GOV 2000 / E-2000 Section 5¹ Sampling, Interval Estimation, and Hypothesis Testing

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October 3, 2013

¹These notes and accompanying code draw on the notes from TF's from previous years.

LOGISTICS - THIS WEEK

Problem Set 4- Due by 1pm on Tuesday to course dropbox.

Problem Set 3 Corrections- Due by 1pm on Tuesday to course dropbox.

No Reading Quiz

Practice Midterms- Not due, but recommended as midterm preparation.

LOGISTICS - NEXT WEEK

Midterm- Posted October 8, due by 11:59pm on Sunday, October 13 with 5 hours to complete.

Office Hours- Cancelled Friday October 6. Rescheduled for Tuesday October 15 from 9:30am to 12pm.

Problem Set 5- Due by 1pm on Tuesday, October 15 to course dropbox.

Problem Set 4 Corrections- Due by 1pm on Tuesday, October 15 to course dropbox.

Reading Quiz (ALZ ch. 11 & 12)- Due by 9pm on Sunday, October 13 on course website.

Mid-semester Evaluation- Due by 9pm on Monday, October 14 on course website.

PROBLEM SET EXPECTATIONS

- Must be typeset using LATEX or Word and submitted as one document containing graphics and explanation electronically in pdf form
- ► Must be accompanied by source-able, commented code
- Corrections do not need to include any of your original work, just 1-2 pages explaining where you made a mistake and how to correct it

MIDTERM

- ► Window of exam: Tuesday, October 8th (after class) Sunday, October 13th at 11:59pm
- ▶ Needs to be completed in 5 hours
- ▶ Open note / book, but no collaboration allowed

▶ **Parameter-** characteristic of the population distribution (the distribution of X_i)

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ex.
$$\mu$$
 or σ^2

► Estimand- parameter being estimated

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- Estimand- parameter being estimated ex. μ
- ► **Statistic-** function of the sample used to estimate a parameter

▶ **Parameter-** characteristic of the population distribution (the distribution of X_i)

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$$\mu$$
 or σ^2

► Estimand- parameter being estimated

ex.
$$\mu$$

➤ **Statistic-** function of the sample used to estimate a parameter

ex.
$$\bar{X}_n$$
 or $\hat{\mu}$ or S^2

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➤ **Statistic-** function of the sample used to estimate a parameter

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► **Estimator-** a statistic (random variable) that describes the estimation procedure

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Which are random variables?

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Parameter?

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Parameter?No

Which are random variables?

Parameter?No

Estimand?

Which are random variables?

Parameter?No

Estimand?No

Which are random variables?

Parameter?No

Estimand?No

Statistic?

Which are random variables?

Parameter?No

Estimand?No

Statistic?Yes

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Estimator?

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Which are random variables?

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Estimand?No

Statistic?Yes

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Estimator?Yes

Estimate?No

Which are random variables?

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Estimand?No

Statistic?Yes

Estimator?Yes

Estimate?No

Estimand μ

Estimand

Estimator \bar{X}

Estimand

Estimator \bar{X}_i

Estimate \bar{x}_n

Estimand

Estimator \bar{X}_i

Estimate \bar{x}_n

Sampling Distribution of \bar{X}_n with **Known** Population

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Sampling Distribution of X_n with **Known** Population

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- 3. Choose a plausible estimator. For us, it's $\hat{\mu} = \bar{X}_n$.

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- 5. Repeat step 4 many times (we will have s = 1,000).

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Sampling Distribution of \bar{X}_n

The sampling distribution of \bar{X}_n is the distribution of the column vector:

			$\hat{\mu} =$			
		X_1	X_2		X_n	\bar{X}_n
	1	$x_{1,1}$	$x_{2,1}$		$x_{n,1}$	$\bar{x}_{n,1}$
Samples (s)	2	$x_{1,2}$	$x_{2,2}$		$x_{n,2}$	$\bar{x}_{n,2}$
	:	:	:	٠.,	:	:
	10,000	$x_{1,10000}$	$x_{2,10000}$		$x_{n,10000}$	$\bar{x}_{n,10000}$

In reality, we only get to observe one sample (not the entire population), so we cannot directly observe the sampling distribution of \bar{X}_n :

