• Group invariance

**DISCUSSIONS**

1. What are the cons of being data-driven?

*Keywords*: quality of data, correlation  $\neq$  causation, overfitting, privacy concerns, thought experiment.

2. Do priors play different roles in i) estimation or **prediction** and ii) model selection or model comparison?

*Keywords*: conformal prediction, multicollinearity, suppression, confounding, prior-hacking.

3. The prior should ideally be chosen before data collection.

"lwr" = mean(data) - 6 \* sd(data), "upr" = mean(data) + 6 \* sd(data)

<https://discourse.mc-stan.org/u/dexterw/activity/topics>

4. Improper priors for estimation may not guarantee proper posterior.

*Keywords*: multivariate, heteroscedasticity, Jeffreys priors for  $\sigma_i$ , MCMC divergence warning.

$$\pi(\mu, \sigma_1^2, \dots, \sigma_a^2) \propto \prod_{i=1}^a \sigma_i^{-2}$$

5. Effect-size parametrization and measurement invariance,  $Y_{ij} = \mu + \sigma_\epsilon(t_i + b_j) + \epsilon_{ij}$ ,  $\epsilon_{ij} \stackrel{\text{i.i.d.}}{\sim} \mathcal{N}(0, \sigma_\epsilon^2)$ .

<https://rpubs.com/sherloconan/1024130>

$$\mathcal{M}_{\text{full}}^* : Y_{ijk} = \mu + \alpha_i + \beta_j + (\alpha\beta)_{ij} + \epsilon_{ijk}^*$$

$$\pi(\mu) \propto 1$$

$$(\alpha_1^*, \dots, \alpha_{a-1}^*) = (\alpha_1, \dots, \alpha_a) \cdot \mathbf{Q}, \quad \mathbf{I}_a - a^{-1} \mathbf{J}_a = \mathbf{Q} \cdot \mathbf{Q}^\top, \quad \dots$$

$$\alpha_i^* | g_A \stackrel{\text{i.i.d.}}{\sim} \mathcal{N}(0, g_A), \quad \beta_j^* | g_B \stackrel{\text{i.i.d.}}{\sim} \mathcal{N}(0, g_B), \quad (\alpha\beta)_{ij}^* | g_{AB} \stackrel{\text{i.i.d.}}{\sim} \mathcal{N}(0, g_{AB})$$

$$g_A \sim \text{Scale-inv-}\chi^2(1, h_A^2), \quad g_B \sim \text{Scale-inv-}\chi^2(1, h_B^2), \quad g_{AB} \sim \text{Scale-inv-}\chi^2(1, h_{AB}^2)$$

$$\epsilon_k = (\epsilon_{11k}^*, \epsilon_{21k}^*, \epsilon_{12k}^*, \epsilon_{22k}^*)^\top \sim \mathcal{N}_4(\mathbf{0}, \mathbf{\Sigma}), \quad \mathbf{\Sigma} \sim \mathcal{W}^{-1}(\mathbf{I}_{4 \times 4}, 5)$$