

Estimate Dynamic Models with CCP

Problem Set 1

Empirical Industrial Organization 2025 Spring

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Utility (Profit) The profit function of the Harold Zurcher is

$$u(s, Y) = (-RC)Y + (-\mu s)(1 - Y) = (-RC + \mu s)Y - \mu s$$

The realized profit is

$$v(s, Y, \nu_0, \nu_1) = u(s, Y) + \nu_Y$$

Variables For **observed variables**, we have

- s : the state variable, the mileage at the end of period t
- Y : the decision variable, the decision to replace the bus at the end of period t

For **unobserved variables**, we have ν_0, ν_1 .

State transition The state variable s evolves according to this transition function

$$s_{t+1} = s_t + \epsilon_{t+1} \quad \epsilon_{t+1} \sim N(\rho, \sigma_\rho^2)$$

Parameters The profit function parameters:

- μ : the cost of maintaining the bus
- RC : the replacement cost of the bus

The state transition parameters: ρ and σ_ρ^2 . Discount factor $\beta = 0.99$ is given.

Value function The value function is

$$\begin{aligned} V(s, \nu) &= \max_Y \{v(s, Y, \nu) + \beta \mathbb{E}[V(s', \nu')|s, Y]\} \\ &= \max_Y \left\{ u(s, Y) + \nu_Y + \beta \int \mathbb{E}_{\nu'}[V(s', \nu')]f(s'|s)ds' \right\} \end{aligned}$$

To understand this value function, notice that at the end of each period t Harold Zurcher is essentially comparing the **continuation value** from replacing $Y = 1$ and that from not replacing $Y = 0$. We define the **continuation value** as $\tilde{v}(s, Y, \nu)$. The value function is the maximum of them.

$$\begin{aligned} \tilde{v}(s, 1, \nu) &= \tilde{u}(s, 1) + \nu_1 = (-RC + \mu s) + \nu_1 + \beta \int \mathbb{E}_{\nu'}[V(s', \nu')|s, 1]f(s'|s)ds' \\ \tilde{v}(s, 0, \nu) &= \tilde{u}(s, 0) + \nu_0 = \nu_0 + \beta \int \mathbb{E}_{\nu'}[V(s', \nu')|s, 0]f(s'|s)ds' \end{aligned}$$

Therefore the value function can be rewritten as

$$V(s, \nu) = \max_Y \{\tilde{v}(s, 1, \nu), \tilde{v}(s, 0, \nu)\}$$

If we know the $\tilde{u}(s, 1)$ and $\tilde{u}(s, 0)$, and assume ν are iid from T1EV, we are back to the static logit discrete choice model such that

$$\Pr(Y = 1|s) = \frac{\exp(\tilde{u}(s, 1))}{\exp(\tilde{u}(s, 1)) + \exp(\tilde{u}(s, 0))}$$

We follow Bajari et al. (2007) to approximate the continuation value \tilde{v} by

$$\begin{aligned} \tilde{v}(s, 1, \nu) &= \tilde{u}(s, 1) + \nu_1 = x_{i1}^1 + x_{i1}^2\beta_2 + x_{i1}^3 + \nu_1 \\ \tilde{v}(s, 0, \nu) &= \tilde{u}(s, 0) + \nu_0 = x_{i0}^1\beta_1 + x_{i0}^2\beta_2 + x_{i0}^3 + \nu_0 \end{aligned}$$

Where

- $x_i^1 = -\frac{1}{K} \sum_{k=1}^K \sum_{t=1}^T Y_{it}$
- $x_i^2 = -\frac{1}{K} \sum_{k=1}^K \sum_{t=1}^T s_{it}(1 - Y_{it})$
- $x_i^3 = \frac{1}{K} \sum_{k=1}^K \sum_{t=1}^T (\gamma - \ln(\Pr(Y_{it}|s_{it})))$: **I was tripped by this term.** It is $\ln(\Pr(Y = 1|s))$ when $Y_{it} = 1$ and $\ln(\Pr(Y = 0|s))$ when $Y_{it} = 0$.

