

3D Point Clouds Lecture 3 – Clustering



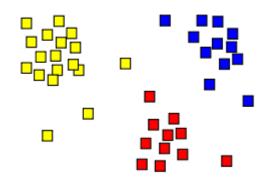
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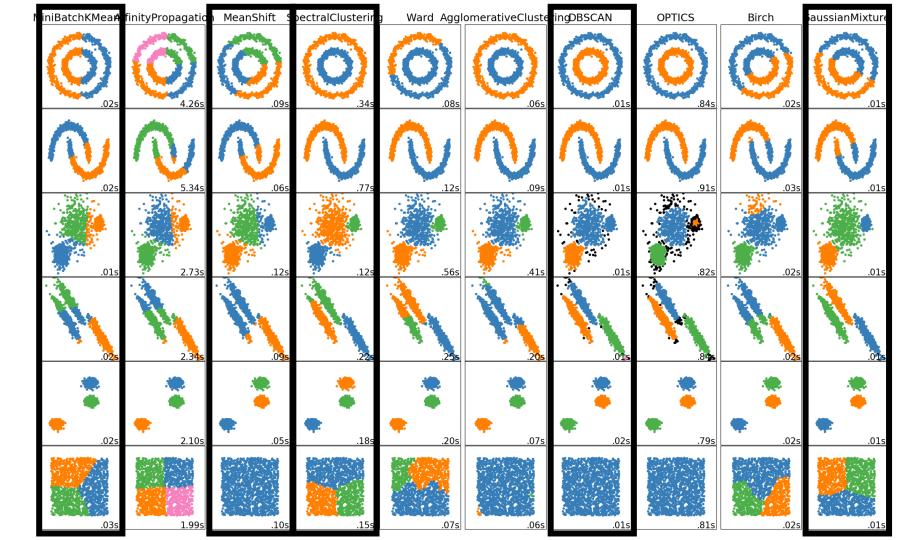


- 1. Math Prerequisite
- 2. K-Means
- 3. Gaussian Mixture Models (GMM)
- 4. Expectation Maximization (EM)
- 5. Spectral Clustering
- 6. Others Mean Shift, DBSCAN

Introduction Clustering

- Clustering is to group objects such that objects in the same group (called a cluster) are similar, while objects belong to different groups are dis-similar.
- There are many clustering algorithms
- We focus on two types:
 - Theoretically profound
 - K-Means,
 - GMM (Gaussian Mixture Model)
 - EM (Expectation-Maximation)
 - Spectral Clustering
 - Practically useful
 - Mean-Shift
 - DBSCAN (Density-Based Spatial Clustering of Applications with Noise)





S Math Prerequisite

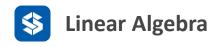
- Need some background knowledge to understand K-Means, GMM, EM, Spectral Clustering
 - Linear Algebra
 - SVD
 - Rayleigh Quotient
 - Probability
 - Bayes rule
 - Conditional Probability
 - Graphical Modeling
 - Directed graph
 - Undirected graph
 - Lagrange Multiplier



Spectral Theorem

Let $A \in \mathbb{R}^{n,n}$ be symmetric, and $\lambda_i \in \mathbb{R}, i = 1, 2, \dots, n$ be the eigenvalues of A. There exists a set of orthonormal vectors $u_i \in \mathbb{R}_n, i = 1, 2, \dots, n$, such that $Au_i = \lambda_i u_i$. Equivalently, there exists an orthogonal matrix $U = [u_1, \dots, u_n]$ (i.e., $UU^T = U^TU = I_n$), such that,

$$A = U\Lambda U^T = \sum_{i=1}^n \lambda_i u_i u_i^T, \Lambda = ext{diag}(\lambda_1, \cdots, \lambda_n)$$



Rayleigh Quotients

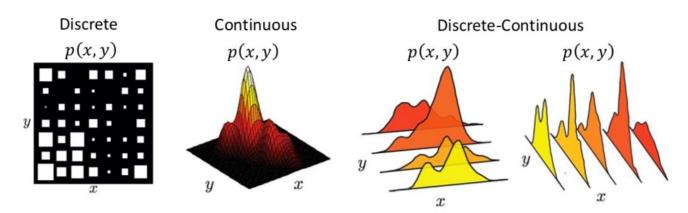
Given a symmetric matrix $A \in S^n$,

Physical meaning of SVD!

$$\lambda_{\min}(A) \leq rac{x^T A x}{x^T x} \leq \lambda_{\max}(A), orall x
eq 0$$
 $\lambda_{\max}(A) = \max_{x:\|x\|_2 = 1} x^T A x$
 $\lambda_{\min}(A) = \min_{x:\|x\|_2 = 1} x^T A x$

The maximum and minimum are attained for $x = u_1$ and for $x = u_n$, respectively, where u_1 and u_n are the largest and smallest eigenvector of A, respectively.

- Consider all combinations events of two random variables X and Y
- This is captured in the *joint probability* distribution p(x, y)
- Read as "probability of X and Y"
- Can be more than two random variables, i.e., p(a, b, c, ...)



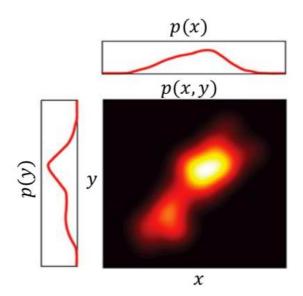


Probability - Marginalization

- Recover probability distribution of any variable in a joint distribution by integrating (or summing) over all other variables.
- sum rule" of probability

Continuous:

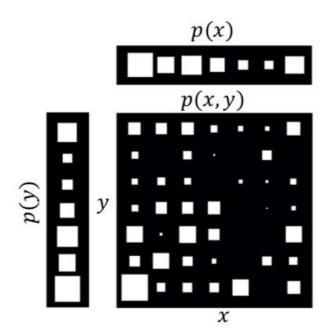
$$p(x) = \int p(x, y) dy$$
$$p(y) = \int p(x, y) dx$$



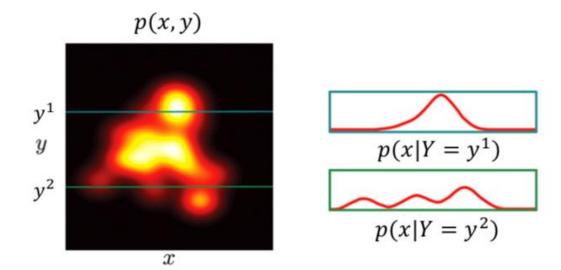
Discrete:

$$p(x) = \sum_{y} p(x, y)$$

$$p(y) = \sum_{x} p(x, y)$$



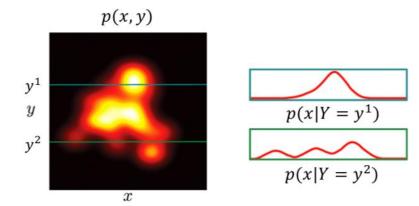
 $p(x|Y=y^*)$: probability of X given $Y=y^*$





- Conditional probability can be extracted from joint probability
- Extract appropriate slice and normalize (sum to 1)

$$P(x|Y = y^*) = \frac{p(x, Y = y^*)}{\int p(x, Y = y^*) dx} = \frac{p(x, Y = y^*)}{p(Y = y^*)}$$



Conditional probability:

$$P(x|Y = y^*) = \frac{p(x, Y = y^*)}{\int p(x, Y = y^*) dx} = \frac{p(x, Y = y^*)}{p(Y = y^*)}$$

Can be written in compact form:

$$p(x|y) = \frac{p(x,y)}{p(y)}$$

Which leads to:

$$p(x,y) = p(x|y)p(y)$$

$$p(x,y) = p(y|x)p(x)$$

Hence, the name "product rule"!

