(10 Pts)

1. Bias-Variance Decomposition

State and prove the bias-variance decomposition for the <u>predictive risk</u> of a regression estimator $\widehat{m}_n(x)$ at x,

$$R_x(\widehat{m}_n, m) = \mathbb{E}_n \mathbb{E}(Y - \widehat{m}_n(x))^2$$

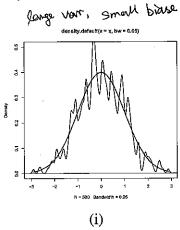
under the model $Y = m(x) + \epsilon$ where $\epsilon \sim N(0, \sigma^2)$. The expectation \mathbb{E} is over the noise ϵ and the expectation \mathbb{E}_n is over the training data $(X_1, Y_1), \dots, (X_n, Y_n)$.

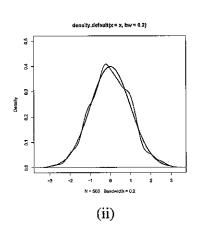
$$\begin{split} R_{n}(\widehat{m}_{n}, m) &= \mathbb{E}_{n} \mathbb{E}(Y - \widehat{m}_{n}(x))^{2} = \mathbb{E}_{n} \mathbb{E}(m(x) + \varepsilon - \widehat{m}_{n}(x))^{2} \\ &= \mathbb{E}_{n} \mathbb{E}\left[(m(x) - \widehat{m}_{n}(x))^{2} + 2\varepsilon(m(x) - \widehat{m}_{n}(x)) + \varepsilon^{2}\right] \\ &= \mathbb{E}_{n}\left((m(x) - \widehat{m}_{n}(x))^{2} + 2\mathbb{E}(\mathbb{E})\mathbb{E}(m(x) - \widehat{m}_{n}(x)) + \mathbb{E}_{8}^{2} \\ &= \mathbb{E}_{n}\left[(m(x) - \mathbb{E}_{n}\widehat{m}_{n}(x)) + (\mathbb{E}\widehat{m}_{n}(x) - \widehat{m}_{n}(x))^{2} + \sigma^{2} \\ &= (m(x) - \mathbb{E}_{n}\widehat{m}_{n}(x))^{2} + 2(m(x) - \mathbb{E}_{n}\widehat{m}_{n}(x)) + \mathbb{E}_{n}\left[\mathbb{E}_{n}\widehat{m}_{n}(x) - \widehat{m}_{n}(x)\right]^{2} + \mathbb{E}_{n}\left[(\mathbb{E}_{n}\widehat{m}_{n}(x) - \widehat{m}_{n}(x))^{2}\right] \\ &= (m(x) - \mathbb{E}_{n}\widehat{m}_{n}(x))^{2} + \mathbb{E}_{n}\left[(\mathbb{E}_{n}\widehat{m}_{n}(x) - \widehat{m}_{n}(x))^{2}\right] + \sigma^{2} \\ &= (m(x) - \mathbb{E}_{n}\widehat{m}_{n}(x))^{2} + \mathbb{E}_{n}\left[(\mathbb{E}_{n}\widehat{m}_{n}(x) - \widehat{m}_{n}(x))^{2}\right] + \sigma^{2} \\ &= (m(x) - \mathbb{E}_{n}\widehat{m}_{n}(x))^{2} + \mathbb{E}_{n}\left[(\mathbb{E}_{n}\widehat{m}_{n}(x) - \widehat{m}_{n}(x))^{2}\right] + \sigma^{2} \\ &= (m(x) - \mathbb{E}_{n}\widehat{m}_{n}(x))^{2} + \mathbb{E}_{n}\left[(\mathbb{E}_{n}\widehat{m}_{n}(x) - \widehat{m}_{n}(x))^{2}\right] + \sigma^{2} \\ &= (m(x) - \mathbb{E}_{n}\widehat{m}_{n}(x))^{2} + \mathbb{E}_{n}\left[(\mathbb{E}_{n}\widehat{m}_{n}(x) - \widehat{m}_{n}(x))^{2}\right] + \sigma^{2} \\ &= (m(x) - \mathbb{E}_{n}\widehat{m}_{n}(x))^{2} + \mathbb{E}_{n}\left[(\mathbb{E}_{n}\widehat{m}_{n}(x) - \widehat{m}_{n}(x))^{2}\right] + \sigma^{2} \\ &= (m(x) - \mathbb{E}_{n}\widehat{m}_{n}(x))^{2} + \mathbb{E}_{n}\left[(\mathbb{E}_{n}\widehat{m}_{n}(x) - \mathbb{E}_{n}\widehat{m}_{n}(x))^{2}\right] + \sigma^{2} \\ &= (m(x) - \mathbb{E}_{n}\widehat{m}_{n}(x))^{2} + \mathbb{E}_{n}\left[(\mathbb{E}_{n}\widehat{m}_{n}(x) - \mathbb{E}_{n}\widehat{m}_{n}(x))^{2}\right] + \sigma^{2} \\ &= (m(x) - \mathbb{E}_{n}\widehat{m}_{n}(x))^{2} + \mathbb{E}_{n}\left[(\mathbb{E}_{n}\widehat{m}_{n}(x) - \mathbb{E}_{n}\widehat{m}_{n}(x))^{2}\right] + \sigma^{2} \\ &= (m(x) - \mathbb{E}_{n}\widehat{m}_{n}(x))^{2} + \mathbb{E}_{n}\left[(\mathbb{E}_{n}\widehat{m}_{n}(x) - \mathbb{E}_{n}\widehat{m}_{n}(x))^{2}\right] + \sigma^{2} \\ &= (m(x) - \mathbb{E}_{n}\widehat{m}_{n}(x)^{2}) + \mathbb{E}_{n}\left[(\mathbb{E}_{n}\widehat{m}_{n}(x) - \mathbb{E}_{n}\widehat{m}_{n}(x))^{2}\right] + \sigma^{2} \\ &= (m(x) - \mathbb{E}_{n}\widehat{m}_{n}(x)^{2}) + \mathbb{E}_{n}\left[(\mathbb{E}_{n}\widehat{m}_{n}(x) - \mathbb{E}_{n}\widehat{m}_{n}(x)\right] + \mathbb{E}_{n}\left[(\mathbb{E}_{n}\widehat{m}_{n}(x) - \mathbb{E}$$

