

# 信息学中的概率统计

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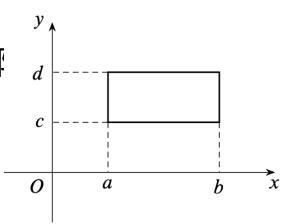
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# 多维连续随机变量

- 1. 多维连续随机变量的分布函数和密度函数
- 2. 多维连续随机变量的独立性
- 3. 多维连续随机变量的特征数
- 4. 多维连续随机变量函数的分布
- 5. 多维高斯分布



- ▶ 给定二维随机变量X,Y和实数x,y,定义 $F(x,y) = P(X \le x,Y \le y)$ 为二维随机变量X,Y的**联合分布函数**
- ▶ 性质1(**有界性**):  $0 \le F(x,y) \le 1, F(-\infty,y) = 0, F(x,-\infty) = 0, F(+\infty,+\infty) = 1$
- ▶ 性质2(**单调性**):  $x_1 < x_2 \Rightarrow F(x_1, y) \le F(x_2, y), y_1 < y_2 \Rightarrow F(x, y_1) \le F(x, y_2)$
- ▶ 性质3(**右连续**): F(x+0,y) = F(x,y), F(x,y+0) = F(x,y)
- ▶ 性质4(**非负性**):  $P(a < X \le b, c < Y \le d) = F(b,d) F(a,d) F(b,c) + F(a,c) \ge 0$
- ▶ F(x,y)满足上述四条性质**等价于**F(x,y)是某个二维
- ▶ 性质4是否被性质1-3蕴含?
  - ▶  $F(x,y) = 1 \, \text{当} x + y \ge 0$ , 否则F(x,y) = 0





- ▶ 给定二维随机变量X,Y,若存在 $f(x,y) \ge 0$ 使得分布函数 F(x,y)可表示为  $F(x,y) = \int_{-\infty}^{x} \int_{-\infty}^{y} f(u,v) du dv$ ,则称二维随机变量X,Y为**二维连续随机变**量, 称f(x,y)为X,Y的**联合密度函数**
- ▶ 联合密度函数的性质
- ▶ 性质1(**非负性**):  $f(x,y) \ge 0$
- ▶ 性质2(**正则性**):  $\int_{-\infty}^{+\infty} \int_{-\infty}^{+\infty} f(u,v) du dv = 1$
- ► 在F(x,y) 偏导数存在的点,  $\frac{\partial^2 F(x,y)}{\partial x \partial y} = f(x,y)$
- ▶ 对于区域G,  $P((X,Y) \in G) = \iint_G f(x,y) dxdy$



- ▶ 例1: X, Y的联合密度函数满足

  - ▶ 否则f(x,y)=0
- ▶ 这里*c*是某个常数
- ▶ 计算常数c, 并计算P(X < 1, Y > 1), P(X > Y)
- ▶ 由正则性 $\int_{-\infty}^{+\infty} \int_{-\infty}^{+\infty} f(x,y) dx dy = c \cdot \frac{1}{2} \cdot \frac{1}{3} = 1 \Rightarrow c = 6$
- $P(X < 1, Y > 1) = \int_0^1 \int_1^{+\infty} 6 \cdot e^{-2x 3y} dx dy = (1 e^{-2})e^{-3}$
- $P(X > Y) = \int_0^{+\infty} \int_0^x 6 \cdot e^{-2x 3y} dy dx = \int_0^{+\infty} 6 \cdot e^{-2x} \cdot \frac{1}{3} \cdot (1 e^{-3x}) dx = \frac{4}{5}$



- ▶ 例2:给定 $\mathbb{R}^2$ 中的一个有界区域 D。 随机变量(X,Y)表示从 D中均匀取一点的坐标。写出X,Y的联合密度函数。
  - ▶  $f(x,y) = 1/S_D$ 若 $(x,y) \in D$
  - ▶ 否则f(x,y) = 0
- ▶ m(X,Y) 服从D上的**二维均匀分布**,记为 $(X,Y) \sim U(D)$



- ▶ 给定二维随机变量X,Y的**联合分布函数**F(x,y),如何计算X的分布函数?
  - $P(X \le x) = P(X \le x, Y < +\infty) = F(x, +\infty)$
- ▶ 定义 $F_X(x) = F(x, +\infty) = \lim_{y \to +\infty} F(x, y)$ 为 X的**边际分布函数**
- ▶ 类似有  $F_Y(y) = F(+\infty, y) = \lim_{x \to +\infty} F(x, y)$ 为 Y的**边际分布函数**
- ▶ 例: X,Y的联合分布函数满足
  - ►  $F(x,y) = 1 e^{-x} e^{-y} + e^{-x-y-\lambda xy} \stackrel{\text{def}}{=} x > 0, y > 0$ ,
  - ▶ 否则F(x,y)=0
- ▶ 计算 $F_X(x)$ 和  $F_Y(y)$ 
  - $F_X(x) = 1 e^{-x} \stackrel{\text{def}}{=} x > 0$
  - $F_Y(y) = 1 e^{-y} \stackrel{\text{def}}{=} y > 0$

- ▶ 给定二维连续随机变量X,Y, f(x,y)为X,Y的联合密度函数
- ▶ 如何计算X的密度函数 $f_X(x)$ ?

$$P(X \le x) = F(x, +\infty) = \int_{-\infty}^{x} \int_{-\infty}^{+\infty} f(u, v) \, dv du = \int_{-\infty}^{x} f_X(u) du$$

- ►  $f_X(x) = \int_{-\infty}^{+\infty} f(x,y) dy$ 为X的**边际密度函数**
- ▶ 类似 $f_Y(y) = \int_{-\infty}^{+\infty} f(x,y) dx$ 为Y的**边际密度函数**
- ▶ 例1: X,Y的联合密度函数满足

  - ▶ 否则f(x,y) = 0
- ▶ 计算X,Y的边际密度函数
  - $f_X(x) = \int_{-\infty}^{+\infty} f(x, y) dy = 2 \cdot e^{-2x}$
  - $f_Y(y) = \int_{-\infty}^{+\infty} f(x, y) dx = 3 \cdot e^{-3y}$



- ►  $f_X(x) = \int_{-\infty}^{+\infty} f(x,y) dy$ 为X的边际密度函数
- ►  $f_Y(y) = \int_{-\infty}^{+\infty} f(x,y) dx$ 为Y的边际密度函数
- ▶ 例2: X,Y的联合密度函数满足
  - ▶ f(x,y) = 1 若 0 < x < 1, |y| < x
  - ▶ 否则f(x,y)=0
- ▶ 计算Y的边际密度函数
  - ▶ ||y| > 1,  $f_Y(y) = 0$
  - ▶  $\dot{\underline{}} = 1 < y < 0$ , -y < x < 1,  $f_Y(y) = \int_{-\infty}^{+\infty} f(x, y) dx = \int_{-\gamma}^{1} f(x, y) dx = 1 + y$
  - ▶  $\triangleq 0 < y < 1$ , y < x < 1,  $f_Y(y) = \int_{-\infty}^{+\infty} f(x, y) dx = \int_{y}^{1} f(x, y) dx = 1 y$