The product rule works for higher dimensions as well

$$p(w,x,y,z) = p(w,x,y|z)p(z)$$

$$= p(w,x|y,z)p(y|z)p(z)$$

$$= p(w|x,y,z)p(x|y,z)p(y|z)p(z)$$



Graphical Modeling - Directed Graphical Model (DGM)

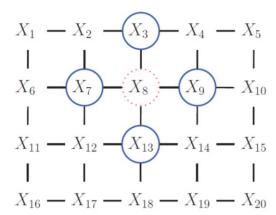


- DGM is used to represent conditional independence
- DGM is represented by G(V, E) V is a set of nodes. One node is one variable. E is a set of oriented edges. One edge is one conditional relationship.
- Markov assumption a random variable X is independent of its non-descendants given its parents
 - E.g., the joint distribution is p(x, z) = p(z)p(x|z)
 - E.g., $z \rightarrow$ "age of a person", $x \rightarrow$ "hair white or black" p(x|z) means the probability of white/black hair given a person's age It is different to p(x), which means the probability of white/black hair without other info



Graphical Modeling - Undirected Graphical Model (UGM)

- It is called Markov Random Field (MRF) or Markov Network as well
- UGM is represented by G(V, E), where V is a set of nodes, E is a set of undirected edges
 - E.g., below can be a map, X_i are cities, the edges are roads.
 - E.g., with UGM, I know the shortest path from X_1 to X_8





Optimization – Lagrange multiplier

Consider the optimization problem

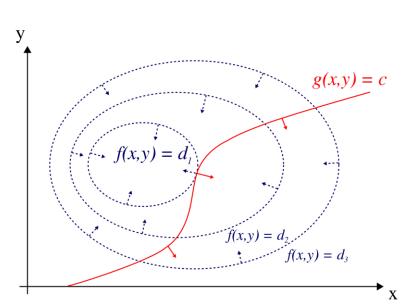
$$\max f(x, y), \ s.t.: \ g(x, y) = 0$$

lacktriangle Assume that both f and g have continuous first partial derivatives

Intuition

- The circles are level sets of f(x, y)
- g(x,y) = 0 means x, y have to be on the red line
- At a (local) maximum, f(x, y) can NOT be increasing along the direction (tangent) of g(x, y) = 0

$$\nabla_{x,y} f = \lambda \, \nabla_{x,y} g$$



Optimization – Lagrange multiplier

- Introduce a new variable λ called Lagrange multiplier
- Let's study the Lagrange function:

$$\mathcal{L}(x, y, \lambda) = f(x, y) - \lambda g(x, y)$$

And solve:

$$\nabla_{x,y,\lambda}\mathcal{L}(x,y,\lambda)=0.$$

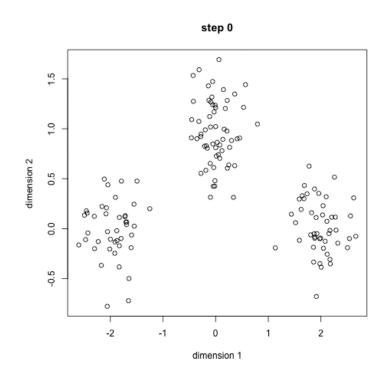
- The above equation means
 - (1) g(x,y) = 0
 - (2) gradient of f and g are parallel

$$abla_{x,y,\lambda}\mathcal{L}(x,y,\lambda) = 0 \iff egin{cases}
abla_{x,y}f(x,y) = \lambda \,
abla_{x,y}g(x,y) \ g(x,y) = 0 \end{cases}$$



N input data points, find *K* clusters.

- 1. Randomly select *K* center points
- 2. Each data point is assigned to one of the *K* centers.
- 3. Re-compute the *K* centers by the mean of each group
- 4. Iterate step 2 & 3.



Data set
$$\{x_1, \cdots, x_N\}, x_n \in \mathbb{R}^D$$

- Cluster center μ_k , $k = 1, \dots K$
 - μ_k is the center of k^{th} cluster
- Binary variable $r_{nk} \in \{0, 1\}$
 - x_n belong to which cluster
- Objective of K-Means is to minimize the distortion measure

$$J = \sum_{n=1}^{N} \sum_{k=1}^{K} r_{nk} \|\mathbf{x}_n - \boldsymbol{\mu}_k\|^2$$

\$ K-Means

- Iteratively optimize r_{nk} and μ_k
- $loodsymbol{\circ}$ For some fixed values for μ_k , minimize J with respect to r_{nk}
 - This is the E (expectation) step of the EM algorithm
- $loodsymbol{\circ}$ For some fixed values of r_{nk} , minimize J with respect to μ_k
 - This is the M (maximization) step of the EM algorithm

$$J = \sum_{n=1}^{N} \sum_{k=1}^{K} r_{nk} \|\mathbf{x}_n - \boldsymbol{\mu}_k\|^2$$



$$J = \sum_{n=1}^{N} \sum_{k=1}^{K} r_{nk} ||\mathbf{x}_n - \boldsymbol{\mu}_k||^2$$

- $lue{lue}$ N data points are independent, so we can optimize for each n separately.
- Simply assign the n^{th} data point to the closest cluster center, which will minimize $||x_n \mu_k||^2$
- Formally

$$r_{nk} = \begin{cases} 1 & \text{if } k = \arg\min_{j} \|\mathbf{x}_n - \boldsymbol{\mu}_j\|^2 \\ 0 & \text{otherwise.} \end{cases}$$



$$J = \sum_{n=1}^{N} \sum_{k=1}^{K} r_{nk} \|\mathbf{x}_n - \boldsymbol{\mu}_k\|^2$$

- $lue{f v}$ With r_{nk} fixed, the objective function J is a quadratic function of μ_k
- Compute its first order derivative and make it to 0
- Consider each center μ_k separately

$$2\sum_{n=1}^{N} r_{nk}(\mathbf{x}_n - \boldsymbol{\mu}_k) = 0$$

$$oldsymbol{\mu}_k = rac{\sum_n r_{nk} \mathbf{x}_n}{\sum_n r_{nk}}$$



- Illustration of the K=2
 - a) Initialization
 - b) Expectation find r_{nk}
 - c) Maximization find μ_k

Other figures are repeating (b)(c)

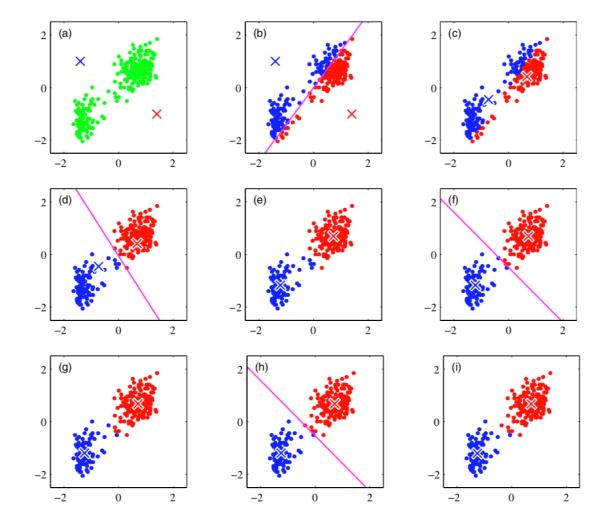


Figure source: Pattern Recognition and Machine Learning, Bishop Dataset: re-scaled Old Faithful dataset

- $lue{f O}$ Randomly select K points as the initial μ_k
- Run K-Means multiple times with random initialization, select the one with the lowest cost J
- In E-step, it is necessary to compute distance between x_n and μ_k
 - Nearest Neighbor algorithms in Lecture 2
- Mini-batch K-Means: In each E&M iteration, select a subset of the data points.
 - Reduce the compute time
 - Converges faster
 - The final result may be slightly worse.

