

Final BUEC 333. Version D.

August 11, 2014, 12:00-15:00

Questions

1. Consider the example used in the chapter on panel data. We have a panel data set on $n = 48$ U.S. states during $T = 7$ periods, from 1982 up to and including 1988. The total number of observations is 336.
 - (a) Is this a balanced panel? Explain. **Yes: 7*48=336, so there are not gaps.**
 - (b) For each state, in each time period, let Y_{it} denote the number of annual traffic deaths per 10000 in the population. Let X_{it} denote the beer tax in 1988 U.S. dollars. Temporarily ignore the data after 1982, so that we have a cross-section of 48 states. The estimated regression line gives

$$\hat{Y}_{i,1982} = 2.01 + 0.13X_{i,1982}.$$

Alternatively, we can use fixed effects regression to estimate the effect fixed effects regression line is

$$\hat{Y}_{i,t} = \hat{\alpha}_i - 0.66X_{i,t}. \quad (0.29)$$

Do you think that the Least Square assumptions hold, i.e. do you believe that the 0.13 in the first result comes from an unbiased estimator? If YES: explain what causes the difference between 0.13 and -0.66. If NO: explain why the Least Square assumptions are unlikely to hold. Include in your answer: “tax on beer”. **They do not hold: there are omitted variables that the FE accounts for.**

2. [Based on SW, 14.7] Suppose that Y_t follows the stationary AR(1) model

$$Y_t = 3.9 + 0.5Y_{t-1} + u_t$$

where u_t is i.i.d. with $E(u_i) = 0$ and $\text{var}(u_i) = 9$.

- (a) Compute the mean and variance of Y_t . **$E(Y_t) = 3.9/(1-0.5)=7.8$. $\text{Var}(Y_t) = 9 / (3/4) = 12$**
 - (b) Compute the first autocovariance of Y_t . **$\text{Cov}(Y_t, Y_{t-1})=0.5*12=6$**
 - (c) Compute the second autocovariance of Y_t . **$\text{Cov}(Y_t, Y_{t-2})=0.5^2*12 = 3$**
3. [Stock and Watson, 2.5] “In September, Seattle’s daily high temperature has a mean of 59 degrees Fahrenheit. The standard deviation is 9 degrees Fahrenheit.” Remember that, to convert from degrees Fahrenheit to degrees Celsius, we need to subtract 32 and then multiply by 5/9, so

$$T_C = \frac{5}{9} \times (T_F - 32).$$

For the questions that follow, indicate which formulas you are using.

- (a) What is the mean of Seattle’s daily high temperature in September in degrees Celsius? **$5/9*(59-32)=15$**
 - (b) What is the standard deviation of Seattle’s daily high temperature in September in degrees Celsius? **$5/9*9=5$**
 - (c) What is the variance of Seattle’s daily high temperature in September in degrees Celsius? **$5*5=25$**
4. Suppose you want to estimate the population mean of Y , $E(Y) = \mu_Y$. You have a random sample of size n , $\{Y_1, \dots, Y_n\}$. For simplicity, assume that $n = 2$. Then, the sample average $\bar{Y} = \frac{1}{n} \sum_{i=1}^n Y_i$, simplifies to $\bar{Y} = (Y_1 + Y_2)/2$.
 - (a) Is the sample average \bar{Y} unbiased for μ_Y ? Explain. **$E(1/2(Y_1+Y_2))=1/2 E(Y_1+Y_2) = 1/2 * (E(Y_1) + E(Y_2)) = 1/2 * 2 \mu = \mu$**