- ▶ 回顾: 若 $P(Y = y_j) > 0$ ,则称 $p_{i|j} = P(X = x_i | Y = y_j) = P(X = x_i, Y = y_i)/P(Y = y_i)$ 为给定  $Y = y_i$ 条件下 X的条件分布列
- ▶ 对于连续随机变量,如何定义条件分布函数和条件密度函数?
- ►  $P(X \le x | Y = y)$ 可定义为  $\lim_{\Delta \to 0} P(X \le x | y \le Y \le y + \Delta)$
- $P(X \le x | Y = y) = \lim_{\Delta \to 0} P(X \le x | y \le Y \le y + \Delta) = \lim_{\Delta \to 0} \frac{P(X \le x, y \le Y \le y + \Delta)}{P(y \le Y \le y + \Delta)}$

- ▶ 当密度函数连续
  - $\lim_{\Delta \to 0} \frac{1}{\Delta} \int_{y}^{y+\Delta} f(u, v) dv = f(u, y)$
  - $\lim_{\Delta \to 0} \frac{1}{\Delta} \int_{y}^{y+\Delta} f_{Y}(v) dv = f_{Y}(y)$
- $P(X \le x | Y = y) = \lim_{\Delta \to 0} \frac{\int_{-\infty}^{x} \left(\frac{1}{\Delta} \int_{y}^{y+\Delta} f(u,v) dv\right) du}{\frac{1}{\Delta} \int_{y}^{y+\Delta} f_{Y}(v) dv} = \frac{\int_{-\infty}^{x} f(u,y) du}{f_{Y}(y)} = \int_{-\infty}^{x} \frac{f(u,y)}{f_{Y}(y)} du$
- ▶  $F(x|y) = P(X \le x|Y = y) = \int_{-\infty}^{x} \frac{f(u,y)}{f_Y(y)} du$  为给定Y = y条件下 X的条件分布函数
- ►  $f(x|y) = \frac{f(x,y)}{f_Y(y)}$ 为给定Y = y条件下 X的**条件密度函数** 
  - ▶ 正则性?

$$\int_{-\infty}^{+\infty} \frac{f(x,y)}{f_{Y(y)}} dx = \frac{\int_{-\infty}^{+\infty} f(x,y) dx}{f_{Y}(y)} = \frac{f_{Y}(y)}{f_{Y}(y)} = 1$$

- ▶ 例: 随机变量X,Y服从单位圆  $(x^2 + y^2 \le 1)$ 上的二维均匀分布。计算f(x|y)
- $f(x|y) = \frac{f(x,y)}{f_Y(y)}$
- ►  $f(x,y) = \frac{1}{\pi}$ 若 (x,y)在单位圆内
- ▶ 否则 f(x,y) = 0

$$f_Y(y) = \int_{-\infty}^{+\infty} f(x, y) dx = \int_{-\sqrt{1 - y^2}}^{+\sqrt{1 - y^2}} f(x, y) dx = \frac{2\sqrt{1 - y^2}}{\pi}$$

- ►  $f(x|y) = \frac{1}{\pi} / \frac{2\sqrt{1-y^2}}{\pi} = \frac{1}{2\sqrt{1-y^2}}$   $= \frac{1}{2\sqrt{1-y^2}}$   $= \frac{1}{2\sqrt{1-y^2}}$
- ▶ 否则 f(x|y) = 0

- ▶ 回顾: 给定二维离散随机变量(X,Y), 若对于任意实数x,y均有  $P(X = x,Y = y) = P(X = x) \cdot P(Y = y)$ , 则称X,Y**相互独立**
- ▶ 给定二维随机变量(X,Y), 分布函数为F(x,y), 边际分布函数为 $F_X(x)$ 和 $F_Y(y)$ 。 若对于任意实数x,y均有  $F(x,y) = F_X(x) \cdot F_Y(y)$ , 则称X,Y**相互独立**
- ▶ 对于离散随机变量,相互独立**等价于**对于任意实数x, y均有  $P(X = x, Y = y) = P(X = x) \cdot P(Y = y)$
- ▶ 对于连续随机变量,相互独立**等价于**对于任意实数x,y均有密度函数  $f(x,y) = f_X(x) \cdot f_Y(y)$

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- ▶ 例1: X, Y的联合密度函数满足

  - ▶ 否则f(x,y) = 0
- ▶ 判断 X, Y 是否相互独立
  - $f_X(x) = 2 \cdot e^{-2x}$
  - $f_{V}(y) = 3 \cdot e^{-3y}$
  - $f(x,y) = f_X(x) \cdot f_Y(y)$

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- ▶ 例2: 令随机变量 X某服务器第一次发生故障的时间, Y表示另一台服务器第一次发生故障的时间。已知则  $X \sim \text{Exp}(\lambda_1)$ ,  $Y \sim \text{Exp}(\lambda_2)$ , 且  $X \hookrightarrow Y$ 相互独立。
- ▶ 计算P(X < Y)</p>

► 
$$P(X < Y) = \int_{x=0}^{+\infty} \int_{y=x}^{+\infty} f(x, y) dx dy = \int_{x=0}^{+\infty} \int_{y=x}^{+\infty} f_X(x) \cdot f_Y(y) dx dy$$

- $f_X(x) \cdot f_Y(y) = \lambda_1 e^{-\lambda_1 x} \cdot \lambda_2 e^{-\lambda_2 y}$
- $P(X < Y) = \int_0^{+\infty} \lambda_1 e^{-\lambda_1 x} \cdot \int_x^{+\infty} \lambda_2 e^{-\lambda_2 y} dy dx = \int_0^{+\infty} \lambda_1 e^{-\lambda_1 x} \cdot P(Y \ge x) dx$
- $P(Y \ge x) = e^{-\lambda_2 x}$
- $P(X < Y) = \int_{x=0}^{+\infty} \lambda_1 e^{-\lambda_1 x} \cdot e^{-\lambda_2 x} dx = \frac{\lambda_1}{\lambda_1 + \lambda_2}$