K-Means: Sequential Update

- $lue{}$ The E / M step are performed for each data point x_n
 - 1. For each data point x_n , find the nearest center μ_k^{old}
 - 2. Update the center into

$$\boldsymbol{\mu}_k^{ ext{new}} = \boldsymbol{\mu}_k^{ ext{old}} + \eta_n (\mathbf{x}_n - \boldsymbol{\mu}_k^{ ext{old}})$$

where η_n is the learning rate

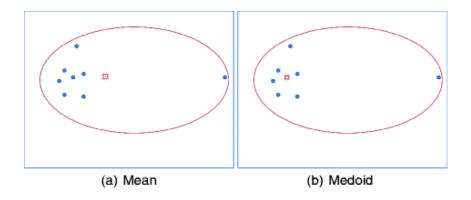
Mini-batch version of this is the mini-batch K-Means mentioned in previous slice

S K-Medoids

- Standard K-Means with Euclidean distance is problematic in some cases.
 - Outlier outliers will drag the centers away
 - Limits the type of data points E.g., Euclidean distance is not for categorical labels
- Apply general dissimilarity metric

K-Medoids
$$\widetilde{J} = \sum_{n=1}^N \sum_{k=1}^K r_{nk} \mathcal{V}(\mathbf{x}_n, \pmb{\mu}_k)$$
 K-Means
$$J = \sum_{n=1}^N \sum_{k=1}^K r_{nk} \|\mathbf{x}_n - \pmb{\mu}_k\|^2$$

 $n=1 \ k=1$

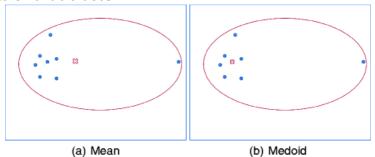




- **E** step
 - Same, just need to evaluate the dissimilarity measure $V(\cdot,\cdot)$ instead of Euclidean distance
 - O(KN) complexity
- M step

$$\widetilde{J} = \sum_{n=1}^{N} \sum_{k=1}^{K} r_{nk} \mathcal{V}(\mathbf{x}_n, \boldsymbol{\mu}_k)$$

- First order derivative over $V(\cdot,\cdot)$ can be intractable
- Therefore, assign μ_k to be one of the data point that has smallest \tilde{I}
- Can be implemented as discrete search over the N_k points of that cluster
- $O(N_k^2)$ complexity



K-Means Application – Compression

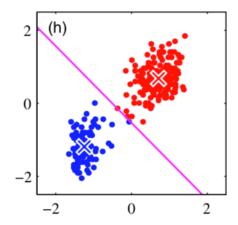
- Illustrated as image compression, can be easily applied to point cloud (XYZRGB) compression
 - 1. A $H \times W \times 3$ image is represented as $N \times 3$ points
 - 2. Cluster them into *K* clusters
 - 3. Now the $N \times 3$ points can be represented by N labels.
- Similarly, a [XYZRGB] point cloud can be compressed into [XYZL], where L is the cluster label



Figure source: Pattern Recognition and Machine Learning, Bishop

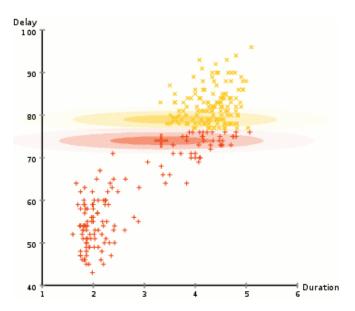
- Original Image have N pixels, 3 channels
 - Each channel is 8bits
 - In total 24N bits
- Compressed Image have N labels + cluster centers
 - $N \text{ labels } \rightarrow \lceil N \log_2 K \rceil$ bits (Note that $M \text{ bits can represent } 2^M \text{ integers})$
 - Cluster centers -> 24K bits
 - *K* clusters, each is [R,G,B]
 - In total $[N \log_2 K] + 24K$ bits
- \bigcirc $K \ll N$ therefore compression is effective

- - Sometimes a data point lies around the border line between centers
 - Lack of uncertainty measure for the assignments
- K is unknown



Gaussian Mixture Model (GMM)

- Represent a cluster by a Gaussian Distribution $\mathcal{N}(\mu, \sigma)$
- GMM tells the probability of a point belonging to each cluster



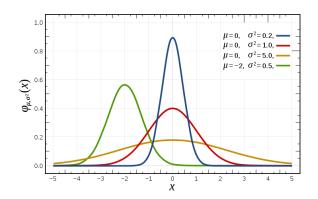
Gaussian Distribution

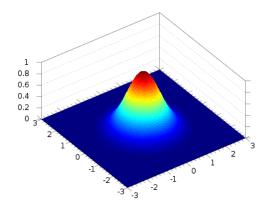
Single variable Gaussian

$$\mathcal{N}(x|\mu,\sigma^2) = \frac{1}{(2\pi\sigma^2)^{1/2}} \exp\left\{-\frac{1}{2\sigma^2}(x-\mu)^2\right\}$$

Multivariate Gaussian with D-dimensional vector

$$\mathcal{N}(\mathbf{x}|\boldsymbol{\mu}, \boldsymbol{\Sigma}) = \frac{1}{(2\pi)^{D/2}} \frac{1}{|\boldsymbol{\Sigma}|^{1/2}} \exp\left\{-\frac{1}{2} (\mathbf{x} - \boldsymbol{\mu})^{\mathrm{T}} \boldsymbol{\Sigma}^{-1} (\mathbf{x} - \boldsymbol{\mu})\right\}$$





Formally, a Gaussian mixture distribution can be written as a linear combination of single Gaussians $\mathcal{N}(\mu_k, \Sigma_k)$ using weights π_k

$$p(x) = \sum_{k=1}^{K} \pi_k \mathcal{N}(x|\mu_k, \Sigma_k)$$

- The $\{\mu_k, \Sigma_k, \pi_k\}$, $k = 1, \dots, K$ describes the GMM built from the data points. In another word, we can "generate" points using this p(x)
- But we want clustering! That is, the probability of each data point belongs to each single Gaussian $\mathcal{N}(\mu_k, \Sigma_k)$

\$ GMM

Let's introduce a K-dimensional binary variable z having a 1-of-K representation.

