

<< Algorithm Design >> by Jon Kleinberg and Éva Tardos

§ Chapter 13: Randomized Algorithms

When one thinks about random process, it is usually in one of two distinct ways.

- 1) average-case analysis (randomly generated input)
- 2) consider algorithms that behave randomly \rightarrow randomized algorithm

13.1 A First Application: Contention Resolution

Suppose n processes P_1, P_2, \dots, P_n , each competing for access to a single shared database. Time is being divided into discrete rounds. Database can be accessed by at most one process in a single round. Processes can't communicate with one another at all.

Algorithm

each process will attempt to access the database in each round with probability p , independently of the decisions of the other processes.

Analyzing the Algorithm

let $A[i, t]$ denote the event that P_i attempts to access database in round t . $\Pr[A[i, t]] = p$
let $S[i, t]$ denote the event that P_i succeeds in accessing the database in round t . $\Pr[\overline{A[i, t]}] = 1 - p$

$$\Pr[S[i, t]] = \Pr[A[i, t]] \cdot \prod_{j \neq i} \Pr[\overline{A[j, t]}] = p (1-p)^{n-1}$$

$$f(p) = p (1-p)^{n-1}$$

$$f'(p) = (1-p)^{n-1} - p \cdot (n-1) (1-p)^{n-2}$$

$$(1-p)^{n-1} - (n-1) \cdot p (1-p)^{n-2} \geq 0$$

$$(1-p) \geq (n-1)p$$

$$\frac{1}{p} - 1 \geq n-1$$

$$p \leq \frac{1}{n}$$

we set $p = \frac{1}{n}$, $\Pr[S[i, t]] = \frac{1}{n} (1 - \frac{1}{n})^{n-1}$

(13.1) (a) The function $(1 - \frac{1}{n})^n$ converges monotonically from $\frac{1}{4}$ up to $\frac{1}{e}$ as n increases from 2.

(b) The function $(1 - \frac{1}{n})^{n-1}$ converges monotonically from $\frac{1}{2}$ down to $\frac{1}{e}$ as n increases from 2.

$$\frac{1}{en} \leq \Pr[S[i, t]] \leq \frac{1}{2n}$$

$\sim \Theta(\frac{1}{n})$

let $F[i, t]$ denote the "failure event" that process P_i does not succeed in any of the rounds 1 through t .

$$\Pr[F[i, t]] = \Pr[\bigcap_{r=1}^t \overline{S[i, r]}] = \left[1 - \frac{1}{n} (1 - \frac{1}{n})^{n-1}\right]^t \stackrel{\text{set } t = \lceil en \rceil}{\leq} \left(1 - \frac{1}{en}\right)^t \leq \left(1 - \frac{1}{en}\right)^{en} \leq \frac{1}{e}$$

$$t = \lceil \epsilon n \rceil \cdot c \ln n$$

$$\Pr[F(i, t)] \leq \left(1 - \frac{1}{\epsilon n}\right)^t = \left(\left(1 - \frac{1}{\epsilon n}\right)^{\lceil \epsilon n \rceil}\right)^{c \ln n} \leq e^{-c \ln n} = n^{-c}$$

Conclusion: After $\Theta(n)$ rounds ($t = \lceil \epsilon n \rceil$), the probability that P_i has not succeeded is bounded by a constant ($\frac{1}{e}$); and between then and $\Theta(n \ln n)$, this probability drops to a quantity that is quite small, bounded by an inverse polynomial in n .

Waiting for All Processes to Get Through

$$F_i = \bigcup_{i=1}^n F(i, t)$$

(13.2) [The Union Bound] Given events E_1, E_2, \dots, E_n , we have

$$\Pr\left[\bigcup_{i=1}^n E_i\right] \leq \sum_{i=1}^n \Pr[E_i].$$

$$\therefore \Pr[F_i] \leq \sum_{i=1}^n \Pr[F(i, t)]$$

choose $t = \lceil \epsilon n \rceil (c \ln n)$, $\Pr[F(i, t)] \leq n^{-c}$

choose $t = 2 \lceil \epsilon n \rceil \ln n$

$$\Pr[F_i] \leq \sum_{i=1}^n \Pr[F(i, t)] \leq n \cdot n^{-2} = n^{-1}$$

(13.3) With probability at least $1 - n^{-1}$, all processes succeed in accessing the database at least once within $t = 2 \lceil \epsilon n \rceil \ln n$ rounds.