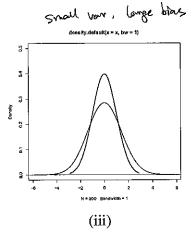
2. True or False?

Indicate whether each of the following statements is true or false, by circling your answer.

1. Consider the three density estimates below. The true density is Gaussian, and is plotted together with the kernel density estimate \hat{f}_n for different bandwidths.







TRUE FALSE

(a) The estimate in (ii) has a higher variance than the estimate in (i).

TRUE) FALSE

- (b) The estimate in (iii) has a higher bias than the estimate in (ii).
- 2. As the bandwidth h of a local linear regression estimator increases:

TRUE FALSE

(a) the variance decreases.

TRUE FALSE

(b) the squared bias decreases.

TRUE FALSE

- (c) the risk increases.
- 3. Suppose that $Y_i = m(x_i) + \sigma \epsilon_i$ where $\epsilon_i \sim N(0,1)$, $x_i \in [0,1]$ and m has bounded second derivatives on [0,1]. Let $\widehat{m}_n(x) = \ell(x)^T Y$ be a local linear smoother with bandwidth chosen by leave-one-out cross validation. Define the confidence interval

$$I_n(x) = (\widehat{m}_n(x) - z_{\alpha/2}\sigma \|\ell(x)\|, \ \widehat{m}_n(x) + z_{\alpha/2}\sigma \|\ell(x)\|)$$

where $z_{\alpha/2} = \Phi^{-1}(1 - \alpha/2)$. All of the following statements are with respect to the sample size n increasing to infinity.

TRUE (FALSE)

(a) $\mathbb{P}(m(x) \in I_n(x)) \longrightarrow 1 - \alpha$

TRUE FALSE

(b) $\mathbb{P}(\mathbb{E}\widehat{m}_n(x) \in I_n(x)) \longrightarrow 1-\alpha$

TRUE (FALSE)

(c) $\mathbb{P}(m(x) \in I_n(x) \text{ for all } x \text{ in } [0,1]) \longrightarrow 1-\alpha$

TRUE FALSE

(d) $\mathbb{P}\big(\mathbb{E}\widehat{m}_n(x)\in I_n(x) \text{ for all } x \text{ in [0,1]}\big)\longrightarrow 1-\alpha$

TRUE (FALSE)

(e) $\sup_{m} \mathbb{P}(m(x) \in I_n(x)) \longrightarrow 1 - \alpha$

(d) is probably the friteiest one.

The frey is to see that In(n) is a good CI 2

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only at fixed values of x6[0,1].

To obtain confidence bounds uniform in x6[0,1]

To obtain confidence bounds uniform to the example

To obtain confidence bounds constructions constant to the AdNS, 85.7.