# Novel R tools for analysis of genome-wide population genetic data with emphasis on clonality

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## ABSTRACT

To gain a detailed understanding of how plant microbes evolve and adapt to hosts, pesticides, and other factors, knowledge of the population dynamics and evolutionary history of populations is crucial. Plant pathogen populations are often clonal or partially clonal which requires different analytical tools. With the advent of high throughput sequencing technologies, obtaining genome-wide population genetic data has become easier than ever before. We previously contributed the R package *poppr* specifically addressing issues with analysis of clonal populations. In this paper we provide several significant extensions to *poppr* with a focus on large, genome-wide SNP data. Specifically, we provide several new functionalities including the new function `mlg.filter` to define clone boundaries allowing for inspection and definition of what is a clonal lineage, a sliding-window analysis of the index of association, modular bootstrapping of any genetic distance, and analyses across any level of hierarchies.

**Keywords: clonality, population genomics, bootstrap, index of association, hierarchical analysis, sliding window**

## INTRODUCTION

To paraphrase Dobzhansky, nothing in the field of plant-microbe interactions makes sense except in the light of population genetics (Dobzhansky, 1973). Genetic forces such as selection and drift act on alleles in a population. Thus, a true understanding of how plant pathogens emerge, evolve and adapt to crops, fungicides, or other factors, can only be elucidated in the context of population level phenomena given the demographic history of populations (McDonald and Linde, 2002; Grünwald and Goss, 2011; Milgroom et al., 1989). The field of population genetics, in the era of whole genome resequencing, provides unprecedented power to describe the evolutionary history and population processes that drive coevolution between pathogens and hosts. This powerful field thus critically enables effective deployment of R genes, design of pathogen informed plant resistance breeding programs, and implementation of fungicide rotations that minimize emergence of resistance.

Most computational tools for population genetics are based on concepts developed for sexual model organisms. Populations that reproduce clonally or are polyploid are thus difficult to characterize using

classical population genetic tools because theoretical assumptions underlying the theory are violated. Yet, many plant pathogen populations are at least partially clonal if not completely clonal (Milgroom, 1996; Anderson and Kohn, 1995). Thus, development of tools for analysis of clonal or polyploid populations is needed.

Genotyping by sequencing and whole genome resequencing provide the unprecedented ability to identify thousands of single nucleotide polymorphisms (SNPs) in populations (Elshire et al., 2011; Luikart et al., 2003; Davey et al., 2011). With traditional marker data (e.g., SSR, AFLP) a clone was typically defined as a unique multilocus genotype (MLG) (Grünwald and Hoheisel, 2006; Falush et al., 2003; Goss et al., 2009; Cooke et al., 2012; Taylor and Fisher, 2003). Availability of large SNP data sets provides new challenges for data analysis. These data are based on reduced representation libraries and high throughput sequencing with moderate sequencing depth which invariably results in substantial missing data, error in SNP calling due to sequencing error, lack of read depth or other sources of spurious allele calls (Mastretta-Yanes et al., 2015). It is thus not clear what a clone is in large SNP data sets and novel tools are required for definition of clone boundaries.

The research community using the R statistical and computing language (R Core Team, 2015) has developed a plethora of new resources for population genetic analysis (Paradis, 2010; Jombart, 2008). R is particularly appealing because all code is open source and functions can be evaluated and modified by any user. Recently, we introduced the R package *poppr* specifically developed for analysis of clonal populations (Kamvar et al., 2014b). *Poppr* previously introduced several novel features including the ability to conduct a hierarchical analysis across unlimited hierarchies, test for linkage association, graph minimum spanning networks or provide bootstrap support for Bruvo's distance in resulting trees. *Poppr* has been rapidly adopted and applied to a range of studies including for example horizontal transmission in leukemia of clams (Metzger et al., 2015), study of the vector-mediated parent-to-offspring transmission in an avian malaria-like parasite (Chakarov et al., 2015), and characterization of the emergence of the invasive forest pathogen *Hymenoscyphus pseudoalbidus* (Gross et al., 2014). It has also been used to implement real-time, online shinyR based tools for visualizing relationships among unknown MLGs in reference databases (http ://phytophthora-id.org/) (Grünwald et al., 2011).

Here, we introduce *poppr* 2.0, which provides a major update to *poppr* (Kamvar et al., 2014b) including novel tools for analysis of clonal populations specifically addressing large SNP data. Significant novel tools include functions for calculating clone boundaries and collapsing individuals into clonal groups based on a user-specified genetic distance threshold, sliding window analyses, genotype accumulation curves, reticulations in minimum spanning networks, and bootstrapping for any genetic distance.

# IMPLEMENTATIONS AND EXAMPLES

## CLONAL IDENTIFICATION

As highlighted in previous work, clone correction is an important component of population genetic analysis of organisms that are known to reproduce asexually (Kamvar et al., 2014b; Milgroom, 1996; Grünwald et al., 2003). This method is a partial correction for bias that affects metrics that rely on allele frequencies assuming panmixia. It was initially designed for data with only a handful of markers. With the advent of large-scale sequencing and reduced- representation libraries, it has become easier to sequence tens of thousands of markers from hundreds of individuals (Elshire et al., 2011; Davey et al., 2011; Davey and Blaxter, 2010). With this larger number of markers, the genetic resolution is much greater, but the chance of genotyping error is also greatly increased and missing data is frequent (Mastretta-Yanes et al., 2015). Taking this fact and occasional somatic mutations into account, it would be impossible to separate true clones from independent individuals by just comparing what MLGs are different. We introduce a new method for collapsing unique multilocus genotypes determined by naive string comparison into multilocus lineages utilizing any genetic distance given three different clustering algorithms: farthest neighbor, nearest neighbor, and UPGMA (average neighbor) (Sokal, 1958).