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- ▶ 二维连续随机变量X,Y的联合密度函数满足
- $f(x,y) = \frac{1}{2\pi\sigma_1\sigma_2\sqrt{1-\rho^2}} \exp\left[-\frac{1}{2(1-\rho^2)} \cdot \left(\frac{(x-\mu_1)^2}{\sigma_1^2} + \frac{(y-\mu_2)^2}{\sigma_2^2} \frac{2\rho(x-\mu_1)(y-\mu_2)}{\sigma_1\sigma_2}\right)\right]$
- ▶ 其中  $\mu_1, \mu_2 \in \mathbb{R}, \sigma_1, \sigma_2 > 0, |\rho| \le 1$
- ▶ 称 X, Y 服从参数为 $\mu_1, \mu_2, \sigma_1, \sigma_2, \rho$  的二维正态 (高斯) 分布
- ▶ 记号:  $X, Y \sim N(\mu_1, \mu_2, \sigma_1^2, \sigma_2^2, \rho)$
- ▶ 验证正则性:  $\int_{-\infty}^{+\infty} \int_{-\infty}^{+\infty} f(x,y) dx dy = 1$
- ▶ 计算边际密度函数
- ▶ 计算条件密度函数
- ▶ 判断 X, Y 是否相互独立

- $f(x,y) = \frac{1}{2\pi\sigma_1\sigma_2\sqrt{1-\rho^2}} \exp\left[-\frac{1}{2(1-\rho^2)} \cdot \left(\frac{(x-\mu_1)^2}{\sigma_1^2} + \frac{(y-\mu_2)^2}{\sigma_2^2} \frac{2\rho(x-\mu_1)(y-\mu_2)}{\sigma_1\sigma_2}\right)\right]$
- ▶ 计算 $\int_{-\infty}^{+\infty} \int_{-\infty}^{+\infty} f(x,y) dx dy$ ?
- ▶ 換元法: 定义u',v', 使得 $\frac{(x-\mu_1)^2}{\sigma_1^2} + \frac{(y-\mu_2)^2}{\sigma_2^2} \frac{2\rho(x-\mu_1)(y-\mu_2)}{\sigma_1\sigma_2} = (u')^2 + (v')^2$ 
  - ▶ 配方得:  $u' = \frac{x-\mu_1}{\sigma_1} \rho \cdot \frac{y-\mu_2}{\sigma_2} \Rightarrow v' = \frac{y-\mu_2}{\sigma_2} \cdot \sqrt{1-\rho^2}$
- ► 若  $u = \frac{\left(\frac{x-\mu_1}{\sigma_1} \rho \cdot \frac{y-\mu_2}{\sigma_2}\right)}{\sqrt{1-\rho^2}}$ ,  $v = \frac{y-\mu_2}{\sigma_2}$ , 则有 $f(x,y) = \frac{1}{2\pi\sigma_1\sigma_2\sqrt{1-\rho^2}} \exp\left[-\frac{1}{2}\left(u^2 + v^2\right)\right]$

$$\left| \frac{\partial(u,v)}{\partial(x,y)} \right| = \begin{vmatrix} \frac{1}{\sigma_1\sqrt{1-\rho^2}} & -\frac{\rho}{\sigma_2\sqrt{1-\rho^2}} \\ 0 & \frac{1}{\sigma_2} \end{vmatrix} = \frac{1}{\sigma_1\sigma_2\sqrt{1-\rho^2}} \Rightarrow \left| \frac{\partial(x,y)}{\partial(u,v)} \right| = \sigma_1\sigma_2\sqrt{1-\rho^2}$$

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# 2. 多维连续随机变量的独立性

$$f(x,y) = \frac{1}{2\pi\sigma_1\sigma_2\sqrt{1-\rho^2}} \exp\left[-\frac{1}{2} (u^2 + v^2)\right]$$

$$\left| \frac{\partial(x,y)}{\partial(u,v)} \right| = \sigma_1 \sigma_2 \sqrt{1 - \rho^2}$$

$$= \iint \frac{1}{2\pi} \exp\left[-\frac{1}{2}(u^2 + v^2)\right] du dv$$

$$= \int_{-\infty}^{-\infty} \frac{1}{\sqrt{2\pi}} e^{-\frac{u^2}{2}} du \cdot \int_{-\infty}^{-\infty} \frac{1}{\sqrt{2\pi}} e^{-\frac{v^2}{2}} dv = 1$$

▶ 思考: 如何从随机变量的角度理解换元  $u = \frac{\left(\frac{x-\mu_1}{\sigma_1} - \rho \cdot \frac{y-\mu_2}{\sigma_2}\right)}{\sqrt{1-\rho^2}}$ ,  $v = \frac{y-\mu_2}{\sigma_2}$ ?

- ▶ 边际密度函数 $f_X(x) = \int_{-\infty}^{+\infty} f(x,y) dy$
- $f(x,y) = \frac{1}{2\pi\sigma_1\sigma_2\sqrt{1-\rho^2}} \exp\left[-\frac{1}{2(1-\rho^2)} \cdot \left(\frac{(x-\mu_1)^2}{\sigma_1^2} + \frac{(y-\mu_2)^2}{\sigma_2^2} \frac{2\rho(x-\mu_1)(y-\mu_2)}{\sigma_1\sigma_2}\right)\right]$
- $f(x,y) = \frac{1}{2\pi\sigma_1\sigma_2\sqrt{1-\rho^2}} \exp\left(-\frac{u^2}{2}\right) \cdot \exp\left(-\frac{v^2}{2}\right)$
- $f_X(x) = \int_{-\infty}^{+\infty} f(x, y) dy = \sqrt{1 \rho^2} \cdot \sigma_2 \cdot \int_{-\infty}^{+\infty} f(x, y) du$
- $f_X(x) = \frac{1}{2\pi\sigma_1} \int_{-\infty}^{+\infty} \exp\left(-\frac{u^2}{2}\right) du \cdot \exp\left(-\frac{v^2}{2}\right) = \frac{1}{\sqrt{2\pi}\cdot\sigma_1} \exp\left(-\frac{(x-\mu_1)^2}{2\sigma_1^2}\right)$
- ► 类似有  $f_Y(y) = \frac{1}{\sqrt{2\pi} \cdot \sigma_2} \exp\left(-\frac{(y-\mu_2)^2}{2\sigma_2^2}\right)$

