Dinur's Proof of the PCP Theorem

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- ▶ Dinur presents a new proof that NP = $PCP_{\frac{1}{2}}[c \log n, q, 1]$ where c and q are constants.

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 - ► CG ∈ NP. Coloring is the proof, verification complexity is linear in number of edges

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- ▶ Boost the accuracy of the solution to show that NP⊆ $PCP_{(1-\epsilon)^c}[O(\log n), 8c, 1]$

Regarding Expanders

▶ Edge expansion property for a graph $G = \langle V, E \rangle$: $\varphi(G) = \min_{|S| \leq \frac{n}{2}} \left\{ \frac{E(S,\bar{s})}{|S|} \right\}$ such that $S \subseteq V$.

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Step 2: Power Step

Assuming that the degree of G is constant, we amplify UNSAT(G) while blowing up |G| and $|\Sigma|$ by a constant factor.

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 - Now, we have that $|V'| = \sum_{v \in V} d_v = 2|E|$ and |E'| = frac(d+1)|V'|2 = (d+1)|E|, so |G'| is $\Theta(|G|)$.

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- ▶ $S = \{v \in V' \mid \sigma'(v) \text{ is not the popular color}\}$

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$$|B'| \geq \frac{\sum |E(S_a^{\mathsf{v}}, \bar{S_a^{\mathsf{v}}})|}{2} \geq \frac{\varphi_0|S|}{2} \geq \frac{\mu\varphi_0}{4}|E| \geq \frac{\mu\varphi_0}{4(d+1)}|E'| = \mathsf{UNSAT}(G)\frac{\varphi_0}{4(d+1)}$$

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There exists as constant $\beta>0$ such that if UNSAT(G) $\leq \frac{1}{t}$ then UNSAT(G') $\geq \beta \sqrt{t}$ UNSAT(G)

Proof:

- ▶ Let $\sigma': V \mapsto \Sigma'$ be the best coloring for G', hence $\alpha' = \mathsf{UNSAT}(G')$ is the fraction of walk constraints violated by σ'
- Let $\sigma: V \mapsto \Sigma$ be $\sigma(v) = \arg\max_{a \in \Sigma} \Pr[X_{v,t/2} = a]$ where $X_{v,i}$ is the opinion that a vertex has of v that is i random steps away from v
- If $\sigma(v) = a$ then $\Pr[X_{v,t/2} = a] \ge \frac{1}{\Sigma}$
- Let B be the set of edges that are violated by σ in G, then $\frac{|B|}{|E|} \ge \text{UNSAT}(G) = \alpha$



Proof of Main Lemma

We want to show that bad walks do not overlap very much

Let w be a random t-walk in G', we now define the following random variable:

$$N = egin{cases} ext{number of bad edges in } I & C_w ext{ is violated by } \sigma' \ 0 & ext{otherwise} \end{cases}$$

- ▶ UNSAT(G') ≥ Pr[N > 0], we wish to lower bound Pr[N > 0]
- ▶ Claim 1: $\exists \mu > 0$ such that $E[N] \geq \frac{2\mu\sqrt{t}|B|}{|E|}$
- ▶ Claim 2: $\exists C > 0$ such that $E[N^2] \leq \frac{C\sqrt{t|B|}}{|E|}$
- ▶ Chosing $\beta = \frac{4\mu^2}{C}$ completes the proof:

$$\begin{split} \Pr[N>0] \geq & \frac{E[N]^2}{E[N^2]} \text{ (By the second moments inequality)} \\ & = \frac{\frac{4\mu^2t|B|^2}{|E|^2}}{\frac{C\sqrt{t}|B|}{|E|}} = \frac{4\mu^2\sqrt{t}}{C} \frac{|B|}{|E|} \geq \beta \text{UNSAT}(G)\sqrt{t} \end{split}$$



 \triangleright For a random walk w, define 2 random variables:

$$Z_i = egin{cases} 1, & i^{ ext{th}} ext{ edge of } w ext{ is in } B \ 0, & ext{otherwise} \end{cases}$$
 $Y_i = egin{cases} 1, & w ext{ is a rejecting } t ext{-walk and } Z_i = 1 \ 0, & ext{otherwise} \end{cases}$

▶ Note: $\forall i, Y_i \leq Z_i$

▶ Let $N = \sum_{i \in I} Y_i$, then we have:

$$E[N] = \sum_{i \in I} E[Y_i]$$

$$= \sum_{i \in I} \Pr[Y_i = 1]$$

$$= \sum_{i \in I} \Pr[Y_i = 1 \mid Z_i = 1] \Pr[Z_i = 1]$$

We know that
$$\Pr[Z_i=1]=\frac{|B|}{|E|}$$
 and $\Pr[Y_i=1\mid Z_i=1]\geq \frac{\tau^2}{|\Sigma|^2}=\mu$, so we get:

$$\geq \sum_{i \in I} \mu \frac{|B|}{|E|}$$
$$= \frac{2\mu \sqrt{t}|B|}{|E|}$$

▶ Since $Y_i \leq Z_i$, we know that $N \leq \sum_{i \in I} Z_i$, therefore:

$$\begin{split} E[N^2] &\leq E[(\sum_{i \in I} Z_i)^2] \\ &\leq 2 \sum_{i \in I} \sum_{j \geq i} E[Z_i Z_j] \\ &= 2 \sum_{i \in I} \Pr[Z_i = 1] \sum_{j \geq i} \Pr[Z_j = 1 \mid Z_i = 1] \\ &= 2 \frac{|B|}{|E|} \sum_{i \in I} \sum_{j \geq i} \underbrace{\Pr[Z_j = 1 \mid Z_i = 1]}_{\downarrow} \end{split}$$

Probability that a random walk has its $(j - i + 1)^{\text{th}}$ edge in B given that the first edge in the walk is in B

$$E[N^{2}] \leq 2 \frac{|B|}{|E|} \sum_{i \in I} \sum_{j \geq i} \Pr[Z_{j} = 1 \mid Z_{i} = 1]$$

$$\leq 2 \frac{|B|}{|E|} \sum_{i \in I} \sum_{j \geq i} \left(\frac{|B|}{|E|} + \lambda^{j-i} \right)$$

$$\leq C \frac{|B|}{|E|} 2 \sqrt{t} \sqrt{t} \left(\frac{1}{\sqrt{t}} \right)$$

$$\leq \frac{C \sqrt{t}|B|}{|E|}$$