- (b) Now, consider $\tilde{Y} = \frac{1}{4}Y_1 + \frac{1}{4}Y_2$. What is the variance of \tilde{Y} ? On the basis of the variances, do you prefer \bar{Y} or \tilde{Y} ? **$1/8 \sigma_Y^2$, which is smaller than $\text{var}(\bar{Y})$, so you prefer \tilde{Y}**
- (c) What is wrong with \tilde{Y} ? **It is biased**
- (d) Now, consider the estimator $\tilde{Y} = (\mu_Y + \bar{Y})/2$. Why is this not a good estimator? (Hint: the answer has nothing to do with efficiency, unbiasedness, or consistency.) **The estimator is based on the unknown quantity that you are trying to estimate!**
5. Consider the following estimated regression equation that describes the relationship between a student's weight and height:

$$\widehat{WEIGHT} = 100 + 6.0 HEIGHT$$

- (a) A student has height 5. What is the regression's prediction for that student's weight? **$100 + 6 \cdot 5 = 130$**
- (b) In the sample, the sample average of HEIGHT is 4. What can you say about the sample average for WEIGHT? **124**

Now, an additional variable is included, is ID, a student's SFU ID. Obviously, this is a nonsensical variable to include: it is not in any way related to a student's weight. The new estimated regression equation is

$$\widehat{WEIGHT} = 101.5 + 5.98 HEIGHT + 0.02 ID$$

- (c) Someone's weight has nothing to do with their SFU ID. Still, the R^2 went up from 0.74 to 0.75. How is this possible? **The R^2 never decreases**
- (d) If the post office box number is not related to a student's weight, should the estimated coefficient not be equal to 0? How could it be that it is 0.02? **Sampling variability. Even if $\beta_{ID} = 0$, $\hat{\beta}_{ID} \sim \mathcal{N}(0, \sigma_{\hat{\beta}}^2)$**
6. Let D_i be a dummy variable. Consider the model that consists of the equation

$$Y_i = \beta_0 + \beta_1 X_i + \beta_2 D_i + \beta_3 D_i X_i + u_i$$

and the standard OLS assumptions

- (a) Draw a graph to visualize this model. **One graph, two lines. One is labelled $E(Y_i|X_i, D_i = 1)$, the other $E(Y_i|X_i, D_i = 0)$. The horizontal axis is labelled X_i . The two lines have different slopes, and different intercepts. The intercepts $(\beta_0, \beta_0 + \beta_2)$ and slopes $(\beta_1, \beta_1 + \beta_3)$ are clearly marked in the graph.**
- (b) What is the interpretation of β_3 ? **Difference in slopes $\partial E(Y_i|X_i, D_i)/\partial X_i$ between $E(Y_i|X_i, D_i = 1)$, $E(Y_i|X_i, D_i = 0)$.**
7. [Based on SW, Exercise 8.7] This problem is inspired by the study of the gender gap in top corporate jobs in Bertrand and Hallock (2001). The study compares total compensation among top executives in a large set of U.S. public corporations in the 1990s.

- (a) Let *Female* be an indicator variable that is equal to 1 for females and to 0 for males. A regression of the logarithm of earnings onto *Female* yields

$$\log(\widehat{Earnings}) = 6.48 - 0.44 Female$$

where the estimated regression coefficient -0.44 has a standard error of 0.05. Explain what the -0.44 means. **In expectation, c.p., women earn 44% less.**

- (b) Does this regression suggest that there is gender discrimination? Explain. **No: omitted variables.**
- (c) Two new variables are added to the regression: $\log(MarketValue)$, where *MarketValue* is a measure of firm size, in millions; and *Return*, the stock return, in percentage points. The resulting estimated regression line is

$$\log(\widehat{Earnings}) = 3.68 - 0.28 Female + 0.37 \log(MarketValue) + 0.004 Return$$

where the standard errors for the three regressors are 0.04, 0.004, and 0.003, respectively. The coefficient estimate for *Female* has changed from -0.44 to -0.28 . Why has it changed? **Omitted variables in the first regression**

8. A mixed bag of questions:
- (a) Describe the difference between “internal validity” and “external validity”. **Book**
 - (b) List two threats to internal validity. (The book lists five). **Book**
 - (c) What are the two conditions an instrumental variable must satisfy? **(1) Relevance and validity, OR (2) $E(u_i|Z_i) = 0$ and $Cov(Z_i, X_i) \neq 0$.**
9. In the context of your second hand-in assignment, consider the following code and output. Why are there “NA”s in the row for “occ9”? **Multicollinearity / dummy variable trap.**