- - ▶ 边际分布与参数 ρ无关 ⇒ 具有相同边际分布的多维联合分布可以不同
- ▶ 计算条件密度函数  $f(x|y) = \frac{f(x,y)}{f_{Y(y)}}$ 
  - $f(x,y) = \frac{1}{2\pi\sigma_1\sigma_2\sqrt{1-\rho^2}} \exp\left[-\frac{1}{2(1-\rho^2)} \cdot \left(\frac{(x-\mu_1)^2}{\sigma_1^2} + \frac{(y-\mu_2)^2}{\sigma_2^2} \frac{2\rho(x-\mu_1)(y-\mu_2)}{\sigma_1\sigma_2}\right)\right]$
  - $f_Y(y) = \frac{1}{\sqrt{2\pi} \cdot \sigma_2} \exp\left(-\frac{(y \mu_2)^2}{2\sigma_2^2}\right)$
  - $f(x|y) = \frac{1}{\sqrt{2\pi}\sigma_1\sqrt{1-\rho^2}} \exp\left(-\frac{1}{2(1-\rho^2)} \cdot \left(\frac{x-\mu_1}{\sigma_1} \rho \cdot \frac{y-\mu_2}{\sigma_2}\right)^2\right)$
  - $= \frac{1}{\sqrt{2\pi}\sigma_1\sqrt{1-\rho^2}} \exp\left(-\frac{1}{2(1-\rho^2)} \cdot \left(\frac{x-\mu_1-\rho \cdot \frac{\sigma_1}{\sigma_2} \cdot (y-\mu_2)}{\sigma_1}\right)^2\right)$
- ▶ 给定Y = y条件下, X的条件分布服从 $N(\mu_1 + \rho \cdot \frac{\sigma_1}{\sigma_2} \cdot (y \mu_2), \sigma_1^2 (1 \rho^2))$

- ▶ 给定Y = y条件下, X的条件分布服从 $N(\mu_1 + \rho \cdot \frac{\sigma_1}{\sigma_2} \cdot (y \mu_2), \sigma_1^2 (1 \rho^2))$
- ▶ 给定X = x条件下, Y的条件分布服从 $N(\mu_2 + \rho \cdot \frac{\sigma_2}{\sigma_1} \cdot (x \mu_1), \sigma_2^2 (1 \rho^2))$
- ▶ 何时有 *X,Y*相互独立?
  - $f(x|y) = \frac{f(x,y)}{f_Y(y)}$
  - ▶ 若X,Y相互独立,  $f(x,y) = f_X(x) \cdot f_Y(y)$ , 也即 $f(x|y) = f_X(x)$
  - $X \sim N(\mu_1, \sigma_1^2), Y \sim N(\mu_2, \sigma_2^2)$
  - ▶ X, Y相互独立等价于 $\rho = 0$

- ▶ 回顾: 给定离散随机变量X,Y和函数g,Z=g(X,Y)。
- ▶ 定理:  $E(Z) = \sum_{i} \sum_{j} P(X = x_i, Y = y_j) \cdot g(x_i, y_j)$
- ▶ 给定连续随机变量X,Y和函数g,Z=g(X,Y)。如何计算E(Z)?
- ▶ 定理:  $E(Z) = \int_{-\infty}^{+\infty} \int_{-\infty}^{+\infty} f(x,y) \cdot g(x,y) \, dxdy$
- - $E(|X Y|) = \int_0^1 \int_0^1 |x y| dx dy = \frac{1}{3}$

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- ▶ 数学期望的线性性: E(X + Y) = E(X) + E(Y)
  - $\int_{-\infty}^{+\infty} \int_{-\infty}^{+\infty} f(x,y) \cdot (x+y) dx dy$
  - $= \int_{-\infty}^{+\infty} \int_{-\infty}^{+\infty} f(x, y) \cdot x \, dx dy + \int_{-\infty}^{+\infty} \int_{-\infty}^{+\infty} f(x, y) \cdot y \cdot dx dy$
  - $= \int_{-\infty}^{+\infty} x \cdot \int_{-\infty}^{+\infty} f(x, y) \, dy dx + \int_{-\infty}^{+\infty} y \cdot \int_{-\infty}^{+\infty} f(x, y) \cdot dx dy$
  - $= \int_{-\infty}^{+\infty} x \cdot f_X(x) dx + \int_{-\infty}^{+\infty} y \cdot f_Y(y) dy = E(X) + E(Y)$
- ▶ 推广:  $E(X_1 + X_2 + \dots + X_n) = E(X_1) + E(X_2) + \dots + E(X_n)$

- ▶ 定理: 若连续随机变量X和Y相互独立,则有 $E(XY) = E(X) \cdot E(Y)$
- $= \int_{-\infty}^{+\infty} \int_{-\infty}^{+\infty} f_X(x) \cdot x \cdot f_Y(y) \cdot y \, dx dy$
- $= \int_{-\infty}^{+\infty} x \cdot f_X(x) dx \cdot \int_{-\infty}^{+\infty} y \cdot f_Y(y) dx$
- $ightharpoonup = E(X) \cdot E(Y)$
- ▶ 推广: 若连续随机变量 $X_1, X_2, ..., X_n$ 相互独立,则有 $E(X_1X_2 ... X_n) = E(X_1) ...$  $E(X_2) ... E(X_n)$
- ▶ 推论: 若连续随机变量 $X_1, X_2, ..., X_n$ 相互独立,则有 $Var(X_1 \pm X_2 \pm ... \pm X_n) = Var(X_1) + Var(X_2) + ... + Var(X_n)$

- ▶ 例1: 随机向量 $X = (X_1, X_2, ..., X_n)$ 满足 $X_i \sim N(0,1)$
- ▶ 随机变量Y表示X的模长。计算  $E(Y^2)$
- $E(Y^2) = E(\sum_{i=1}^n X_i^2) = \sum_{i=1}^n E(X_i^2) = n$

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- ▶ 例2: 随机向量 $X = (X_1, X_2, ..., X_n)$ 满足 $X_i \sim N(0,1)$ ,且 $X_i$ 相互独立
- ▶ 给定固定向量 $a = (a_1, a_2, ..., a_n)$ 。 令随机变量Y表示 X与a的内积。 计算 E(Y)和 Var(Y)
- $Y = \sum_{i=1}^{n} a_i X_i$
- $\blacktriangleright E(Y) = E(\sum_{i=1}^{n} a_i X_i) = \sum_{i=1}^{n} a_i E(X_i) = 0$

- ▶ 例3:  $n \times n$ 矩阵 A每个元素均服从 N(0,1),且不同元素相互独立
- ▶ 计算  $E(\det(A))$ ,  $E(\operatorname{trace}(A))$ ,  $E(\operatorname{trace}(A^2))$
- $det(A) = \sum_{\sigma} (sgn(\sigma) \cdot \prod_{i=1}^{n} A_{i,\sigma(i)})$
- $E(\det(A)) = \sum_{\sigma} \operatorname{sgn}(\sigma) \cdot E(\prod_{i=1}^{n} A_{i,\sigma(i)}) = \sum_{\sigma} \operatorname{sgn}(\sigma) \cdot \prod_{i=1}^{n} E(A_{i,\sigma(i)}) = 0$
- $\blacktriangleright$  trace $(A) = \sum_{i=1}^{n} A_{i,i}$
- $\blacktriangleright E(\operatorname{trace}(A)) = E(\sum_{i=1}^{n} A_{i,i}) = 0$
- $A_{i,i}^2 = \sum_{j=1}^n A_{i,j} \cdot A_{j,i}$
- $E(A_{i,i}^2) = \sum_{i=1}^n E(A_{i,i} \cdot A_{i,i}) = 1$
- $\blacktriangleright E(\operatorname{trace}(A^2)) = E(\sum_{i=1}^n A_{i,i}^2) = n$