72  These clustering algorithms act on a distance matrix that is either provided by the user or generated via
73  a function that will calculate a distance from genclone objects such as `bruvo.dist`, which in particular
74  applies to any level of ploidy (Bruvo et al., 2004). All algorithms have been implemented in C and
75  utilize the OpenMP framework for optional parallel processing (Dagum and Menon, 1998). Default is the
76  conservative farthest neighbor algorithm (Fig. 1A), which will only cluster samples together if all samples
77  in the cluster are at a distance less than the given threshold. By contrast, the nearest neighbor algorithm
78  will have a chaining effect that will cluster samples akin to adding links on a chain where a sample can
79  be included in a cluster if all of the samples have at least one connection below a given threshold (Fig.
80  1C). The UPGMA, or average neighbor clustering algorithm is the one most familiar to biologists as it is
81  often used to generate ultra-metric trees based on genetic distance (Fig. 1B). This algorithm will cluster by
82  creating a representative sample per cluster and joining clusters if these representative samples are closer
83  than the given threshold.

84  We utilize data from the microbe *Phytophthora infestans* to show how the `mlg.filter` function
85  collapses multilocus genotypes with Bruvo's distance assuming a genome addition model (Bruvo et
86  al., 2004). *P. infestans* is the causal agent of potato late blight originating from Mexico that spread to
87  Europe in the mid 19th century (Goss et al., 2014; Yoshida et al., 2013). *P. infestans* reproduces both
88  clonally and sexually. The clonal lineages of *P. infestans* have been formally defined into 18 separate
89  clonal lineages using a combination of various molecular methods including AFLP and microsatellite
90  markers (Lees et al., 2006; Li et al., 2013). For these data, we used `mlg.filter` to detect all of the
91  distance thresholds at which 18 multilocus lineages would be resolved. We used these thresholds to define
92  multilocus lineages and create contingency tables and dendrograms to determine how well the multilocus
93  lineages were detected.

94  For the *P. infestans* population, the three algorithms were able to detect 18 multilocus lineages at
95  different distance thresholds (Fig. 2). Contingency tables between the described multilocus genotypes
96  and the genotypes defined by distance show that most of the 18 lineages were resolved, except for US-8,
97  which is polytomic (Table 1).

98  We utilized simulated data to evaluate the effect of sequencing error and missing data on MLG calling.
99  We constructed the data using the `glSim` function in *adegenet* (Jombart and Ahmed, 2011) to obtain a
100 SNP data set for demonstration. Two diploid data sets were created, each with 10k SNPs (25% structured
101 into two groups) and 200 samples with 10 ancestral populations of even sizes. Clones were created in one
102 data set by marking each sample with a unique identifier and then randomly sampling with replacement.
103 It is well documented that reduced- representation sequencing can introduce several erroneous calls and
104 missing data (Mastretta-Yanes et al., 2015). To reflect this, we mutated SNPs at a rate of 10% and inserted
105 an average of 10% missing data for each sample after clones were created, ensuring that no two sequences
106 were alike. The number of mutations and missing data per sample were determined by sampling from
107 a Poisson distribution with $\lambda = 1000$. After pooling, 20% of the data set was randomly sampled for
108 analysis. Genetic distance was obtained with the function `bitwise.dist`, which calculates the fraction
109 of different sites between samples equivalent to Provesti's distance, counting missing data as equivalent
110 in comparison (Prevosti et al., 1975).

111 All three filtering algorithms were run with a threshold of 1, returning a numeric vector of length $n - 1$
112 where each element represented a threshold at which two samples/clusters would join. Since each data set
113 would have varying distances between samples, the clonal boundary threshold was defined as the midpoint
114 of the largest gap between two thresholds that collapsed less than 50% of the data.

115 Out of the 100 simulations run, we found that across all methods, detection of duplicated samples had
116 $\sim 98\%$ true positive fraction and $\sim 0.8\%$ false positive fraction indicating that this method is robust to
117 simulated populations (supplementary information).

## MINIMUM SPANNING NETWORKS WITH RETICULATION

118 In its original iteration, *poppr* introduced minimum spanning networks that were based on the *igraph*
119 function `minimum.spanning.tree` (Csardi and Nepusz, 2006). This algorithm produces a minimum
120 spanning tree with no reticulations where nodes represent individual MLGs. In other minimum spanning
121 network programs, reticulation is obtained by calculating the minimum spanning tree several times and
122 returning the set of all edges included in the trees. Due to the way *igraph* has implemented Prim's
123 algorithm, it is not possible to utilize this strategy, thus we implemented an internal C function to walk
124 the space of minimum spanning trees based on genetic distance to connect groups of nodes with edges of
125 equal weight.

126 To demonstrate the utility of minimum spanning networks with reticulation, we used two clonal data
127 sets: the H3N2 flu virus data from the *adegenet* package using years of each epidemic as the population
128 factor, and *Phytophthora ramorum* data from Nurseries and Oregon forests (Jombart et al., 2010; Kamvar
129 et al., 2014a). Minimum spanning networks were created with and without reticulation using the *poppr*
130 functions `diss.dist` and `bruvo.msn` for the H3N2 and *P. ramorum* data, respectively (Kamvar et
131 al., 2014b; Bruvo et al., 2004). To detect mlg clusters, the infoMAP community detection algorithm was
132 applied with 10,000 trials as implemented in the R package *igraph* version 0.7.1 utilizing genetic distance
133 as edge weights and number of samples in each MLG as vertex weights (Csardi and Nepusz, 2006; Rosvall
134 and Bergstrom, 2008).

