

# Incremental Causal Effect for Time to Treatment Initialization

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## Background

- Causal questions often involve **time to treatment initialization**, such as when individuals begin vaccination or undergo screening. These decisions are not always deterministic.
- A motivating example from (Bonvini et al., 2023) is about behavioral health services for probationers, in order to reduce their chances of re-arrest. This can be achieved via affecting probationers' *likelihood* of attending services by, for example, providing transportation stipends.
- Standard causal effects (like ATE) often require the **positivity assumption**: everyone must have a chance to receive every treatment level.
- Positivity is often violated in practice (ineligibility, mandatory treatment). **Many traditional causal quantities of interest might be vetoed by absence of positivity alone after checking the data.**
- On the other hand, we introduced the incremental causal effect, bridging this gap by **not requiring positivity**.

## Incremental Intervention

- We are interested in answering the question like, “What would happen to outcomes if the **intensity of treatment initialization** were increased or decreased?”
- This leads to the idea of an **incremental intervention**, which shifts the *intensity* (hazard function) of treatment initialization.
- Let  $\lambda(t|L)$  be the baseline hazard. Intervention shifts it to  $\theta(t, L)\lambda(t|L)$  ( $\theta(t, L) > 0$ ).
- Let  $T(\theta)$  be the time-to-treatment under the shifted hazard.
- **Estimand**: Incremental causal effect  $\psi(\theta) = \mathbb{E}[Y_{T(\theta)}]$  (mean potential outcome  $Y$  under shifted treatment hazard).
- **Key Advantage**: Identification does NOT require positivity.

## Identification

### Assumptions:

- Consistency:  $Y = Y_{T \wedge \tau}$ .
- No Unmeasured Confounding (NUC):  $T \perp Y_t | L$ .

**Theorem 1 (IPW Identification)**: Under Consistency and NUC,

$$\psi(\theta) = \mathbb{E}[W(\theta, T, \Delta, L)Y],$$

where the weight  $W(\theta, T, \Delta, L)$  is:

$$\theta(T, L)^\Delta \exp \left\{ - \int_0^{T \wedge \tau} (\theta(t, L) - 1) d\Lambda(t|L) \right\}$$

( $\Delta = 1(T < \tau)$ ,  $\Lambda(t|L)$  is cumulative hazard).

- **Avoids the need for positivity.**

## Estimation

- 1 Estimate cumulative hazard  $\Lambda(t|L)$  using standard methods (e.g., Cox model), get  $\hat{\Lambda}(t|L)$ .
- 2 Compute plug-in IPW estimator:

$$\hat{\psi}(\theta) = \frac{1}{n} \sum_{i=1}^n \theta(T_i, L_i)^{\Delta_i} \times e^{-\int_0^{T_i \wedge \tau} \{\theta(t, L_i) - 1\} d\hat{\Lambda}(t|L)} Y_i.$$

- 3 Use bootstrap (e.g., multiplier) for confidence intervals.

**Theorem 2 (Consistency and Asymptotic Normality)**: Under Assumptions Consistency, NUC, and some regularity conditions, we have that  $\hat{\psi}(\theta)$  converges to  $\psi(\theta)$  in probability, that is, for any constant  $\varepsilon > 0$ ,

$$\mathbb{P}(|\hat{\psi}(\theta) - \psi(\theta)| > \varepsilon) \rightarrow 0,$$

when  $n \rightarrow \infty$ , and the root- $n$  scaled centered difference  $\sqrt{n}\{\hat{\psi}(\theta) - \psi(\theta)\}$  is asymptotically linear and thus converges to a normal variable weakly, that is,

$$\sqrt{n}\{\hat{\psi}(\theta) - \psi(\theta)\} = \frac{1}{\sqrt{n}} \sum_{i=1}^n \Xi_i + o_P(1) \rightarrow \mathcal{N}(0, \text{Var}(\Xi_1)),$$

for some random variable  $\Xi_1$ , in distribution.

## Simulation Results

Setup:

- 1  $L_i \sim \text{Unif}(0, 1)$ .
- 2  $\mathbb{P}(T_i > t | L_i) = \exp\{-\exp(0.25L_i)t\}$ .
- 3  $Y_i \sim \mathcal{N}(\exp(1 - 1.5L_i - (2 - T_i \wedge 2)), 0.5^2)$ .

$\hat{\Lambda}(t|L)$  from Cox model.

Results for  $n = 1000$ ,  $R = 1000$ :

Table 1: Simulation results of the IPW estimator. We report bias, percent bias (%Bias), standard error of estimate (SEE), average estimated standard deviation (SD) and coverage probability of Wald type 95% confidence intervals (95% CP) of  $\hat{\psi}(\theta)$  by  $B = 200$  multiplier bootstrap, for sample size  $n = 1000$  and  $R = 1000$  Monte Carlo samples.

