```
#net1 <- network(Adj, directed=F)</pre>
### library(latentnet)
#net2 <- ergmm(net1~euclidean(d=2), tofit=c("mle"))</pre>
#coords <- net2$mle$Z</pre>
### library(MCMCpack)
#coords = procrustes(coords, Z, translation = T, dilation = F)$X.new
\#plot(coords, xlab = "Z_1", ylab = "Z_2", type = "p")
#points(Z, col = "red") # true position
(case 2. dilation = T)
net1 <- network(Adj, directed=F)</pre>
# library(latentnet)
net2 <- ergmm(net1~euclidean(d=2), tofit=c("mle"))</pre>
coords <- net2$mle$Z</pre>
# library(MCMCpack)
 \textit{\# coords = procrustes(coords, Z, translation=T, dilation=F)$X.new } \\
coords = procrustes(coords, Z, translation=T, dilation=T)$X.new
plot(coords, xlab = "Z_1", ylab = "Z_2", col = cl, pch = 19)
points(Z, col = cl) # true position
      \alpha
      0
     7
                                                                              0
      9
                            0
                                 -2
                                               0
                                                             2
                                                                          4
                                                                                        6
                    -4
                                                Z_1
```

```
msq_dev1 = sum((coords - Z)^2) / N
msq_dev1
```

sum of squared deviations from the truth:

## [1] 4.784991

Implement JAGS:

Step 1: Define the model using the BUGS language:

```
### JAGS, in BUGS language
func1 <- function(){</pre>
 for (i in 1:N){
        for (j in 1:N){
            dist[i,j] \leftarrow pow(pow((z[i,1] - z[j,1]), 2) + pow((z[i,2] - z[j,2]), 2), 0.5)
            M[i,j] <- sigmasq * exp(- phi * dist[i,j])</pre>
            probit(padj[i,j]) <- a + b * abs(x[i]-x[j]) - exp(theta) * dist[i,j]</pre>
            adj[i,j] ~ dbern(padj[i,j])
        }
  }
  for (i in 1:N){
       mu x[i] <- beta</pre>
        z[i, 1:2] ~ dmnorm(mu_z[categ, 1:K], inv_kappasq[categ] * I[1:K, 1:K])
   }
   x[1:N] ~ dmnorm(mu_x[1:N], error_prec[1:N, 1:N]) ### x is the attributes
    ### Spatial Priors
   for (h in 1:H){
      mu_z[h, 1:2]
                        ~ dmnorm(mu_prior1[1:K], (1/mu_prior2) * I[1:K, 1:K])
                          - dgamma(kappa_prior1, kappa_prior2)
      inv_kappasq[h]
                        ~ ddirch(alpha[1:H])
  omega[1:H]
                        ~ dcat(omega[1:H])
  categ
                                               ### categorical distribution
                          ~ dnorm(0.0, 0.0001)
   beta
   inv_sigmasq
                         ~ dgamma(sigmasq_prior1, sigmasq_prior2)
                         <- 1 / inv_sigmasq
   sigmasq
                         ~ dgamma(tausq_prior1, tausq_prior2)
   inv_tausq
   tausq
                         <- 1 / inv_tausq
                         ~ dunif(phi_lower, phi_upper)
   phi
   var_all[1:N, 1:N] <- M[1:N, 1:N] + tausq * I[1:N, 1:N]</pre>
   error_prec[1:N, 1:N] <- inverse(var_all[1:N, 1:N])</pre>
                          ~ dnorm(amean_prior, 1/avar_prior)
   a
                          ~ dnorm(bmean_prior, 1/bvar_prior)
```

```
### specify independent priors parameters
data1 <- list(adj</pre>
                    = Adj,
                                 ### adjacent matrix N by N
                         = as.vector(x), ### the observed N attributes s(z_i)
            mu_prior1 = c(0, 0),
            mu_prior2
                        = 4,
            kappa_prior1 = 2,
            kappa_prior2 = 2,
                     = c(2, 2, 2),
            alpha
            sigmasq_prior1 = 2,
            sigmasq_prior2 = 0.1,
            tausq_prior1 = 2,
            tausq_prior2 = 0.02,
                        = 0.01,
            phi_lower
            phi_upper
                        = 1,
            amean_prior = 5,
            avar_prior = 1,
            bmean_prior = 13,
            bvar prior = 2,
            tmean_prior = 1.5,
            tvar_prior = 2,
            N = N,
                   ### latent space dim
            K = K
            H = H, ### assume N > H.
            I = diag(rep(1, N))
if (length(data1$alpha) != H){
 print("dimension error: alpha dimention does not comform")
```

Determine the parameters in the priors:

## Step 2: Read in the model file using the "jags.model()" function.

This creates an object of class "jags".

```
## Observed stochastic nodes: 10001
## Unobserved stochastic nodes: 115
## Total graph size: 170267
##
## Initializing model
```

Step 3: Update the model using the update method for "jags" objects.