135 To evaluate the results, we compared the number, size, and entropy ($H$) of the resulting communities
136 as we expect a highly clonal organism with low genetic diversity to result in a few, large communities.
137 We also created contingency tables of the community assignments with the defined populations and used
138 those to calculate entropy using Shannon's index with the function `diversity` from the R package
139 *vegan* version 2.2-1 (Oksanen et al., 2015; Shannon, 2001). A low entropy indicates presence of a few
140 large communities whereas high entropy indicates presence of many small communities.

141 The infoMAP algorithm revealed 63 communities with a maximum community size of 77 and $H = 3.56$
142 for the reticulate network of the H3N2 data and 117 communities with a maximum community size of
143 26 and $H = 4.65$ for the minimum spanning tree. The entropy across years was greatly decreased for
144 all populations with the reticulate network compared to the minimum spanning tree (Fig. 3). Note that
145 the reticulated network (Fig. 3B) showed patterns corresponding with those resulting from a discriminant
146 analysis of principal components (Fig. 3D) (Jombart et al., 2010).

147 Graph walking of the reticulated minimum spanning network of *P. ramorum* by the infoMAP algorithm
148 revealed 16 communities with a maximum community size of 13 and $H = 2.60$. The un-reticulated
149 minimum spanning tree revealed 20 communities with a maximum community size of 7 and $H = 2.96$.
150 In the ability to predict Hunter Creek as belonging to a single community, the reticulated network was
151 successful whereas the minimum spanning tree separated one genotype from that community. The entropy
152 for the reticulated network was lower for all populations except for the coast population (supplementary
153 information).

## BOOTSTRAPPING

154 Assessing population differentiation through methods such as $G_{st}$, AMOVA, and Mantel tests relies on
155 comparing samples within and across populations (Nei, 1973; Excoffier et al., 1992; Mantel, 1967).
156 Confidence in distance metrics is related to the confidence in the markers to accurately represent the
157 diversity of the data. Especially true with microsatellite markers, a single hyper-diverse locus can make a
158 population appear to have more diversity based on genetic distance. Using a bootstrapping procedure of
159 randomly sampling loci with replacement when calculating a distance matrix provides support for clades
160 in hierarchical clustering.

161 Data in genind and genpop objects are represented as matrices with individuals in rows and alleles in
162 columns (Jombart, 2008). This gives the advantage of being able to use R's matrix algebra capabilities

163 to efficiently calculate genetic distance. Unfortunately, this also means that bootstrapping is a non- trivial
164 task as all alleles at a single locus need to be sampled together. To remedy this, we have created an internal
165 S4 class called "bootgen", which extends the internal "gen" class from *adegenet*. This class can be created
166 from any genind, genclone, or genpop object, and allows loci to be sampled with replacement. To further
167 facilitate bootstrapping, a function called `aboot`, which stands for "any boot", is introduced that will
168 bootstrap any genclone, genind, or genpop object with any genetic distance that can be calculated from it.

169     To demonstrate calculating a dendrogram with bootstrap support, we used the *poppr* function `aboot`
170 on population allelic frequencies derived from the data set `microbov` in the *adegenet* package with
171 1000 bootstrap replicates (Jombart, 2008; Laloë et al., 2007). The resulting dendrogram shows bootstrap
172 support values $> 50\%$ (Fig. 4) and used the following code:

```
library("poppr")
data("microbov", package = "adegenet")
strata(microbov) <- data.frame(other(microbov))
setPop(microbov) <- ~coun/spe/breed
bov_pop <- genind2genpop(microbov)

set.seed(20150428)
pop_tree <- aboot(bov_pop, sample = 1000, cutoff = 50)
```

## GENOTYPE ACCUMULATION CURVE

173 Analysis of population genetics of clonal organisms often borrows from ecological methods such as
174 analysis of diversity within populations (Milgroom, 1996; Arnaud-Hanod et al., 2007; Grünwald et al.,
175 2003). When choosing markers for analysis, it is important to make sure that the observed diversity in your
176 sample will not appreciably increase if an additional marker is added (Arnaud-Hanod et al., 2007). This
177 concept is analogous to a species accumulation curve, obtained by rarefaction. The genotype accumulation
178 curve in *poppr* is implemented in the function `genotype_curve`. The curve is constructed by randomly
179 sampling $x$ loci and counting the number of observed MLGs. This repeated $r$ times for 1 locus up to $n-1$
180 loci, creating $n-1$ distributions of observed MLGs.

181     The following code example demonstrates the genotype accumulation curve for data from Everhart and
182 Scherm (2015) showing that these data reach a small plateau and have a greatly decreased variance with
183 12 markers, indicating that there are enough markers such that adding more markers to the analysis will
184 not create very many new genotypes (Fig. 5).

```
library("poppr")
library("ggplot2")
data("monpop", package = "poppr")

set.seed(20150428)
genotype_curve(monpop, sample = 1000)
p <- last_plot() + theme_bw()   # get the last plot
p + geom_smooth(aes(group = 1)) # plot with a trendline
```

## INDEX OF ASSOCIATION

185 The index of association ($I_A$) is a measure of multilocus linkage disequilibrium that is most often used
186 to detect clonal reproduction within organisms that have the ability to reproduce via sexual or asexual
187 processes (Brown et al., 1980; Smith et al., 1993; Milgroom, 1996). It was standardized in 2001 as
188 $\bar{r}_d$ by Agapow and Burt (2001) to address the issue of scaling with increasing number of loci. This
189 metric is typically applied to traditional dominant and co-dominant markers such as AFLPs, SNPs, or

190  microsatellite markers. With the advent of high throughput sequencing, SNP data is now available in a
191  genome-wide context and in very large matrices including thousands of SNPs. For this reason, we devised
192  two approaches using the index of association for large numbers of markers typical for population genomic
193  studies. Both functions utilize *adegenet*'s "genlight" object class, which efficiently stores 8 binary alleles
194  in a single byte (Jombart and Ahmed, 2011). As calculation of the $\bar{r}_d$ requires distance matrices of absolute
195  number of differences, we utilize a function that calculates these distances directly from the compressed
196  data called `bitwise.dist`.

