

# Multifidelity write-up

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## Abstract

We derived the explicit sample size estimation in terms of the desired accuracy requirement.

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## 1. Model problem

We have a high fidelity model denoted as  $f_1 : \Omega \rightarrow Z$  we desire, and several low fidelity models (surrogates)  $f_k$  for  $k \in \mathbb{N}$ . Our objective is to approximate

$$\mathbb{E}(f_1(\omega)).$$

Note for each  $f_i(\omega)$ , its variance and Pearson product-moment correlation coefficient are

$$\sigma_i^2 = \mathbb{V}(f_i(\omega)), \quad \rho_{i,j} = \frac{\text{Cov}(f_i(\omega), f_j(\omega))}{\sigma_i \sigma_j}, \quad i, j = 1, \dots, K,$$

where  $\mathbb{V}(f) := \mathbb{E}(\|f - \mathbb{E}(f)\|_Z^2)$ .

## 2. Monte Carlo estimator

The Monte Carlo estimator for the expectation of each  $f_k$  is defined as the sample mean of  $N_k$  i.i.d realizations  $\omega_1, \dots, \omega_N$

$$A_{k,N_k}^{\text{MC}} := \frac{1}{N_k} \sum_{i=1}^{N_k} f_k(\omega_i), \quad \forall k = 1, \dots, K, \quad (1)$$

where  $\mathbb{E}(A_{k,N_k}^{\text{MC}}) = \mathbb{E}(f_k)$ ,  $\mathbb{V}(A_{k,N_k}^{\text{MC}}) = \mathbb{V}(f_k)/N_k$ . Let  $C_k$  denote the average evaluation cost per sample for  $f_k$ , then the total sampling cost for each Monte Carlo estimator  $A_{k,N_k}^{\text{MC}}$  is

$$\mathcal{W}_{\text{MC}}^k = C_k N_k.$$

We define the *normalized mean squared error* (nMSE), denoted as  $\mathcal{E}_A^2$ , with normalizing factor  $\|\mathbb{E}(f)\|_Z^2$  for estimator  $A$  as

$$\mathcal{E}_A^2 := \frac{\mathbb{E}[\|\mathbb{E}(f) - A\|_Z^2]}{\|\mathbb{E}(f)\|_Z^2}.$$

If we use a Monte Carlo estimator to estimate  $\mathbb{E}(f_1(\omega))$ ,

$$\mathcal{E}_{A_{1,N_1}^{\text{MC}}}^2 = \frac{\mathbb{E}[\|\mathbb{E}(f_1) - A_{1,N_1}^{\text{MC}}\|_Z^2]}{\|\mathbb{E}(f_1)\|_Z^2} = \frac{\|\mathbb{E}(f_1) - \mathbb{E}(A_{1,N_1}^{\text{MC}})\|_Z^2 + \mathbb{E}[\|\mathbb{E}(A_{1,N_1}^{\text{MC}}) - A_{1,N_1}^{\text{MC}}\|_Z^2]}{\|\mathbb{E}(f_1)\|_Z^2} = \frac{\mathbb{V}(f_1)}{N_1 \|\mathbb{E}(f_1)\|_Z^2}.$$

Given a target tolerance  $\epsilon^2$  for the nMSE, the sample size  $N_1$  is estimated as

$$N_1 = \left\lceil \frac{\mathbb{V}(f_k)}{\epsilon^2 \|\mathbb{E}(f_k)\|_Z^2} \right\rceil \simeq \epsilon^{-2}.$$

So the total expense of sampling with  $N_1$  samples using Monte Carlo method for  $\mathbb{E}(f_1(\omega))$  is

$$\mathcal{W}_{\text{MC}}^1 = C_1 N_1.$$

### 3. Multifidelity Monte Carlo

The Multifidelity Monte Carlo (MFMC) estimator is defined as

$$A_{\text{MFMC}} := A_{1,N_1}^{\text{MC}} + \sum_{k=2}^K \alpha_k (A_{k,N_k}^{\text{MC}} - A_{k,N_{k-1}}^{\text{MC}}), \quad (2)$$

where  $\alpha_k$  are the coefficients to weight the correction term. In each correction term, the two Monte Carlo estimators are dependent in the sense that  $A_{k,N_{k-1}}^{\text{MC}}$  recycles the first  $N_{k-1}$  samples of  $A_{k,N_k}^{\text{MC}}$  so we require  $N_{k-1} \leq N_k$  for  $k = 2, \dots, K$ . Using this property and (1), we can remove the dependent samples and rewrite (2) as

$$A^{\text{MFMC}} = A_{1,N_1}^{\text{MC}} + \sum_{k=2}^K \alpha_k \left[ \left( \frac{N_{k-1}}{N_k} - 1 \right) A_{k,N_{k-1}}^{\text{MC}} + \left( 1 - \frac{N_{k-1}}{N_k} \right) A_{k,N_k-N_{k-1}}^{\text{MC}} \right],$$