			$\hat{\mu} =$			
		X_1	X_2		X_n	\bar{X}_n
Samples (s)	1	$x_{1,1}$	$x_{2,1}$		$x_{n,1}$	$\bar{x}_{n,1}$
	2	$x_{1,2}$	$x_{2,2}$		$x_{n,2}$	$\bar{x}_{n,2}$
	:	:	:	٠	:	:
	10,000	$x_{1,10000}$	$x_{2,10000}$		$x_{n,10000}$	$\bar{x}_{n,10000}$

CENTRAL LIMIT THEOREM

Remember that if our samples are sufficiently large ($n \ge 30$), the Central Limit Theorem states that the sample mean will be distributed as:

$$\bar{X}_n \sim N(\mu, \frac{\sigma^2}{n})$$

We could apply the CLT to get the sampling distribution of our estimator, except σ^2 is **unkown**.

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We can estimate σ^2 using the sample variance!

		Observations			$\hat{\mu} =$	$\hat{\sigma^2} =$		
		X_1	X_2		X_n	\bar{X}_n	S_n^2	$\hat{SE}[\bar{X}_n]$
	1	$x_{1,1}$	$x_{2,1}$		$x_{n,1}$	$\bar{x}_{n,1}$	$s_{n,1}^2$	$\frac{s_{n,1}}{\sqrt{n}}$
s	2	$x_{1,2}$	$x_{2,2}$		$x_{n,2}$	$\bar{x}_{n,2}$	$s_{n,2}^{2}$	$\frac{S_{n,2}}{\sqrt{n}}$
	:	:	:	٠	:	:	:	:
	1000	$x_{1,1000}$	$x_{2,1000}$		$x_{n,1000}$	$\bar{x}_{n,1000}$	$s_{n,1000}^2$	$\frac{s_{n,1000}}{\sqrt{n}}$
	E[·]	μ	μ		μ	μ	σ^2	
	$V[\cdot]$	σ^2	σ^2		σ^2	$\frac{\sigma^2}{n}$		

We could apply the CLT to get the sampling distribution of our estimator, except σ^2 is **unkown**.

We can estimate σ^2 using the sample variance!

			Observa	ations		$\hat{\mu} =$	$\hat{\sigma^2} =$	
		X_1	X_2		X_n	\bar{X}_n	S_n^2	$\hat{SE}[\bar{X}_n]$
	1	$x_{1,1}$	$x_{2,1}$		$x_{n,1}$	$\bar{x}_{n,1}$	$s_{n,1}^{2}$	$\frac{s_{n,1}}{\sqrt{n}}$
s	2	$x_{1,2}$	$x_{2,2}$		$x_{n,2}$	$\bar{x}_{n,2}$	$s_{n,2}^{2}$	$\frac{s_{n,2}}{\sqrt{n}}$
	:	:	:	٠.	:	:	:	:
	1000	$x_{1,1000}$	$x_{2,1000}$		$x_{n,1000}$	$\bar{x}_{n,1000}$	$s_{n,1000}^2$	$\frac{s_{n,1000}}{\sqrt{n}}$
	E[·]	μ	μ		μ	μ	σ^2	
	$V[\cdot]$	σ^2	σ^2		σ^2	$\frac{\sigma^2}{n}$		

CENTRAL LIMIT THEOREM

So, by the CLT again with *n* large we can estimate the sampling distribution of our estimator as:

$$\bar{X}_n \sim_{approx} N(\bar{x}_n, \frac{s_n^2}{n})$$

Generally, we report our point estimate \bar{x}_n and the estimated standard error of our estimate $\frac{S_n}{\sqrt{n}}$.

Or we can estimate the sampling distribution using boostrapping.

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Estimand [LB, UB] s.t.
$$P(LB \le \mu \le UB) = 100(1 - \alpha)\%$$

Hypothesis Testing

Interval Estimation for Large Samples and σ KNOWN

Estimand [LB, UB] s.t.
$$P(LB \le \mu \le UB) = 100(1 - \alpha)\%$$

Estimator
$$\bar{X}_n \pm z_{1-\alpha/2} \cdot \frac{\sigma}{\sqrt{n}}$$

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Estimate
$$\bar{x}_n \pm z_{1-\alpha/2} \cdot \frac{\sigma}{\sqrt{n}}$$