197      The first approach is a sliding window analysis implemented in the function `win.ia`. It utilizes the
198  position of markers in the genome to calculate $\bar{r}_d$ among any number of SNPs found within a user-
199  specified windowed region. It is important that this calculation utilize $\bar{r}_d$ as the number of loci will be
200  different within each window (Agapow and Burt, 2001). This approach would be suited for a quick
201  calculation of linkage disequilibrium across the genome that can detect potential hotspots of LD that
202  could be investigated further with more computationally intensive methods assuming that the number of
203  samples $<<$ the number of loci.

204      As it would necessarily focus on loci within a short section of the genome that may or may not
205  be recombining, a sliding window approach would not be good for utilizing $\bar{r}_d$ as a test for clonal
206  reproduction. A remedy for this is implemented in the function `samp.ia`, which will randomly sample
207  $m$ loci, calculate $\bar{r}_d$, and repeat $r$ times, thus creating a distribution of expected values of $\bar{r}_d$.

208      To demonstrate the sliding window and random sampling of $\bar{r}_d$ with respect to clonal populations, we
209  simulated two populations containing 1,100 neutral SNPs for 100 diploid individuals under the same
210  initial seed. One population had individuals randomly sampled with replacement, representing the clonal
211  population. After sampling, both populations had 5% random error and 1% missing data independently
212  propagated across all samples. On average, we obtained a higher value of $\bar{r}_d$ for the clonal population
213  compared to the sexual population for both methods (Fig. 6).


## DATA FORMAT UPDATES: POPULATION STRATA AND HIERARCHIES

214  Assessments of population structure through methods such as hierarchical $F_{st}$ (Goudet, 2005) and
215  AMOVA (Michalakis and Excoffier, 1996) require hierarchical sampling of populations across space
216  or time (Linde et al., 2002; Everhart and Scherm, 2015; Grünwald and Hoheisel, 2006). With clonal
217  organisms, basic practice has been to clone-censor data to avoid downward bias in diversity due to
218  duplicated genotypes that may or may not represent different samples (Milgroom, 1996). This correction
219  should be performed with respect to a population hierarchy to accurately reflect the biology of the
220  organism. Traditional data structures for population genetic data in most analysis tools allow for only
221  one level of hierarchical definition. The investigator thus had to provide the data set for analysis at each
222  hierarchical level.

223      To facilitate handling hierarchical and mutlilocus genotypic metadata, *poppr* version 1.1 introduced
224  a new S4 data object called "genclone", extending *adegenet*'s "genind" object (Kamvar and Grünwald,
225  unpublished). The genclone object formalized the definitions of multilocus genotypes and population
226  hierarchies by adding two slots called "mlg" and "hierarchy" that carried a numeric vector and a data
227  frame, respectively. These new slots allow for increased efficiency and ease of use by allowing these
228  metadata to travel with the genetic data. The hierarchy slot in particular contains a data frame where each
229  column represents a separate hierarchical level. This is then used to set the population factor of the data by
230  supplying a hierarchical formula containing one or more column names of the data frame in the hierarchy
231  slot.

232      The functionality represented by the hierarchy slot has now been migrated from the *poppr* to the
233  *adegenet* package version 2.0 to allow hierarchical analysis in *adegenet*, *poppr*, and other dependent
234  packages. The prior *poppr* `hierarchy` slot and methods have now been renamed `strata` in *adegenet*.
235  A short example of the utility of these methods can be seen in the code segment under **Bootstrapping**,

236  above. This migration provides end users with a broader ability to analyze data hierarchically in R across
237  packages.

## AVAILABILITY

238  As of this writing, the *poppr* R package version 2.0 containing all of the features described here
239  is located at `https://github.com/grunwaldlab/poppr/tree/2.0-rc`. It is necessary
240  to install *adegenet* 2.0 before installing *poppr*. It can be found at `https://github.com/`
241  `thibautjombart/adegenet`. Both of these can be installed via the R package *devtools* (Wickham
242  and Chang, 2015).

## REQUIREMENTS

243  • R version 3.0 or better
244  • A C compiler. For windows, this can be obtained via Rtools (`http://cran.r-project.org/`
245  `bin/windows/Rtools/`). On OSX, this can be obtained via Xcode. For parallel support, gcc
246  version 4.6 or better is needed.

## INSTALLATION

247  From within R, *poppr* can be installed via:

```
install.packages("devtools")
library("devtools")
install_github("thibautjombart/adegenet")
install_github("grunwaldlab/poppr@2.0-rc")
```

248  Several population genetics packages in R are currently going through a major upgrade
249  following the 2015 R hackathon on population genetics (`https://github.com/NESCent/`
250  `r-popgen-hackathon`) and have not yet been updated in CRAN. We will upload *poppr* 2.0 to CRAN
251  once all other reverse dependent packages have been updated.