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- ▶ 回顾:
  - ►  $F(x|y) = \int_{-\infty}^{x} \frac{f(u,y)}{f_Y(y)} du$  为给定Y = y条件下 X的条件分布函数
  - ►  $f(x|y) = \frac{f(x,y)}{f_Y(y)}$ 为给定Y = y条件下X的条件密度函数
  - ▶ 对于离散随机变量,  $E(X|Y=y_i) = \sum_i x_i \cdot P(X=x_i \mid Y=y_i)$
- ▶ 对于二维连续随机变量X,Y,**定义条件数学期望**  $E(X|Y=y)=\int_{-\infty}^{+\infty} f(x|y) \cdot x \cdot dx$
- ▶ 回顾: E(E(X|Y)) = E(X)

- ▶ 条件数学期望  $E(X|Y=y) = \int_{-\infty}^{+\infty} f(x|y) \cdot x \cdot dx$
- E(E(X|Y)) = E(X)
- ▶ 例:  $X \sim \text{Exp}(\lambda)$ ,  $Y \sim U(0,X)$ 。计算E(X)和 Var(X)
- $E(Y|X=x) = \frac{x}{2}$
- $E(Y) = E(E(Y|X)) = \frac{E(X)}{2} = \frac{1}{2\lambda}$
- $E(Y^2|X=x) = \frac{x^2}{3}$
- $E(Y^2) = E(E(Y^2|X)) = \frac{E(X^2)}{3} = \frac{2}{3\lambda^2}$

- ▶ 给定随机变量X和Y, 定义X和Y的**协方差**
- ▶ 性质回顾:
- $\operatorname{Cov}(X, X) = \operatorname{Var}(X) = E(X^2) (E(X))^2$
- $ightharpoonup \operatorname{Cov}(X,Y) = \operatorname{Cov}(Y,X)$
- $ightharpoonup \operatorname{Cov}(aX, bY) = ab \cdot \operatorname{Cov}(X, Y)$
- $ightharpoonup Cov(X_1 + X_2, Y) = Cov(X_1, Y) + Cov(X_2, Y)$
- ▶ 若X和Y相互独立,则Cov(X,Y)=0
- $Var(X_1 + X_2 + \dots + X_n) = \sum_i \sum_j Cov(X_i, X_j) = \sum_i Var(X_i) + 2\sum_i \sum_{j < i} Cov(X_i, X_j)$
- ightharpoonup Var(X + Y) = Var(X) + Var(Y) + 2Cov(X, Y)

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- ▶ 例: X,Y的联合密度函数满足

  - ▶ 否则f(x,y)=0
- ▶ 计算Cov(X,Y), Var(X), Var(Y)
- $E(X) = \int_0^1 x \cdot \int_0^2 f(x, y) \cdot dy \, dx = \int_0^1 \frac{2x + 2}{3} \cdot x \cdot dx = \frac{5}{9}$
- $E(X^2) = \int_0^1 x^2 \cdot \int_0^2 f(x, y) \cdot dy \, dx = \int_0^1 \frac{2x + 2}{3} \cdot x^2 \cdot dx = \frac{7}{8}$
- $E(Y) = \int_0^2 y \cdot \int_0^1 f(x, y) \cdot dx \, dy = \int_0^2 \frac{2y+1}{6} \cdot y \cdot dx = \frac{11}{9}$
- $E(Y^2) = \int_0^2 y^2 \cdot \int_0^1 f(x, y) \cdot dx \, dy = \int_0^2 \frac{2y+1}{6} \cdot y^2 \cdot dx = \frac{16}{9}$
- $E(XY) = \int_0^1 x \cdot \int_0^2 y \cdot f(x, y) \cdot dy \, dx = \int_0^1 x \cdot \frac{2x + 8/3}{3} \cdot dx = \frac{2}{3}$
- ►  $Var(X) = E(X^2) (E(X))^2 = \frac{13}{162}, Var(Y) = E(Y^2) (E(Y))^2 = \frac{23}{81}$
- ►  $Cov(X,Y) = E(XY) E(X)E(Y) = -\frac{1}{81}$

### 

- ►  $X, Y \sim N(\mu_1, \mu_2, \sigma_1^2, \sigma_2^2, \rho)$ , 计算Cov(X, Y)
- $E(X) = \mu_1, E(Y) = \mu_2$
- $f(x,y) = \frac{1}{2\pi\sigma_1\sigma_2\sqrt{1-\rho^2}} \exp\left[-\frac{1}{2(1-\rho^2)} \cdot \left(\frac{(x-\mu_1)^2}{\sigma_1^2} + \frac{(y-\mu_2)^2}{\sigma_2^2} \frac{2\rho(x-\mu_1)(y-\mu_2)}{\sigma_1\sigma_2}\right)\right]$
- $u = \frac{\left(\frac{x-\mu_1}{\sigma_1} \rho \cdot \frac{y-\mu_2}{\sigma_2}\right)}{\sqrt{1-\rho^2}}, v = \frac{y-\mu_2}{\sigma_2}$
- $y = \sigma_2 v + \mu_2, x = \left(\sqrt{1 \rho^2} \cdot u + \rho \cdot v\right) \cdot \sigma_1 + \mu_1$
- $(x \mu_1)(y \mu_2) = \sigma_1 \sigma_2 v \left( \sqrt{1 \rho^2} \cdot u + \rho \cdot v \right)$
- $f(x,y) = \frac{1}{2\pi\sigma_1\sigma_2\sqrt{1-\rho^2}} \exp\left[-\frac{1}{2}\left(u^2 + v^2\right)\right], \left|\frac{\partial(x,y)}{\partial(u,v)}\right| = \sigma_1\sigma_2\sqrt{1-\rho^2}$