Therefore, we obtain $RC = \beta_1, \mu = \beta_2$.

The optimal θ^* is the one that maximizes the likelihood function

$$\max_{\theta} \sum_{i=1}^N \sum_{t=1}^T \ln \Pr(Y_{it}|s_{it}, \theta) \tag{1}$$

Question 1 Reduced Form CCP (Probit)

Estimating a reduced form conditional choice probability is the preparation step for simulation. Experimenting with 4 different specifications.

- **probit1**: $\Pr(Y = 1|s) = \Phi(\gamma_0 + \gamma_1 s)$
- **probit2**: $\Pr(Y = 1|s) = \Phi(\gamma_0 + \gamma_1 s + \gamma_2 s^2)$
- **probit3**: $\Pr(Y = 1|s) = \Phi(\gamma_0 + \gamma_1 s + \gamma_2 s^2 + \gamma_3 s^3)$
- **probitlog**: $\Pr(Y = 1|s) = \Phi(\gamma_0 + \gamma_1 \ln(s))$

Then, we plot the estimated CCPs against a range of value for $s \in [0, 40]$. Results are shown in Figure 1. I chose the **probit1** for later simulation.

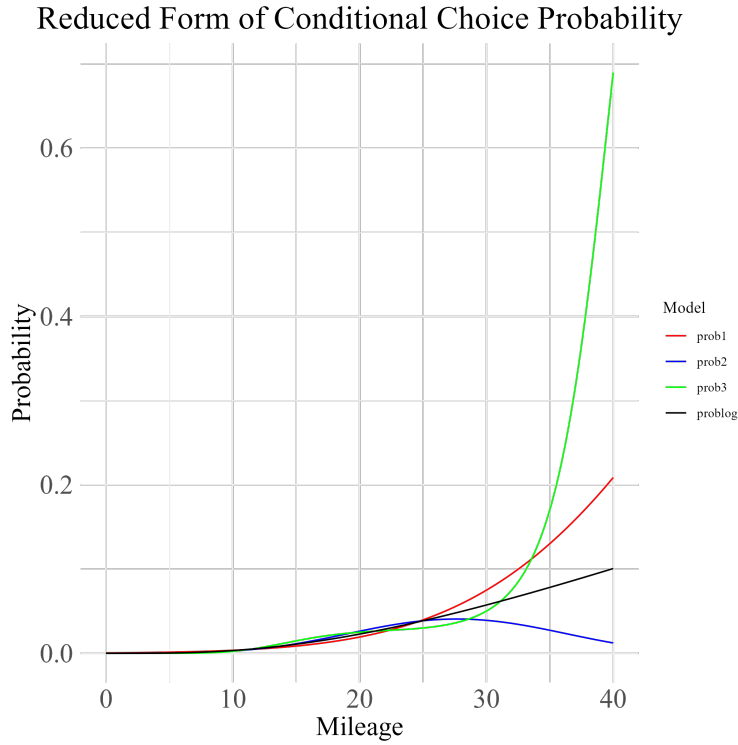


Figure 1: Estimated Reduced Form CCPs from Probit Models

Question 2 State Transition

The transition of states is guided by $\epsilon \in N(\rho, \sigma_\rho^2)$. This is a somewhat unrealistic assumption since the mileage can only increase. However, we proceed as instructed.