$$z_k \in \{0, 1\}, \qquad \Sigma_k z_k = 1$$

- This $p(z_k=1)$ is the prior probability of Gaussian distribution $\mathcal{N}(\mu_k,\Sigma_k)$ $p(z_k=1)=\pi_k$
- The conditional probability p(z|x) is the clustering "label probability" we want, given a data point x

$$p(z|x) = \frac{p(x|z)p(z)}{p(x)}, \qquad p(x) = \sum_{z} p(z)p(x|z)$$

- Gaussian Mixture Model is represented by a graphical model, where the joint distribution p(x,z) = p(z)p(x|z)
- $p(z_k=1)=\pi_k \text{ , } z=[z_1,\cdots,z_k,\cdots,z_K] \text{, 1-of-K representation Constraints: } 0\leq \pi_k \leq 1 \text{, } \sum_{k=1}^K \pi_k=1$

Alternatively,

$$p(z) = \prod_{k=1}^{K} \pi_k^{z_k}$$





 $p(x|z_k = 1)$ is a single Gaussian distribution

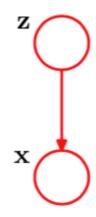
$$p(x|z_k = 1) = \mathcal{N}(x|\mu_k, \Sigma_k)$$

Alternatively,

$$p(x|z) = \prod_{k=1}^{K} \mathcal{N}(x|\mu_k, \Sigma_k)^{z_k}$$

The marginal distribution of x is given by $p(x) = \sum_{z} p(x, z)$

$$p(x) = \sum_{z} p(z)p(x|z) = \sum_{k=1}^{K} \pi_k \mathcal{N}(x|\mu_k, \Sigma_k)$$

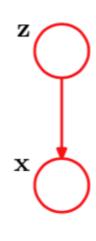


\$ GMM

- Now we can get the posterior p(z|x)
 - Given a data point x, which cluster it belongs to?
- Denote $p(z_k = 1|x)$ as $\gamma(z_k)$
- Use Bayes' rule and Total Probability Rule

$$p(\mathbf{z}|\mathbf{x}) = \frac{p(\mathbf{z}|\mathbf{x})}{p(\mathbf{x})} = \frac{p(\mathbf{z})p(\mathbf{x}|\mathbf{z})}{\sum_{j=1}^{K} p(\mathbf{x}|z_j)p(z_j)}$$

$$\gamma(z_k) \equiv p(z_k = 1 | \mathbf{x}) = \frac{p(z_k = 1)p(\mathbf{x}|z_k = 1)}{\sum_{j=1}^{K} p(z_j = 1)p(\mathbf{x}|z_j = 1)}$$
$$= \frac{\pi_k \mathcal{N}(\mathbf{x}_n | \mu_k, \Sigma_k)}{\sum_{j=1}^{K} \pi_j \mathcal{N}(\mathbf{x}_n | \mu_j, \Sigma_j)}$$



GMM – Maximum Likelihood (MLE)

- Now we know how to get $\gamma(z_k)$, given GMM parameters $\{\pi_k, \mu_k, \Sigma_k\}$
- But how do we estimate $\{\pi_k, \mu_k, \Sigma_k\}, \gamma(z_k)$ given data points $\{x_1, \dots, x_N\}$
- Maximum likelihood!

$$p(\mathbf{x}) = \sum_{k=1}^{K} \pi_k \mathcal{N}(\mathbf{x}|\mu_k, \Sigma_k)$$

$$\ln p(\mathbf{X}|\pi, \mu, \Sigma) = \sum_{n=1}^{N} \ln \left\{ \pi_k \mathcal{N}(\mathbf{x_n}|\mu_k, \Sigma_k) \right\}$$

GMM – Maximum Likelihood (MLE)

- Maximum likelihood formulation suffers from singularity problem
- Assume that the GMM has the covariance matrix represented as $\Sigma_k = \sigma_k^2 I$, where I is identity matrix
- Assume that there is a data point $x_n = \mu_j$, the likelihood contributed by x_n is:

$$\mathcal{N}(\mathbf{x}_n | \mathbf{x}_n, \sigma_j^2 \mathbf{I}) = \frac{1}{(2\pi)^{1/2}} \frac{1}{\sigma_j}$$

- What if σ_j is very small?
- So in GMM maximum likelihood formulation, we can put a Gaussian over a single point, and this leads to very large likelihood.



- There are methods to avoid singularity
- Tricks
 - If a Gaussian component is collapsing into very small Σ_k , reset its mean μ_k to randomly chosen values, and Σ_k to a large value
- Systematic methods
 - Maximum-a-Posterior (MAP) Add prior constraints to the parameters
 - Bayesian approach Prior and proper uncertainty estimation of the parameters

Solve the MLE

Maximum Likelihood Formulation of GMM

$$\pi, \mu, \mathbf{\Sigma} = rg \max_{\pi, \mu, \mathbf{\Sigma}} \ln p(\mathbf{X}|\pi, \mu, \mathbf{\Sigma}) = rg \max_{\pi, \mu, \mathbf{\Sigma}} \sum_{n=1}^{N} \ln \left\{ \sum_{k=1}^{K} \pi_k \mathcal{N}(\mathbf{x}_n | \mu_k, \mathbf{\Sigma}_k) \right\}$$

- Solve π , μ , Σ iteratively, fixed the others, solve one
- $lue{}$ Let's look at μ_k first
- Compute the first order derivatives with respect to μ_k

$$0 = -\sum_{n=1}^{N} \frac{\pi_k \mathcal{N}(\mathbf{x}_n | \boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k)}{\sum_j \pi_j \mathcal{N}(\mathbf{x}_n | \boldsymbol{\mu}_j, \boldsymbol{\Sigma}_j)} \boldsymbol{\Sigma}_k(\mathbf{x}_n - \boldsymbol{\mu}_k)$$

$$\gamma(z_{nk})$$

- Here N_k can be interpreted as the effective number of points assigned to cluster k
- μ_k is the weighted average of all point in the data set
- The weight is the posterior probability $\gamma(z_{nk})$
- The denominator is the effective number N_k

$$0 = -\sum_{n=1}^{N} \gamma(z_{nk}) \mathbf{\Sigma}_k (\mathbf{x}_n - \mu_k)$$

这里公式不全,对于Nk等参数 的介绍,参考该ppt的后面部分 Compute the first order derivatives with respect to Σ_k , we can solve Σ_k similarly,

$$\Sigma_k = \frac{1}{N_k} \sum_{n=1}^{N} \gamma(z_{nk}) (\mathbf{x}_n - \mu_k) (\mathbf{x}_n - \mu_k)^T$$

- Σ_k is the weighted average of all point's variance centered by μ_k
- The weight is the posterior probability $\gamma(z_{nk})$
- The denominator is the effective number N_k

$$\ln p(\mathbf{X}|\pi,\mu,\Sigma) = \sum_{n=1}^{N} \ln \left\{ \pi_k \mathcal{N}(\mathbf{x_n}|\mu_k,\Sigma_k) \right\}$$

- Now let's solve for π_k
- However, there is a constraint $\sum_{k=1}^K \pi_k = 1$
- lacksquare Lagrange Multiplier λ

$$\ln p(\mathbf{X}|\pi,\mu,\Sigma) + \lambda \left(\sum_{k=1}^{K} \pi_k - 1\right)$$

