L'Hôpital's (Selection) Rule
An Empirical Bayes Application to French Hospitals

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Outline

Introduction

Data and Estimation

Empirical Bayes Compound decision: The Selection Problem

Conclusion

Motivation: Invidious decision

- *League table mentality*: Ranking & Selection.[Gu and Koenker, 2023]
- *Noisy estimates*: Unobserved heterogeneity. [Chetty et al., 2014]
- *Bayesian view*: Prior distribution.[Gu and Koenker, 2017]

Motivation: French Hospitals

- *Productivity/Efficiency*: Factories, Schools, Hospitals etc.
- *Methodology*:
  1. (Parametric) Stochastic Frontier
     [Aigner et al., 1977, Meeusen and van Den Broeck, 1977]$\rightarrow$ How far it is to the frontier.
  2. (Non-Parametric): Data Envelopment [Charnes et al., 1978]$\rightarrow$ Compare with other units.
  3. Input demand function: [Croiset and Gary-Bobo, 2024].
- *Ownership*: Public (Teaching, Ordinary) vs. Private (For profit, Non-profit).

Questions

- Out of the top 20% hospitals in France[1], how many of them are public hospitals/private hospitals?
- What would be the selection outcome if I also control for the **False Discovery Rate**?
- Does different ranking/selection rule produce contradicting results? And to what degree?

[1] in terms of labor (nurses) employment efficiency

Roadmap

1. Data: The Annual Statistics of Health Establishments (SAE), from 2013 to 2022.
2. Estimation of efficiency
   - Y: Labor input (number of full time equivalent nurses).
   - X: Hospital output (e.g., inpatient/outpatient stays, medical sessions).
3. Selection problem: Compound decision framework and optimal decision rule
4. Comparison of outcomes under different decision rules
5. Conclusion

Outline

Hospital Types

Year	Teaching	Normal Public	Private FP	Private NP	Total
2013	198	1312	1305	1382	4197
2014	201	1274	1293	1349	4117
2015	211	1275	1297	1349	4132
2016	212	1266	1297	1313	4088
2017	211	1249	1297	1306	4063
2018	214	1247	1296	1288	4045
2019	214	1236	1287	1281	4018
2021	219	1222	1293	1264	3998
2022	220	1220	1296	1259	3995

Output

Output	Normal Public	Private Non Profit	Private For Profit	Teaching
STAC[2] inpatient	8.08%	5.66%	16.3%	7.9%
STAC oupatient	2.26%	4.02%	22.61%	3.59%
Sessions	4.34%	23.31%	27.17%	4.8%
Outpatient Consultations	58.23%	43.55%	0.8%	69.18%
Emergency	21.14%	6.78%	17.3%	12.64%
Follow-up care and Long-term care	1.67%	11.26%	12.16%	1.09%
Home hospitalization	0.06%	0.76%	0.17%	0.08%
Psychiatry stays	4.22%	4.66%	3.49%	0.72%

[2] Short term acute care

First glance

Dependent Variable:	Nurses	
	OLS	Lagged IV
Model:	(1)	(2)
Variables		
Constant	1.59***	1.58***
	(0.067)	(0.069)
STAC inpatient	0.278***	0.277***
	(0.012)	(0.013)
...	...	...
Private Forprofit	-0.303***	-0.280***
	(0.061)	(0.065)
Private Nonprofit	-0.215***	-0.188***
	(0.056)	(0.055)
Teaching	0.717***	0.709***
	(0.056)	(0.056)
Fit statistics		
Observations	15,335	13,402
R^2	0.835	0.837

Clustered (FI) standard-errors in parentheses
*Signif. Codes: ***: 0.01, **: 0.05, *: 0.1*

Panel data Estimator

- Strict exogeneity: Within Group/First Diffrence

$$E[\epsilon_{it}|x_{i1},\ldots,x_{iT},\theta_i]=0$$

- Relaxed: First Difference GMM [Arellano and Bond, 1991], System GMM [Arellano and Bover, 1995, Blundell and Bond, 1998a].

$$E[\epsilon_{it}|x_{i1},\ldots,x_{it-p},\theta_i]=0$$

Issues: Weak instruments [Blundell and Bond, 1998b] and the proliferation of instruments [Roodman, 2007].

$$\mathbb{E}[\Delta x_{i,t-1}(y_{it}-\beta x_{it})] \quad \text{if} \quad \mathbb{E}\left[\Delta x_{i,t-1}(\theta_i+\varepsilon_{i,t})\right]=0$$