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$$(x - \mu_1)(y - \mu_2) = \sigma_1 \sigma_2 v \left( \sqrt{1 - \rho^2} \cdot u + \rho \cdot v \right)$$

$$f(x,y) = \frac{1}{2\pi\sigma_1\sigma_2\sqrt{1-\rho^2}} \exp\left[-\frac{1}{2}\left(u^2 + v^2\right)\right], \left|\frac{\partial(x,y)}{\partial(u,v)}\right| = \sigma_1\sigma_2\sqrt{1-\rho^2}$$

$$= \iint \frac{1}{2\pi} \exp\left[-\frac{1}{2} (u^2 + v^2)\right] \cdot (x - \mu_1) \cdot (y - \mu_2) du dv$$

$$= \iint \frac{1}{2\pi} \exp\left[-\frac{1}{2} \left(u^2 + v^2\right)\right] \cdot \sigma_1 \sigma_2 v \left(\sqrt{1 - \rho^2} \cdot u + \rho \cdot v\right) \cdot du dv$$

$$= \sigma_1 \sigma_2 \iint \frac{1}{2\pi} \exp\left[-\frac{1}{2} \left(u^2 + v^2\right)\right] \cdot v\left(\sqrt{1 - \rho^2} \cdot u + \rho \cdot v\right) \cdot du dv$$

$$\iint \frac{1}{2\pi} \exp\left[-\frac{1}{2}\left(u^2 + v^2\right)\right] \cdot vududv = 0$$

$$\iint \frac{1}{2\pi} \exp\left[-\frac{1}{2}\left(u^2 + v^2\right)\right] \cdot v^2 du dv = 1$$

- ▶ 给定随机变量X和Y,若 $\sigma(X)$ ,  $\sigma(Y) > 0$ ,定义X和Y的**相关系数**  $Corr(X,Y) = \frac{Cov(X,Y)}{\sigma(X)\sigma(Y)}$
- ▶ 回顾: Cov(X,Y) = E((X-E(X))(Y-E(Y)))
- ▶ 回顾:  $\tilde{X} = \frac{X E(X)}{\sigma(X)}$ 为X的标准化随机变量
- - ▶ Corr(X,Y) > 0 (或Cov(X,Y) > 0) : X和Y**正相关**
  - ▶ Corr(X,Y) < 0 (或Cov(X,Y) < 0) : X和Y**负相关**
  - ▶ Corr(X,Y) = 0 (或Cov(X,Y) = 0) : X和Y**不相关**
- ▶ 若相互独立,一定有不相关
  - $E\left(\left(X E(X)\right)\left(Y E(Y)\right)\right) = E\left(X E(X)\right) \cdot E\left(Y E(Y)\right) = 0$
- ▶ 不相关,是否一定有相互独立?

# 3. 多维连续随机变量的特征数

- ▶ 给定随机变量X和Y,若 $\sigma(X)$ ,  $\sigma(Y) > 0$ ,定义X和Y的**相关系数**  $Corr(X,Y) = \frac{Cov(X,Y)}{\sigma(X)\sigma(Y)}$
- ▶ 性质1: |Corr(*X*, *Y*)| ≤ 1
- ▶ 证明:

$$g(t) \ge 0, \sigma(X), \sigma(Y) > 0 \Rightarrow \left(2\operatorname{Cov}(X, Y)\right)^2 - 4\sigma(X)^2\sigma(Y)^2 \le 0$$

▶ 性质2:  $Corr(X,Y) = \pm 1$ 当且仅当存在 $a \neq 0$ 与b, P(Y = aX + b) = 1

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- ▶ 例1: X, Y的联合密度函数满足

  - ▶ 否则f(x,y)=0
- ▶ 计算Corr(X,Y)并判断相关性
- ►  $Var(X) = E(X^2) (E(X))^2 = \frac{13}{162}$ ,  $Var(Y) = E(Y^2) (E(Y))^2 = \frac{23}{81}$
- ►  $Cov(X,Y) = E(XY) E(X)E(Y) = -\frac{1}{81}$
- ►  $Corr(X,Y) = \frac{Cov(X,Y)}{\sigma(X)\sigma(Y)} = -\frac{\frac{1}{81}}{\sqrt{\frac{13}{162}} \cdot \sqrt{\frac{23}{81}}} = -\sqrt{\frac{2}{299}}$ , 负相关

# 3. 多维连续随机变量的特征数

- ▶ 例2:  $X,Y \sim N(\mu_1,\mu_2,\sigma_1^2,\sigma_2^2,\rho)$ 证明 X,Y相互独立当且仅当 X,Y 不相关
- ▶ X, Y相互独立等价于  $\rho = 0$
- ▶ X, Y不相关等价于 $\rho = 0$

- ▶ 给定连续随机变量X,Y和函数g(x,y), 求Z = g(X,Y)的概率密度函数
- ▶ **卷积公式**: 若X,Y相互独立, Z = X + Y, 则  $f_Z(z) = \int_{-\infty}^{+\infty} f_X(z y) f_Y(y) dy$
- ▶ 证明:
- $P(Z \le z) = \iint_{x+y \le z} f_X(x) f_Y(y) dx dy = \int_{-\infty}^{+\infty} \int_{-\infty}^{z-y} f_X(x) dx \cdot f_Y(y) dy$
- ▶  $P(Z \le Z) = \int_{-\infty}^{+\infty} P(X \le Z y) \cdot f_Y(y) dy$ , 两边对Z求导

- ▶ 例1:  $X \sim N(0, \sigma_1), Y \sim N(0, \sigma_2), X, Y$ 相互独立,求Z = X + Y的概率密度函数
- $f_Z(z) = \int_{-\infty}^{+\infty} f_X(z y) f_Y(y) dy = \frac{1}{2\pi\sigma_1\sigma_2} \int_{-\infty}^{+\infty} \exp\left(-\frac{1}{2} \left(\frac{(z y)^2}{\sigma_1^2} + \frac{y^2}{\sigma_2^2}\right)\right) dy$

- $f_Z(z) = \frac{1}{2\pi\sigma_1\sigma_2} \exp\left(-\frac{1}{2} \cdot \frac{z^2}{\sigma_1^2 + \sigma_2^2}\right) \cdot \int_{-\infty}^{+\infty} \exp\left(-\frac{A}{2}\left(y \frac{z}{\sigma_1^2 A}\right)^2\right) dy$

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- ▶ 例1:  $X \sim N(0, \sigma_1), Y \sim N(0, \sigma_2), X, Y$ 相互独立,求Z = X + Y的概率密度函数
- $f_Z(z) = \frac{1}{2\pi\sigma_1\sigma_2} \exp\left(-\frac{1}{2} \cdot \frac{z^2}{\sigma_1^2 + \sigma_2^2}\right) \cdot \int_{-\infty}^{+\infty} \exp\left(-\frac{A}{2}\left(y \frac{z}{\sigma_1^2 A}\right)^2\right) dy$
- $f_Z(z) = \frac{1}{2\pi\sigma_1\sigma_2} \exp\left(-\frac{1}{2} \cdot \frac{z^2}{\sigma_1^2 + \sigma_2^2}\right) \cdot \sqrt{2\pi} \cdot \sqrt{1/A}$
- $A = \frac{1}{\sigma_1^2} + \frac{1}{\sigma_2^2} \Rightarrow f_Z(z) = \frac{1}{\sqrt{2\pi} \cdot \sqrt{\sigma_1^2 + \sigma_2^2}} \exp\left(-\frac{1}{2} \cdot \frac{z^2}{\sigma_1^2 + \sigma_2^2}\right)$
- ▶ 也即  $Z \sim N(0, \sigma_1^2 + \sigma_2^2)$