- $\rho = 0.19$
- $\sigma_\rho^2 = 0.0129$

Question 3 Simulation

For each observation bus \times period, we perform simulation to get the 6 variables $x_{i1}^1, x_{i1}^2, x_{i1}^3$ and $x_{i0}^1, x_{i0}^2, x_{i0}^3$. Then, we take average over all observations to get Table 1.

| | Replace | Non Replace |
|---------|---------|-------------|
| x_i^1 | -1.19 | -0.57 |
| x_i^2 | -436.09 | -702.38 |
| x_i^3 | 37.81 | 39.42 |

Table 1: Average of Simulation Results

Question 4 Estimation

We estimate by Equation 1 using the simulated approximation. Results are shown in Table 2.

| | <i>Dependent variable:</i> |
|-------------------|----------------------------|
| | choice |
| RC | 8.812*** (0.700) |
| μ | 0.005*** (0.001) |
| Observations | 7,250 |
| Log Likelihood | -319.721 |
| Akaike Inf. Crit. | 643.442 |

Note: *p<0.1; **p<0.05; ***p<0.01

Table 2: Estimation Results

Question 5 Comparison

In order to plot the CCPs estimated from Question 4 against mileage s , we first simulate for each $s \in [0, 40]$ exactly as in Question 3. Then we use the estimated parameters θ^* to compute the CCPs. The results are shown in Figure 2. Alternatively, one could take the

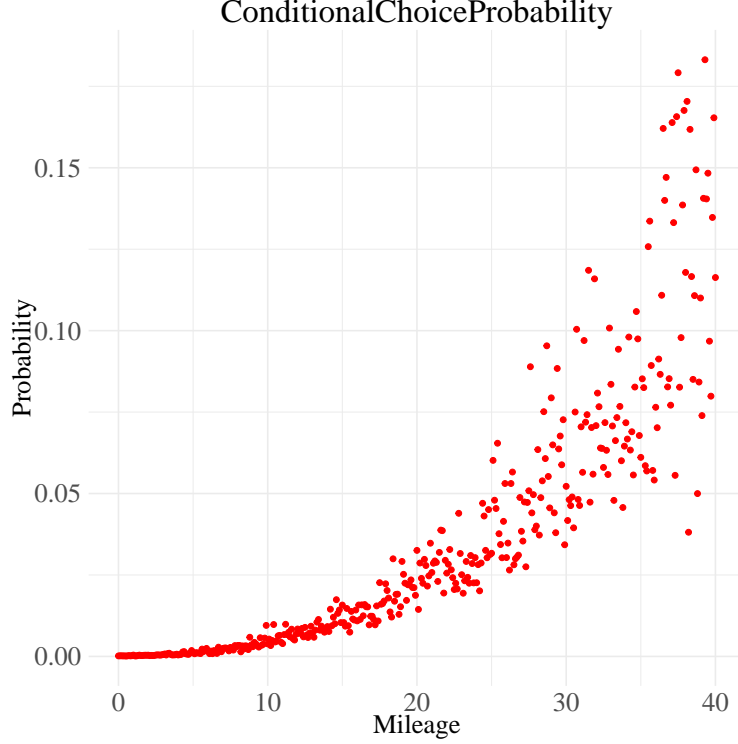


Figure 2: Estimated dynamic CCPs following Bajari et al. (2007).

parameters θ^* and solve the dynamic programming problem analytically as in Rust (1987). Solving for $\bar{V}(s)$ from Equation 2 and then compute CCP from Equation 3.

$$\bar{V}(s) = \gamma + \ln \left(\sum_Y \exp \left(u(s, Y) + \beta \int \bar{V}(s') f(s'|s) ds' \right) \right) \quad (2)$$

$$\Pr(Y = 1|s) = \frac{\exp \{u(s, Y = 1) + \beta \sum_{s'} \bar{V}(s') \Pr(s'|s, Y)\}}{\sum_{Y \in \{0,1\}} \exp \{u(s, Y) + \beta \sum_{s'} \bar{V}(s') \Pr(s'|s, Y)\}} \quad (3)$$

References

- Bajari, P., Benkard, C. L., and Levin, J. (2007). Estimating dynamic models of imperfect competition. *Econometrica*, 75(5):1331–1370.
- Rust, J. (1987). Optimal replacement of gmc bus engines: An empirical model of harold zurcher. *Econometrica: Journal of the Econometric Society*, pages 999–1033.