Panel Data Estimation

Dependent Variable:		log(ETP_INF)	
	Within Group	First Difference	*System GMM*
Model:	(1)	(2)	(3)
Variables			
log(SEJHC_MCO)	0.10^{***}	0.07^{***}	0.70^{***}
	(0.00)	(0.01)	(0.05)
log(SEJHP_MCO)	0.02^{***}	0.01^{***}	-0.05
	(0.00)	(0.00)	(0.04)
...	...	...	...
log(SEANCES_MED)	0.02^{***}	0.02^{***}	0.07^{***}
	(0.00)	(0.00)	(0.03)
log(SEJ_PSY)	0.00	0.00	0.07^{***}
	(0.00)	(0.00)	(0.01)
Fit statistics			
Observations	15335	13502	11536
n	1833	1833	1833
T	9	9	9

*Signif. Codes: ***: 0.01, **: 0.05, *: 0.1*

Outline

Compound Decision Framework

Observe:

$$\hat{\boldsymbol{\theta}} = (\hat{\theta}_1, \ldots, \hat{\theta}_n)$$

$$\text{where} \quad \hat{\theta}_i | \theta_i \sim P_{\theta_i}$$

Decision:

$$\delta(\hat{\boldsymbol{\theta}}) = (\delta_1(\hat{\boldsymbol{\theta}}), \ldots, \delta_n(\hat{\boldsymbol{\theta}}))$$

Compound Loss and Risk

Loss:

$$L_n(\theta, \delta(\hat{\boldsymbol{\theta}})) = \sum_{i=1}^{n} L(\theta_i, \delta_i(\hat{\theta})).$$

Risk (Expectation of loss):

$$\begin{aligned}
R_n(\theta, \delta(\hat{\boldsymbol{\theta}})) &= \mathbb{E}[L_n(\theta, \delta(\hat{\boldsymbol{\theta}}))] \\
&= \frac{1}{n} \sum_{i=1}^{n} \mathbb{E}[L(\theta_i, \delta_i(\hat{\boldsymbol{\theta}}))] \\
&= \frac{1}{n} \sum_{i=1}^{n} \int \dots \int L(\theta_i, \delta_i(\hat{\theta}_1, \dots, \hat{\theta}_n)) dP_{\theta_1}(\hat{\theta}_1) \dots dP_{\theta_n}(\hat{\theta}_n).
\end{aligned}$$

The Selection Task

- Select the bottom 20% [3] out of the population of θ_i, those i whose $\theta_i < G^{-1}(0.2)$
- Control the overall false discovery rate at 20%,

$$\frac{\mathbb{E}_G\left[1\left\{\theta_i > \theta_\alpha, \delta_i = 1\right\}\right]}{\mathbb{E}_G\left[\delta_i\right]} \leq \gamma$$

[3] The most efficient 20%.

Problem Formulation

The **loss** function is $(h_i = 1\{\theta_i < \theta_\alpha = G^{-}1(\alpha)\})$

$$L(\delta,\theta) = \sum h_i(1-\delta_i) + \tau_1\left(\sum(1-h_i)\delta_i - \gamma\delta_i\right) + \tau_2\left(\sum\delta_i - \alpha n\right)$$

Therefore, the problem is to find δ such that

$$\begin{aligned}\min_{\delta} \quad & \mathbb{E}_G\mathbb{E}_{\theta|\hat{\theta}}\left[L(\delta,\theta)\right] \\ =& \mathbb{E}_G \ \sum \mathbb{E}(h_i)(1-\delta_i) + \tau_1\left(\sum(1-\mathbb{E}(h_i))\delta_i - \gamma\delta_i\right) \\ & + \tau_2\left(\sum\delta_i - \alpha n\right) \\ =& \mathbb{E}_G\sum v_\alpha(\hat{\theta})(1-\delta_i) + \tau_1\left(\sum(1-v_\alpha(\hat{\theta}))\delta_i - \gamma\delta_i\right) + \tau_2\left(\sum\delta_i - \alpha n\right)\end{aligned}$$

where $v_\alpha(\hat{\theta}) = \mathbb{P}(\theta < \theta_\alpha|\hat{\theta})$ is the **posterior tail probability**.