## DISCUSSION

252  Given low cost and high throughput of current sequencing technologies we are entering a new era of
253  population genetics where large SNP data sets with thousands of markers are becoming available for large
254  populations in a genome- wide context. This data provides new possibilities and challenges for population
255  genetic analyses. We provide novel tools that enable analysis of this data in R with a particular emphasis
256  on clonal organisms.

257  Particularly useful is the implementation of $\bar{r}_d$ in a genomic context (Agapow and Burt, 2001). Random
258  sampling of loci across the genome can give an expected distribution of $\bar{r}_d$, which is expected to have
259  a mean of zero for panmictic populations. This metric is not affected by the number of loci sampled, is
260  model free, and has the ability to detect population structure. $\bar{r}_d$ is also implemented for sliding window
261  analyses that are useful to detect candidate regions of linkage disequilibrium for further analysis.

262  Clustering multilocus genotypes into multilocus lineages based on genetic distances is a non-trivial task
263  given large SNP data sets. Moreover, this has not previously been implemented for genomic data for
264  clonal populations. Clonal assignment has previously been available in the programs GENCLONE and
265  GENODIVE for classical markers (Arnaud-Hanod et al., 2007; Meirmans and Van Tienderen, 2004). Our
266  method with `mlg.filter` builds upon this idea and allows the user to choose between three different

267 approaches for clustering MLGs. The choice of clustering algorithm has an impact on the data (Fig 1, 2),
268 where for example a genetic distance cutoff of 0.1 would be the difference between 14 multilocus lineages
269 (MLLs) and 17 MLLs for nearest neighbor and UPGMA clustering, respectively (Fig. 2). The option to
270 choose the clustering algorithm gives the user the ability to choose what is biologically relevant to their
271 populations. While there is not one optimal procedure for defining boundaries in clonal lineages, our tool
272 provides a means of exploring the potential MLG or MLL boundary space.

273      Minimum spanning networks are a useful tool to analyze the relationships between individuals in a
274 population, because it reduces the complexity of a distance matrix to the connections that are strongest.
275 By default, these networks are drawn without reticulations, but for clonal organisms where many of the
276 connections between samples are equivalent, the minimum spanning network appears as a chain and
277 reduces the information that can be communicated. This is problematic because the ability to detect
278 population structure with a simplistic minimum spanning network is limited. Adding reticulation into
279 the minimum spanning network thus presents all equivalent connections and allows population structure
280 to be more readily detectable. As shown in Fig. 3 population structure is apparent both visually and
281 by graph community detection algorithms such as the infoMAP algorithm (Rosvall and Bergstrom,
282 2008). Additionally, the current implementation in *poppr* has been successfully used in analyses such
283 as reconstruction of the *P. ramorum* epidemic in Curry County, OR (Kamvar et al., 2014a, 2015).

284      *Poppr* 2.0 is open source and available on GitHub. Members of the community are invited to contribute
285 by raising issues or pull requests on our repository at `https://github.com/grunwaldlab/`
286 `poppr/issues`.

## ACKNOWLEDGEMENTS

## CONFLICT OF INTEREST STATEMENT

294 The authors declare no observable conflict of interest.

## AUTHOR CONTRIBUTIONS

295 ZNK and JCB wrote and tested the code. ZNK maintains the code. ZNK and NJG conceived, discussed
296 implications, and wrote the manuscript. NJG coordinated the collaborative effort.

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## FIGURES AND TABLES

**FIGURE 1**



427

**FIGURE 2**



*P. infestans* reference isolates (12 SSR loci)

**FIGURE 3**



A

B

● 2001  ● 2004

● 2002  ● 2005

● 2003  ● 2006

C

No Reticulation   ◇ With Reticulation

entropy (H)