$\theta(t l) \equiv$	1/2.5	1/2	1/1.5	1.5	2	2.5
$\psi(\theta)$	0.893	0.808	0.694	0.404	0.336	0.297
Bias ( $\times 10^{-2}$ )	-0.388	-0.253	-0.139	0.031	0.043	0.039
%Bias	-0.434	-0.313	-0.201	0.077	0.128	0.132
SEE ( $\times 10^{-2}$ )	3.035	2.778	2.508	2.055	2.033	2.096
SD ( $\times 10^{-2}$ )	2.982	2.745	2.495	2.059	2.032	2.088
95% CP (%)	93.9	93.7	94.0	94.5	95.0	95.2

## Application: Methotrexate for Rheumatoid Arthritis

**Goal**: Evaluate effect of changing Methotrexate (MTX) initialization rate on joint pain ( $Y$ ) at 1 year. Such an analysis can illustrate how varying levels of aggressiveness or conservatism in prescribing MTX might influence average disease progression.

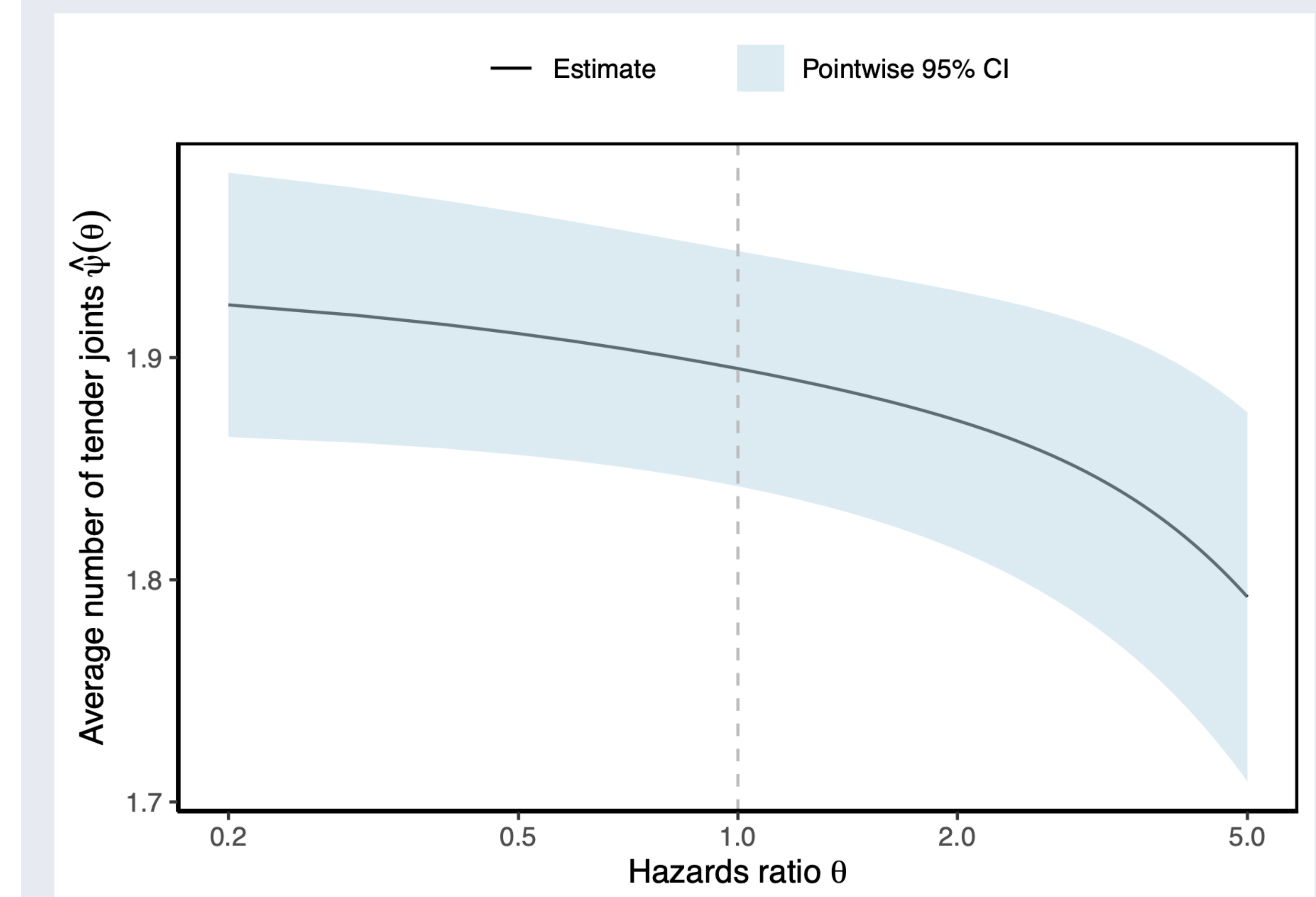


Figure 1: Estimated effect  $\hat{\psi}(\theta)$  vs. hazard multiplier  $\theta$  (constant). Increasing the hazard of starting MTX (higher  $\theta$ ) reduces average joint pain at 1 year.

### Findings:

- Doubling the hazard ( $\theta = 2$ ) reduces avg. joint pain by 1.23%.
- Multiplying hazard by 5 ( $\theta = 5$ ) reduces avg. joint pain by 5.42%.
- Consistent with known protective effect of MTX.

## Conclusion

- Defined and identified **incremental causal effect** for time-to-treatment initialization.
- Key contribution: Identification **without positivity assumption**.
- Provided practical IPW estimation framework using standard survival analysis tools.
- Useful for evaluating policy-style interventions affecting treatment timing.
- Future work: Efficiency improvements (IPW is known to be inefficient), extensions to complex longitudinal settings, relaxing NUC.