Prior Distribution G

Observe[4]

$$Y_{it} = \theta_i + \varepsilon_{it} \quad \varepsilon_{it} \sim \mathcal{N}(0, \sigma_i^2) \quad (\theta_i, \sigma_i^2) \sim G$$

Neither θ_i nor σ_i^2 is known. But the sufficient statistics are

$$Y_i = \frac{1}{T_i} \sum_{t=1}^{T_i} Y_{it} \quad \text{where} \quad Y_i | \theta_i, \sigma_i^2 \sim \mathcal{N}(\theta_i, \sigma_i^2 / T_i)$$

$$S_i = \frac{1}{T_i - 1} \sum_{t=1}^{T_i} (Y_{it} - Y_i)^2 \quad \text{where} \quad S_i | \sigma_i^2 \sim \Gamma(r_i = (T_i - 1)/2, 2\sigma_i^2/(T_i - 1))$$

Appendix

[4] $Y_{it} = \theta_i + \varepsilon_{it} + x_{it}(\beta - \hat{\beta})$

Tail probability

Given the two sufficient statistics, the posterior tail probability is

$$v_\alpha(Y_i, S_i) = P(\theta_i < \theta_\alpha | Y_i, S_i)$$
$$= \frac{\int_{-\infty}^{\theta_\alpha} \Gamma(s_i | r_i, \sigma_i^2) f(y_i | \theta_i, \sigma_i^2) dG(\theta_i, \sigma_i^2)}{\int_{-\infty}^{\infty} \Gamma(s_i | r_i, \sigma_i^2) f(y_i | \theta_i, \sigma_i^2) dG(\theta_i, \sigma_i^2)}$$

We want to find a cutoff λ such that both constraints are satisfied [5]:

- Capacity: $\int\int P(v_\alpha(Y_i, S_i) > \lambda) dG(\theta_i, \sigma_i^2) \leq \alpha$
- FDR: $\int\int \frac{E[1\{v_\alpha(Y_i,S_i)>\lambda\}(1-v_\alpha(Y_i,S_i))]}{E[1\{v_\alpha(Y_i,S_i)>\lambda\}]} dG(\theta_i, \sigma_i^2) \leq \gamma$

[5] Relaxed discrete optimization problem, following [Basu et al., 2018]

Estimate G

Following [Koenker and Mizera, 2014, Andersen and Andersen, 2010] The primal problem:

$$\min_{f=dG} \left\{ -\sum_i \log g(y_i) \middle| g_i = T(f_i),\ K(f_i) = 1,\ \forall i \right\}$$

where $T(f_i) = \int p(y|\alpha) f_i d\alpha$ and $K(f_i) = \int f_i d\alpha$.
Discretize the support:

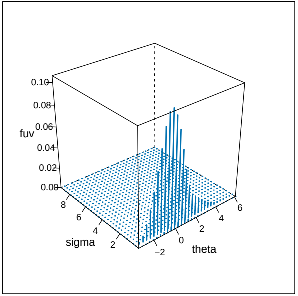
$$\min_{f=dG} \left\{ -\sum_i \log g(y_i) \middle| g = Af,\ 1^T f = 1 \right\}$$

where $A_{ij} = p(y_i|\alpha_j)$ and $f = (f(\alpha_1), f(\alpha_2), \ldots, f(\alpha_m))$.
The dual problem:

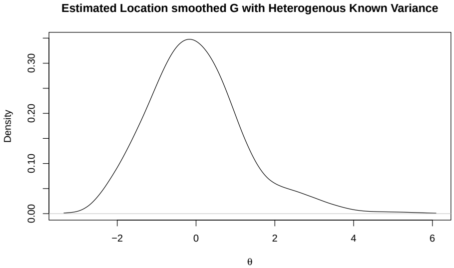
$$\max_{\lambda,\mu} \left\{ \sum_i \log \lambda_1(i) \middle| A^T \lambda_1 < \lambda_2 1,\ (\lambda_1 > 0) \right\}$$

The estimated $\hat{G}$

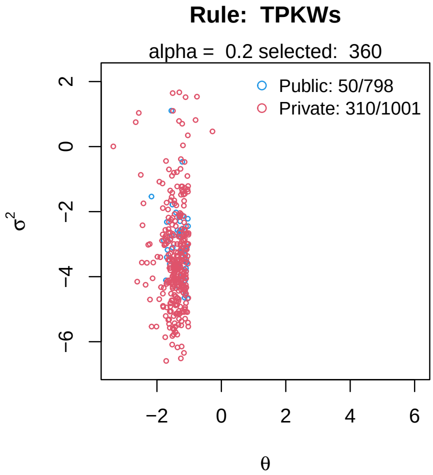
Case 1: σ_i unknown, only S_i observed.