		Observations						
		X_1	X_2		X_n	$\bar{X}_n \pm z_{1-\alpha/2} \cdot \frac{\sigma}{\sqrt{n}}$		
	1	$x_{1,1}$	$x_{2,1}$		$x_{n,1}$	$\bar{x}_{n,1} \pm z_{1-\alpha/2} \cdot \frac{\sigma}{\sqrt{n}}$		
s	2	$x_{1,2}$	$x_{2,2}$		$x_{n,2}$	$\bar{x}_{n,2} \pm z_{1-\alpha/2} \cdot \frac{\sigma}{\sqrt{n}}$		
	:	:	:	٠.	÷	:		
	1000	$x_{1,1000}$	$x_{2,1000}$		$x_{n,1000}$	$\bar{x}_{n,1000} \pm z_{1-\alpha/2} \cdot \frac{\sigma}{\sqrt{n}}$		

How did we choose our estimator?

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By the CLT, if *n* is large, then $\bar{X}_n \sim N(\mu, \frac{\sigma^2}{n})$.

Interval Estimation

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If we standardize, we get that $Z = \frac{X_n - \mu}{\frac{\sigma}{G}} \sim N(0, 1)$

Interval Estimation for Large Samples and σ KNOWN

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By the CLT, if *n* is large, then $\bar{X}_n \sim N(\mu, \frac{\sigma^2}{n})$.

If we standardize, we get that $Z = \frac{\bar{X}_n - \mu}{\frac{\sigma}{C}} \sim N(0, 1)$

Which implies that:

$$P(-z_{1-\alpha/2} \le \frac{\bar{X}_n - \mu}{\frac{\sigma}{\sqrt{n}}} \le z_{1-\alpha/2}) = 1 - \alpha$$

Interval Estimation for Large Samples and σ KNOWN

How did we choose our estimator?

By the CLT, if *n* is large, then $\bar{X}_n \sim N(\mu, \frac{\sigma^2}{\pi})$.

If we standardize, we get that $Z = \frac{X_n - \mu}{\frac{\sigma}{C}} \sim N(0, 1)$

Which implies that:

$$P(-z_{1-\alpha/2} \le \frac{\bar{X}_n - \mu}{\frac{\sigma}{\sqrt{n}}} \le z_{1-\alpha/2}) = 1 - \alpha$$

$$\Rightarrow P(\bar{X}_n - z_{1-\alpha/2} \cdot \frac{\sigma}{\sqrt{n}} \le \mu \le \bar{X}_n + z_{1-\alpha/2} \cdot \frac{\sigma}{\sqrt{n}}) = 1 - \alpha$$

Interval Estimation for Large Samples and σ KNOWN

How did we choose our estimator?

By the CLT, if *n* is large, then $\bar{X}_n \sim N(\mu, \frac{\sigma^2}{n})$.

If we standardize, we get that $Z = \frac{X_n - \mu}{\frac{\sigma}{c}} \sim N(0, 1)$

Which implies that:

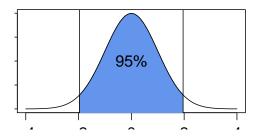
$$P(-z_{1-\alpha/2} \le \frac{\bar{X}_n - \mu}{\frac{\sigma}{\sqrt{n}}} \le z_{1-\alpha/2}) = 1 - \alpha$$

$$\Rightarrow P(\bar{X}_n - z_{1-\alpha/2} \cdot \frac{\sigma}{\sqrt{n}} \le \mu \le \bar{X}_n + z_{1-\alpha/2} \cdot \frac{\sigma}{\sqrt{n}}) = 1 - \alpha$$

How do we get $z_{1-\alpha/2}$?

How do we get $z_{1-\alpha/2}$?

```
## alpha = .05
qnorm(1-.025, mean = 0, sd=1)
```



How do we interpret our estimated confidence interval?

How do we interpret our estimated confidence interval? The true population mean μ will lie within the estimated intervals in $100(1-\alpha)\%$ of many repeated samples.

QUIZ BREAK!

What is random in our estimator: $\bar{X}_n \pm z_{1-\alpha/2} \cdot \frac{\sigma}{\sqrt{n}}$?

Quiz Break!

What is random in our estimator: $\bar{X}_n \pm z_{1-\alpha/2} \cdot \frac{\sigma}{\sqrt{n}} ? \overline{\bar{X}}_n$

What is random in our estimator: $\bar{X}_n \pm z_{1-\alpha/2} \cdot \frac{\sigma}{\sqrt{n}} ? \bar{X}_n$

What is random in our estimate: $\bar{x}_n \pm z_{1-\alpha/2} \cdot \frac{\sigma}{\sqrt{n}}$?

