Efficient Policy Learning

Machine Learning and Causal Inference, 2017

Based on Athey and Wager (2017)

Drug Facts

Ranitidine 150 mg (as ranitidine hydrochloride 168 mg). Uses · relieves heartburn associated with acid indigestion and sour stomach

Active ingredient (in each tablet)

 prevents heartburn associated with acid indigestion and sour stomach brought on by eating and drinking certain foods or beverages Warnings Allergy alert: Do not use if you are allergic to ranitidine or other acid reducers

Do not use . if you have trouble or pain swallowing food, vomiting with blood, or bloody or black stools. These may be signs of a serious condition. See your doctor.

· with other acid reducers . If you have kidney disease, except under the advice and supervision of a doctor

Ask a doctor before use if you have . had heartburn over 3 months. This may be a sign of a more serious condition. heartburn with lightheadedness, sweating or dizziness

. chest pain or shoulder pain with shortness of breath; sweating; pain spreading to arms, neck or shoulders; or lightheadedness frequent chest pain . frequent wheezing, particularly with heartburn

· unexplained weight loss · nausea or vomiting · stomach pain Stop use and ask a doctor if • your heartburn continues or worsens • you need to take this product for more than 14 days

If pregnant or breast-feeding, ask a health professional before use.

Directions · adults and children 12 years and over:

that cause hearthurn

. children under 12 years; ask a doctor

. to relieve symptoms, swallow 1 tablet with a glass of water . to prevent symptoms, swallow 1 tablet with a glass of water 30 to 60 minutes before eating food or drinking beverages

Keep out of reach of children. In case of overdose, get medical help or contact a Poison Control Center right away.

Purpose

Acid reducer

 can be used up to twice daily (do not take more than 2 tablets in 24 hours) Other information • do not use if printed foil under bottle cap is open or torn • store at 20°-25°C (68°-77°F) · avoid excessive heat or humidity . this product is sodium and sugar free

Inactive inaredients hypromellose, magnesium stearate, microcrystalline cellulose, synthetic red iron oxide, titanium dioxide, triacetin Questions? call 1-888-285-9159 (English/Spanish) M - F, 8:30 - 5 EST, or visit www.zantacotc.com

Stanford Hospital and Clinics: Discharge criteria for post-anesthesia care.

- ightharpoonup Consciousness score: ≥ 1 out of 2.
- ► Respiration score: 2 out of 2.
- ▶ Blood pressure score: ≥ 1 out of 2.
- **...**

Total score must be ≥ 10 out of 12.

Statistical Setup

We want to **learn a policy** π that can be applied in the future:

$$\pi: \mathcal{X} \to \{\pm 1\}, \quad \pi \in \Pi.$$

To do so, we have access to **observational data** collected in the past. In order to predict the effect of policy changes, we need to identify and estimate the **causal effect** of the treatment.

- ▶ We have **i.i.d.** observations $(X_i, Y_i, W_i) \in \mathcal{X} \times \mathbb{R} \times \{\pm 1\}$ for i = 1, ..., n, where W_i is the treatment assignment.
- Following the Neyman-Rubin model, we posit **potential outcomes** $\{Y_i(\pm 1)\}$ corresponding to how *i*-th subject would have responded to different W_i , such that $Y_i = Y_i(W_i)$.
- ▶ To identify treatment effects, we assume **unconfoundedness** (Rosenbaum and Rubin, 1983), $\{Y_i(-1), Y_i(+1)\} \perp W_i \mid X_i$, and **overlap**.

What is Policy Learning?

The **optimal policy** is $\pi^* := \operatorname{argmax}\{\mathbb{E}\left[Y(\pi(X))\right] : \pi \in \Pi\}$, or,

$$\begin{split} \pi^* &= \operatorname{argmax} \left\{ Q(\pi) : \pi \in \Pi \right\}, \\ Q(\pi) &= \mathbb{E} \left[Y(\pi(X)) - \frac{Y(-1) + Y(+1)}{2} \right]. \end{split}$$

In some cases, policy learning reduces to classical statistical tasks:

- If Π has **no structure**, e.g., \mathcal{X} is discrete, Π contains all $2^{|\mathcal{X}|}$ assignments, finding π^* is just **non-parametric regression**.
- If Π is a **doubleton** $\{\pi_+(x) = +1, \pi_-(x) = -1\}$, then $Q(\pi_+)$ is (half) the average treatment effect; finding π^* reduces to **ATE estimation** in observational studies (Hahn, 1998; Heckman et al., 1998; Hirano et al., 2003; Robins et al., 1995; Rosenbaum, 2002; Rubin, 1974; van der Laan and Rose, 2011;...).
- ▶ If Π has **structure**, e.g., if Π consists of linear rules, then... ?