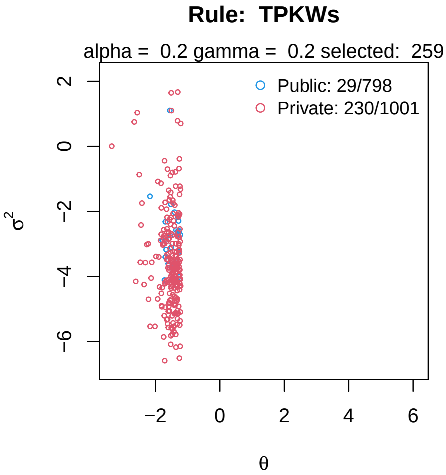


Case 2: σ_i known.

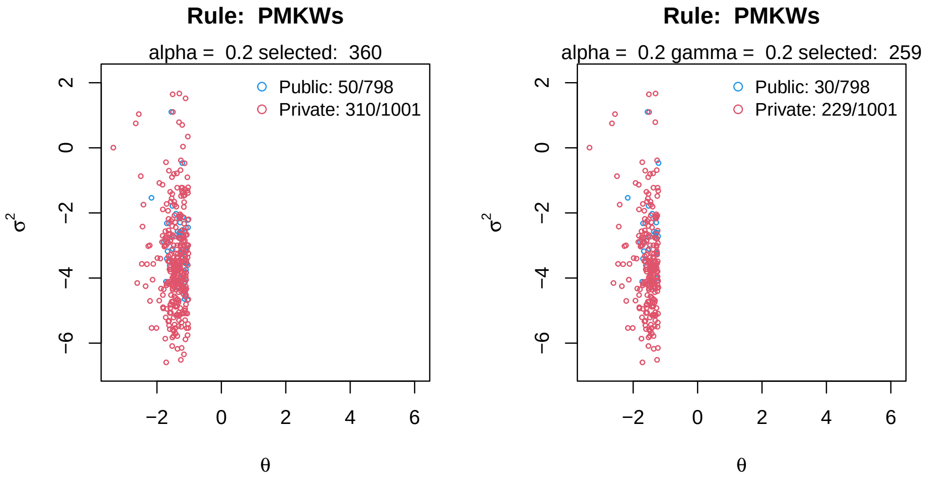


Unknown σ_i: Posterior Tail probability

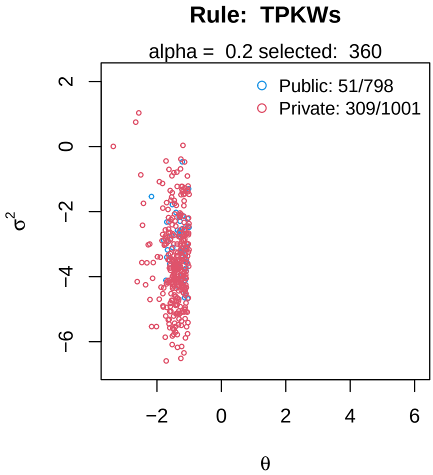


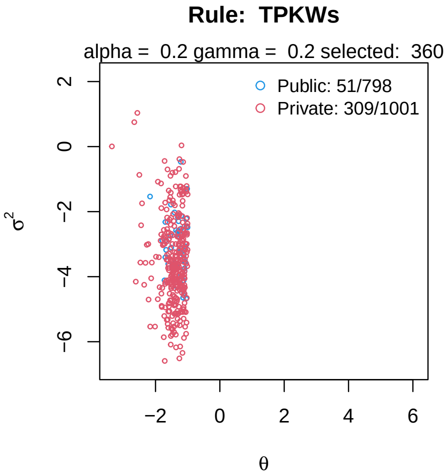


Unknown σ_i: Posterior Mean

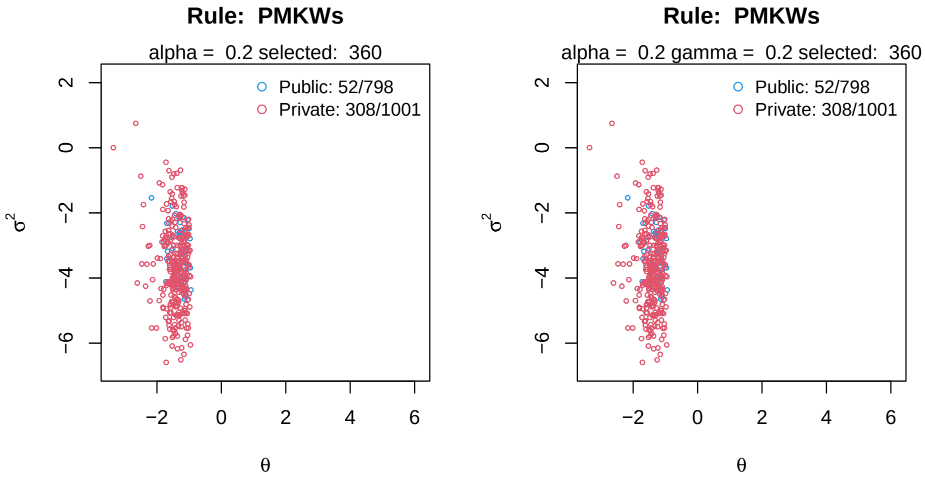


Known σ_i: Posterior Tail probability

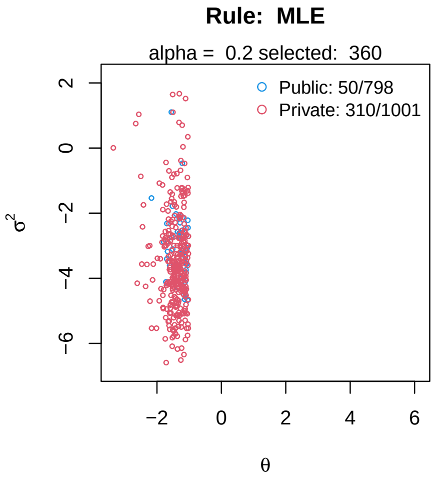




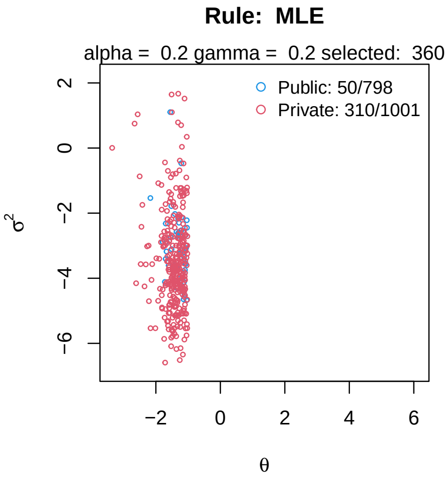
Known σ_i: Posterior Mean



"Face value"



Rule: MLE
alpha = 0.2 selected: 360
Public: 50/798