This constitutes a burn-in period.

```
update(jags1, n.iter = 5000)
```

Step 4: Extract samples from the model object using the "coda.samples()" function.

## PART III:

Diagnostics-

```
<- jags2$z
z_res
omega_res <- jags2$omega
mu_z_res <- jags2$mu_z
inv_kappasq_res <- jags2$inv_kappasq</pre>
tausq_res <- jags2$tausq
sigmasq_res <- jags2$sigmasq
phi_res <- jags2$phi
beta_res <- jags2$beta
                 <- jags2$a
a_res
                <- jags2$b
b_res
theta_res
                <- jags2$theta
# dim(z_new)
# dim(omega_res)
# dim(mu_z_res)
# dim(tausq_res)
kappasq_res <- 1 / inv_kappasq_res</pre>
true_para = c(phi, sigmasq, tausq, beta,
                a, b, theta,
                omega[1], omega[2],
                kappasq[])
```

```
mcmc_para = c(mean(phi_res), mean(sigmasq_res), mean(tausq_res), mean(beta_res),
              mean(a_res), mean(b_res), mean(theta_res),
              mean(omega_res[1,,1]), mean(omega_res[2,,1]),
              mean(kappasq_res[1,,1]), mean(kappasq_res[2,,1]), mean(kappasq_res[3,,1]))
compare_para = data.frame(true_para, mcmc_para)
rownames(compare para) <- c("phi", "sigmasq", "tausq", "beta",</pre>
                             "a", "b", "theta",
                             "omega[1]", "omega[2]",
                             "kappasq[1]", "kappasq[2]", "kappasq[3]")
compare_para
##
              true_para
                          mcmc_para
## phi
                  0.50 0.635259024
## sigmasq
                   0.10 0.125845514
                  0.02 0.009188966
## tausq
## beta
                  1.00 1.028035571
## a
                  5.00 1.207795152
## b
                 13.00 8.347637485
## theta
                  1.50 0.802770004
## omega[1]
                  0.30 0.347527929
## omega[2]
                  0.40 0.287694991
                  3.00 4.119292285
## kappasq[1]
                  4.00 1.975491796
## kappasq[2]
## kappasq[3]
                   5.00 4.870626849
compare mu:
mu
##
        [,1] [,2]
## [1,]
          -2
## [2,]
## [3,]
               -3
mu_z_res
## mcarray:
              [,1]
                           [,2]
##
## [1,] 0.5567184 0.012135555
## [2,] -0.0199149 0.005530977
## [3,] 0.3726574 0.225281425
## Marginalizing over: iteration(10000),chain(1)
compare Z:
z_new <- array(NA, dim = dim(z_res))</pre>
for(i in 1:dim(z_res)[3]){
    z_new[, , i, 1] <- procrustes(z_res[,,i,1], Z, translation = T,dilation = T)$X.new</pre>
}
```

```
dim(z_new)
Procrustes transformations on locations Z:
##
                       iteration
                                     chain
##
         100
                     2
                           10000
z_{post_mean} = apply(z_{new}[, , , 1], 1:2, mean)
z_post_mean
##
                             [,2]
                 [,1]
##
     [1,] 1.96755262 1.16053786
     [2,] -0.66920957 -0.17380971
##
##
     [3,] 0.79104920 3.54408877
##
     [4,] 0.96240748 -3.45116231
     [5,] -0.15101875 -5.03770368
##
##
     [6,] -3.11935599 1.14650434
##
     [7,] 3.07685335 -0.47421366
##
     [8,] -2.58795722 1.36059286
##
     [9,] -2.09114949 0.64883743
    [10,] 1.29338431
##
                      1.42641583
##
    [11,] 3.46666331
                      2.40372788
    [12,] -2.79387484 1.35409349
##
    [13,] -0.28666017
                       2.53187221
    [14,] -3.64765577 -2.89982811
##
##
    [15,] -1.70902041 2.06482950
   [16,] 4.23219389 0.25159089
##
    [17,] -3.62858893 -1.47488869
##
    [18,] 6.54042520 0.94182935
##
    [19,] -1.91578192 2.38386619
##
    [20,] -1.43086918 2.75110072
##
    [21,] -0.39413128 2.62859808
    [22,] -0.14773322 0.74182166
##
    [23,]
         1.02432918 3.19094134
##
    [24,] 0.57764144
                      0.44573109
##
    [25,] 3.72768599 0.07738893
##
    [26,] -4.29343865 1.11631425
    [27,] 1.72446216 -3.50269136
##
    [28,] -0.21718160 2.71031145
    [29,] -1.66210211 4.53507784
##
##
    [30,] 1.73003510 -3.39302616
          1.22114705 -0.01653950
##
    [31,]
    [32,] 4.29706338 1.64148450
##
##
    [33,] -2.47686287 2.41749243
##
    [34,] -2.01599852 -1.38061029
##
    [35,] 1.26212648 0.90776816
    [36,] 3.04049356 -2.48757164
##
##
    [37,] 4.28701818 1.18101286
##
   [38,] -2.91756738 -0.28009995
```