Compute first order derivate with respect to π_k and make it 0

$$0 = \sum_{n=1}^{N} \frac{\mathcal{N}(\mathbf{x}_n | \mu_k, \Sigma_k)}{\sum_{j=1}^{K} \pi_j \mathcal{N}(\mathbf{x}_n | \mu_k, \Sigma_k)} + \lambda$$

Solve the MLE

Solve λ first

$$0 = \sum_{n=1}^{N} \frac{\mathcal{N}(\mathbf{x}_{n}|\mu_{k}, \Sigma_{k})}{\sum_{j=1}^{K} \pi_{j} \mathcal{N}(\mathbf{x}_{n}|\mu_{k}, \Sigma_{k})} + \lambda$$

$$0 = \sum_{k=1}^{K} \sum_{n=1}^{N} \frac{\pi_{k} \mathcal{N}(\mathbf{x}_{n}|\mu_{k}, \Sigma_{k})}{\sum_{j=1}^{K} \pi_{j} \mathcal{N}(\mathbf{x}_{n}|\mu_{k}, \Sigma_{k})} + \sum_{k=1}^{K} \pi_{k} \lambda$$

$$0 = \left(\sum_{n=1}^{N} 1\right) + \lambda$$

$$\lambda = -N$$

$$\gamma(z_{nk}) = \frac{\pi_{k} \mathcal{N}(\mathbf{x}_{n}|\mu_{k}, \Sigma_{k})}{\sum_{j=1}^{K} \pi_{j} \mathcal{N}(\mathbf{x}_{n}|\mu_{j}, \Sigma_{j})}$$

$$N_{k} = \sum_{n=1}^{N} \gamma(z_{nk})$$

$$\gamma(z_{nk}) = \frac{\pi_k \mathcal{N}(\mathbf{x}_n | \boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k)}{\sum_{j=1}^K \pi_j \mathcal{N}(\mathbf{x}_n | \boldsymbol{\mu}_j, \boldsymbol{\Sigma}_j)}$$
$$N_k = \sum_{n=1}^N \gamma(z_{nk})$$

\$ Solve the MLE

- $lue{}$ Solve π_k
- Intuitively, π_k is the effective cluster member number over data set size.

$$0 = \sum_{n=1}^{N} \frac{\mathcal{N}(\mathbf{x}_n | \mu_k, \Sigma_k)}{\sum_{j=1}^{K} \pi_j \mathcal{N}(\mathbf{x}_n | \mu_k, \Sigma_k)} + \lambda$$

$$0 = \frac{1}{\pi_k} \sum_{n=1}^{N} \frac{\pi_k \mathcal{N}(\mathbf{x}_n | \mu_k, \Sigma_k)}{\sum_{j=1}^{K} \pi_j \mathcal{N}(\mathbf{x}_n | \mu_k, \Sigma_k)} - N$$

$$0 = \frac{1}{\pi_k} \sum_{n=1}^{N} \gamma_{nk} - N$$

$$N_k$$

$$\gamma(z_{nk}) = \frac{\pi_k \mathcal{N}(\mathbf{x}_n | \boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k)}{\sum_{j=1}^K \pi_j \mathcal{N}(\mathbf{x}_n | \boldsymbol{\mu}_j, \boldsymbol{\Sigma}_j)}$$
$$N_k = \sum_{n=1}^N \gamma(z_{nk})$$

Summary for GMM MLE

- 1. Initialize the means μ_k , covariances Σ_k and weights π_k
- 2. E-step. Evaluate the posterior $p(z_{nk} = 1 | x_n)$, intuitively this is the probability of x_n being assigned to each of the K clusters.

$$\gamma(z_{nk}) = \frac{\pi_k \mathcal{N}(\mathbf{x}_n | \boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k)}{\sum_{j=1}^K \pi_j \mathcal{N}(\mathbf{x}_n | \boldsymbol{\mu}_j, \boldsymbol{\Sigma}_j)}$$

上式表示N个点中某一个点n,属于第k类的概率

- 3. M-Step. Estimate the parameters using MLE.
- 4. Evaluate the log likelihood, if converges, stop. Otherwise go back to E-step

$$\mu_k^{\text{new}} = \frac{1}{N_k} \sum_{n=1}^N \gamma(z_{nk}) \mathbf{x}_n$$

$$\Sigma_k^{\text{new}} = \frac{1}{N_k} \sum_{n=1}^N \gamma(z_{nk}) (\mathbf{x}_n - \boldsymbol{\mu}_k^{\text{new}}) (\mathbf{x}_n - \boldsymbol{\mu}_k^{\text{new}})^{\text{T}}$$

$$\pi_k^{\text{new}} = \frac{N_k}{N}$$

$$\ln p(\mathbf{X}|\boldsymbol{\mu}, \boldsymbol{\Sigma}, \boldsymbol{\pi}) = \sum_{n=1}^{N} \ln \left\{ \sum_{k=1}^{K} \pi_k \mathcal{N}(\mathbf{x}_n | \boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k) \right\}$$

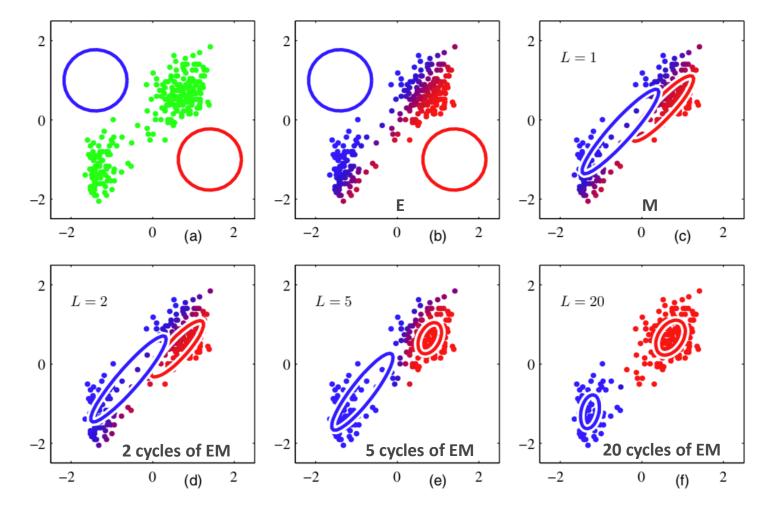


Figure source: Pattern Recognition and Machine Learning, Bishop

Expectation – Maximization (EM)

- Given a joint distribution $p(X,Z|\theta)$ over observed variables X and latent variables Z, governed by parameters θ , the goal is to maximize the likelihood function $p(X|\theta)$ with respect to θ
 - 1. Choose an initial parameter θ^{old}
 - 2. E-step. Evaluate $p(Z|X, \theta^{old})$

- $p(\mathbf{X}|\boldsymbol{\theta}) = \sum_{\mathbf{Z}} p(\mathbf{X}, \mathbf{Z}|\boldsymbol{\theta})$
- 3. M-step. Evaluate θ^{new} by solving the optimization problem

$$egin{aligned} heta^{ ext{new}} &= rg\max_{ heta} \mathcal{Q}(heta, heta^{ ext{old}}) \ \mathcal{Q}(heta, heta^{ ext{old}}) &= \sum_{\mathbf{z}} p(\mathbf{Z}|\mathbf{X}, heta^{ ext{old}}) \ln p(\mathbf{X}, \mathbf{Z}| heta) = \mathbb{E}_{\mathbf{z}} [\ln p(\mathbf{X}, \mathbf{Z}| heta)] \end{aligned}$$

