

Trace and Schatten norm estimation using randomized sampling

1 Preliminaries

Definition. For a matrix $\mathbf{B} \in \mathbb{R}^{m \times n}$ with singular values $\sigma_1(\mathbf{B}) \geq \dots \geq \sigma_r(\mathbf{B}) > 0$, we define its Schatten p -norm as:

$$\|\mathbf{B}\|_p = \begin{cases} (\sum_{i=1}^r \sigma_i(\mathbf{B})^p)^{\frac{1}{p}}, & p \in [1, \infty) \\ \sigma_1(\mathbf{B}), & p = \infty \end{cases}$$

2 Trace estimation

- Given: Non-zero PSD matrix $\mathbf{A} \in \mathbb{R}^{n \times n}$ via the matrix-vector product $\mathbf{u} \mapsto \mathbf{A}\mathbf{u}$
- Goal: Estimate $\text{tr}(\mathbf{A})$ without computing $\delta_i^\top (\mathbf{A}\delta_i)$ for $i = 1, \dots, n$

2.1 Estimator based on randomized sampling

- For any random test vector $\boldsymbol{\omega} \in \mathbb{R}^n$ that is isotropic ($\mathbb{E}[\boldsymbol{\omega}\boldsymbol{\omega}^\top] = \mathbf{I}_n$), the random variable $X = \boldsymbol{\omega}^\top (\mathbf{A}\boldsymbol{\omega})$ is an unbiased estimator of $\text{tr}(\mathbf{A})$.
- Average over k independent copies of X :

$$\boldsymbol{\omega}_i \sim \boldsymbol{\omega} \text{ i.i.d.}, \quad X_i = \boldsymbol{\omega}_i^\top (\mathbf{A}\boldsymbol{\omega}_i), \quad \bar{X}_k = \frac{1}{k} \sum_{i=1}^k X_i$$

to obtain an unbiased and consistent estimator of $\text{tr}(\mathbf{A})$:

$$\mathbb{E}[\bar{X}_k] = \text{tr}(\mathbf{A}), \quad \text{Var}[\bar{X}_k] = \frac{1}{k} \text{Var}[X]$$

- Runtime advantageous for $k \ll n$
- If the coordinates of $\boldsymbol{\omega}$ are independent samples from a standardized random variable Z , then $\mathbb{E}[\boldsymbol{\omega}\boldsymbol{\omega}^\top] = \mathbf{I}_n$ and $\text{Var}[X] = 2 \sum_{i \neq j} (\mathbf{A})_{ij}^2 + (\mathbb{E}[Z^4] - 1) \sum_i (\mathbf{A})_{ii}^2$.

Important examples:

- Girard estimator $\boldsymbol{\omega} \sim \mathcal{N}(\mathbf{0}, \mathbf{I}_n)$: $\text{Var}[X] = 2 \sum_{i,j} (\mathbf{A})_{ij}^2 = 2\|\mathbf{A}\|_F^2 \leq 2\|\mathbf{A}\|_s \text{tr}(\mathbf{A})$
- Hutchinson estimator $\boldsymbol{\omega} \sim \mathcal{U}\{-1, 1\}^n$: $\text{Var}[X] = 2 \sum_{i \neq j} (\mathbf{A})_{ij}^2 < 2\|\mathbf{A}\|_F^2$

2.2 A priori error estimates

We can employ concentration inequalities to derive probability bounds for relative errors. They indicate how many samples k are sufficient to achieve a certain error tolerance. However, such prior bounds depend on information about the input matrix, e.g. $\text{intdim}(\mathbf{A})$, which we usually have no access to. Therefore, it is wiser to rely on a posteriori bounds in practice.