3

2

1

0

2001  2002  2003  2004  2005  2006

year

D

429

**FIGURE 4**



430

**FIGURE 5**



431
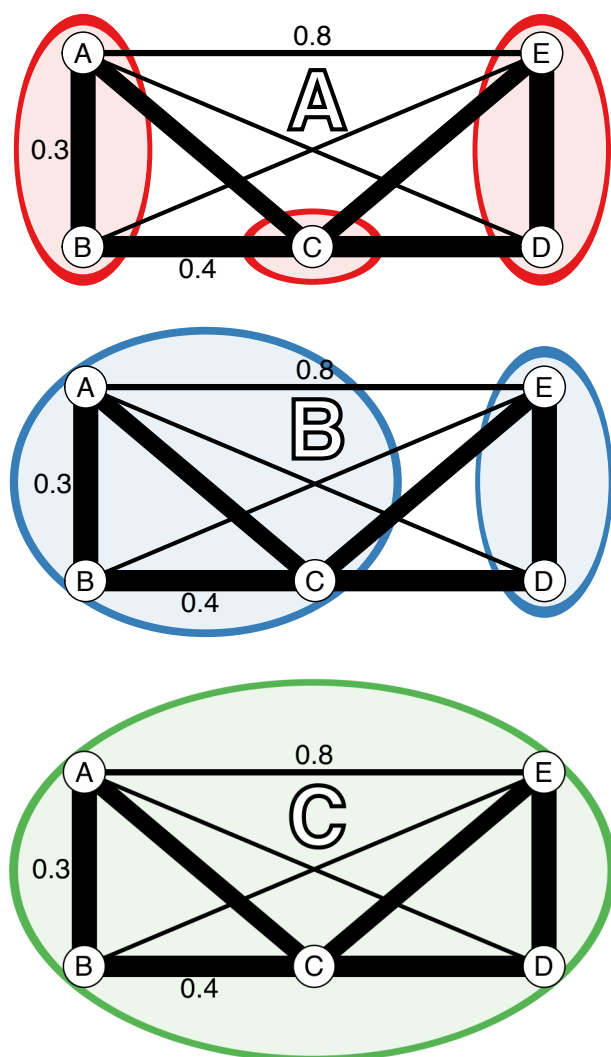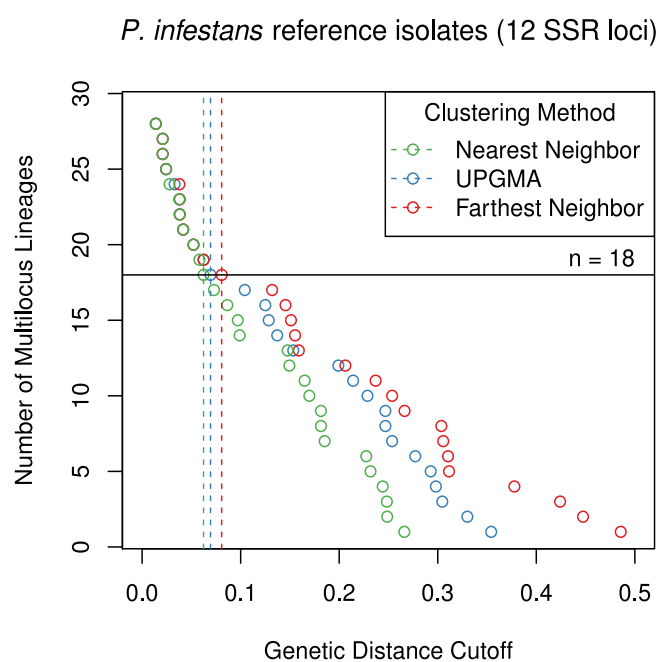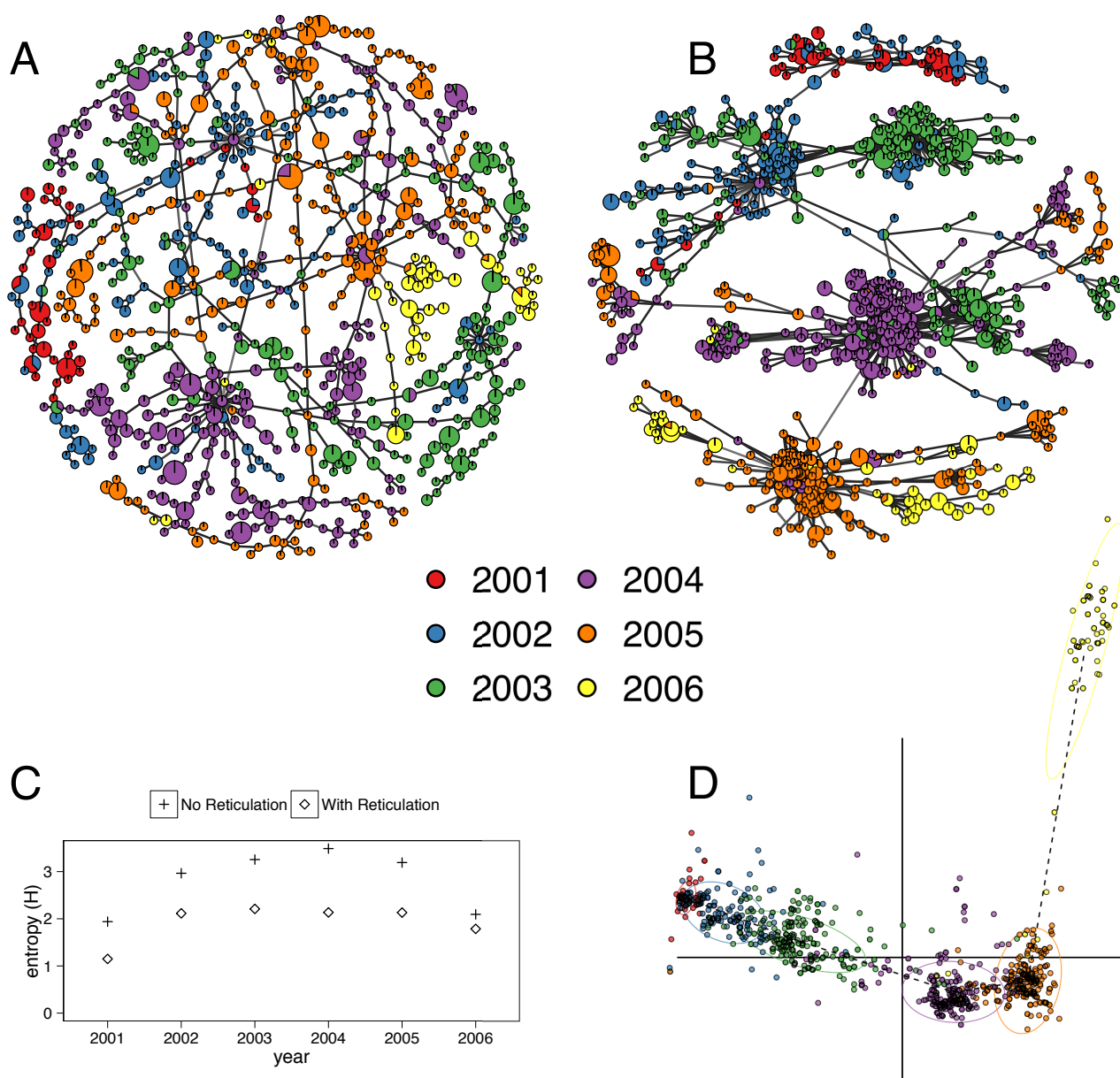
**FIGURE 6**

433     **Table 1** Contingency table comparing multilocus lineages assigned based on average neighbor clustering
434     (columns) vs. multilocus lineages defined in Li et al. (2013) and Lees et al. (2006).