Now, the samples in  $A_{k,N_{k-1}}^{\text{MC}}$  and  $A_{k,N_k-N_{k-1}}^{\text{MC}}$  are independent with each other. Define  $Y_1 := A_{1,N_1}^{\text{MC}}$ ,  $Y_k := \left( \frac{N_{k-1}}{N_k} - 1 \right) A_{k,N_{k-1}}^{\text{MC}} + \left( 1 - \frac{N_{k-1}}{N_k} \right) A_{k,N_k-N_{k-1}}^{\text{MC}}$  for  $k = 2, \dots, K$ . Note that  $Y_k$  are independent with each other for  $k = 2, \dots, K$ . Then

$$A^{\text{MFMC}} = Y_1 + \sum_{k=2}^K \alpha_k Y_k.$$

So  $\mathbb{E}(Y_k) = 0$  for  $k \geq 2$  and  $\mathbb{E}(A^{\text{MFMC}}) = \mathbb{E}(f_1)$ . For independent random variable, since each realization is uncorrelated to each other, this indicates that the sum of sample realizations and variance are interchangeable, therefore

$$\mathbb{V}(Y_1) = \frac{\sigma_1^2}{N_1}, \quad \mathbb{V}(Y_k) = \left( \frac{N_{k-1}}{N_k} - 1 \right)^2 \frac{\sigma_k^2}{N_{k-1}} + \left( 1 - \frac{N_{k-1}}{N_k} \right)^2 \frac{\sigma_k^2}{N_k - N_{k-1}} = \left( \frac{1}{N_{k-1}} - \frac{1}{N_k} \right) \sigma_k^2.$$

So,

$$\begin{aligned} \mathbb{V}(A^{\text{MFMC}}) &= \mathbb{V}(Y_1) + \mathbb{V}\left(\sum_{k=2}^K \alpha_k Y_k\right) + 2 \text{Cov}\left(Y_1, \sum_{k=2}^K \alpha_k Y_k\right), \\ &= \mathbb{V}(Y_1) + \sum_{k=2}^K \alpha_k^2 \mathbb{V}(Y_k) + 2 \sum_{2 \leq k < j \leq K} \alpha_k \alpha_j \text{Cov}(Y_k, Y_j) + 2 \sum_{k=2}^K \alpha_k \text{Cov}(Y_1, Y_k), \\ &= \frac{\sigma_1^2}{N_1} + \sum_{k=2}^K \left( \frac{1}{N_{k-1}} - \frac{1}{N_k} \right) (\alpha_k^2 \sigma_k^2 - 2\alpha_k \rho_{1,i} \sigma_1 \sigma_i), \end{aligned} \quad (3)$$

where we use the fact that  $Y_k$  and  $Y_j$  are independent with each other for  $k \geq 2$  and [1, Lemma 3.2]

$$\text{Cov}(Y_1, Y_k) = \text{Cov}(A_{1,N_1}^{\text{MC}}, A_{k,N_k}^{\text{MC}}) - \text{Cov}(A_{1,N_1}^{\text{MC}}, A_{k,N_{k-1}}^{\text{MC}}) = -2 \sum_{k=2}^K \alpha_k \left( \frac{1}{N_{k-1}} - \frac{1}{N_k} \right) \rho_{1,i} \sigma_1 \sigma_i.$$