What is random in our estimator: $\bar{X}_n \pm z_{1-\alpha/2} \cdot \frac{\sigma}{\sqrt{n}} ? \bar{X}_n$

What is random in our estimate: $\bar{x}_n \pm z_{1-\alpha/2} \cdot \frac{\sigma}{\sqrt{n}}$? Nothing!

Interval Estimation for Large Samples and σ Unknown

Point Estimation

Estimand [LB, UB] s.t.
$$P(LB \le \mu \le UB) = 100(1 - \alpha)\%$$

Interval Estimation for Large Samples and σ UNKNOWN

Estimand [LB, UB] s.t.
$$P(LB \le \mu \le UB) = 100(1 - \alpha)\%$$

Estimator
$$\bar{X}_n \pm z_{1-\alpha/2} \cdot \frac{S_n}{\sqrt{n}}$$

Interval Estimation for Large Samples and σ UNKNOWN

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Estimate
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			Observa	ations		
		X_1	X_2		X_n	$\bar{X}_n \pm z_{1-\alpha/2} \cdot \frac{S_n}{\sqrt{n}}$
s	1	$x_{1,1}$	$x_{2,1}$		$x_{n,1}$	$\bar{x}_{n,1} \pm z_{1-\alpha/2} \cdot \frac{s_{n,1}}{\sqrt{n}}$
	2	$x_{1,2}$	$x_{2,2}$		$x_{n,2}$	$\bar{x}_{n,2} \pm z_{1-\alpha/2} \cdot \frac{s_{n,2}}{\sqrt{n}}$
	:	:	:	٠.	:	:
	1000	$x_{1,1000}$	$x_{2,1000}$		$x_{n,1000}$	$\bar{x}_{n,1000} \pm z_{1-\alpha/2} \cdot \frac{s_{n,1000}}{\sqrt{n}}$

Interval Estimation for Large Samples and σ Unknown

How did we choose our estimator?

Interval Estimation for Large Samples and σ UNKNOWN

How did we choose our estimator?

Same as before except we *estimated* σ/\sqrt{n} with $\hat{SE}[\bar{X}_n] = \frac{S_n}{\sqrt{n}}$.

Interval Estimation for Large Samples and σ UNKNOWN

How did we choose our estimator?

Same as before except we *estimated*
$$\sigma/\sqrt{n}$$
 with $\hat{SE}[\bar{X}_n] = \frac{S_n}{\sqrt{n}}$.

Are we still getting 95% coverage over repeated samples?

Interval Estimation for Large Samples and σ Unknown

Interval Estimation

How did we choose our estimator?

Same as before except we estimated
$$\sigma/\sqrt{n}$$
 with $\hat{SE}[\bar{X}_n] = \frac{S_n}{\sqrt{n}}$.

Are we still getting 95% coverage over repeated samples? No, but as n increases, this matters less.

What is random in our estimator: $\bar{X}_n \pm z_{1-\alpha/2} \cdot \frac{S_n}{\sqrt{n}}$?

What is random in our estimator: $\bar{X}_n \pm z_{1-\alpha/2} \cdot \frac{S_n}{\sqrt{n}} ? \bar{X}_n$ and S_n

What is random in our estimate: $\bar{x}_n \pm z_{1-\alpha/2} \cdot \frac{s_n}{\sqrt{n}}$?

What is random in our estimator: $\bar{X}_n \pm z_{1-\alpha/2} \cdot \frac{S_n}{\sqrt{n}} ? \bar{X}_n$ and S_n

What is random in our estimate: $\bar{x}_n \pm z_{1-\alpha/2} \cdot \frac{s_n}{\sqrt{n}}$? Nothing!

Interval Estimation for **Small** Samples and σ **Unknown**

Interval Estimation for **Small** Samples and σ UNKNOWN

Assuming $X_i \sim_{i,i,d} N(\mu, \sigma^2)$,

[LB, UB] s.t. $P(LB \le \mu \le UB) = 100(1 - \alpha)\%$ Estimand

Interval Estimation for **Small** Samples and σ Unknown

Assuming
$$X_i \sim_{i.i.d} N(\mu, \sigma^2)$$
,

Estimand [LB, UB] s.t.
$$P(LB \le \mu \le UB) = 100(1 - \alpha)\%$$

Estimator
$$\bar{X}_n \pm t_{n-1,1-\alpha/2} \cdot \frac{S_n}{\sqrt{n}}$$

Interval Estimation for **Small** Samples and σ Unknown

Assuming
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,

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Estimate
$$\bar{x}_n \pm t_{n-1,1-\alpha/2} \cdot \frac{s_n}{\sqrt{n}}$$