Private: 310/1001
σ^2
θ
2
0
−2
−4
−6
−2
0
2
4
6



Rule: MLE
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Outline

Conclusion

- Difference in whether to assume known σ_i.
- Control for the False Discovery Rate.
- Private (FP and NP) hospitals are indeed more "efficient".

Limitation

- Interpretation of the θ_i.
- Specification, endogeneity *etc.*
- Normality assumption on ε_{it}. Next

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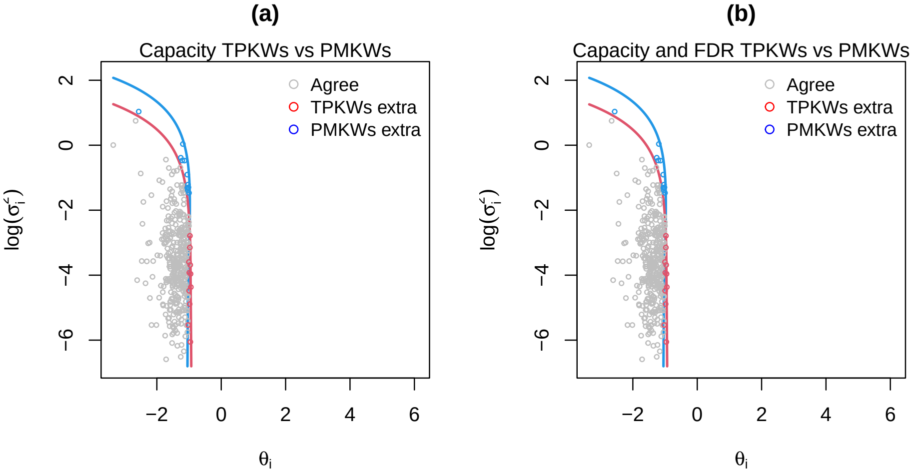
Normality assumption on ε_{it}