- ▶ 推广:  $X_i \sim N(\mu_i, \sigma_i)$ , 且相互独立,则 $\sum_{i=1}^n a_i X_i \sim N(\mu_0, \sigma_0^2)$
- $\blacktriangleright \mu_0 = \sum_{i=1}^n a_i \mu_i$  ,  $\sigma_0^2 = \sum_{i=1}^n a_i^2 \sigma_i^2$
- ▶ 特别有,若 $X_i$ 独立同分布,且 $X_i \sim N(0,1)$ ,则 $\sum_{i=1}^n a_i X_i \sim N(0,|a|^2)$ 
  - $\implies a_i = \frac{1}{n}, \quad \frac{1}{n} \sum_{i=1}^n X_i \sim N\left(0, \frac{1}{n}\right)$

# 4. 多维连续随机变量函数的分布

▶ 例2:  $X \sim \text{Exp}(1)$ ,  $Y \sim \text{Exp}(1)$ , X, Y相互独立,  $\bar{X}Z = X + Y$ 的概率密度函数

$$f_Z(z) = \int_{-\infty}^{+\infty} f_X(z - y) f_Y(y) dy = \int_0^z e^{-(z - y) - y} dy = z e^{-z}$$

- ▶ 也即 Z ~ Γ(2,1)
- ▶ 回顾:对于 $\alpha,\lambda>0$ ,定义概率密度函数
  - $f(x) = \frac{\lambda^{\alpha}}{\Gamma(\alpha)} x^{\alpha 1} e^{-\lambda x}, \quad \exists x \ge 0$
- ▶ 推广:  $X \sim \Gamma(\alpha_1, \lambda), Y \sim \Gamma(\alpha_2, \lambda), X, Y$ 相互独立,  $X + Y \sim \Gamma(\alpha_1 + \alpha_2, \lambda)$

- ▶  $X \sim U(0,1), Y \sim U(0,1), X, Y$ 相互独立,求 $Z = \max\{X,Y\}$ 的概率密度函数
  - $P(Z \le z) = P(X \le z)P(Y \le z) = z^2$
  - ►  $f_Z(z) = 2z \not\equiv z \in (0,1)$
- ▶  $X \sim U(0,1), Y \sim U(0,1), X, Y$ 相互独立,求 $Z = \min\{X,Y\}$ 的概率密度函数
  - ►  $P(Z \ge z) = P(X \ge z)P(Y \ge z) = (1 z)^2$
  - $P(Z \le z) = 1 (1 z)^2$
  - ►  $f_Z(z) = 2(1-z) \stackrel{\text{def}}{=} z \in (0,1)$

- ▶ 若 $X_1, X_2, ..., X_n$ 相互独立,则 $Y = \max\{X_1, X_2, ..., X_n\}$ 的分布函数F(y)满足
  - $F_Y(y) = F_{X_1}(y) \cdot F_{X_2}(y) \cdot \cdots \cdot F_{X_n}(y)$
- ▶ 若  $X_1, X_2, ..., X_n$ 相互独立,则 $Y = \min\{X_1, X_2, ..., X_n\}$ 的分布函数F(y)满足
  - $F_Y(y) = 1 (1 F_{X_1}(y)) \cdot (1 F_{X_2}(y)) \cdot \dots \cdot (1 F_{X_n}(y))$

- ▶ 回顾: 设X为连续随机变量, 若函数y = g(x)严格单调, 其反函数h(y)有连续导数, 则Y = g(X)的概率密度函数为
  - ►  $f_Y(y) = f_X(h(y)) \cdot |h'(y)| \stackrel{\text{def}}{=} y \in (\alpha, \beta)$
  - ►  $f_Y(y) = 0 \stackrel{\text{def}}{=} y \notin (\alpha, \beta)$
- ▶ 若连续随机变量X,Y的联合密度函数为f(x,y)。函数u = u(x,y),v = v(x,y)有连续偏导数且x = x(u,v),y = y(u,v)为唯一的反函数
- ▶ 则U = u(X,Y), V = v(X,Y)的联合密度函数为  $f(x(u,v),y(u,v))\cdot |J|$ , 其中

- $u = x + y, v = x y \Rightarrow x = \frac{u + v}{2}, y = \frac{u v}{2}$

- ▶ 也即  $U \sim N(2\mu, 2), V \sim N(0, 2), 且U, V$ 相互独立

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- ▶ 例2: 若连续随机变量X,Y相互独立, 计算U = XY的概率密度函数
- $\blacktriangleright u = xy, v = y \Rightarrow x = u/v, y = v$

- ▶ U,V的联合密度函数为  $f\left(\frac{u}{v},v\right)\cdot |J| = f_X\left(\frac{u}{v}\right)\cdot f_Y(v)\cdot \frac{1}{|v|}$
- ▶ U的边际密度函数为  $\int_{-\infty}^{+\infty} f_X\left(\frac{u}{v}\right) \cdot f_Y(v) \cdot \frac{1}{|v|} \cdot dv$

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- ▶ 例2: 若连续随机变量X,Y相互独立, 计算U = XY的概率密度函数
- ▶ U的概率密度函数为  $\int_{-\infty}^{+\infty} f_X\left(\frac{u}{v}\right) \cdot f_Y(v) \cdot \frac{1}{|v|} \cdot dv$
- ▶ 验证:
- $P(U \le u) = \iint_{xy \le u} f_X(x) f_Y(y) \, dx dy$
- $= \int_0^{+\infty} \int_{-\infty}^{u/y} f_X(x) dx \cdot f_Y(y) dy + \int_{-\infty}^0 \int_{u/y}^{+\infty} f_X(x) dx \cdot f_Y(y) dy$
- ▶ 対u求导,  $f_U(u) = \int_0^{+\infty} f_X(u/y) \cdot f_Y(y) \cdot \frac{1}{y} \cdot dy \int_{-\infty}^0 f_X\left(\frac{u}{y}\right) \cdot f_Y(y) \cdot \frac{1}{y} \cdot dy$
- $= \int_{-\infty}^{+\infty} f_X(u/y) \cdot f_Y(y) \cdot \frac{1}{|y|} \cdot dy$