Interval Estimation

Interval Estimation for **Small** Samples and σ UNKNOWN

			Observa	ations		
		X_1	X_2		X_n	$\bar{X}_n \pm t_{n-1,1-\alpha/2} \cdot \frac{S_n}{\sqrt{n}}$
s	1	$x_{1,1}$	$x_{2,1}$		$x_{n,1}$	$\bar{x}_{n,1} \pm t_{n-1,1-\alpha/2} \cdot \frac{s_{n,1}}{\sqrt{n}}$
	2	$x_{1,2}$	$x_{2,2}$		$x_{n,2}$	$\bar{x}_{n,2} \pm t_{n-1,1-\alpha/2} \cdot \frac{s_{n,2}}{\sqrt{n}}$
	:	:	:	٠.	:	:
	1000	$x_{1,1000}$	$x_{2,1000}$		$x_{n,1000}$	$\bar{x}_{n,1000} \pm t_{n-1,1-\alpha/2} \cdot \frac{s_{n,1000}}{\sqrt{n}}$

Interval Estimation for **Small** Samples and σ **Unknown**

Why $t_{n-1,1-\alpha/2}$ instead of $z_{1-\alpha/2}$?

Interval Estimation for **Small** Samples and σ UNKNOWN

Why $t_{n-1,1-\alpha/2}$ instead of $z_{1-\alpha/2}$?

When we standardize \bar{X}_n and estimate σ with S_n ,

$$\frac{\bar{X}_n - \mu}{\frac{S_n}{\sqrt{n}}} \sim t_{n-1}$$

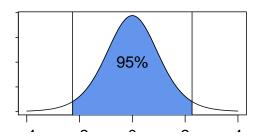
Interval Estimation for Small Samples and σ Unknown

How do we get $t_{n-1,1-\alpha/2}$?

Interval Estimation for **Small** Samples and σ **Unknown**

How do we get $t_{n-1,1-\alpha/2}$?

```
## alpha = .05
## 9 degrees of freedom
qt(1-.025, df=9)
```



Quiz Break

You are told that your estimand is [LB, UB] s.t. $P(LB \le \mu_1 - \mu_2 \le UB) = 100(1 - \alpha)\%$. Assuming n_1 and n_2 are both large and σ_1 and σ_2 are both unknwn, what is your estimator and estimate?

Estimand [*LB*, *UB*] s.t.
$$P(LB \le \mu_1 - \mu_2 \le UB) = 100(1 - \alpha)\%$$

Estimator

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Estimate

Ouiz Break

Logistics

You are told that your estimand is [LB, UB] s.t. $P(LB \le \mu_1 - \mu_2 \le UB) = 100(1 - \alpha)\%$. Assuming n_1 and n_2 are both large and σ_1 and σ_2 are both unknown, what is your estimator and estimate?

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Reject null if
$$\frac{\bar{x}_n - \mu_0}{\frac{s_n}{\sqrt{n}}} \in [-\infty, -1.96] \cup [1.96, \infty]$$

Fail to reject null if $\frac{\bar{x}_n - \mu_0}{\frac{s_n}{\sqrt{n}}} \notin [-\infty, -1.96] \cup [1.96, \infty]$

▶ p-value- Assuming that the null hypothesis is true, the probability of getting something at least as extreme as our observed test statistic, where extreme is defined in terms of the alternative hypothesis.

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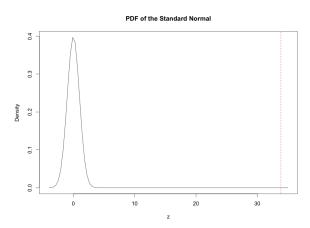
How do we find p-values?

```
# create a population with mean 12, sd 4
pop <- rnorm(n=1e6, mean=12, sd=4)

# draw a single sample of size 100 from population
my.sample <- sample(pop, size=100, replace=T)

# calculate our test statistic
test.statistic <- mean(my.sample) /(sd(my.sample)/10)

# find the p-value
p.value <- 2*(pnorm(test.statistic))</pre>
```



QUESTIONS

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