4. Check for convergence, if not, go back to E-step

$$\theta^{old} \leftarrow \theta^{new}$$

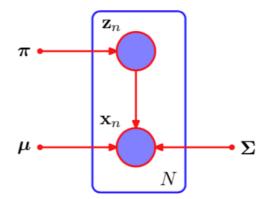
\$ GMM revisited

In GMM, we have the joint distribution p(x,z) = p(z)p(x|z)

$$p(\mathbf{z}) = \prod_{k=1}^{K} \pi_k^{z_k}$$
 $p(\mathbf{x}|\mathbf{z}) = \prod_{k=1}^{K} \mathcal{N}(\mathbf{x}|\boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k)^{z_k}$

Alternatively

$$p(\mathbf{X}, \mathbf{Z} | \boldsymbol{\mu}, \boldsymbol{\Sigma}, \boldsymbol{\pi}) = \prod_{n=1}^{N} \prod_{k=1}^{K} \pi_k^{z_{nk}} \mathcal{N}(\mathbf{x}_n | \boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k)^{z_{nk}}$$



where z_{nk} denotes the k^{th} component of \mathbf{z}_n

Take the logarithm, we have

$$\ln p(\mathbf{X}, \mathbf{Z} | \boldsymbol{\mu}, \boldsymbol{\Sigma}, \boldsymbol{\pi}) = \sum_{n=1}^{N} \sum_{k=1}^{K} z_{nk} \left\{ \ln \pi_k + \ln \mathcal{N}(\mathbf{x}_n | \boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k) \right\}$$

Let's recall the M-step of EM algorithm

$$egin{aligned} heta^{ ext{new}} &= rg\max_{ heta} \mathcal{Q}(heta, heta^{ ext{old}}) \ \mathcal{Q}(heta, heta^{ ext{old}}) &= \sum_{\mathbf{z}} p(\mathbf{Z}|\mathbf{X}, heta^{ ext{old}}) \ln p(\mathbf{X}, \mathbf{Z}| heta) = \mathbb{E}_{\mathbf{z}} [\ln p(\mathbf{X}, \mathbf{Z}| heta)] \end{aligned}$$

In the context of GMM, we have

$$\ln p(\mathbf{X}, \mathbf{Z} | \boldsymbol{\mu}, \boldsymbol{\Sigma}, \boldsymbol{\pi}) = \sum_{n=1}^{N} \sum_{k=1}^{K} z_{nk} \left\{ \ln \pi_k + \ln \mathcal{N}(\mathbf{x}_n | \boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k) \right\}$$

$$\mathbb{E}_{z}[\ln p(\mathbf{X}, \mathbf{Z} | \mu, \Sigma, \pi)] = \sum_{n=1}^{N} \sum_{k=1}^{K} \mathbb{E}[z_{nk}] \{\ln \pi_k + \ln \mathcal{N}(\mathbf{x}_n | \mu_k, \Sigma_k)\}$$

Recall the posterior in GMM

$$p(z_{nk} = 1 | \mathbf{x}_n, \mu_k, \Sigma_k, \pi_k) = \gamma(z_{nk}) = \frac{\pi_k \mathcal{N}(\mathbf{x}_n | \mu_k, \Sigma_k)}{\sum_{j=1}^K \pi_j \mathcal{N}(\mathbf{x}_n | \mu_j, \Sigma_j)}$$

So we have

$$\mathbb{E}[z_{nk}] = \sum_{z_{nk}} z_{nk} p(z_{nk} | \mathbf{x}_n, \mu_k, \Sigma_k, \pi_k)$$

$$= 0 \cdot p(z_{nk} = 0 | \mathbf{x}_n, \mu_k, \Sigma_k, \pi_k) + 1 \cdot p(z_{nk} = 1 | \mathbf{x}_n, \mu_k, \Sigma_k, \pi_k)$$

$$= \gamma(z_{nk})$$

$$egin{aligned} heta^{ ext{new}} &= rg\max_{ heta} \mathcal{Q}(heta, heta^{ ext{old}}) & ext{M-step} \ \mathcal{Q}(heta, heta^{ ext{old}}) &= \sum_{ heta} p(\mathbf{Z}|\mathbf{X}, heta^{ ext{old}}) \ln p(\mathbf{X}, \mathbf{Z}| heta) = \mathbb{E}_{\mathbf{z}} [\ln p(\mathbf{X}, \mathbf{Z}| heta)] \end{aligned}$$

 $lue{}$ Now we have Q function for the GMM problem

$$\mathbb{E}_{\mathbf{Z}}[\ln p(\mathbf{X}, \mathbf{Z} | \boldsymbol{\mu}, \boldsymbol{\Sigma}, \boldsymbol{\pi})] = \sum_{n=1}^{N} \sum_{k=1}^{K} \gamma(z_{nk}) \left\{ \ln \pi_k + \ln \mathcal{N}(\mathbf{x}_n | \boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k) \right\}$$

Solving for the maximum we have the same result as the previous GMM solution

$$\boldsymbol{\mu}_{k}^{\text{new}} = \frac{1}{N_{k}} \sum_{n=1}^{N} \gamma(z_{nk}) \mathbf{x}_{n}$$

$$\boldsymbol{\Sigma}_{k}^{\text{new}} = \frac{1}{N_{k}} \sum_{n=1}^{N} \gamma(z_{nk}) \left(\mathbf{x}_{n} - \boldsymbol{\mu}_{k}^{\text{new}}\right) \left(\mathbf{x}_{n} - \boldsymbol{\mu}_{k}^{\text{new}}\right)^{\text{T}}$$

$$\boldsymbol{\pi}_{k}^{\text{new}} = \frac{N_{k}}{N}$$

- In summary, we derivate the GMM result using two methods
 - 1. Standard MLE formulation

$$egin{aligned} \pi, \mu, oldsymbol{\Sigma} &= rg \max_{\pi, \mu, oldsymbol{\Sigma}} \ln p(\mathbf{X}|\pi, \mu, oldsymbol{\Sigma}) = rg \max_{\pi, \mu, oldsymbol{\Sigma}} \sum_{n=1}^N \ln iggl\{ \sum_{k=1}^K \pi_k \mathcal{N}(\mathbf{x}_n|\mu_k, oldsymbol{\Sigma}_k) iggr\} \end{aligned}$$

2. EM algorithm

$$\mathbb{E}_{\mathbf{Z}}[\ln p(\mathbf{X}, \mathbf{Z} | \boldsymbol{\mu}, \boldsymbol{\Sigma}, \boldsymbol{\pi})] = \sum_{n=1}^{N} \sum_{k=1}^{K} \gamma(z_{nk}) \left\{ \ln \pi_k + \ln \mathcal{N}(\mathbf{x}_n | \boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k) \right\}$$

