

Problem 1

By definition of marginal probability,

$$P(X) = \sum_Y P(X, Y) = \sum_Y P(X|Y) P(Y) \quad ①$$

$$P(Y) = \sum_X P(X, Y) = \sum_X P(Y|X) P(X) \quad ②$$

And we know, $P(X, Y) = f(x)g(Y)$

thus, from ①, $\sum_Y f(x)g(Y) = \sum_Y P(X|Y) P(Y)$

from ②, $\sum_X f(x)g(Y) = \sum_X P(Y|X) P(X)$

$$f(x) = P(X|Y) = P(X)$$

$$g(Y) = P(Y) = P(Y|X)$$

$$P(X) = P(X|Y), \quad P(Y) = P(Y|X)$$

therefore, X and Y are independent.

$$P(X)P(Y) = \sum_y P(X, y) = \sum_x P(x, Y)$$

$$= \sum_y f(x)g(y) = \sum_x f(x)g(Y)$$

$$= f(x)g(Y) = g(Y) \sum_x f(x) = f(x)g(Y) \sum_{x,y} f(x)g(y)$$

$$= f(x)g(Y) \sum_{x,y} P(x, y) = f(x)g(Y) = P(X, Y)$$

Problem 2

$$\begin{aligned} & E_{x \sim p(x)} [af(x) + bg(x)] \\ &= \sum_x p(x) [af(x) + bg(x)] \\ &= \sum_x [ap(x)f(x) + b p(x)g(x)] \\ &= \sum_x a p(x)f(x) + \sum_x b p(x)g(x) \\ &= a \sum_x p(x)f(x) + b \sum_x p(x)g(x) \\ &= a E_{x \sim p(x)}[f(x)] + b E_{x \sim p(x)}[g(x)] \end{aligned}$$

Thus, the expectation is a linear system.

Problem 3.

(a) likelihood: $L(\theta = [a, b] ; D = \{x^{(m)}\}_{m=1}^M)$

$$= \prod_{m=1}^M f(x^{(m)}; [a, b])$$

$$= \prod_{m=1}^M \frac{1}{b-a} = \frac{1}{(b-a)^M}$$

(b) maximum likelihood estimate of parameters:

$$\begin{aligned}\hat{\theta}_{MLE} &= [a, b]_{MLE} = \operatorname{argmax}_{[a, b]} L(\theta) \\ &= \operatorname{argmax}_{[a, b]} L([a, b]) \\ &= \operatorname{argmax}_{[a, b]} \log L([a, b]) \\ &= \operatorname{argmax}_{[a, b]} \log \left[\prod_{m=1}^M f(x^{(m)}; [a, b]) \right] \\ &= \operatorname{argmax}_{[a, b]} \log \left[\prod_{m=1}^M \frac{1}{b-a} \right] \\ &= \operatorname{argmax}_{[a, b]} \log \left[\frac{1}{(b-a)^M} \right] \\ &= \operatorname{argmax}_{[a, b]} (-M) \log (b-a)\end{aligned}$$

$$\frac{\partial}{\partial a} \left(-m \log(b-a) \right) = -\left(\frac{-m}{b-a} \right) = \frac{m}{b-a}$$

We want to maximum the $\hat{\theta}_{MLE}$, therefore,

$$\hat{a}_{MLE} = \min(X_1, X_2, X_3, \dots, X_m)$$

$$\hat{b}_{MLE} = \max(X_1, X_2, X_3, \dots, X_m)$$

Problem 4

(c) For uniformly distributed,

$$\begin{aligned} E(x) &= \int_a^b \frac{x}{b-a} dx \\ &= \frac{1}{b-a} \int_a^b x dx \\ &= \frac{1}{b-a} \cdot \left[\frac{x^2}{2} \right]_a^b \\ &= \frac{1}{b-a} \cdot \left(\frac{b^2}{2} - \frac{a^2}{2} \right) \\ &= \frac{1}{2} \cdot \frac{b^2 - a^2}{b-a} \\ &= \frac{1}{2} \cdot (a+b) = \frac{a+b}{2} \end{aligned}$$

$$E[\hat{a}_{MLE}] = \frac{a+b}{2} \quad E[\hat{b}_{MLE}] = \frac{a+b}{2}$$

By Thm, $E[\hat{w}_{MLE}] = w$

$$\text{However, } \mathbb{E}[\hat{a}_{MLE}] = \frac{a+b}{2} \neq a = \min(x_1, x_2, x_3, \dots, x_m)$$

$$\mathbb{E}[\hat{b}_{MLE}] = \frac{a+b}{2} \neq b = \max(x_1, x_2, x_3, \dots, x_m)$$

Thus, MLE is biased in this case.

(d) From (a), $L(\theta = [a, b] ; D = \{x^{(m)}\}_{m=1}^M) = \frac{1}{(b-a)^M}$
 $M \rightarrow +\infty, (b-a)^M \rightarrow +\infty, \frac{1}{(b-a)^M} \rightarrow 0$

$\Rightarrow L$ is a constant function

$$\Rightarrow x_1 = x_2 = x_3 = x_4 = \dots = x_m$$

$$\Rightarrow \frac{a+b}{2} = a = \min(x_1, x_2, x_3, \dots, x_m)$$

$$\frac{a+b}{2} = b = \max(x_1, x_2, x_3, \dots, x_m)$$

Thus, MLE is asymptotically unbiased
if $M \rightarrow +\infty$.