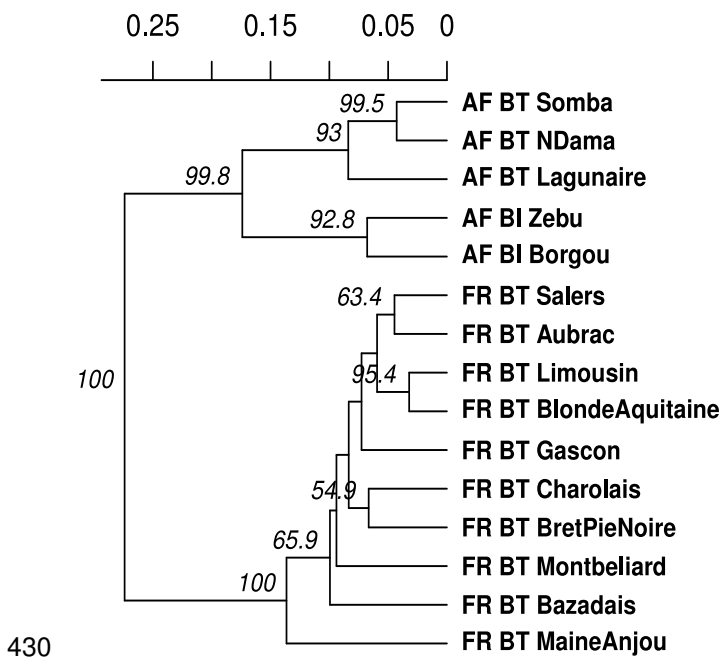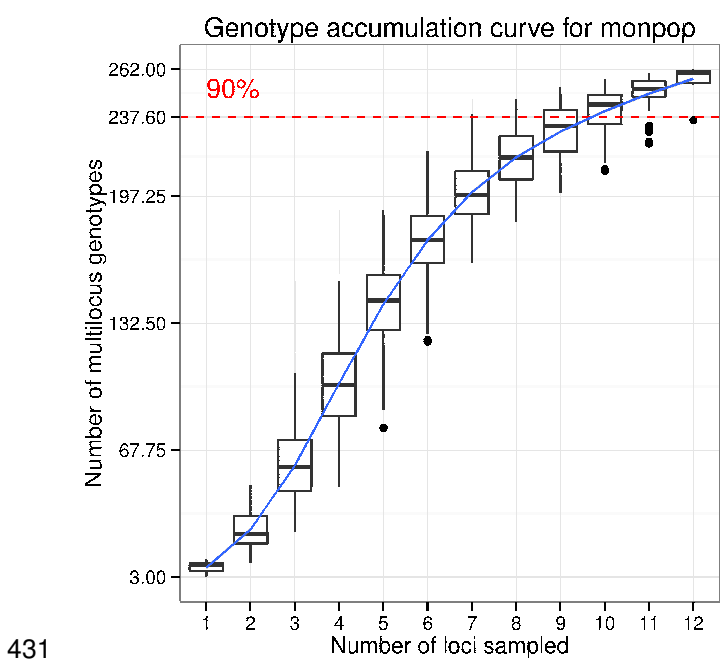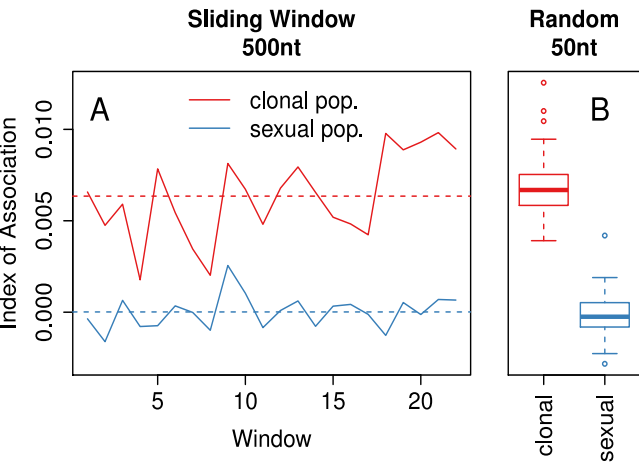|        | 3 | 4 | 5 | 6 | 8 | 10 | 12 | 15 | 16 | 17 | 18 | 20 | 21 | 22 | 24 | 25 | 27 | 28 |
|--------|---|---|---|---|---|----|----|----|----|----|----|----|----|----|----|----|----|----|
| B      | . | . | . | . | . | .  | .  | .  | .  | .  | .  | .  | .  | .  | .  | 1  | .  | .  |
| C      | . | . | . | . | . | .  | .  | .  | .  | .  | .  | .  | .  | .  | 1  | .  | .  | .  |
| D.1    | . | . | . | . | . | .  | .  | .  | .  | .  | .  | .  | .  | 1  | .  | .  | .  | .  |
| D.2    | . | . | . | . | . | .  | .  | .  | .  | .  | .  | .  | .  | 1  | .  | .  | .  | .  |
| EU-13  | . | . | . | . | . | .  | .  | 1  | .  | .  | .  | .  | .  | .  | .  | .  | .  | .  |
| EU-4   | . | . | . | . | . | .  | .  | .  | 1  | .  | .  | .  | .  | .  | .  | .  | .  | .  |
| EU-5   | . | . | . | . | . | .  | .  | .  | .  | .  | 2  | .  | .  | .  | .  | .  | .  | .  |
| EU-8   | . | . | . | . | . | .  | 1  | .  | .  | .  | .  | .  | .  | .  | .  | .  | .  | .  |
| US-11  | . | . | . | . | . | .  | .  | .  | .  | .  | .  | .  | .  | .  | .  | .  | .  | 2  |
| US-12  | . | 1 | . | . | . | .  | .  | .  | .  | .  | .  | .  | .  | .  | .  | .  | .  | .  |
| US-14  | . | . | . | . | . | 1  | .  | .  | .  | .  | .  | .  | .  | .  | .  | .  | .  | .  |
| US-17  | . | . | . | . | . | .  | .  | .  | .  | .  | .  | 1  | .  | .  | .  | .  | .  | .  |
| US-20  | 2 | . | . | . | . | .  | .  | .  | .  | .  | .  | .  | .  | .  | .  | .  | .  | .  |
| US-21  | . | . | . | . | . | .  | .  | .  | .  | .  | .  | .  | .  | .  | .  | .  | 2  | .  |
| US-22  | . | . | . | . | . | .  | .  | .  | .  | .  | .  | .  | 2  | .  | .  | .  | .  | .  |
| US-23  | . | . | . | . | . | .  | .  | 3  | .  | .  | .  | .  | .  | .  | .  | .  | .  | .  |
| US-24  | . | . | . | . | 3 | .  | .  | .  | .  | .  | .  | .  | .  | .  | .  | .  | .  | .  |
| US-8   | . | . | 1 | 1 | . | 2  | .  | .  | .  | .  | .  | .  | .  | .  | .  | .  | .  | .  |