13.2 Finding the Global Minimum Cut

undirected graph $G = (V, E)$, define a cut of G to be a partition of V into two nonempty sets A and B . For a cut (A, B) , the size of (A, B) is the number of edges with one end in A and the other in B . A global minimum cut is a cut of minimum size.

(13.4) There is a polynomial-time algorithm to find a global min-cut in an undirected graph G .
 convert to directed graph
 fix s . for every $t \in V - \{s\}$, run push-relabel.
 the best among these will be a global min-cut of G .

David Karger 1992.

Algorithm. (Contraction Algorithm)

↳ works with connected multigraph.

The Contraction Algorithm applied to a multigraph $G = (V, E)$:

For each node v , we will record the set $S(v)$ of the nodes that have been contracted into v .

Initially $S(v) = \{v\}$ for each v

If G has two nodes v_1 and v_2 , then return the cut $(S(v_1), S(v_2))$.

Else choose an edge $e = (u, v)$ of G uniformly at random

Let G' be the graph resulting from the contraction of e ,
with a new node z_{uv} replacing u and v .

Define $S(z_{uv}) = S(u) \cup S(v)$

Apply the Contraction Algorithm recursively to G'

Endif

Analyzing the Algorithm

(13.5) The Contraction Algorithm returns a global min-cut of G with probability at least $\frac{1}{\binom{n}{2}}$

Suppose the global min-cut has size k , a set F of k edges with one end in A and the other in B . \hookrightarrow every node in G has degree at least k .

We want an upper bound on the probability that an edge in F is contracted, and for this we need a lower bound on the size of E .

$$|E| \geq \frac{1}{2} kn$$

Hence the probability that an edge in F is contracted is at most

$$\frac{k}{\frac{1}{2} kn} = \frac{2}{n}.$$

Consider the situation after j iterations, there are $n-j$ supernodes in G' .

Thus G' has at least $\frac{1}{2} k(n-j)$ edges. So the probability that an edge of F is contracted in the next iteration $j+1$ is at most.

$$\frac{k}{\frac{1}{2} k(n-j)} = \frac{2}{n-j}$$

We write E_j for the event that an edge of F is not contracted in iteration j , then we have shown $\Pr[E_1] \geq 1 - \frac{2}{n}$ and $\Pr[E_{j+1} | E_1 \cap E_2 \cap \dots \cap E_j] \geq 1 - \frac{2}{n-j}$

We are interested in lower-bounding the quantity $\Pr[E_1 \cap E_2 \dots \cap E_{n-2}]$.

$$\begin{aligned} & \Pr[E_1] \cdot \Pr[E_2 | E_1] \cdots \Pr[E_{j+1} | E_1 \cap E_2 \cdots \cap E_j] \cdots \Pr[E_{n-2} | E_1 \cap E_2 \cdots \cap E_{n-3}] \\ & \geq \left(1 - \frac{2}{n}\right) \cdot \left(1 - \frac{2}{n-1}\right) \cdots \left(1 - \frac{2}{n-j}\right) \cdots \left(1 - \frac{2}{3}\right) \\ & = \left(\frac{n-2}{n}\right) \cdot \left(\frac{n-3}{n-1}\right) \left(\frac{n-4}{n-2}\right) \cdots \left(\frac{2}{5}\right) \left(\frac{2}{4}\right) \left(\frac{1}{3}\right) \\ & = \frac{2}{n(n-1)} = \frac{1}{\binom{n}{2}} \end{aligned}$$

So we know that a single run of the Contraction Algorithm fails to find a global min-cut with probability at most $(1 - 1/\binom{n}{2})$.

If we run the algorithm $\binom{n}{2}$ times,

$$\left(1 - \frac{1}{\binom{n}{2}}\right)^{\binom{n}{2}} \leq \frac{1}{e}$$

If we run the algorithm $\binom{n}{2} \ln n$ times

$$\left[\left(1 - \frac{1}{\binom{n}{2}}\right)^{\binom{n}{2}}\right]^{\ln n} \leq e^{-\ln n} = n^{-1}$$

Further Analysis: The Number of Global Minimum Cuts

Given an undirected graph $G = (V, E)$, what is the maximum number of global min-cuts it can have (as a function of n)?

undirected graph has $\binom{n}{2}$ global min-cuts.