Policy Learning: Take 1

The **Bayes-optimal** policy is clearly

$$\pi_{\mathsf{bayes}}(x) = 1(\{\tau(x) > 0\}).$$

Suggests a simple plug-in strategy for policy learning:

- 1. Estimate the CATE function as $\hat{\tau}(\cdot)$, and then
- 2. Deploy a policy $\hat{\pi}(x) = 1(\{\hat{\tau}(x) > 0\}).$

Given unconfoundedness (Rosenbaum & Rubin, 1983),

$$\{Y_i(0), Y_i(1)\} \perp W_i \mid X_i,$$

there are **numerous methods** available for this (Athey & Imbens, 2016; Hill, 2011; Imai & Ratkovic, 2013; Künzel & al., 2017; Powers & al., 2017; Shalit & al., 2017; W. & Athey, 2017).

This plug-in is strategy is simple (and popular)... but does it solve our problem?

Policy Learning: Take 1

Remarkably, in properly specified **non-parametric setups**, this is optimal, and the best strategies for policy learning take the form

$$\hat{\pi}^*(x) = \mathbf{1}(\{\hat{\tau}(x) > 0\}),$$

where $\hat{\tau}(\cdot)$ is an efficient estimate of $\tau(\cdot)$. In particular,

- ► Manski (2004) considers the case where x is discrete, and studies conditional empirical success rules from an asymptotic perspective.
- ▶ Hirano and Porter (2009) show that, under LeCam local asymptotics where effect sizes shrink as $1/\sqrt{n}$, such thresholding rules are optimal in a broad class of problems.
- **Stoye (2009)** derives **exact minimax** rules when x is discrete and $Y_i \mid X_i = x$ is bounded, and shows that thresholding rules are optimal with matching; with randomization, intriguing small-sample phenomena appear.

Imposing **structure** on Π is essential in many applications (see also Kitagawa & Tetenov, 2015). We use many features with a non-parametric specification to make **unconfoundedness plausible**,

$$\{Y_i(0), Y_i(1)\} \perp W_i \mid X_i.$$

Conversely, the policy $\pi(\cdot)$ must be **implementable in practice**. Features that should not be used in $\pi(\cdot)$ include:

- ▶ Unreliably available features (e.g., collected by specialist).
- ► Gameable features (e.g., self-reported preferences).
- Legally protected classes (e.g., religion, national origin).

Moreover, we may want Π to encode constraints on:

- ► Total budget or marginal subgroup treatment rates (e.g., Bhattacharya and Dupas, 2012).
- Functional form for easier implementation or audit.

We study policy learning in a way that is aware of such constraints.

A First Solution

A natural approach is to optimize an estimated value function,

$$\hat{\pi} = \operatorname{argmax} \left\{ \widehat{Q}(\pi) : \pi \in \Pi \right\}.$$

A simple, **unbiased estimate** of $Q(\pi)$ is (remarkably?) available:

$$1/2 \mathbb{E} \left[\pi(X_i) W_i Y_i / \mathbb{P} \left[W = W_i \mid X = X_i \right] \right]$$

$$= 1/2 \left(\mathbb{E} \left[Y(\pi(X_i)) \right] - \mathbb{E} \left[Y(-\pi(X_i)) \right] \right)$$

$$= \mathbb{E} \left[Y(\pi(X_i)) - \left(Y_i (+1) - Y_i (-1) \right) / 2 \right] = Q(\pi).$$

This insight, along with the induced policy learner,

$$\hat{\pi}_{IPW} = \operatorname{argmax} \left\{ \frac{1}{2} \sum_{i=1}^{n} \frac{\pi(X_i) W_i Y_i}{\mathbb{P}\left[W = W_i \,\middle|\, X = X_i\right]} : \pi \in \Pi \right\},\,$$

has been independently studied across several fields, including statistics (Zhao, Zeng, Rush and Kosorok, 2014), machine learning (Beygelzimer and Langford, 2009; Swaminathan and Joachims, 2015), and economics (Kitagawa and Tetenov, 2015).