- K-Means is a special of the GMM
 - The soft assignment becomes hard $\gamma(z_{nk}) \rightarrow r_{nk}$
 - The variance is zero $\Sigma_k \to 0$
- In the case that $\Sigma_k = \epsilon I$, the original $\gamma(z_{nk})$ becomes

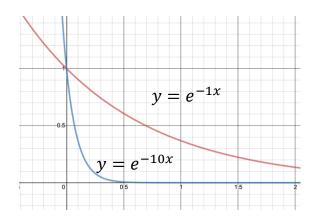
$$\gamma(z_{nk}) = \frac{\pi_k \exp\left\{-\|\mathbf{x}_n - \boldsymbol{\mu}_k\|^2 / 2\epsilon\right\}}{\sum_j \pi_j \exp\left\{-\|\mathbf{x}_n - \boldsymbol{\mu}_j\|^2 / 2\epsilon\right\}}$$

Why? Let's look at the two equations below

$$\gamma(z_{nk}) = \frac{\pi_k \mathcal{N}(\mathbf{x}_n | \mu_k, \Sigma_k)}{\sum_{j=1}^K \pi_j \mathcal{N}(\mathbf{x}_n | \mu_j, \Sigma_j)} \qquad p(\mathbf{x} | \boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k) = \frac{1}{(2\pi\epsilon)^{1/2}} \exp\left\{-\frac{1}{2\epsilon} \|\mathbf{x} - \boldsymbol{\mu}_k\|^2\right\}$$

$$\gamma(z_{nk}) = \frac{\pi_k \exp\left\{-\|\mathbf{x}_n - \boldsymbol{\mu}_k\|^2 / 2\epsilon\right\}}{\sum_j \pi_j \exp\left\{-\|\mathbf{x}_n - \boldsymbol{\mu}_j\|^2 / 2\epsilon\right\}}$$

- Let's consider the limit $\epsilon \to 0$
- Look at the denominator
- The j that $\|x_n \mu_j\|^2$ is smallest goes to zero most slowly
- All other j goes to zero quickly
- Therefore $\gamma(z_{nk}) \to r_{nk}$



$$r_{nk} = \begin{cases} 1 & \text{if } k = \arg\min_{j} \|\mathbf{x}_n - \boldsymbol{\mu}_j\|^2 \\ 0 & \text{otherwise.} \end{cases}$$

The M-step of GMM/EM gives the following

$$\mathbb{E}_{\mathbf{Z}}[\ln p(\mathbf{X}, \mathbf{Z} | \boldsymbol{\mu}, \boldsymbol{\Sigma}, \boldsymbol{\pi})] = \sum_{n=1}^{N} \sum_{k=1}^{K} \gamma(z_{nk}) \left\{ \ln \pi_k + \ln \mathcal{N}(\mathbf{x}_n | \boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k) \right\}$$

- Given $\Sigma_k = \epsilon I$, $\epsilon \to 0$, $\gamma(z_{nk}) \to r_{nk}$
- It can be written as

$$\mathbb{E}_{\mathbf{Z}}[\ln p(\mathbf{X}, \mathbf{Z} | \boldsymbol{\mu}, \boldsymbol{\Sigma}, \boldsymbol{\pi})] \rightarrow -\frac{1}{2} \sum_{n=1}^{N} \sum_{k=1}^{K} r_{nk} \|\mathbf{x}_n - \boldsymbol{\mu}_k\|^2 + \text{const}$$

Which is the same as the distortion measure J in the K-means algorithm

$$J = \sum_{n=1}^{N} \sum_{k=1}^{K} r_{nk} \|\mathbf{x}_n - \boldsymbol{\mu}_k\|^2$$

Conclusion for K-Means, GMM, EM

- GMM is a special application of EM
 - The latent variable $\{z_{nk}\}$ and the model parameters $\{\mu_k, \Sigma_k, \pi_k\}$ are estimated in the E-step and M-step
 - The clustering probability is given by $\gamma(z_{nk})=p(z_{nk}=1|x_n,\mu_k,\Sigma_k,\pi_k)$

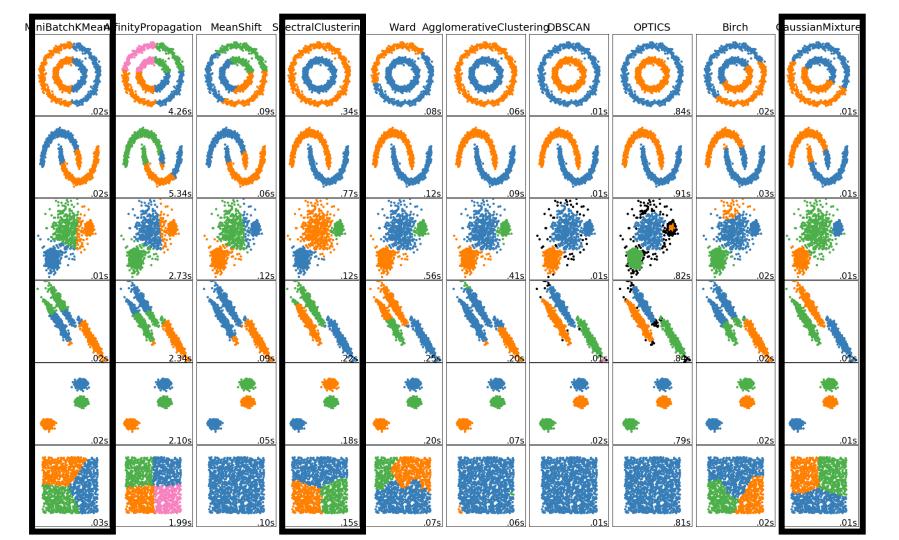
- K-Means is a special case of GMM, where
 - The soft assignment becomes hard $\gamma(z_{nk}) \to r_{nk}$
 - The variance is zero $\Sigma_k \to 0$

S K-Means Summary

- $\bigcirc Complexity O(t \cdot k \cdot n \cdot d)$
 - t- number of iteration, k- number of clusters, n- number of points, d- number of dimension
- Advantage
 - Simple
 - Fast
- Disadvantage
 - Assumes circle/sphere-like data same variance in all directions
 - Need to provide number of clusters
 - Sensitive to initialization
 - Sensitive to outlier
 - Can be alleviated by K-Medoids

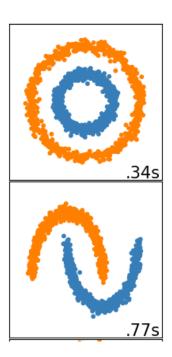
\$ GMM Summary

- $\bigcirc Complexity O(t \cdot k \cdot n \cdot d)$
 - t- number of iteration, k- number of clusters, n- number of points, d- number of dimension
- Advantage
 - Provides uncertainty estimation soft assignment
 - More robust to outlier
- Disadvantage
 - Need to provide number of clusters
 - Mainly works with convex clusters
 - · Gaussian distributions are in the shape of ellipsoid
 - Singularity problem



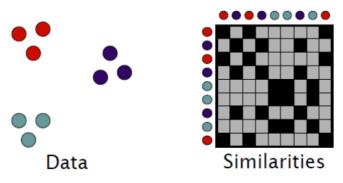
\$ Spectral Clustering

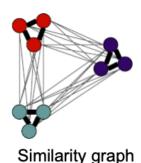
- K-Means fails row 1, 2, 3, 4
- GMM fails row 1, 2
- Why? They are based on Euclidean distance!
- How about other metric? E.g., connectivity?
- Spectral Clustering works with connectivity!