Estimate the fixed effect θ_i by

$$
\begin{aligned}
\hat{\theta}_i =& \frac{1}{T} \sum (\theta_i + \varepsilon_{it} + x_{it}(\beta - \hat{\beta})) \\
\xrightarrow{N \to \infty}& \theta_i + \frac{1}{T} \sum_t \varepsilon_{it}
\end{aligned}
$$

When T is relatively small (or even fixed), can't use central limit theorem to claim that $\hat{\theta}_i \xrightarrow{d} \mathcal{N}(\theta_i, \frac{\sigma_i^2}{T})$. $\longrightarrow$ Assume that $\varepsilon_{it} \sim \mathcal{N}(0, \sigma_i^2)$. Back (main) Back (end)

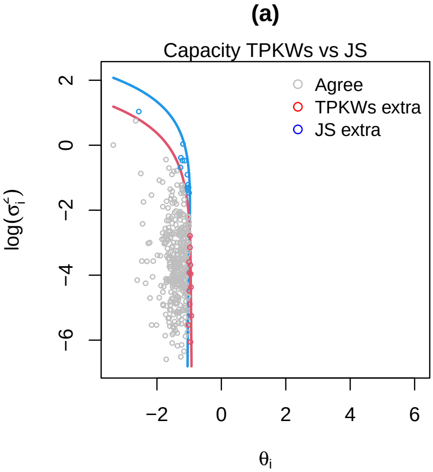
TP vs PM



(a)
Capacity TPKWs vs PMKWs
Agree
TPKWs extra
PMKWs extra
(b)
Capacity and FDR TPKWs vs PMKWs
Agree
TPKWs extra
PMKWs extra

Back

TP vs JS



(a)
Capacity TPKWs vs JS
Agree
TPKWs extra
JS extra
$\log(\sigma_i^2)$
θ_i
2
0
−2
−4
−6
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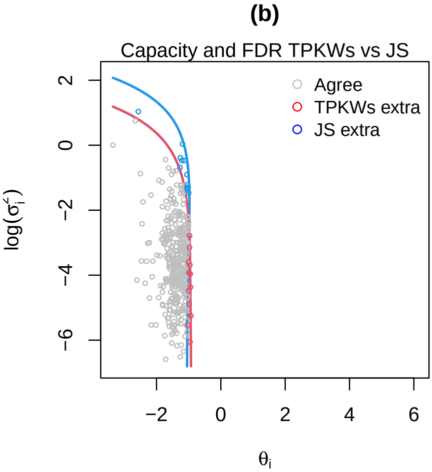
2
4
6



(b)
Capacity and FDR TPKWs vs JS
Agree
TPKWs extra
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$\log(\sigma_i^2)$
θ_i
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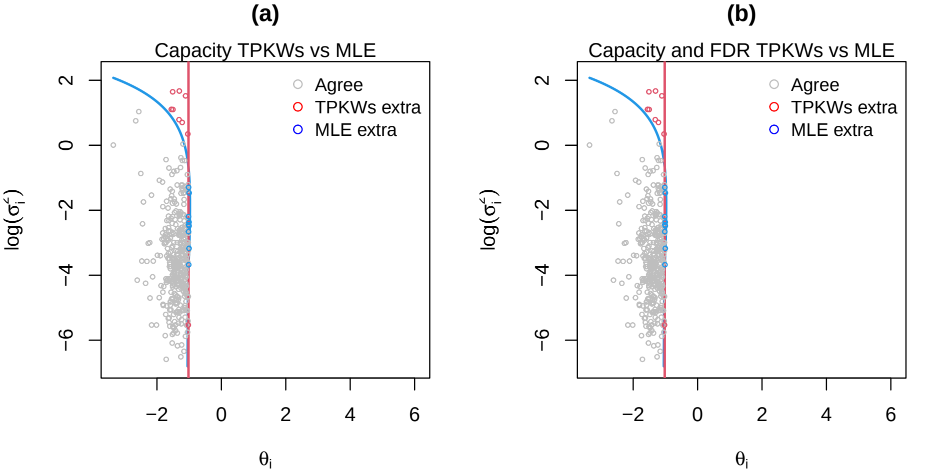
−2
−4
−6
−2
0
2
4
6

TP vs MLE


(a)
Capacity TPKWs vs MLE
Agree
TPKWs extra
MLE extra
$\log(\sigma_i^2)$
θ_i
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Agree
TPKWs extra
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θ_i