2.3 A posteriori error estimates

- Confidence intervals based on t -distribution:

$$\text{– Sample variance: } S_k^2 = \frac{1}{k-1} \sum_{i=1}^k (X_i - \bar{X}_k)^2, \quad \mathbb{E}[S_k^2] = \text{Var}[X]$$

$$\text{– For moderate sample sizes } (k \geq 30): \frac{\bar{X}_k - \text{tr}(\mathbf{A})}{S_k/\sqrt{k}} \sim t(k-1)$$

– $(1 - 2\alpha)$ confidence interval: $\text{tr}(\mathbf{A}) \in [\bar{X}_k \pm t_{1-\alpha, k-1} \frac{S_k}{\sqrt{k}}]$

- Alternative method: Confidence intervals based on bootstrapping

2.4 Application to Frobenius and Schatten 4-norm

- Given: Matrix $\mathbf{B} \in \mathbb{R}^{m \times n}$ accessed via $\mathbf{u} \mapsto \mathbf{B}\mathbf{u}$
- Draw i.i.d. test vectors $\boldsymbol{\omega}_1, \dots, \boldsymbol{\omega}_k \sim \mathcal{N}(\mathbf{0}, \mathbf{I}_n)$ and compute:

$$X_i = \boldsymbol{\omega}_i^\top \mathbf{B}^\top \mathbf{B} \boldsymbol{\omega}_i, \quad \bar{X}_k = \frac{1}{k} \sum_{i=1}^k X_i$$

- Since \bar{X}_k is the Girard trace estimator for $\mathbf{B}^\top \mathbf{B}$, we have: $\mathbb{E}[\bar{X}_k] = \text{tr}(\mathbf{B}^\top \mathbf{B}) = \|\mathbf{B}\|_F^2$
- Use the sample variance S_k^2 to estimate $\|\mathbf{B}\|_4^2$: $\mathbb{E}[S_k^2] = \text{Var}[X] = 2\|\mathbf{B}^\top \mathbf{B}\|_F^2 = 2\|\mathbf{B}\|_4^4$

3 Schatten norm estimation

- Goal: Estimate Schatten $2p$ -norm $\|\mathbf{B}\|_{2p}$ for natural numbers $p \geq 3$
- Idea: Sample i.i.d. isotropic test vectors $\boldsymbol{\omega}_1, \dots, \boldsymbol{\omega}_k \in \mathbb{R}^n$ to extract linear information $\mathbf{Y}_i = \mathbf{B}\boldsymbol{\omega}_i$, form the sample matrix $\mathbf{Y} = [\mathbf{Y}_1, \dots, \mathbf{Y}_k]$ and consider $\mathbf{X} = \mathbf{Y}^\top \mathbf{Y} \in \mathbb{R}^{k \times k}$
- For distinct indices $1 \leq i_1, \dots, i_p \leq k$, we can use the “trace trick” as well as independence and isotropy to conclude:

$$\mathbb{E}[(\mathbf{X})_{i_1 i_2} (\mathbf{X})_{i_2 i_3} \cdots (\mathbf{X})_{i_p i_1}] = \text{tr}((\mathbf{B}^\top \mathbf{B})^p) = \|\mathbf{B}\|_{2p}^{2p}$$

3.1 Approach from classical statistics

- Average over all sequences of distinct indices to obtain an unbiased estimator of $\|\mathbf{B}\|_{2p}^{2p}$:

$$U_p = \frac{(k-p)!}{k!} \sum_{1 \leq i_1, \dots, i_p \leq k} (\mathbf{X})_{i_1 i_2} (\mathbf{X})_{i_2 i_3} \cdots (\mathbf{X})_{i_p i_1}$$

- Computationally demanding (almost k^p summands) and highly variable
 \implies Only feasible for small values of p

3.2 Computationally efficient method

- Average only over strictly increasing sequences of indices:

$$V_p = \binom{k}{p}^{-1} \sum_{1 \leq i_1 < \dots < i_p \leq k} (\mathbf{X})_{i_1 i_2} (\mathbf{X})_{i_2 i_3} \cdots (\mathbf{X})_{i_p i_1} = \binom{k}{p}^{-1} \text{tr}[\mathbf{X}_U^{p-1} \mathbf{X}]$$

(\mathbf{X}_U denotes the strict upper triangular part of \mathbf{X})

- Even higher variance than U_p , but cheaper in computation
 \implies Method is feasible for larger values of p and one can deploy a larger sample size k to account for the high variance