Spectral Clustering

- Works for a (Weighted) Undirected Graph G=(V,E), where $V=\{v_1,\cdots,v_n\}$,
- The edge between two vertices v_i, v_j carries weight $w_{ij} \ge 0$.
 - Two vertices are not connected if $w_{ij} = 0$
 - Specially, $w_{ii} = 0$
- The weighted adjacency matrix / similarity matrix is represented by $W=\left(w_{ij}\right)_{i,j=1,\cdots,n}$







Spectral Clustering – How to build the Graph

- Every point is a vertex
- There are 3 common graphs built for Spectral Clustering
 - 1. ϵ -neighborhood graph. Radius nearest neighbor search. $w_{ij} = d(v_i, v_j)$
 - 2. k-neighborhood graph. KNN search. $w_{ij} = d(v_i, v_j)$
 - a) Build an edge if v_i is one of the knn of v_i OR v_i is one of the knn of v_i
 - b) Build an edge if v_i is one of the knn of v_i AND v_i is one of the knn of v_i
 - 3. Fully connected graph. Connect every point.

Spectral Clustering – Laplacian Matrix

- Degree matrix D
 - A diagonal matrix with degrees d_1, \dots, d_n on the diagonal
 - $d_i = \sum_{j=1}^n w_{ij}$ is the row sum of adjacency matrix W
- $loodsymbol{loodsymbol{loodsymbol{loodsymbol{eta}}}}$ Unnormalized graph Laplacian matrix L=D-W
- Normalized graph Laplacian matrix

•
$$L_{sym} = D^{-1/2}LD^{-1/2} = I - D^{-1/2}WD^{-1/2}$$

•
$$L_{rw} = D^{-1}L = I - D^{-1}W$$

Unnormalized Spectral Clustering

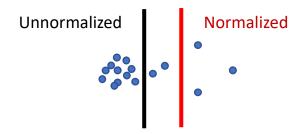
- 1. Build the graph to get adjacency matrix $W \in \mathbb{R}^{n \times n}$
- 2. Compute unnormalized Laplacian *L*
- 3. Compute the first (smallest) k eigenvectors v_1, \dots, v_k of L
- 4. Let $V \in \mathbb{R}^{n \times k}$ be the matrix contraining the vectors v_1, \dots, v_k as columns
- 5. For $i=1,\cdots n$, let $y_i\in\mathbb{R}^k$ be the vector corresponding to the i-th row of V
- 6. Cluster the points $\{y_i \in \mathbb{R}^k\}$ with k-means algorithm into clusters C_1, \dots, C_k
- 7. The final output clusters are A_1, \dots, A_k where $A_i = \{j | y_j \in C_i\}$

Normalized Spectral Clustering

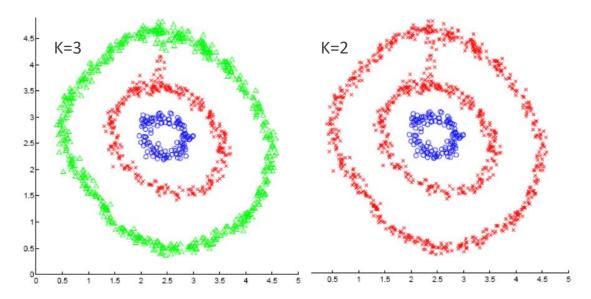
- 1. Build the graph to get adjacency matrix $W \in \mathbb{R}^{n \times n}$
- 2. Compute normalized Laplacian $L' = L_{rw}$
- 3. Compute the first (smallest) k eigenvectors v_1, \dots, v_k of L'
- 4. Let $V \in \mathbb{R}^{n \times k}$ be the matrix contraining the vectors v_1, \dots, v_k as columns
- 5. For $i=1,\cdots n$, let $y_i\in\mathbb{R}^k$ be the vector corresponding to the i-th row of V
- 6. Cluster the points $\{y_i \in \mathbb{R}^k\}$ with k-means algorithm into clusters C_1, \dots, C_k
- 7. The final output clusters are A_1, \dots, A_k where $A_i = \{j | y_j \in C_i\}$

\$ Unnormalized vs. Normalized

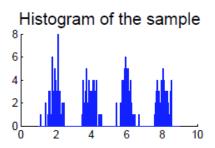
- Unnormalized spectral clustering tends to make clusters have equal number of vertices
- Normalized tends to group same density clusters
- Why? Lecture-4

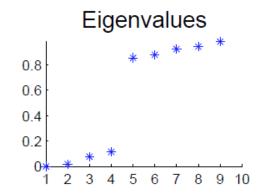


- Usually normalized spectral clustering works well with un-uniform point density.
- $lue{}$ Selection of number of clusters k



- loop Selection of k can be done by eigenvalue analysis
- lacktriangle Most stable clustering is given by the value of k that maximizes the eigen-gap
 - Eigengap is the difference between consecutive eigenvalues
 - $\Delta_k = |\lambda_k \lambda_{k-1}|$





Spectral Clustering

- Complexity: $O(n^3)$
 - This is the complexity of eigen decomposition
 - K-means complexity is $O(t \cdot k \cdot n \cdot d)$
- Advantage
 - No assumption on cluster shape
 - · Works with similarity, including Euclidean, connectivity
 - Works with any dimensional data
 - Able to estimate the number of clusters
- Disadvantage
 - Computational expensive
 - Can be alleviated using sparse similarity matrix and sparse eigen solver

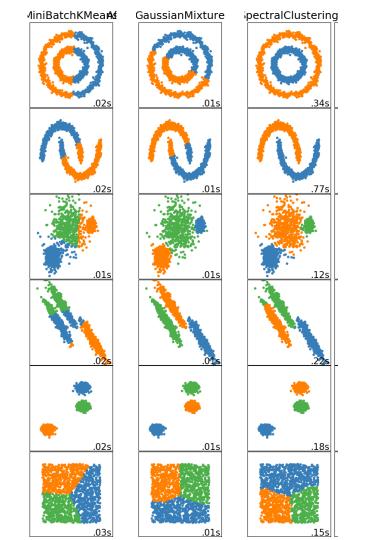




- Euclidean distance
- Hard assignment
- No modeling for a cluster
- Pre-defined cluster number k



- Euclidean distance
- Probability formulation soft clustering
- Mean and variance estimation for each cluster
- Pre-defined cluster number *k*
- Spectral Clustering
 - Works with connectivity
 - Heuristic to determine cluster number k



S Next Lecture

- First half of next lecture covers other Clustering algorithms:
 - Proof and intuition of Spectral Clustering
 - DBSCAN
 - Mean-Shift
- Second half covers Model Fitting:
 - Hough Transform
 - Sample Consensus

\$ Homework

- Generate clustering dataset using sklearn
 - https://scikit-learn.org/stable/auto-examples/cluster/plot-cluster-comparison.html
- Implement your own version of
 - K-Means
 - GMM
 - Spectral Clustering
- Visualize and compare the results with the standard results
 - https://scikit-learn.org/stable/modules/clustering.html#overview-of-clustering-methods



感谢各位聆听 Thanks for Listening