(13.6) An undirected graph $G = (V, E)$ on n nodes has at most $\binom{n}{2}$ global min-cuts.

Let G be a graph, and let C_1, \dots, C_r denote all its global min-cuts.

Let E_i denote the event that C_i is returned by the Contraction Algorithm,

let $E = \bigcup_{i=1}^r E_i$ denote the event that the algorithm returns any global min-cut.

$$\Pr[E] = \Pr\left[\bigcup_{i=1}^r E_i\right] = \sum_{i=1}^r \Pr[E_i] \geq r / \binom{n}{2}$$

$$\Pr[E] \leq 1 \Rightarrow r \leq \binom{n}{2}$$

13.3 Random Variables and Their Expectations

random variable: Given a probability space, a random variable X is a function from the underlying sample space to the natural numbers, such that for each natural number j , the set $X^{-1}(j)$ of all sample points taking value j is an event.

$\Pr[X=j]$ as loose shorthand for $\Pr[X^{-1}(j)]$.

expectation: the "average value" assumed by X .

$$E[X] = \sum_{j=0}^{\infty} j \cdot \Pr[X=j]$$

$$S = \sum_{j=1}^{\infty} j (1-p)^j = 1(1-p)^1 + 2(1-p)^2 + 3(1-p)^3 + \dots$$

$$(1-p)S = 1(1-p)^2 + 2(1-p)^3 + 3(1-p)^4 + \dots$$

$$S - (1-p)S = (1-p)^1 + (1-p)^2 + (1-p)^3 + \dots = \frac{(1-p)}{p}$$

$$pS = \frac{(1-p)}{p} \Rightarrow S = \frac{(1-p)}{p^2}$$

$$E[X] = \sum_{j=0}^{\infty} j \cdot \Pr[X=j] = \sum_{j=1}^{\infty} j (1-p)^{j-1} p = \frac{p}{1-p} \sum_{j=1}^{\infty} j (1-p)^{j-1} = \frac{p}{1-p} \cdot \frac{(1-p)}{p^2} = \frac{1}{p}$$

(13.8) Linearity of Expectation. Given two random variables X and Y defined over the same probability space, we can define $X+Y$ to be the random variable equal to $X(\omega) + Y(\omega)$ on a sample point ω . For any X and Y , we have

$$E[X+Y] = E[X] + E[Y].$$

(13.10) $H(n) = \sum_{i=1}^n \frac{1}{i} = 1 + \frac{1}{2} + \frac{1}{3} + \dots + \frac{1}{n}$ harmonic number $H(n)$.

$$\int_1^{n+1} \frac{1}{x} dx = \ln(n+1)$$

$$\ln(n+1) < H(n) < 1 + \ln n, \text{ and more loosely } H(n) = \Theta(\log n).$$

Conditional Expectation

Suppose we have a random variable X and an event E of positive probability. Then we define the conditional expectation of X , given E , to be the expected value of X computed only over the part of sample space corresponding to E .

$$E[X|E] = \sum_{j=0}^{\infty} j \cdot P[X=j|E].$$

13.4 A Randomized Approximation Algorithm for MAX 3-SAT

Algorithm: set each variable x_1, \dots, x_n independently to 0 or 1 with probability $\frac{1}{2}$ each. Let $Z_i = 1$ if clause C_i is satisfied, and 0 otherwise.

Thus $Z = Z_1 + Z_2 + \dots + Z_k$. $E[Z_i]$ is equal to the probability that C_i is satisfied.

In order for C_i not to be satisfied, each of its three variables must be assigned the value that fails to make it true, $(\frac{1}{2})^3 = \frac{1}{8}$, so $E[Z_i] = \frac{7}{8}$.

$$E[Z] = E[Z_1] + E[Z_2] + \dots + E[Z_k] = \frac{7}{8}k.$$

(13.14) Consider a 3-SAT formula, where each clause has three different variables. The expected number of clauses satisfied by a random assignment is within an approximation factor $\frac{7}{8}$ of optimal.

(13.15) For every instance of 3-SAT, there is a truth assignment that satisfies at least a $\frac{7}{8}$ fraction of all clauses.

Corollary: Every instance of 3-SAT with at most seven clauses is satisfiable.

(13.16) There is a randomized algorithm with polynomial expected running time that is guaranteed to produce a truth assignment satisfying at least $\frac{7}{8}$ fraction of all clauses. $P \geq \frac{1}{8k}$ $n \leq 8/k$