[39,] -4.19059228 -2.27890674

[40,] 3.47616809 -0.82657554

[41,] -3.02182371 -4.21277527

[42,] -2.36649683 1.59006289 [43,] -4.39958297 0.33925877

## ##

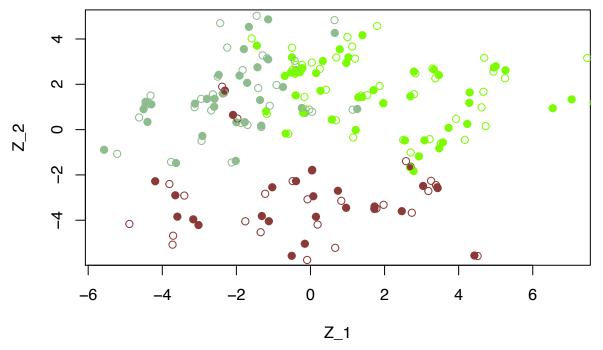
##

##

```
[44,] -1.44527720 3.70768320
    [45,] 0.96791423 2.94141590
    [46,] -1.31250825 -3.81426310
   [47,] -0.69225782 2.36661872
    [48,] 2.46600389 -3.59975774
##
    [49,] 4.42508074 -5.55735478
    [50,] -0.38975192 -2.27490219
    [51,] -4.50465357 0.89817945
##
    [52,] 3.42215783 -2.54479796
##
    [53,] 1.40297144 4.16825048
    [54,] -3.59276472 -3.84240851
    [55,] 3.57464847 -0.56488120
##
    [56,] -2.00060067 0.32370660
##
    [57,] 1.39491758 1.43509089
    [58,] -0.49906785 3.18965560
##
    [59,] -0.39858660 1.51174137
##
    [60,] -0.41086495 2.51974014
##
    [61,] 5.26457230 2.61761517
    [62,] 2.79253915 2.55934985
    [63,] -1.14143852 3.09967485
##
    [64,] 0.03992572 -1.80309706
    [65,] 0.33202026 3.02981404
##
    [66,] 5.00033253 2.79599266
    [67.] -0.53253916 2.52136127
##
    [68,] -0.19836617 0.73925825
    [69,] 7.04981160 1.33354187
##
    [70,] 0.14775805 -3.84662342
    [71,] 2.55064807 -0.47856544
##
    [72,] 3.43263368 -2.58313021
    [73,] 0.65536297 4.26712993
    [74,] -5.57098151 -0.89088271
##
    [75,] -2.30959731 1.72189301
    [76,] 0.27041317 1.71374308
##
    [77,] -1.14209806 4.86602689
##
    [78,] 1.70216082 1.74077294
##
    [79,] -1.77271104 0.33009723
    [80,] -1.03138653 -2.54570527
##
    [81,] -4.45250046 1.22297006
##
    [82,] -1.79757568 3.55362473
##
    [83,] -1.32895936 0.17447894
    [84,] 2.78370387 -1.83000161
    [85,] 4.95384277 2.74057737
##
    [86,] 0.74118147 -2.69960676
##
    [87,] 0.07694291 -2.94185121
    [88,] 2.70005121 -1.63455583
    [89,] -3.16366033 -3.96174108
##
##
    [90,] 3.33262344 2.64955624
##
    [91,] 2.92796602 -1.18154135
    [92,] -1.36075485 1.30416847
##
    [93,] -0.23145893 0.95503036
##
    [94,] -1.19771491 0.80663606
##
   [95,] -1.12337553 -4.04291223
   [96,] -0.50577405 -5.57026824
##
## [97,] 2.69591149 1.47211401
```

```
## [98,] 0.15211075 2.49627598
## [99,] -2.59965803 1.01018427
## [100,] -0.74378214 1.87956379

plot(z_post_mean, xlab = "Z_1", ylab = "Z_2", col = cl, pch = 19)
points(Z, col = cl) # true position
```



```
msq_dev2 = sum((z_post_mean - Z)^2) / N
msq_dev2
```

sum of squared deviations from the truth "Z":

```
## [1] 0.519649
```

```
### compare with latent position model without spatial effect:
msq_dev1
```

## [1] 4.784991

## Markov chain trace plot:

```
# dim(phi_res)
phi_mcmc <- phi_res[1, , 1]
plot(phi_mcmc, type = "1")</pre>
```