Inverse-Propensity Policy Learning: Pros

The inverse-propensity weighted method uses

$$\hat{\pi}_{IPW} = \operatorname{argmax} \left\{ \frac{1}{2} \sum_{i=1}^{n} \frac{\pi(X_i) W_i Y_i}{\mathbb{P}\left[W = W_i \mid X = X_i\right]} : \pi \in \Pi \right\}.$$

In general, the resulting procedure is **consistent**. Moreover:

We can establish policy regret bounds (Kitagawa and Tetenov, 2015; Swaminathan and Joachims, 2015):

$$Q(\pi^*) - Q(\hat{\pi}_{IPW}) = \mathcal{O}_P\left(\frac{\sup\left\{|Y|\right\}}{\inf\left\{\mathbb{P}\left[W = w \mid X
ight]
ight\}} \sqrt{rac{VC(\Pi)}{n}}
ight).$$

Can be implemented as a weighted classification problem:

$$\hat{\pi}_{IPW} = \operatorname{argmax} \left\{ \sum_{i=1}^{n} \pi(X_i) \operatorname{sign}(\Gamma_i) |\Gamma_i| : \pi \in \Pi, \ \Gamma_i := \cdots \right\}.$$

Inverse-Propensity Policy Learning: Cons

The inverse-propensity weighted method uses

$$\hat{\pi}_{IPW} = \operatorname{argmax} \left\{ \frac{1}{2} \sum_{i=1}^{n} \frac{\pi(X_i) W_i Y_i}{\mathbb{P}\left[W = W_i \,\middle|\, X = X_i\right]} : \pi \in \Pi \right\}.$$

In general, the resulting procedure is **consistent**. However:

- ▶ The resulting estimator is not **translation invariant** in Y_i .
- ► The corresponding **regret bounds** are not translation invariant either.
- **>** ...

There are several proposals for improvement, including Dudík et al. (2011), Zhang et al. (2012) and Zhou et al. (2015); however, existing theory gives no guidance on which method to prefer.

► The goal of this talk is to develop an efficiency theory for policy learning.

Statistical Setup, Revisited

We posit **unconfounded** observations via **potential outcomes** $(X_i, Y_i(-1), Y_i(+1), W_i)$, with $Y_i = Y_i(W_i)$, and write

$$\mu_w(x) = \mathbb{E}\left[Y_i(w) \mid X_i = x\right], \quad e_w(x) = \mathbb{P}\left[W_i = w \mid X_i = x\right].$$

Throughout, we will assume that $\mu_w(\cdot)$ and $e(\cdot)$ belong to a **non-parametric class** that allows for $o(n^{-1/4})$ -consistent estimation under L_2 error.

We want to learn a **policy** $\pi: \mathcal{X} \to \{\pm 1\}$ such that $\pi \in \Pi$, where Π is a "simple" class of functions. We will assume that Π has a **finite VC-dimension** or, more generally, a finite entropy integral.

Statistical Setup, Revisited

Considering different function classes for $\mu_w(\cdot)$ and $e(\cdot)$ versus $\pi(\cdot)$ may appear strange, but is essential in many applications.

The functions $\mu_w(\cdot)$ and $e(\cdot)$ need to **describe nature**. Using more pre-treatment features (usually) helps unconfoundedness,

$$\{Y_i(-1), Y_i(+1)\} \perp W_i \mid X_i.$$

Conversely, the policy $\pi(\cdot)$ need to be **implementable in practice**. Features we can use for $\mu_w(\cdot)$ and $e(\cdot)$ but not $\pi(\cdot)$ include:

- Unreliably available features (e.g., collected by specialist).
- ► Gameable features (e.g., self-reported preferences).
- ▶ **Legally protected classes** (e.g., religion, national origin).

Average treatment effect estimation is **policy learning with a doubleton** Π ; now, we let Π be a finite-dimensional continuum.

Efficient Treatment Effect Estimation

Recall that inverse-propensity weighted policy learning uses

$$\widehat{\pi}_{IPW} = \operatorname{argmax} \left\{ \widehat{Q}_{IPW}(\pi) : \pi \in \Pi \right\}, \ \widehat{Q}_{IPW}(\pi) := \frac{1}{2} \sum_{i=1}^{n} \frac{\pi(X_i) W_i Y_i}{\widehat{e}_w(X_i)}.$$

Note that $\widehat{Q}(\pi)$ estimates (half) an average treatment effect,

$$Q(\pi) = \frac{1}{2} \left(\mathbb{E} \left[Y(\pi(X)) \right] - \mathbb{E} \left[Y(-\pi(X)) \right] \right),$$

where "treated" people get policy $\pi(\cdot)$ and controls get $-\pi(\cdot)$. The efficient estimator for $Q(\pi)$ in this setup is well known in the **semiparametric efficiency** literature (Bickel et al., 1998; Hahn, 1998; Hirano et al., 2003; Robins and Rotnitzky, 1995):

$$\widehat{Q}_{DR}(\pi) = \frac{1}{2} \sum_{i=1}^{n} \pi(X_i) \left(\widehat{\mu}_{+}(X_i) - \widehat{\mu}_{-}(X_i) + W_i \frac{Y_i - \widehat{\mu}_{W_i}(X_i)}{\widehat{e}_{W_i}(X_i)} \right).$$