## FIGURE LEGENDS

435     **Figure 1** Diagrammatic representation of the three clustering algorithms implemented in `mlg.filter`.
436     (**A-C**) Represent different clustering algorithms on the same imaginary network with a threshold of 0.451.
437     Edge weights are represented in arbitrary units noted by the line thickness and numerical values next to
438     the lines. All outer angles are 90 degrees, so the un-labeled edge weights can be obtained with simply
439     geometry. Colored circles represent clusters of genotypes. (**A**) Farthest neighbor clustering does not cluster
440     nodes B and C because nodes A and C are more than a distance of 0.451 apart. (**B**) UPGMA (average
441     neighbor) clustering clusters nodes A, B, and C together because the average distance between them and

442  C is < 0.451. (**C**) Nearest neighbor clustering clusters all nodes together because the minimum distance
443  between them is always < 0.451.

444    **Figure 2** Graphical representation of three different clustering algorithms collapsing multilocus
445  genotypes for 12 SSR loci from *Phytophthora infestans* representing 18 clonal lineages. The horizontal
446  axis is Bruvo's genetic distance assuming the genome addition model. The vertical axis represents the
447  number of multilocus lineages observed. Each point shows the threshold at which one would observe a
448  given number of multilocus genotypes. The horizontal black line represents 18 multilocus genotypes and
449  vertical dashed lines mark the thresholds used to collapse the multilocus genotypes into 18 multilocus
450  lineages.

451    **Figure 3** (**A-B**) Minimum spanning networks of the hemagglutinin (HA) segment of H3N2 viral DNA
452  from the *adegenet* package representing flu epidemics from 2001 to 2006 without reticulation (**A**) and with
453  reticulation (**B**) (Jombart, 2008; Jombart et al., 2010). Each node represents a unique multilocus genotype,
454  colors represent epidemic year, and edge color represents absolute genetic distance. (**C**) Shannon entropy
455  values for population assignments compared with communities determined by the infoMAP algorithm on
456  (**A**) and (**B**). (**D**) Graphic reproduced from Jombart et al. (2010) showing that the 2006 epidemic does not
457  cluster neatly with the other years.

458    **Figure 4** UPGMA dendrogram generated from Nei's genetic distance on 15 breeds of *Bos taurus* (BT)
459  or *Bos indicus* (BI) from Africa (AF) or France (FR). These data are from Laloë et al. (2007). Node labels
460  represent bootstrap support > 50% out of 1,000 bootstrap replicates.

461    **Figure 5** Genotype accumulation curve for 694 isolates of the peach brown rot pathogen, *Monilinia*
462  *fructicola* genotyped over 13 loci from Everhart and Scherm (2015). The horizontal axis represents the
463  number of loci randomly sampled without replacement up to $n-1$ loci, the vertical axis shows the number
464  of multilocus genotypes observed, up to 262, the number of unique multilocus genotypes in the data set.
465  The red dashed line represents 90% of the total observed multilocus genotypes. A trendline (blue) has
466  been added using the *ggplot2* function `stat_smooth`.

467    **Figure 6** (**A**) Sliding window analysis of the standardized index of association ($\bar{r}_d$) across a simulated
468  $1.1 \times 10^4$ nt chromosome containing 1,100 variants among 100 individuals. Each window analyzed
469  variants within 500nt chunks. The black line refers to the clonal and the blue line to the sexual populations.
470  (**B**) boxplots showing 100 random samples of 50 variants to calculate a distribution of $\bar{r}_d$ for the clonal
471  (red) and sexual (blue) populations. Each box is centered around the mean, with whiskers extending out
472  to 1.5 times the interquartile range. The median is indicated by the center line. (**A**) and (**B**) are plotted on
473  the same y-axis.