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### 5. 多维高斯分布

▶ 随机向量 $X = (X_1, X_2, ..., X_n)$ 。 定义 $E(X) = (E(X_1), E(X_2, ), ..., E(X_n))$ 为 X的数学期望向量,  $Cov(X) = E(X - E(X))(X - E(X))^T$ 为X的协方差矩阵

$$\triangleright \operatorname{Cov}(\boldsymbol{X}) = \begin{pmatrix} \operatorname{Var}(X_1) & \operatorname{Cov}(X_1, X_2) & \dots & \operatorname{Cov}(X_1, X_n) \\ \operatorname{Cov}(X_2, X_1) & \operatorname{Var}(X_2) & \dots & \operatorname{Cov}(X_2, X_n) \\ \dots & \dots & \dots & \dots \\ \operatorname{Cov}(X_n, X_1) & \operatorname{Cov}(X_n, X_2) & \dots & \operatorname{Var}(X_n) \end{pmatrix}$$

▶ 性质: Cov(X)是半正定的对阵矩阵

$$\qquad \qquad \boldsymbol{\alpha}^T \operatorname{Cov}(\boldsymbol{X}) \boldsymbol{\alpha} = E\left(\left(\boldsymbol{\alpha}^T \big(\boldsymbol{X} - \operatorname{E}(\boldsymbol{X})\big)\right)^2\right) \geq 0$$

# 5. 多维高斯分布

- ▶  $X_1, X_2 \sim N(\mu_1, \mu_2, \sigma_1^2, \sigma_2^2, \rho)$ , 计算协方差矩阵和其逆矩阵
  - ▶ 协方差矩阵为 $B = \begin{pmatrix} \sigma_1^2 & \rho \sigma_1 \sigma_2 \\ \rho \sigma_1 \sigma_2 & \sigma_2^2 \end{pmatrix}$
  - $\rightarrow$  det(B) =  $(1 \rho^2)\sigma_1^2\sigma_2^2$
  - 逆矩阵为  $B^{-1} = \frac{1}{1-\rho^2} \cdot \begin{pmatrix} \frac{1}{\sigma_1^2} & -\frac{\rho}{\sigma_1 \sigma_2} \\ -\frac{\rho}{\sigma_1 \sigma_2} & \frac{1}{\sigma_2^2} \end{pmatrix}$
- $f(x) = \frac{1}{2\pi\sigma_1\sigma_2\sqrt{1-\rho^2}} \exp\left[-\frac{1}{2(1-\rho^2)} \cdot \left(\frac{(x_1-\mu_1)^2}{\sigma_1^2} + \frac{(x_2-\mu_2)^2}{\sigma_2^2} \frac{2\rho(x_1-\mu_1)(x_2-\mu_2)}{\sigma_1\sigma_2}\right)\right]$
- $= \frac{1}{2\pi \cdot (\det(B))^{\frac{1}{2}}} \exp\left(-\frac{1}{2} \cdot (\boldsymbol{x} \boldsymbol{\mu})^T B^{-1} (\boldsymbol{x} \boldsymbol{\mu})\right)$

### 5. 多维高斯分布

► n维高斯分布的联合密度函数为

$$f(x) = \frac{1}{(2\pi)^{\frac{n}{2}} \cdot (\det(B))^{\frac{1}{2}}} \exp\left(-\frac{1}{2} \cdot (x - \mu)^T B^{-1} (x - \mu)\right)$$

- ▶ 数学期望向量:  $\mu \in \mathbb{R}^n$
- ▶ 协方差矩阵:  $B \in \mathbb{R}^{n \times n}$
- ▶ 记号:  $X \sim N(\mu, B)$
- $f(x) = (2\pi)^{-\frac{n}{2}} \cdot \exp\left(-\frac{1}{2}|x-\mu|^2\right) = \prod_{i=1}^n \frac{1}{\sqrt{2\pi}} \cdot \exp\left(-\frac{(x_i \mu_i)^2}{2}\right)$
- ▶ 也即  $X_1, X_2, ..., X_n$ 相互独立且  $X_i \sim N(0,1)$

# 5. 多维高斯分布

► n维高斯分布的联合密度函数为

$$f(\mathbf{x}) = \frac{1}{(2\pi)^{\frac{n}{2}} \cdot \left(\det(B)\right)^{\frac{1}{2}}} \exp\left(-\frac{1}{2} \cdot (\mathbf{x} - \boldsymbol{\mu})^T B^{-1} (\mathbf{x} - \boldsymbol{\mu})\right)$$

### ▶ 性质:

- ▶ 边际分布  $X_i \sim N(\mu_i, B_{i,i})$
- ▶  $X_1, X_2, ..., X_n$ 相互独立等价于  $X_1, X_2, ..., X_n$ 两两不相关,也即B为对角矩阵
- ▶  $X \sim N(0, I_n)$ , U为正交矩阵, 则 $UX \sim N(0, I_n)$
- $X \sim N(\boldsymbol{\mu}, B) \Rightarrow B^{-\frac{1}{2}}(X \mu) \sim N(0, I_n)$

### 5. 多维高斯分布

- ▶ n维高斯分布的联合密度函数为
- $f(x) = \frac{1}{(2\pi)^{\frac{n}{2}} \cdot (\det(B))^{\frac{1}{2}}} \exp\left(-\frac{1}{2} \cdot (x \mu)^T B^{-1} (x \mu)\right)$
- $\blacktriangleright X \sim N(\mu, B), \ \ \diamondsuit Y = AX + b$
- $f_{Y}(y) = f_{X}(A^{-1}(y b)) \cdot \left| \frac{\partial x}{\partial y} \right| = f_{X}(A^{-1}(y b)) \cdot \left| \frac{\partial y}{\partial x} \right|^{-1} = f_{X}(A^{-1}(y b)) \frac{1}{|\det(A)|}$
- $= \frac{1}{(2\pi)^{\frac{n}{2}} \cdot (\det(B))^{\frac{1}{2}}} \exp\left(-\frac{1}{2} \cdot (A^{-1}(y-b) \mu)^T B^{-1}(A^{-1}(y-b) \mu)\right) \cdot \frac{1}{|\det(A)|}$
- $= \frac{1}{(2\pi)^{\frac{n}{2}} \cdot \left(\det(ABA^T)\right)^{\frac{1}{2}}} \exp\left(-\frac{1}{2} \cdot (\boldsymbol{y} A\boldsymbol{\mu} \boldsymbol{b})^T (ABA^T)^{-1} (\boldsymbol{y} A\boldsymbol{\mu} \boldsymbol{b})\right)$