Efficient Treatment Effect Estimation

Let $\hat{\pi}_{DR}$ be the maximizer of \widehat{Q}_{DR} over Π , where

$$\widehat{Q}_{DR}(\pi) = \frac{1}{2} \sum_{i=1}^{n} \pi(X_i) \left(\widehat{\mu}_{+}(X_i) - \widehat{\mu}_{-}(X_i) + W_i \frac{Y_i - \widehat{\mu}_{W_i}(X_i)}{\widehat{e}_{W_i}(X_i)} \right).$$

We can immediately note the following:

For a single, deterministic policy π , we know that

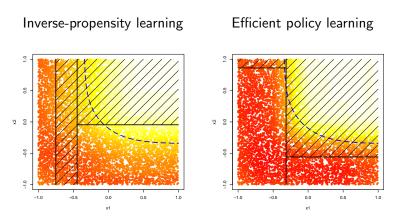
$$\begin{split} &\sqrt{n}\left(\widehat{Q}_{DR}(\pi)-Q(\pi)\right)\Rightarrow\mathcal{N}\left(0,\ V(\pi)\right),\\ &4V(\pi):=\operatorname{Var}\left[\pi(X)\left(\mu_{+}(X)-\mu_{-}(X)\right)\right]\\ &+\mathbb{E}\left[\operatorname{Var}\left[Y(-)\left|X\right|\middle/e_{-}(x)\right]+\mathbb{E}\left[\operatorname{Var}\left[Y(+)\left|X\right|\middle/e_{+}(x)\right], \right. \end{split}$$

and this is the **efficient asymptotic variance**.

Again, can be implemented via **weighted classification**.

What kind of asymptotic regret guarantees can we get?

Simulation Example



Here, we took Π to be the set of **depth-2 decision trees**; the optimal treatment boundary is the blue curve. The colors depict average decisions across many simulations

The **policy regret** of IPW was $2.3 \times$ higher than our method's.

Efficient Treatment Effect Estimation

Theorem. (Athey and Wager, 2017) Let $\hat{\pi}_{DR}$ be the maximizer of \widehat{Q}_{DR} over Π , where

$$\widehat{Q}_{DR}(\pi) = \frac{1}{2} \sum_{i=1}^{n} \pi(X_i) \left(\widehat{\mu}_{+}(X_i) - \widehat{\mu}_{-}(X_i) + W_i \frac{Y_i - \widehat{\mu}_{W_i}(X_i)}{\widehat{e}_{W_i}(X_i)} \right),$$

and let π^* be the best policy in Π . Assume that $\hat{\mu}_{\pm}(\cdot)$ and $\hat{e}(\cdot)$ are estimated via an $o(n^{-1/4})$ -consistent method with cross-fitting.

Then, if Π has a finite VC-dimension, the **policy regret** of $\hat{\pi}_{DR}$ decays as

$$Q(\pi^*) - Q(\hat{\pi}_{DR}) = \mathcal{O}_P\left(\sqrt{V(\pi^*)\log\left(\frac{V_{\mathsf{max}}}{V(\pi^*)}\right)\frac{VC(\Pi)}{n}}\right),$$

where V_{π^*} is the semiparametric **efficient variance** for estimating $Q(\pi^*)$, and V_{max} is a bound for sup $\{V(\pi) : \pi \in \Pi\}$.

Discussion

We found that **policy regret** is controlled as

$$Q(\pi^*) - Q(\hat{\pi}_{DR}) = \mathcal{O}_P\left(\sqrt{V(\pi^*)\log\left(rac{V_{\mathsf{max}}}{V(\pi^*)}
ight)rac{VC(\Pi)}{n}}
ight).$$

If we just has a single policy π , the **optimal confidence intervals** for the improvement of $\pi(\cdot)$ over the opposite policy $-\pi(\cdot)$ scale as

length of conf. interval for
$$Q(\pi) = \mathcal{O}_P\left(\sqrt{V(\pi^*) / n}\right)$$
.

Ignoring constants and log-factors, our regret bounds scale as $\sqrt{VC(\Pi)}$ times the optimal confidence interval length.

Very heuristically, if we think of optimizing over a class of dimensions VC(Π) as picking the best of roughly 2^{VC(Π)} policies, this is essentially the best result we could hope for.