The nMSE error for the multifidelity Monte Carlo estimator is

$$\mathcal{E}_{A^{\text{MFMC}}}^2 = \frac{\mathbb{E} \left[ \left\| \mathbb{E}(f_1) - A^{\text{MFMC}} \right\|_Z^2 \right]}{\left\| \mathbb{E}(f_1) \right\|_Z^2} = \frac{\left\| \mathbb{E}(f_1) - \mathbb{E}(A^{\text{MFMC}}) \right\|_Z^2 + \mathbb{E} \left[ \left\| \mathbb{E}(A^{\text{MFMC}}) - A^{\text{MFMC}} \right\|_Z^2 \right]}{\left\| \mathbb{E}(f_1) \right\|_Z^2} = \frac{\mathbb{V}(A^{\text{MFMC}})}{\left\| \mathbb{E}(f_1) \right\|_Z^2}.$$

The total sampling cost for MFMC estimator is

$$\sum_{k=1}^K C_k N_k.$$

Our next goal is to determine the sample size  $N_k$  such that the MFMC estimation satisfy the accuracy threshold  $\mathcal{E}_{A^{\text{MFMC}}}^2 \leq \epsilon^2$ . We formulate an following optimization problem to minimize the sampling cost subject to a bounded variance of MFMC estimator, and solve for sample size  $N_k \in \mathbb{R}$  for  $k = 1 \dots, K$  and  $\alpha_k \in \mathbb{R}$  for  $k = 2 \dots, K$  as

$$\begin{aligned} \min_{N_1, \dots, N_K \in \mathbb{R}, \alpha_2, \dots, \alpha_K \in \mathbb{R}} \quad & \sum_{k=1}^K C_k N_k, \\ \text{s.t.} \quad & -N_1 \leq 0, \\ & N_{k-1} - N_k \leq 0, \quad k = 2 \dots, K, \\ & \mathbb{V}(A^{\text{MFMC}}) \leq \left\| \mathbb{E}(f_1) \right\|_Z^2 \epsilon^2. \end{aligned} \tag{4}$$

**Theorem 1.** Let  $f_k$  be  $K$  models that satisfy the following conditions

$$(i) \quad |\rho_{1,1}| > \dots > |\rho_{1,K}| \quad (ii) \quad \frac{C_{k-1}}{C_k} > \frac{\rho_{1,k-1}^2 - \rho_{1,k}^2}{\rho_{1,k}^2 - \rho_{1,k+1}^2}, \quad k = 2, \dots, K.$$

Then the global minimizer to (4) is

$$\begin{aligned} \alpha_k^* &= \frac{\rho_{1,k} \sigma_1}{\sigma_k}, \quad k = 2 \dots, K, \\ N_k^* &= \frac{\sigma_1^2}{\left\| \mathbb{E}(f_1) \right\|_Z^2 \epsilon^2} \sqrt{\frac{\rho_{1,k}^2 - \rho_{1,k+1}^2}{C_k}} \sum_{j=1}^K \left( \sqrt{\frac{C_j}{\rho_{1,j}^2 - \rho_{1,j+1}^2}} - \sqrt{\frac{C_{j-1}}{\rho_{1,j-1}^2 - \rho_{1,j}^2}} \right) \rho_{1,j}^2, \quad k = 1 \dots, K, \end{aligned}$$

with  $\rho_{1,0} = \infty$  and  $\rho_{1,K+1} = 0$ .

*Proof.* First note that (ii) indicates a strict increasing sequence of sample size, the minimizer  $\alpha_k^*$  and  $N_k^*$  satisfy the constraints (i) and (ii).

Our objective function

$$J^* = \sum_{k=1}^K C_k N_k^* = \frac{\sigma_1^2}{\left\| \mathbb{E}(f_1) \right\|_Z^2 \epsilon^2} \sum_{k=1}^K \sqrt{(\rho_{1,k}^2 - \rho_{1,k+1}^2) C_k} \sum_{k=1}^K \left( \sqrt{\frac{C_k}{\rho_{1,k}^2 - \rho_{1,k+1}^2}} - \sqrt{\frac{C_{k-1}}{\rho_{1,k-1}^2 - \rho_{1,k}^2}} \right) \rho_{1,k}^2$$

Consider another local minimum  $\alpha_k^+$  and  $N_k^+$  of (4) such that  $N_{k-1}^+ = N_k^+$  for some  $k \geq 2$ . Let index  $j_i$  be the first entry followed by equality for  $N_k$ ,

$$N_{k_i}$$

□

## 4. Appendix

### References

- [1] B. Peherstorfer, K. Willcox, and M. Gunzburger. Optimal model management for multifidelity Monte Carlo estimation. SIAM J. Sci. Comput., 38(5):A3163–A3194, 2016.